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Nomogram for Predicting Risk of Digestive Carcinoma Among Patients with Type 2 Diabetes

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Purpose: Digestive carcinomas remain a major health burden worldwide and are closely related to type 2 diabetes. The aim of this study was to develop and validate a digestive carcinoma risk prediction model to identify high-risk individuals among those with type 2 diabetes.

Patients and Methods: The prediction model was developed in a primary cohort that consisted of 655 patients with type 2 diabetes. Data were collected from November 2013 to December 2018. Clinical parameters and demographic characteristics were analyzed by logistic regression to develop a model to predict the risk of digestive carcinomas; then, a nomogram was constructed. The performance of the nomogram was assessed with respect to calibration, discrimination, and clinical usefulness. The results were internally validated by a bootstrapping procedure. The independent validation cohort consisted of 275 patients from January 2019 to December 2019.

Results: Predictors in the prediction nomogram included sex, age, insulin use, and body mass index. The model showed good discrimination (C-index 0.747 [95% CI, 0.718–0.791]) and calibration (Hosmer–Lemeshow test P=0.541). The nomogram showed similar discrimination in the validation cohort (C-index 0.706 [95% CI, 0.682–0.755]) and good calibration (Hosmer–Lemeshow test P=0.418). Decision curve analysis demonstrated that the nomogram would be clinically useful.

Conclusion: We developed a low-cost and low-risk model based on clinical and demographic parameters to help identify patients with type 2 diabetes who might benefit from digestive cancer screening.

Keywords: type 2 diabetes, digestive cancer, prediction, demographic

Introduction

Both diabetes and cancer are prevalent diseases whose incidence is increasing globally and which seriously threaten human health.^{1,2} More attention is paid to complications of diabetes such as cardiovascular disease, blindness, kidney failure, and lower limb amputation, and the relationship between diabetes and cancer is rarely considered. However, a meta-analytic study in 2009 by Vigneri et al showed that the risk of several cancer types (including cancers of the pancreas, liver, breast, colorectal, urinary tract, and female reproductive organs) was increased in diabetic patients, with a moderate increase in mortality.³ There is now growing recognition that there may be a meaningful association between diabetes and cancer,⁴ although the potential biological links between the two conditions are incompletely understood.¹

Digestive carcinomas remain a major health burden worldwide, especially in Asia.⁵ Studies show that the incidence of digestive carcinomas and type 2 diabetes has

increased exponentially in the past decade.^{6,7} Accumulating evidence indicates that digestive carcinomas are closely associated with type 2 diabetes,^{8,9} especially primary liver cancer, pancreatic cancer, and colorectal cancer.^{10–13} Epidemiological data suggest that diabetes increases the relative risk of digestive cancers more than twofold,¹ and type 2 diabetes mellitus is associated with more aggressive digestive tract cancers.¹⁴ Universal screening using potential biomarkers such as cancer antigen 19–9, carcinoembryonic antigens, or clinical tests (eg, computed tomography, endoscopic ultrasound, or magnetic resonance imaging) is not feasible,^{15–17} as the existing biomarkers have limited diagnostic accuracy. Therefore, there is an urgent clinical need for novel prediction models and screening methods that are both efficient and cost-effective for the detection of asymptomatic cases.¹⁸

The nomogram is considered a reliable tool that can be used to create a simple intuitive graph of a statistical predictive model that quantifies the risk of a clinical event.^{19,20} In this study, we aimed to identify a combination of variables that would enable highly accurate prediction of digestive carcinomas in patients with type 2 diabetes. A nomogram for predicting the risk of digestive carcinomas was then constructed to support recommendations based on physical examination.

Patients and Methods

This retrospective study was performed with approval from the Ethics Committee of the Affiliated Tumor Hospital of Guangxi Medical University (approval no. LW2020010). As this was a retrospective analysis of anonymized clinically obtained data and all patient identifiers were removed, there was no need for patients to sign the informed consent. This study was conducted in accordance with the tenets of the Declaration of Helsinki.

