Meta-analysis provides evidence-based interpretation guidelines for the clinical significance of mean differences for the FACT-G, a cancer-specific quality of life questionnaire

Madeleine T King¹
David Cella²
David Osoba³
Martin Stockler⁴
David Eton⁵
Joanna Thompson⁶
Amy Eisenstein⁷

¹Psycho-oncology Co-operative Research Group School of Psychology, University of Sydney, New South Wales, Australia; ²Department of Medical Social Sciences, Northwestern University, Illinois, USA; ³QOL Consulting, Vancouver, British Columbia, Canada; ⁴NHMRC Clinical Trials Centre, University of Sydney, New South Wales, Australia; ⁵Mayo Clinic, Rochester, Minnesota, USA; ⁶Centre for Health Economics Research and Evaluation, University of Technology, Sydney, New South Wales, Australia; ⁷Center on Outcomes Research and Education (CORE), Evanston Northwestern Healthcare (ENH), Evanston, Illinois, USA

Abstract: Our aim was to develop evidence-based interpretation guidelines for the Functional Assessment of Cancer Therapy-General (FACT-G), a cancer-specific health-related quality of life (HRQOL) instrument, from a range of clinically relevant anchors, incorporating expert judgment about clinical significance. Three clinicians with many years’ experience managing cancer patients and using HRQOL outcomes in clinical research reviewed 71 papers. Blinded to the FACT-G results, they considered the clinical anchors associated with each FACT-G mean difference, predicted which dimensions of HRQOL would be affected, and whether the effects would be trivial, small, moderate, or large. These size classes were defined in terms of clinical relevance. The experts’ judgments were then linked with FACT-G mean differences, and inverse-variance weighted mean differences were calculated for each size class. Small, medium, and large differences (95% confidence interval) from 1,118 cross-sectional comparisons were as follows: physical well-being 1.9 (0.6–3.2), 4.1 (2.7–5.5), 8.7 (5.2–12); functional well-being 2.0 (0.5–3.5), 3.8 (2.0–5.5), 8.8 (4.3–13); emotional well-being 1.0 (0.1–2.6), 1.9 (0.3–3.5), no large differences; social well-being 0.7 (−0.7 to 2.1), 0.8 (−2.9 to 4.5), no large differences. Results from 436 longitudinal comparisons tended to be smaller than the corresponding cross-sectional results. These results augment other interpretation guidelines for FACT-G with information on sample size, power calculations, and interpretation of cancer clinical trials that use FACT-G.

Keywords: health-related quality of life, patient-reported outcomes

Introduction

Health-related quality of life (HRQOL) questionnaires have the potential to play a key role in bringing the patient’s voice to evidence-based health care. But to realize this potential, we need to interpret the relevance of HRQOL outcomes to decisions about treatment. Such decisions are made at both the individual level, when a patient (or patient’s clinician, acting as patient’s agent) chooses among treatment options, and the group level, when clinical research is conducted to test the effectiveness of new treatments relative to current routine treatment.¹² In this article, we focus on the latter. For example, if a clinical trial shows that a new treatment improves the mean HRQOL by 10 points on a 100-point scale at no extra cost and with no adverse effects relative to the current best treatment, which improves mean HRQOL by only 5 points, we need to know whether this difference is big enough to change policy and practice. Investigators planning new studies need this type of knowledge to calculate sample sizes, and end-users need to understand the implications of study results for future clinical practice and policy.
Despite increasing acceptance and use of HRQOL questionnaires as valid and informative measures in clinical research, interpreting the clinical relevance of outcomes measured on their scales remains challenging. HRQOL is multifaceted and subjective, and there are a large number and wide range of measurement scales, each of which has a different and somewhat arbitrary scale. Few people have the requisite understanding of psychometrics or the hands-on experience to confidently interpret specific HRQOL scales. Familiarity with a measurement scale, whether it measures physical phenomena such as temperature or perception-based phenomena such as HRQOL, is the key to intuitive understanding of results recorded on that scale. Such familiarity develops with experience of how the numbers from such scales correspond to events in our lives and the significance of those to our actions and decisions. The multitude of HRQOL instruments compounds the problem of developing familiarity with specific HRQOL scales. Some HRQOL questionnaires are now very widely used. The 2 most commonly used cancer-specific instruments are the European Organisation for Research and Treatment of Cancer’s Quality of Life Questionnaire (QLQ)-C30 and the functional assessment of cancer therapy-general (FACT-G). Collective experience with these 2 instruments has amassed a rich evidence base for developing interpretation guidelines.

