The accuracy assessment of presepsin (sCD14-ST) for mortality prediction in adult patients with sepsis and a head-to-head comparison to PCT: a meta-analysis

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Objective: The soluble cluster of differentiation 14 subtype (sCD14-ST) or presepsin has recently been identified as a promising biomarker in sepsis. The present meta-analysis is performed to assess the prognostic value of presepsin in septic patients. Further, we compare the prognostic performance between presepsin and procalcitonin (PCT) in predicting all-cause mortality in these patients.

Methods: A systematic and comprehensive search was conducted in PubMed, Embase and Cochrane databases by using Exploded Medical Subject Headings and appropriate corresponding keywords. Studies were eligible if they assessed the prognostic value of presepsin in sepsis and provided sufficient information to construct a 2×2 contingency table. A bivariate meta-analysis model was used to calculate the pooled sensitivity, specificity, positive/negative likelihood ratios and diagnostic odds ratio. The Chi-square and I2 index were used to assess the heterogeneity and inconsistency. The Deek’s funnel plot asymmetry test was used to assess the likelihood of publication bias.

Results: Nine publications, comprising 1,561 patients, were included in this study. The overall area under the receiver operating characteristic curve (AUROC) of presepsin was 0.77 (95% CI, 0.73–0.81) with a pooled prognostic sensitivity (SEN) and specificity (SPE) of 0.83 (95% CI, 0.72–0.90) and 0.69 (95% CI, 0.63–0.74), respectively. Additionally, the PLR, NLR and DOR of presepsin were 2.6 (95% CI, 2.1–3.3), 0.25 (95% CI, 0.15–0.44) and 0.44 (95% CI, 0.33–0.94), respectively. Furthermore, the AUROC of PCT was 0.81 (95% CI, 0.78–0.84) with a pooled SEN of 0.76 (95% CI, 0.55–0.89) and SPE of 0.74 (95% CI, 0.33–0.94). There is no statistically significant difference in the performance of pooled SEN and SPE between presepsin and PCT, with a p value of 0.39 and 0.71, respectively.

Conclusions: Based on the results of this meta-analysis, both presepsin and PCT are promising biomarkers for the prognosis of mortality in sepsis.

Keywords: presepsin, sCD14-ST, procalcitonin, prognostic, sepsis, mortality

Introduction
Sepsis is a life-threatening organ dysfunction characterized by a dysregulated host response to infection. Sepsis accounts for about 25% of intensive care unit (ICU) admissions, with a mortality range from 10% to 52%.1–3 Despite advances and breakthroughs in antibiotic treatment or other managements of bundled care for the patients with sepsis, the fatality rate remains unacceptably high and the incidence is
To improve sepsis-related survival, it is imperative to early recognize septic patients at high risk of poor clinical outcomes. However, due to the lack of an ideal prognostic indicator, timely identification of patients at risk of dying from the condition is challenging.

To identify mortality risk and ensure the appropriate therapeutic interventions, clinical scores have been introduced. In clinical practice, the most commonly used clinical scores are the Acute Physiology and Chronic Health Evaluation (APACHE) II score and Sequential Organ Failure Assessment (SOFA) score. They are validated as the most recognized tools to stratify the severity of the condition. However, with the increasing controversies and complicated methods for using these clinical scoring systems, a growing body of evidence has proposed blood biomarkers as promising alternatives.

Hundreds of the circulating biomarkers have been investigated for potential use in improving early diagnosis and monitoring the outcome of septic patients. However, due to the controversial and uncertain factors in clinical value, fewer biomarkers were used in the practical medical work. Among the most extensively studied in recent years, PCT, CRP, pentraxin-3, IL-6 and myeloid cells expressing triggering receptor-1 (TREM-1) have been identified as probable predictors in sepsis.