Patients

The primary cohort of this study comprised 655 patients in our hospital from November 2013 to December 2018 who had been diagnosed with type 2 diabetes. From January 2019 to December 2019, an independent validation cohort consisting of 275 consecutive patients was screened using the same criteria as the primary cohort. The proportions of patients with digestive carcinoma were 37.8% and 40.7% in the primary and validation cohorts, respectively. The inclusion criteria were: (1) a clear diagnosis of type 2 diabetes before the diagnosis of carcinoma; and (2) clinical pathology that supported the diagnosis of carcinoma. Exclusion criteria were: (1) the presence of two or more cancer types; (2) pancreatic cancers diagnosed within 5 years after the type 2 diabetes diagnosis (owing to reverse causation^{1,21}); (3) incomplete patient data.

Data Collection

Participants were divided into digestive carcinoma and nondigestive carcinoma groups; approximately 20% of the latter group had tumors of other systems, including lung, breast, and cervical cancers. Clinical parameters and demographic data were collected, including age, sex, diabetes duration (months), comorbidities (hypertension, coronary artery disease, cerebrovascular disease), use of blood glucoselowering drugs, history of smoking, body mass index (BMI), and family history of cancer. Diabetes duration was defined as the time from first diagnosis of type 2 diabetes to the diagnosis of malignancy; the diagnosis of type 2 diabetes in our institution complies with World Health Organization criteria:²² (1) diabetes symptoms plus random blood glucose \geq 11.1 mmol/L, or (2) fasting blood glucose \geq 7.0 mmol/L, or (3) 2-hour blood glucose oral glucose tolerance test ≥ 11.1 mmol/L. Blood glucose-lowering drugs included insulin, thiazolidinediones, alpha-glucosidase inhibitors, and sulfonylureas. Cerebrovascular disease was defined as transient ischemic attack, stroke, history of carotid artery stent, or cerebral hemorrhage. Coronary artery disease was defined as a history of myocardial infarction, coronary artery bypass grafting, or percutaneous coronary intervention. Family history of cancer was defined as more than one relative with first-degree carcinoma. Smoking was defined as the number of packs of cigarettes (1 pack = 20 cigarettes) smoked per day in a year, including current smokers and ex-smokers.²³ Overweight was defined as 24 < BMI < 28 kg/m², and obese was defined as BMI≥28 kg/m², as is standard for the Chinese population.

Statistical Analysis Construction of the Nomogram

The statistical analyses were performed and graphics were produced with SPSS version 26.0 (IBM Corp, Armonk, NY, USA) and R software (the rms package of the R library) version 3.6.2 (<u>https://www.r-project.org/</u>). For all analyses, P<0.05 was considered to be statistically significant and all tests were two-tailed unless otherwise indicated. For continuous variables, data are presented as median and interquartile range (M (P_{25} , P_{75})) or mean \pm standard deviation. Categorical variables are presented as whole numbers and proportions. Differences between groups were assessed using the chi-squared test or Wilcoxon rank-sum test with SPSS. The significance of each variable in the primary cohort was assessed by univariate logistic regression to investigate the independent risk factors for digestive cancers. All variables with P<0.05 in the univariate logistic analyses were further assessed by multivariable logistic regression with backward step-wise selection, using the likelihood ratio test with Akaike's information criterion as the stopping rule.²⁴ To provide the clinician with a quantitative tool to evaluate digestive carcinoma risk in patients with type 2 diabetes, a nomogram for digestive carcinoma risk was constructed on the basis of the results from the final multivariable logistic regression using the R software.²⁰

Performance of the Nomogram in the Primary Cohort

Calibration curves were plotted to calibrate the nomogram.²⁰ The C-index was used to evaluate discriminative ability, ranging from 0.5 (absence of discrimination) to 1 (perfect discrimination).²⁵ In addition, the nomogram was subjected to 1000 bootstrap resamples for internal validation to assess predictive accuracies.²⁶

External Validation of the Nomogram

The performance of the nomogram was tested in the validation cohort. The logistic regression formula was developed based on the primary cohort and applied to all patients in the validation cohort, and total points were calculated for each patient. Logistic regression of the validation cohort was performed using total points as a factor. The C-index and calibration curve were derived on the basis of the regression analysis.