Clinicians who have specialized in HRQOL research and used particular HRQOL measures over many years in both research and clinical contexts are ideally positioned to develop intuitive interpretations of HRQOL results. Over many years of managing patients, administering HRQOL questionnaires, analyzing and scrutinizing the results, and linking HRQOL scores with the condition of their patients, they understand the variation among patients in their levels of HRQOL, how they react to their disease and treatment, and how their HRQOL scores change over time. They also understand how to interpret the mean scores from groups of such individuals.

Some of the most useful evidence for developing interpretations that are meaningful to clinicians comes from studies that report the HRQOL of patients grouped by established clinical criteria, sometimes called ‘clinical anchors’ or ‘known groups’. This type of data is often collected during the validation of an instrument, with patterns of HRQOL scores across clinical groups providing evidence of clinical validity. Quantitatively, these patterns can be used to develop interpretation guidelines about the relative size and significance of mean differences on HRQOL scales. Similarly, the patterns in longitudinal data collected during conventional treatments with well known and understood clear clinical effects help us understand the relative size and significance of mean changes in HRQOL. This method has been applied to the QLQ-C30, a cancer-specific HRQOL questionnaire. In this article, we further develop this method and apply it to another cancer-specific HRQOL questionnaire, the FACT-G.

The FACT-G forms the central core of a suite of questionnaires, referred to as the Functional Assessment of Chronic Illness Therapy (FACIT). Additional questions focus on specific diagnoses such as lung cancer (FACT-L) or breast cancer (FACT-B), and on symptoms such as fatigue (FACT-F) and anemia (FACT-An). These questionnaires are used increasingly in clinical research, adding to the body of evidence about the FACIT suite. In this article, we focus on the FACT-G, which is summarized as a total score and 4 subscales: physical well-being (PWB), social or family well-being (SWB), emotional well-being (EWB), and functional well-being (FWB).

We have previously described a new approach to synthesizing evidence about HRQOL measures with information on the interpretation of effect sizes derived from the FACT-G. In that paper, we focused on effect sizes, comparing our results with Cohen’s guidelines for small, medium, and large effect sizes, and the proposition that a 0.5 SD is the minimum significant difference. The current article presents analogous results for the raw scores from the 5 scales of the FACT-G.

Methods

A detailed description of the methods in relation to effect sizes is given elsewhere. Condensed extracts included here allow readers to assess the meta-analysis of the 5 FACT-G scale scores reported in this article.

Data sources

Papers that reported on the FACT-G were identified by searching relevant online databases. Unpublished information was identified through the FACIT Projects Register. Papers were included if they reported the mean difference at least between 2 independent groups (a cross-sectional contrast) or the mean change within at least 1 group over time (a longitudinal contrast). Results that represented duplicate publication were excluded, as were results from a total sample of less than 10 patients (potentially unreliable) or repeated measures from a sample with greater than 20% attrition (potentially biased).