Among those, procalcitonin (PCT) is considered as the most preferable biological predictors and is a widely recognized biomarker in predicting outcomes in septic patients. Presepsin, a truncated N-terminal fragment of CD14, also known as soluble cluster of differentiation 14 subtype (sCD14-ST), has recently attracted great clinical interest and risen up as a promising prognostic serum marker in sepsis. The generation and release of the presepsin are cleaved from the monocyte/macrophage-specific CD14 receptor complex which is binding with bacterial lipopolysaccharide (LPS) and LPS binding protein (LPB) during the pro-inflammatory signaling cascade. Previous studies showed that presepsin existed in much higher concentration in septic patients than that in healthy individuals and those with Systemic Inflammatory Response Syndrome (SIRS). Moreover, publications revealed that the level of presepsin increased within 2 hrs after infection, which has proved to be reliable for early stage sepsis diagnosis and prognostic stratification. However, some researchers proposed that presepsin may not outweigh traditional biomarkers in prognosis of survival rate. Yang et al recently published a meta-analysis to assess the prognostic value of presepsin in adult sepsis patients, which demonstrated that presepsin did have some mortality prediction value as evidenced by a significantly higher level of presepsin in non-survivors. We conducted this systematic review of the literature and a meta-analysis to further evaluate the relationship between presepsin and hospital mortality and provide additional information of pooled sensitivity and specificity to determine its prognostic performance. Since PCT is a widely used predictor in mortality, we further performed a head-to-head comparison with this well-known serum marker in the prognostic value of septic patients.

Materials and methods

Search strategy and selection criteria

We systematically searched studies using PubMed, Embase, and the Cochrane Library without language limitation through 15th August, 2018. The search terms were as follows: (presepsin OR “sCD14-ST” OR “sCD14-subtype” OR “soluble cluster of differentiation 14 subtype”) AND (sepsis OR SIRS OR “systemic inflammatory response syndrome” OR “septic shock” OR “severe sepsis” OR “bacteria infection”) AND (mortality OR prediction OR prognosis OR “prognostic value” OR “prognostic biomarker”). Additionally, to obtain potentially relevant articles, we reviewed the reference list of the selected articles, and one was included in this study.

Eligible studies were required to have a well-defined reference standard for patients recruited (sepsis or severe sepsis or septic shock) based on the criteria of the American College of Chest Physicians/Society of Critical Care Medicine in 1991/2001 (sepsis 1.0 and 2.0), International Guidelines for Management of Severe Sepsis and Septic Shock 2012 or the third international consensus definitions for sepsis and septic shock (sepsis 3.0). Further, a 2*2 contingency table should be available. If multiple studies reused the same patient sample, only the most recent article or the most informative article was included. For studies that evaluated the prognostic accuracy of presepsin levels at multiple time points, we chose the data relied on the initial presepsin levels of patients admitted to the ED or ICU. For studies predicting mortality in different follow-up periods across studies, we basically used the data on 28 days, 30 days and in-hospital stays if no exact time course was mentioned. Followed modalities such as reviews, letters, editorials, case reports, meeting abstracts, animal
experiments or meta-analysis, and articles involving pediatric patients were considered ineligible. Two investigators (Zhu Y and Li JD) independently conducted the search strategy and assessed the studies. Any disagreements were resolved by a third reviewer (Tao TZ).

Data extraction and quality evaluation
Descriptive data were extracted from the selected preferences, including the following items: 1) Surname of the first author and year of publication, 2) country and clinical setting, 3) study design, 4) characteristics of participants such as sample size, gender and age, 5) cut-off point, 6) true positives (TP), 7) false positives (FP), 8) false negatives (FN), 9) true negatives (TN), 10) the overall area under the ROC (AUROC), 11) sensitivity (SEN) and 12) specificity (SPE). Quality Assessment of Diagnostic Accuracy Studies (QUADAS-2) was performed to assess the risk of bias in each study, which was recommended by the Cochrane collaboration for the quality assessment of diagnostic studies.\(^{29,30}\)

Statistical analysis
All statistical and meta-analyses were conducted by using the MIDAS module in the software of STATA version 13 (Stata Corporation, College Station, TX, USA) Based on the hierarchical summary receiver operating characteristic (HSROC) method, the pooled sensitivity (SEN), specificity (SPE), diagnostic odds ratio (DOR), positive likelihood ratio (PLR), negative likelihood ratio (NLR) and the area under the receiver operating characteristic curve (AUC) were calculated for meta-analysis in the prognostic value of presepsin and PCT.