Clinical Use of the Nomogram

Decision curve analysis was conducted using the R library rmda package to determine the clinical usefulness of the nomogram by quantifying the net benefit at different threshold probabilities in the primary dataset. The net benefit was calculated by subtracting the proportion of false positives from the proportion of true positives and weighting by the relative harm of forgoing treatment compared with the negative consequences of an unnecessary treatment.²⁷

Results

Characteristics of Patients in the Primary and Validation Cohorts

A total of 930 patients were enrolled in our study, with 655 and 275 patients assigned to the primary and validation cohorts, respectively. Table 1 shows the patient characteristics

Table I	Characteristics	of Patients	in the	Primary	and \	/alidation
Cohorts						

Variable	Primary Cohort	Validation Cohort	P-value
Age, years	59 (53,65)	61 (54,66)	0.054
Comorbidities Hypertension, yes, N (%) Coronary artery disease, yes, N (%) Cerebrovascular disease, yes, N (%) Family history of carcinomas, yes, N (%)	260 (39.7%) 156 (23.8%) 126 (19.2%) 89 (13.6%)	99 (36.0%) 55 (20.0%) 48 (17.5%) 34 (12.4%)	0.291 0.205 0.525 0.615
Diabetes duration, month	69 (24,120)	81 (27,135)	0.962
Sex, N (%) Male Female	380 (58.0%) 275 (42.0%)	64 (59.6%) (40.4%)	0.647
Blood glucose-lowering drugs Insulin, yes, N (%) Thiazolidinediones, yes, N (%) Alpha glucosidase inhibitors, yes, N (%) Sulfonylureas, yes, N (%) Metformin, yes, N (%) Combination of oral drugs and insulin, yes, N (%) Combination of oral drugs, yes, N (%) Smoking, yes, N (%)	125 (19.1%) 8 (1.2%) 72 (11.0%) 85 (13.0%) 109 (16.6%) 26 (4.0%) 66 (10.1%) 223 (34.0%)	40 (14.5%) 8 (2.9%) 36 (13.1%) 44 (16.0%) 57 (20.7%) 8 (2.9%) 42 (15.3%) 95 (34.5%)	0.098 0.071 0.362 0.224 0.138 0.432 0.024 0.883
BMI, N (%) Normal Overweight Obese	385 (58.8%) 211 (32.2%) 59 (9.0%)	45 (52.7%) 0 (36.7%) 29 (0.5%)	0.234
Digestive carcinoma type, N (%) Esophageal cancer Gastric cancer Primary liver cancer Pancreatic cancer Colorectal cancer	9 (1.4%) 23 (3.5%) 114 (17.4) 7 (1.1%) 95 (14.5%)	7 (2.5%) 11 (4.0%) 45 (16.4%) 5 (1.8%) 44 (16.0%)	0.620

Notes: $24 \leq BMI \leq 28 \text{ kg/m}^2$ is defined as overweight and $\geq 28 \text{ kg/m}^2$ is defined as obese for the Chinese population.

Abbreviation: BMI, body mass index.

by cohort. The baseline clinical characteristics were similar between the two cohorts, with digestive carcinoma proportions of 37.8% and 40.7% in the primary and validation cohorts, respectively (P=0.061).