Expert judgment

The clinical relevance of included mean differences was judged by 3 of us (DC, DO, MS). All identifying information
on papers was obscured, as were any results and conclusions about the FACT-G scores. Each expert predicted the degree of difference in HRQOL for each mean difference between groups or within a group over time. Judgments were constrained to 8 options: much better (3), moderately better (2), a little better (1), much the same (0), a little worse (−1), moderately worse (−2), much worse (−3), and “don’t know”. After the initial round of judgments, contrasts for which any 2 judges’ scores differed by 2 or more categories were reconsidered. The 3 experts worked independently in both the initial and the consensus phases. Weighted kappa was calculated as a measure of concordance of the final judgment scores, and we interpreted kappa values after Landis and Koch. The average of all 3 experts’ judgments, rounded to the nearest integer, was used to determine the final size class for each mean difference.

Definitions of size categories

Prior to the panel judging any papers, we defined 4 size groups explicitly in terms of their clinical relevance, where a clinically relevant difference was one that implied a difference in prognosis or clinical management. When circumstances had obvious and unequivocal clinical relevance (eg, patients with asymptomatic, early stage disease vs those with end-stage disease, or the change induced by a treatment well known to markedly improve the health state of most patients treated), group-level HRQOL was expected to be much better or worse (large effect). When circumstances were likely to have clinical relevance but to a lesser extent (eg, for patients with metastatic disease, the contrast of those who were responding to treatment compared with those who were not responding, or a treatment that was known to be effective for half the patients treated), group-level HRQOL was expected to be moderately better or worse (moderate effect). When effects were expected to be subtle but nevertheless clinically relevant (eg, the contrast of patients with regionally advanced cancer vs those with newly diagnosed metastatic disease, or a treatment that was known to improve the health state of only a small proportion of patients treated), group-level HRQOL was expected to be a little better or worse (small effect). When circumstances were unlikely to have any clinical relevance, group-level HRQOL was not expected to be any better or worse (trivial effect).

Data extraction

The following information was extracted from each paper: sample sizes and attrition; standard deviations (required for the inverse-variance weighting factor); or other information from which standard deviations could be derived; the clinical classifications and circumstances of patients (including anchors for mean differences in HRQOL); and mean scores of the PWB, FWB, EWB, and SWB scales. The total score was calculated as the sum of the PWB, FWB, EWB, and SWB scales. Each mean difference was then linked with the corresponding average expert judgment score to allocate a size class for the meta-analysis.

Meta-analysis

Inverse-variance weighted mean differences (IWMDs) were calculated for each of 4 size classes by grouping corresponding negative and positive size classes. Thus, all contrasts with a −1 average judgment score were grouped with those with a +1 score. To maintain the correct relationship between the sign of the reported HRQOL differences and the sign of the experts’ judgment score, the signs of the HRQOL differences with −1 average judgment scores were reversed prior to grouping them with the +1 contrasts. Results for the medium (−2/+2) and large (−3/+3) contrasts were treated analogously. Cross-sectional contrasts were analyzed separately from longitudinal contrasts.

Sensitivity analysis

We assessed the robustness of results based on all mean differences to discordance between experts by considering only those where at least 2 experts were perfectly concordant, and at most only 1 was discordant by 1 point.

Results

Of 210 papers that satisfied the search criteria, 71 were suitable for inclusion. The full citation details and a summary of the characteristics of these papers can be found in the companion paper. The purpose of the most common study was to describe the effect of disease and treatment on HRQOL (44%), followed by developing and validating a FACIT instrument (21%) and phase 2 clinical trial (7%). A wide range of clinical anchors were reported (Table 1). The most common anchor for differences in HRQOL between groups was routine treatment; 25% of the 71 studies reported mean FACT-G scores by this anchor. The second most commonly used cross-sectional anchor was extent of disease (24%), and the third was performance status (24%) (usually Eastern Cooperative Oncology Group, assessed by the clinician). The most common anchor for change in HRQOL over time was time since starting treatment; 30% of the 71 studies reported mean FACT-G scores by this anchor. The second most commonly used longitudinal anchor was change in
The 71 papers and 22 anchors yielded 1,562 mean differences. For 8 of these, none of the 3 experts were able to make a prediction. In the remainder, the experts differed by 2 or more points for only 17% (261) of their initial judgments. The consensus process reduced this to 6% (95 contrasts). A detailed consideration of expert concordance is given in the companion paper.8