To further evaluate the performance of the predicting assessment, a Chi-square (\(x^2\)) test and the inconsistency \((I^2)\) metric were used to check heterogeneity of the included studies. The latter expresses the variability caused by heterogeneity between the studies in the form of a percentage. A value of \(I^2\) greater than 50% was considered sustainable heterogeneity among selected studies.\(^{31}\) The Deek's funnel plot asymmetry test was used to assess the likelihood of publication bias. In addition, Fagan’s nomogram was conducted to calculate the post-test probability (PTP).\(^{32}\)

Results
Characteristics of included studies
A total of nine publications including 1,561 patients were selected from 206 records by searching Pubmed (n=21), Cochrane Library databases (n=10) and Embase (n=175). The included studies were published from 2011 to 2017 (Figure 1). Three studies were conducted in Europe.\(^{22,33,34}\)

Figure 1 Flow chart of identification, inclusion and exclusion of studies for the meta-analysis.
<table>
<thead>
<tr>
<th>Study</th>
<th>Country setting</th>
<th>Study design</th>
<th>Sample size (n)</th>
<th>Mean age</th>
<th>Male/ female</th>
<th>Biomarker</th>
<th>Cut-off (pg/mL)</th>
<th>TP</th>
<th>FP</th>
<th>FN</th>
<th>TN</th>
<th>AUC (95% CI)</th>
<th>Sensitivity (%)</th>
<th>Specificity (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Spanuth, 2011</td>
<td>NA</td>
<td>NA</td>
<td>140</td>
<td>NA</td>
<td>NA</td>
<td>Presepsin</td>
<td>1622</td>
<td>21</td>
<td>32</td>
<td>2</td>
<td>85</td>
<td>0.878 (0.808–0.941)</td>
<td>93.3</td>
<td>72.5</td>
</tr>
<tr>
<td>ED</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>PCT</td>
<td>13.43</td>
<td>14</td>
<td>23</td>
<td>9</td>
<td>94</td>
<td>0.668 (0.496–0.840)</td>
<td>60.0</td>
<td>80.2</td>
</tr>
<tr>
<td>Liu, 2013</td>
<td>China PR</td>
<td>PR</td>
<td>859</td>
<td>71</td>
<td>513/346</td>
<td>Presepsin</td>
<td>556</td>
<td>181</td>
<td>189</td>
<td>110</td>
<td>379</td>
<td>0.658 (0.614–0.703)</td>
<td>62.2</td>
<td>66.8</td>
</tr>
<tr>
<td>ED</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>PCT</td>
<td>1.125</td>
<td>158</td>
<td>170</td>
<td>133</td>
<td>398</td>
<td>0.679 (0.636–0.722)</td>
<td>54.2</td>
<td>70.0</td>
</tr>
<tr>
<td>Masson, 2014</td>
<td>Italy RR</td>
<td>RR</td>
<td>100</td>
<td>71.5</td>
<td>54/46</td>
<td>Presepsin</td>
<td>1631</td>
<td>35</td>
<td>16</td>
<td>15</td>
<td>34</td>
<td>0.72 (0.61–0.82)</td>
<td>70.9</td>
<td>67.4</td>
</tr>
<tr>
<td></td>
<td>ICU</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>PCT</td>
<td>14.27</td>
<td>28</td>
<td>20</td>
<td>22</td>
<td>30</td>
<td>0.55 (0.44–0.67)</td>
<td>56.4</td>
<td>60.5</td>
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<tr>
<td>Carpio, 2015</td>
<td>Peru PR</td>
<td>PR</td>
<td>123</td>
<td>70</td>
<td>56/67</td>
<td>Presepsin</td>
<td>825</td>
<td>21</td>
<td>35</td>
<td>3</td>
<td>64</td>
<td>0.743</td>
<td>86.0</td>
<td>65.0</td>
</tr>
<tr>
<td>Ali, 2016</td>
<td>Egypt ICU</td>
<td>PR</td>
<td>33</td>
<td>55.2</td>
<td>10/23</td>
<td>Presepsin</td>
<td>957.5</td>
<td>18</td>
<td>2</td>
<td>1</td>
<td>12</td>
<td>0.891</td>
<td>94.7</td>