Construction of the Nomogram

The results of the univariate logistic analysis are presented in Table 2. According to the multivariate analysis, with results reported as odds ratio (95% CI), age (1.021 [1.002–1.040]),

Variable	β	P-value	Odds Ratio (95% CI)
Sex	-1.916	<0.001	0.147 (0.099–0.218)
Age	-0.548	0.006	1.024 (1.007–1.042)
Comorbidities			
Hypertension	0.276	0.089	1.318 (0.958–1.814)
Coronary artery disease	0.315	0.093	1.370 (0.949–1.977)
Cerebrovascular disease	0.365	0.070	1.440 (0.971–2.138)
Diabetes duration	0.004	0.004	1.004 (1.001–1.006)
Family history of carcinomas	0.132	0.572	1.141 (0.722–1.805)
Smoking	0.729	<0.001	2.074 (1.487–2.893)
BMI	-0.462	<0.001	0.630 (0.486–0.816)
Use of blood glucose-lowering			
drugs			
Insulin	0.706	<0.001	2.207 (1.366–3.006)
Thiazolidinediones	-0.557	0.497	0.573 (0.115–2.861)
Alpha glucosidase inhibitors	-0.026	0.921	0.975 (0.586–1.622)
Sulfonylureas	-0.248	0.318	0.780 (0.479–1.270)
Metformin	-0.045	0.838	0.956 (0.623–1.468)
Combination of oral drugs	0.572	0.154	1.771 (0.807–3.886)
and insulin			
Combination of oral drugs	-0.238	0.392	0.788 (0.458–1.359)

Table 2 Univariate Logistic Regression Analysis of FactorsRelating to Digestive Carcinomas in the Primary Cohort

Abbreviation: BMI, body mass index.

sex (0.124 [0.078–0.198]), insulin (1.806 [1.162–2.806]), and BMI (0.630 [0.486–0.816]) were independently associated with digestive carcinoma (Table 3). The model that incorporated the above independent predictors was developed as the nomogram (Figure 1).

Performance of the Nomogram in the Primary Cohort

The calibration curve showed that the risk of digestive carcinoma predicted by the nomogram agreed well with the actual probabilities in the primary cohort (Figure 2). The Hosmer– Lemeshow test yielded a nonsignificant statistic (P=0.541), which suggested that our model was well calibrated with no

Table 3MultivariateLogisticRegressionAnalysisofFactorsRelating toDigestiveCarcinomas in the PrimaryCohort

Variable	β	P-value	Odds Ratio (95% CI)		
Sex	-2.089	<0.001	0.124 (0.078–0.198)		
Age	0.021	0.029	1.021 (1.002–1.040)		
Insulin	0.591	0.009	1.806 (1.162–2.806)		
Smoking	-0.393	0.065	0.675 (0.444–1.025)		
ВМІ	-0.364	0.014	0.695 (0.521–0.928)		
Diabetes duration	0.002	0.122	1.002 (0.999–1.005)		

Notes: Hosmer–Lemeshow test: χ^2 =6.962, P=0.541. C-index: 0.747 (0.718–0.791).

departure from perfect fitting. The C-index for the prediction nomogram was 0.747 (95% CI 0.718–0.791) for the primary cohort, corrected to 0.734 via bootstrapping validation (B=1000), which suggested good discrimination by our model.

External Validation of the Nomogram in the Validation Cohort

Good calibration and discriminative ability were also observed for the nomogram in the validation cohort (Figure 3). The Hosmer–Lemeshow test yielded a nonsignificant statistic (P=0.418), and the C-index of the nomogram for the risk of digestive cancers was 0.706 (95% CI 0.682–0.755).

Clinical Use of the Model

The results of the decision curve analysis for the nomogram are presented in Figure 4. This graphically shows the clinical usefulness of the nomogram based on a continuum of potential thresholds for digestive cancer risk (x-axis) and the net benefit of using the nomogram to stratify patients by risk (y-axis), assuming that no patient will have digestive cancer. In this analysis, the final decision curve showed that for a threshold probability between 10% and 65%, the model had positive net benefit. For example, if the personal threshold probability of a patient was 30% (ie, the patient would opt for further detailed screening if his risk of digestive cancer were 30–65%), the net benefit would be 0.2 when using the nomogram to decide whether to conduct further detailed screening, with added benefit compared with screen-all or screen-none schemes.