Table 2 shows the results of meta-analysis of the raw scale scores from 1,118 cross-sectional contrasts, 436 longitudinal contrasts, and the number of mean differences in each size class. Table 3 shows analogous results from 617 (55% of 1,118) cross-sectional contrasts and 216 (50% of 436) longitudinal contrasts, in which at least 2 experts were perfectly concordant and at least only 1 was discordant by 1 point.

The results in Tables 2 and 3 are very similar, indicating that our results are robust to the degree of agreement between experts. Therefore, we focus on Table 2 hereafter. While we cite point estimates here for simplicity, we urge readers to note the range of possible values for each size class and domain suggested by the confidence intervals (CIs) reported in the tables. Across all domains, for both cross-sectional and longitudinal contrasts, IWMDs considered “trivial” by the experts were very small, and their CIs contained zero.

For the other size classes, we consider the cross-sectional results first. The PWB and FWB scales were similar, with IWMDs of 2, 4, and 9 for small, medium, and large effects, respectively. Results for the EWB scale were about half that size, with IWMDs of 1 and 2 for small and medium effects, respectively. Results for the SWB scale were smaller again, there was no gradient from small to medium effects, and all CIs contained zero. For the total FACT-G score, small and medium IWMDs were 6 and 11 points, respectively. Only 2 contrasts were predicted to yield a large effect for the total score, and none was predicted to yield large effects for the EWB or SWB domain.

We now consider the longitudinal results in Table 2. The IWMDs for the PWB and FWB domains were a little less than half the size of the corresponding cross-sectional values, with IWMD for small effects being somewhat less than 1, and medium effects having IWMDs of 1.5. Small effects for the EWB scale had an IWMD of about 1. Very little evidence was available to estimate medium effects for the EWB and SWB domains, and there was virtually no evidence available for large effects, with only 1 predicted in the PWB domain. The expected gradient across size classes was most pronounced in the PWB and FWB domains and not apparent at all in the SWB domain. For the SWB domain, all but one of the contrasts was predicted to yield trivial or small effects, and for the small effects, IWMDs were very small and their CIs contained zero.

**Discussion**

This study (which focuses on FACT-G raw scores) and its companion (which focuses on effect sizes)8 provide the first formal meta-analysis of anchor-based evidence for a HRQOL instrument, the FACT-G. This evidence covers a wide range of clinically meaningful anchors, and is judged by 3 clinicians with many years of experience managing individual cancer patients and using HRQOL outcomes in cancer clinical trials. In summarizing our results for the FACT-G raw scores, we heed the advice of Guyatt et al2 to avoid misleading oversimplifications and overly complex presentations. We believe that interpretation guidelines for HRQOL scales require some flexibility to accommodate different patient groups and clinical circumstances, so we summarize our results for each size class and domain as likely ranges. Thus, for the PWB and FWB scales, cross-sectional anchors suggest that a small effect is likely to be in the vicinity of 1–3 points, a medium effect in the vicinity of 3–5 points, and a
Table 2: Results of meta-analysis of the FACT-G raw scale scores