<td>85.7</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>PCT</td>
<td>2.6</td>
<td>15</td>
<td>0</td>
<td>4</td>
<td>14</td>
<td>(0.765–1.000)</td>
<td>78.9</td>
<td>100.0</td>
</tr>
<tr>
<td>Brodska, 2017</td>
<td>Czech ICU</td>
<td>PR</td>
<td>60</td>
<td>67</td>
<td>38/22</td>
<td>Presepsin</td>
<td>1853</td>
<td>15</td>
<td>10</td>
<td>5</td>
<td>30</td>
<td>0.734</td>
<td>75.0</td>
<td>75.0</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>PCT</td>
<td>1.8</td>
<td>13</td>
<td>0</td>
<td>7</td>
<td>40</td>
<td>0.844</td>
<td>65.0</td>
<td>100.0</td>
</tr>
<tr>
<td>El-Shafie, 2017</td>
<td>Egypt ICU</td>
<td>PR</td>
<td>31</td>
<td>60</td>
<td>16/15</td>
<td>Presepsin</td>
<td>NA</td>
<td>10</td>
<td>6</td>
<td>1</td>
<td>14</td>
<td>0.755 (0.571–0.938)</td>
<td>91.0</td>
<td>70.0</td>
</tr>
<tr>
<td>Kim, 2017</td>
<td>Korea ICU&amp;ED</td>
<td>RR</td>
<td>157</td>
<td>70</td>
<td>95/62</td>
<td>Presepsin</td>
<td>2455</td>
<td>26</td>
<td>57</td>
<td>8</td>
<td>66</td>
<td>0.684 (0.605–0.756)</td>
<td>76.5</td>
<td>53.7</td>
</tr>
<tr>
<td>Matera, 2019</td>
<td>Italy PR</td>
<td>PR</td>
<td>58</td>
<td>70</td>
<td>42/16</td>
<td>Presepsin</td>
<td>689</td>
<td>16</td>
<td>10</td>
<td>1</td>
<td>31</td>
<td>0.933 (0.824–0.985)</td>
<td>93.3</td>
<td>76.5</td>
</tr>
</tbody>
</table>

**Abbreviations:** PR, prospective recruitment; RR, retrospective recruitment; ED, emergency department; ICU, intensive care unit; NA, not available; PCT, procalcitonin; TP, true positives; FP, false positives; FN, false negatives; TN, true negatives; AUC, the overall area under the receiver operating characteristic curve; SEN, sensitivity; SPE, specificity; CI, confidence intervals.
two in Africa,\textsuperscript{35,36} two in Eastern Asia,\textsuperscript{34,37} one in South America\textsuperscript{38} and one was missing this information.\textsuperscript{25} All included studies were published in English. The mean age of the populations varied between 55.2 and 71.5 years, and the proportion of men ranged from 30.3\% to 72.4\%. Five studies were performed in the ICU,\textsuperscript{22,33,35,36,39} three in the ED,\textsuperscript{37,38,40} and one in both the ICU and ED.\textsuperscript{34} There are six prospective studies,\textsuperscript{22,35–39} two were designed as retrospective recruitments\textsuperscript{33,34} and one is unknown.\textsuperscript{25} Of these, diagnosis as sepsis was defined as “gold standard” by the criteria in 1991 ACCP/SCCM consensus conference (1992 criteria) in four studies,\textsuperscript{22,25,38,39} by 2001 ACCP/SCCM consensus conference in two,\textsuperscript{35,37} by the third international consensus definitions for sepsis and septic shock (sepsis-3) in two\textsuperscript{34,36} and one was carried out according to the Surviving Sepsis Campaign: International Guidelines for Management of Severe Sepsis and Septic Shock 2012.\textsuperscript{33} Levels of the circulating presepsin were measured by a chemiluminescent immunoassay on a PATHFAST immunoanalyzer in all studies (Mitsubishi Chemical or LSI Medience Corporation, Tokyo, Japan or Mitsubishi Chemical Europe GmbH, Dusseldorf, Germany).