Discussion

Early detection and treatment of digestive cancers are important to improve patient quality of life. In this study, we developed and validated a simple intuitive statistical predictive model to quantify the risk of digestive carcinoma, which could support clinicians and patients when making screening decisions for patients with type 2 diabetes. The easy-to-use nomogram incorporates four clinical risk factors (sex, age, insulin use, and BMI) to facilitate risk prediction for digestive carcinoma in patients with type 2 diabetes. The prediction model for digestive carcinoma showed good discrimination and calibration in type 2 diabetes; the C-index was 0.747 (0.734 via bootstrapping validation) and the Hosmer– Lemeshow test yielded a nonsignificant statistic (P=0.541) in the primary cohort, with a similar result in the validation cohort.

Although type 2 diabetes and cancer share several risk factors, including obesity, aging, unhealthy food, and

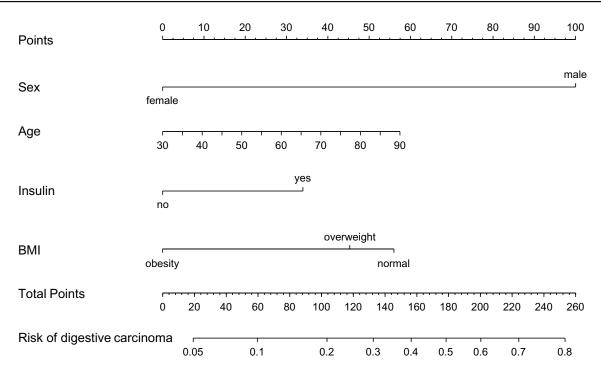


Figure 1 Nomogram developed on the basis of the primary cohort, with sex, age, insulin use, and BMI incorporated.

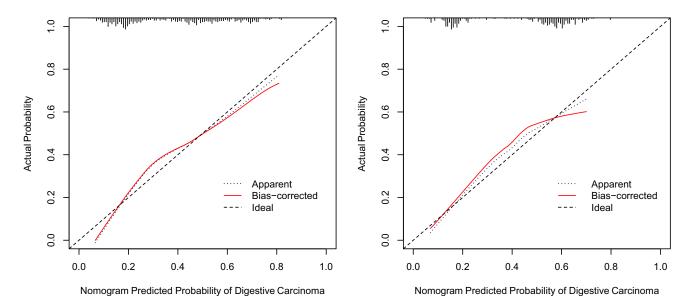


Figure 2 Calibration curves for the nomogram in the primary cohort. The blue dotted line represents the entire cohort (n=655), and the red solid line is the result after bias-correction by bootstrapping (1000 repetitions), indicating nomogram performance.

Figure 3 Calibration curves for the nomogram in the validation cohort. The blue dotted line represents the entire cohort (n=275), and the red solid line is the result after bias-correction by bootstrapping (1000 repetitions), indicating nomogram performance.

physical inactivity,¹ an association between type 2 diabetes and the risk of digestive carcinoma is biologically plausible based on the findings of previous studies.^{1,28,29} To best of our knowledge, few studies have developed or validated risk prediction models for digestive carcinomas in patients with type 2 diabetes. In 2016, Boursi et al³⁰ developed

a model to assess risk of pancreatic cancer among patients with new-onset diabetes, which yielded a good area under the curve (0.82). Hsieh et $al^{31,32}$ used artificial neural network models to assess risk of colorectal and pancreatic cancer among patients with type 2 diabetes mellitus. However, our study differs from previous development

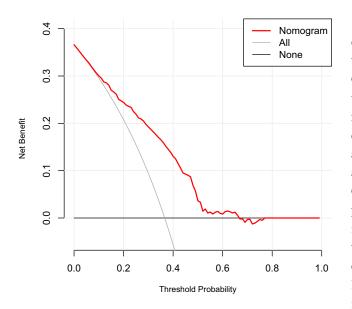


Figure 4 Decision curve analysis for the nomogram.

and validation work as it involves a comprehensive evaluation, whereas other studies developed a single disease prediction model for type 2 diabetes mellitus.