<table>
<thead>
<tr>
<th>Number of contrasts</th>
<th>Physical well-being</th>
<th>Functional well-being</th>
<th>Emotional well-being</th>
<th>Social well-being</th>
<th>Total FACT-G</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>X</td>
<td>L</td>
<td>X</td>
<td>L</td>
<td>X</td>
</tr>
<tr>
<td>Trivial</td>
<td>40</td>
<td>22</td>
<td>46</td>
<td>26</td>
<td>66</td>
</tr>
<tr>
<td>Small</td>
<td>97</td>
<td>32</td>
<td>97</td>
<td>36</td>
<td>122</td>
</tr>
<tr>
<td>Medium</td>
<td>85</td>
<td>33</td>
<td>81</td>
<td>30</td>
<td>44</td>
</tr>
<tr>
<td>Large</td>
<td>13</td>
<td>1</td>
<td>10</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>Cross-sectional (X)</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
</tr>
<tr>
<td>Trivial</td>
<td>−0.09</td>
<td>−2.2, 2.0</td>
<td>0.3</td>
<td>−2.0, 2.7</td>
<td>−0.16</td>
</tr>
<tr>
<td>Small</td>
<td>1.9</td>
<td>0.6, 3.2</td>
<td>2.0</td>
<td>0.5, 3.5</td>
<td>1.0</td>
</tr>
<tr>
<td>Medium</td>
<td>4.1</td>
<td>2.7, 5.5</td>
<td>3.8</td>
<td>2.0, 5.5</td>
<td>1.9</td>
</tr>
<tr>
<td>Large</td>
<td>8.7</td>
<td>5.2, 12</td>
<td>8.8</td>
<td>4.3, 13</td>
<td>–</td>
</tr>
<tr>
<td>Longitudinal (L)</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
</tr>
<tr>
<td>Trivial</td>
<td>0.5</td>
<td>−1.1, 2.0</td>
<td>−0.04</td>
<td>−1.8, 1.7</td>
<td>0.3</td>
</tr>
<tr>
<td>Small</td>
<td>0.8</td>
<td>−0.4, 2.1</td>
<td>0.7</td>
<td>−0.8, 2.2</td>
<td>1.1</td>
</tr>
<tr>
<td>Medium</td>
<td>1.5</td>
<td>0.1, 2.9</td>
<td>1.5</td>
<td>−0.2, 3.2</td>
<td>0.6</td>
</tr>
<tr>
<td>Large</td>
<td>8.2</td>
<td>–</td>
<td>–</td>
<td>–</td>
<td>–</td>
</tr>
</tbody>
</table>

Note: The scores are from all cross-sectional (X, n = 1,118) informative contrasts and longitudinal (L, n = 436) informative contrasts: number of contrasts and IWMD, with 95% CIs for the 5 FACT-G scales and the 4 size classes.

Abbreviations: FACT-G, Functional Assessment of Cancer Therapy-General; IWMDs, inverse-variance weighted mean differences; CIs, confidence intervals.

Table 3: Sensitivity analysis results of meta-analysis of the FACT-G scale scores

<table>
<thead>
<tr>
<th>Number of contrasts</th>
<th>Physical well-being</th>
<th>Functional well-being</th>
<th>Emotional well-being</th>
<th>Social well-being</th>
<th>Total FACT-G</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>X</td>
<td>L</td>
<td>X</td>
<td>L</td>
<td>X</td>
</tr>
<tr>
<td>Trivial</td>
<td>28</td>
<td>14</td>
<td>32</td>
<td>16</td>
<td>51</td>
</tr>
<tr>
<td>Small</td>
<td>56</td>
<td>16</td>
<td>61</td>
<td>20</td>
<td>46</td>
</tr>
<tr>
<td>Medium</td>
<td>41</td>
<td>14</td>
<td>37</td>
<td>13</td>
<td>6</td>
</tr>
<tr>
<td>Large</td>
<td>13</td>
<td>1</td>
<td>10</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>Cross-sectional (X)</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
</tr>
<tr>
<td>Trivial</td>
<td>0.04</td>
<td>−2.4, 2.5</td>
<td>0.7</td>
<td>−2.1, 3.5</td>
<td>−0.06</td>
</tr>
<tr>
<td>Small</td>
<td>2.0</td>
<td>0.3, 3.7</td>
<td>2.2</td>
<td>0.3, 4.1</td>
<td>1.1</td>
</tr>
<tr>
<td>Medium</td>
<td>4.5</td>
<td>2.5, 6.5</td>
<td>4.6</td>
<td>2.1, 7.1</td>
<td>1.8</td>
</tr>
<tr>
<td>Large</td>
<td>8.7</td>
<td>5.2, 12</td>
<td>8.8</td>
<td>4.3, 13</td>
<td>–</td>
</tr>
<tr>
<td>Longitudinal (L)</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
<td>95% CI</td>
<td>IWMD</td>
</tr>
<tr>
<td>Trivial</td>
<td>0.02</td>
<td>−1.9, 2.0</td>
<td>−0.2</td>
<td>−2.2, 1.9</td>
<td>0.2</td>
</tr>
<tr>
<td>Small</td>
<td>0.9</td>
<td>−0.7, 2.5</td>
<td>0.8</td>
<td>−1.4, 3.0</td>
<td>1.7</td>
</tr>
<tr>
<td>Medium</td>
<td>2.4</td>
<td>0.3, 4.5</td>
<td>1.2</td>
<td>−1.3, 3.7</td>
<td>2.1</td>
</tr>
<tr>
<td>Large</td>
<td>8.2</td>
<td>−7.5, 24</td>
<td>–</td>
<td>–</td>
<td>–</td>
</tr>
</tbody>
</table>