Details of included studies involving data of presepsin and PCT are shown in Table 1.

### Assessment of studies quality

Details of the methodological quality assessments of nine included studies are summarized by using QUADAS-2 tool in Table 2 and Figure 2. All included studies showed acceptable quality based on the methodological assessment. All studies scored “low” in the domain of risk of bias in the reference standard since the 1991/2001 ACCP/SCCM consensus conference, sepsis-3.0 and Surviving Sepsis Campaign: International Guidelines for Management of Severe Sepsis and Septic Shock were all recognized for diagnosing sepsis in different periods.\textsuperscript{3,26–28} In patient selection, seven studies providing the criteria of inclusion and exclusion were scored

<table>
<thead>
<tr>
<th>Study</th>
<th>Patient selection</th>
<th>Index test</th>
<th>Reference standard</th>
<th>Flow and timing</th>
<th>Applicability concerns</th>
</tr>
</thead>
<tbody>
<tr>
<td>Spanuth\textsuperscript{25} 2011</td>
<td>?</td>
<td>?</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Liu\textsuperscript{37} 2013</td>
<td>?</td>
<td>?</td>
<td></td>
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<td>Masson\textsuperscript{33} 2014</td>
<td>?</td>
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<td>?</td>
<td>?</td>
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<td></td>
<td></td>
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<td></td>
</tr>
<tr>
<td>Brodsk\textsuperscript{22} 2017</td>
<td>?</td>
<td>?</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>EL-Shafie\textsuperscript{35} 2017</td>
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<td></td>
<td></td>
<td></td>
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<tr>
<td>Kim\textsuperscript{34} 2017</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Matera\textsuperscript{39} 2017</td>
<td></td>
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</tr>
</tbody>
</table>

Notes: \(\bigcirc\) Low Risk, \(\bigstar\) High Risk, \(\bigcirc\) Unclear Risk.

Figure 2 Risk of bias and applicability concerns.
“low” risk, two studies that did not show enough information marked “unclear” and another study that was designed as case-control scored “high”. Regarding the risk of bias of index tests, six studies not pre-specifying a threshold were defined as “high” risk, for two being designed as retrospective recruitment and one missing the information given “unclear”. As for patient flow and timing domain, seven studies scored “low” and two were “high” since some of the patients were lost to follow up.

Egger regression, investigation of heterogeneity and publication bias
The Egger’s regression model was performed to assess the likelihood of publication bias for nine included studies of presepsin and seven of PCT. The Deek’s funnel plot revealed that the slope coefficient for presepsin was associated with a p-value of less than 0.001 and 0.1 for PCT, suggesting the existence of significant publication bias for presepsin but not PCT in this meta-analysis (Figure 3). The overall $I^2$ values of presepsin were 82% (95% CI, 62–100) and 99% (95% CI, 98–99), indicating the significant presence of heterogeneity for both biomarkers.