Furthermore, obesity is generally thought to be associated with the development of many cancers, especially digestive system malignancies,³³ which contradicts our results. However, although a meta-analysis by Dong et al³⁴ showed that it is more appropriate to use abdominal obesity to measure the relationship between cancer and obesity, our study could only measure obesity by BMI owing to its retrospective nature. On the other hand, research on the relationship between obesity and digestive cancer was mainly carried out on populations from Western countries.^{35,36} Eastern and Western countries have very different dietary habits, and there are also certain differences in body shape. This may have led to the conflicts between our results and those of other studies.

In China, 36.4% of cancer-related deaths involve digestive tract cancers (stomach, liver, and esophageal cancers), which have relatively poorer prognoses.³⁷ Currently, surgery and/or chemotherapy are the most effective treatment for digestive carcinomas.³⁸ Some cancers, including primary liver cancer, colorectal cancer, and stomach cancer, can be clinically cured by surgery in the early stages.³⁸ However, most patients with digestive carcinomas are diagnosed at advanced stages of the disease, and the overall 5-year survival rate for these patients is low.^{39–41} There is already evidence that diabetes (mainly type 2) and digestive cancers are closely related;^{1,14} therefore, the implementation of population-based screening programs may significantly reduce the mortality and incidence of digestive cancers in patients with type 2 diabetes.

However, although screening for cancer has many benefits, there are also potential risks. Overscreening may lead to overtreatment and cause more harm than gain. Gastrointestinal endoscopy is an important and painless test for digestive tumors, but it also can delay patient recovery and discharge time and increases the risk of cardiopulmonary complications, which are understood to account for ~50% of mortality and ~60% of morbidity associated with upper gastrointestinal endoscopy.^{42,43} Conducting mass digestive carcinoma screening of all patients with type 2 diabetes patients using costly and/or invasive tests would not be an efficient approach, because the vast majority of these patients do not have digestive carcinomas. Our prediction model provides a low-cost and low-risk solution to this problem by identifying high-risk individuals for definitive diagnostic testing. This easy-touse scoring system could be used to perform an individualized prediction of the risk of digestive cancers.

The most important and final argument for the use of the nomogram is based on individual needs for additional investigation or care. However, despite better risk prediction performance, calibration, and discrimination, the nomogram could not capture the clinical consequences of a particular level of discrimination or degree of miscalibration.44-46 Therefore, to justify its clinical use, decision curve analysis was applied to assess whether the nomogram-based decisions would be helpful. This novel method offers insight into clinical consequences on the basis of a threshold probability from which net benefit can be derived.⁴⁴ The decision curve showed that for a threshold probability between 10% and 65%, the model would result in a positive net benefit. For example, if the personal threshold probability of a patient was 30% (ie, the patient would opt for further detailed screening if his risk of digestive cancer were 30-65%), the net benefit would be 0.2 when using the nomogram to decide whether to conduct further detailed screening, with added benefit compared with a screen-all or screen-none scheme.

The nomogram developed here has certain limitations. First, our prediction nomogram was constructed retrospectively; several known risk factors for digestive cancers, such as race, exercise, and dietary pattern, were not included in the database; and the sample size was small. As such, some analyses may have been limited, and the use of the model may be restricted to certain regions and races. However, the current nomogram model was shown to have excellent predictive power. Second, the model was not intended to be a definitive diagnostic test but rather to identify type 2 diabetes patients at high risk for digestive cancers. Third, a key limitation of this work is that all digestive cancers were combined, when they are in fact very different cancers and have different risks associated with type 2 diabetes. However, its sole reliance on demographic information means that this model could be easily applied in practice to virtually every patient with type 2 diabetes mellitus with negligible cost or risk. Further data collection, wider geographic recruitment, and incorporation of positron emission tomography results and molecular factors could improve the model for future use.

Conclusions

In this work, we have developed a low-cost and low-risk model based on clinical and demographic parameters to identify patients with type 2 diabetes who might benefit from digestive cancer screening. Both physicians and type 2 diabetes patients could perform individualized prediction of the risk of digestive cancers with this easy-to-use scoring system.

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Disclosure

The authors declare that there are no conflicts of interest related to this work.

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