Note: The scores are from the subset of cross-sectional (X, n = 617/1118, 55%) contrasts and longitudinal (L, n = 216/436, 50%) contrasts in which at least 2 experts were perfectly concordant and up to 1 was discordant by at most 1 point: number of contrasts and IWMDs, with 95% CIs for the 5 FACT-G scales and the 4 size classes.

Abbreviations: FACT-G, Functional Assessment of Cancer Therapy-General; IWMDs, inverse-variance weighted mean differences; CIs, confidence intervals.
large effects unlikely. The small effect estimate is similar to the general guideline estimate for the FACT-G minimally important difference (MID), which is 4–7 points.\textsuperscript{15} For the total score, results from longitudinal anchors were smaller than corresponding cross-sectional results and were rather unconvincing as interpretation guidelines as their CI contained zero despite reasonable sample sizes.

It is worth noting that for a given size class, the longitudinal mean differences were smaller than cross-sectional ones for the PWB and FWB scales (and consequently the total score). This pattern has been observed previously for the FACT-An and FACT-F scales.\textsuperscript{16} The responsiveness of the FACT-G scales has been demonstrated in many papers,\textsuperscript{17–26} so we discount lack of responsiveness as an explanation. Several other factors provide more plausible explanations. Response shift, response sets, and other factors related to adjustment may diminish the true size of self-reported change.\textsuperscript{27} Another possibility is that our experts overestimated the magnitude of a health domain change likely to be observed with treatment, change of disease course, or other longitudinal anchor. Yet another is that our clinical experts may not have understood the longitudinal anchors as well as they did the cross-sectional anchors; lower rates of concordance among the experts for the longitudinal contrasts than for the cross-sectional contrasts lend some support to this hypothesis. It is also possible that the longitudinal HRQOL assessments may not have occurred at the best time to capture the effects anticipated by the experts. This is a common problem in longitudinal research since HRQOL is commonly assessed at clinic visits, which is convenient and maximizes response rates but does not necessarily capture the peaks and troughs in HRQOL trajectories.\textsuperscript{28} Finally, there may have been a minimizing bias introduced by sample attrition; patients who were most likely to deteriorate by the largest amounts were also most likely to be lost to follow-up. In this study, we only included within-group contrasts with less than 20% attrition to minimize this problem. All of these factors made the experts’ task of predicting change more difficult than that of separating groups. Some or all of these factors may be working in concert, underlining the challenges in conducting and interpreting longitudinal HRQOL research.

The other finding of interest was that the IWMDs for the social and emotional well-being scales were smaller than those for the physical and functional domains, a pattern evident also for the QLQ-C30.\textsuperscript{7} Three general factors may explain these observations. First, the clinical classifications and circumstances prevalent in cancer research may relate more to physical and functional aspects of HRQOL than to psychosocial domains. Second, following the first, our clinical experts may not have been as accurate in their predictions for psychosocial domains as they were for the physical and functional domains. Lower rates of concordance among the experts in the EWB and SWB domains relative to the PWB and FWB domains lend some support to this hypothesis. Third, scales such as the FACT-G’s EWB and SWB and the QLQ-C30’s emotional, social and cognitive functioning scales may not be as sensitive to real differences in psychosocial aspects of HRQOL as scales such as the FACT-G’s PWB and FWB and the QLQ-C30’s physical and role functioning scales are sensitive to real differences in physical and functional aspects of HRQOL. We believe the first 2 reasons are more likely than the third since several studies have shown change in emotional well-being using the EWB scale.\textsuperscript{5,29–32}

Whatever the reasons for the observation above and if our results generalize to other HRQOL instruments, then the following implications may hold for choice of HRQOL outcomes in cancer trials. HRQOL domains with physical and functional focus may generally yield larger mean differences and hence provide more powerful measures of outcome than do psychosocially focused domains, where at best small effects may be expected. Scales based on social or family well-being may be suitable primary outcomes only for studies of psychosocial interventions targeted specifically at social and family issues in which pilot studies or phase 2 trials have demonstrated an effect for these outcomes.

There was considerable variation in empirical estimates of mean differences within each size category and for each FACT-G scale. There are 2 obvious contributing factors: sampling variation and a degree of mismatch between our experts’ expectations and the actual patterns in the HRQOL data. Our 3 experts collectively had a wealth of clinical experience with cancer patients and with HRQOL assessment, so their judgments should be as good as any available. The influence of sampling variation on the outcomes of individual studies cannot be discounted since it is well documented that individuals vary markedly in HRQOL levels at a particular time and in their trajectories of HRQOL over time. The degree of variation of component estimates within size classes in this meta-analysis highlights the limitations of individual studies for deriving general interpretation guidelines.

Other authors have produced evidence across clinical anchors and studies to develop interpretation guidelines.\textsuperscript{7,10,15,16,25} Our method advances this type of research in 2 important ways. First, we used formal methods of meta-analysis to produce weighted average mean differences
for each size class. Second, clinical meaningfulness was judged by 3 clinicians with many years’ experience managing individual cancer patients and using HRQOL outcomes in cancer clinical trials. Our experts were blinded to the FACT-G scores because we wanted them to place a value on the significance of differences (as determined by the clinical characteristics and circumstances of the patients) rather than to describe the magnitude of differences. Further, our definitions of size classes explicitly address the relevance of HRQOL results to clinical decision making, thereby providing a direct link to Jaeschke et al’s widely cited definition of the minimum clinically important difference,33 more recently modified by Norman et al10 and Schünemann et al.34 Rather than focusing on the MID, we accommodate the possibility that in some circumstances, the MID may be of a moderate absolute size while in others it may be relatively small.

The results presented in this article augment other interpretation guidelines for the FACT-G,15,16,25 adding a substantial evidence base not previously considered for this purpose. We have thereby provided a comprehensive synthesis of anchor-based evidence for the 5 FACT-G scales, incorporating the collective understanding of 3 clinicians with many years’ experience managing individual cancer patients and using HRQOL outcomes in cancer clinical trials.

Acknowledgments
This research was funded by an educational grant from AstraZeneca. We are indebted to Liz Chichen, the librarian at the Centre for Health Economics Research and Evaluation, University of Technology, Sydney, Australia, for developing and testing our electronic search strategy, for identifying and searching all relevant online bibliographic databases, for helping identify potential source papers, and for collecting them. We would also like to thank Julia Brown, Peter Fayers, Kim Hawkins, and Galina Velikova for helpful comments about the results.

Disclosure
The authors report no conflicts of interest in this work.

References