Analysis of the prognostic value of presepsin in predicting mortality
A total of 9 studies with 1,561 patients were included in this group. The pooled SEN and SPE for presepsin were 0.83 (95% CI, 0.72–0.90) and 0.69 (95% CI, 0.63–0.74), respectively (Figure 4) with the overall area under the ROC (AUROC) of 0.77 (95% CI, 0.73–0.81) (Figure 5A). The PLR and NLR were 2.6 (95% CI, 2.1–3.3) and 0.25 (95% CI, 0.15–0.44), respectively. The Diagnostic Odds Ratio (DOR) was 10 (95% CI, 5–22). Fagan’s nomogram for likelihood ratios indicates that predicting mortality in sepsis enhanced the post-probability to 40% by presepsin when the consequences were positive and decreased the post-probability to 6% for presepsin while the results were turning to negative (Figure 6A).

Of all the studies included in our study, only 8 had provided a cut-off value for presepsin and 6 for PCT. Three studies had a cut-off value lower than 1,000 pg/mL, ranging from 556 pg/mL to 957.5 pg/mL. Two studies provided similar concentrations as 1,622 and 1,631 pg/mL, one as 1,853 pg/mL and one was higher than 2,000 pg/mL. As for PCT, the concentrations ranged from 0.16 ng/mL to 14.27 ng/mL (Table 1).

![Figure 3 Deek’s funnel plot asymmetry test for publication bias of presepsin (A) (p=0.001) and PCT (B) (p=0.1).](https://www.dovepress.com/fig3.png)
Performance comparison with PCT in predicting value on mortality

Seven studies with 1,407 patients were recruited for extracting the data of biomarker PCT, and the other 2 of the 9 studies were eliminated for not providing the information. The AUC of PCT was 0.81 (95% CI, 0.78–0.84) with a pooled SEN of 0.76 (95% CI, 0.55–0.89) and SPE of 0.74 (95% CI, 0.33–0.94) (Figures 5B–7). The PLR, NLR and DOR were 2.9 (95% CI, 0.9–9.5), 0.32 (95% CI, 0.17–0.62) and 9 (95% CI, 2–41), accordingly. We found that in contrast to the PCT, the pooled sensitivity of presepsin was slightly higher (0.83, 95% CI, 0.72–0.90 vs 0.76, 95% CI, 0.55–0.89) and the pooled specificity was a bit lower (0.69, 95% CI 0.63–0.74 vs 0.74, 95% CI 0.33–0.94). However, the analytic consequences showed no statistically significant difference between two predictors (p-value=0.39 and 0.71, respectively, for presepsin and PCT). Additionally, Fagan’s nomogram for likelihood ratios indicated that predicting mortality in sepsis enhanced the post-probability to 42% by PCT when the results were positive and decreased the post-probability to 7% as it was turning to negative, which showed the similar predicting ability of mortality with presepsin as well (Figure 6B). The summary analysis of prognostic information between presepsin and PCT is presented in Table 3, indicating the comparable prognostication in sepsis outcome of presepsin with PCT.

Discussion

Sepsis is more a complicated pathophysiological condition rather than a specific syndrome. Despite advanced modern life support, fatality and morbidity remain high in the past two decades.41,42 Early recognition of high-risk-to-die patients may prevent the high morbidity and improve the clinical outcome in this subpopulation. Hence, it is imperative to find a reliable prognostic biomarker with high sensitivity and specificity that can help aid in differentiating septic patients in a fast and cost-effective manner. Yang et al recently reported a meta-analysis showing that presepsin concentrations in decedents outweigh that in survivors.24 However, our research presented additional
evidence with sensitivity and specificity for assessing the accuracy of presepsin in predicting mortality of septic patients and further made a head-to-head comparison with PCT, a widely used biomarker in assessing the predictive value. We found that there is no statistically significant difference between the two predictors, implying that both presepsin and PCT are promising prognostic biomarkers of sepsis.

In our study, the results showed that the presepsin reached a pooled sensitivity of 0.83 and specificity of 0.69, suggesting a relatively high rate of missed prognosis by 17% and misprognosis of 31%. The overall PLR was 2.6, indicating that the likelihood of a positive presepsin test result was almost 3-fold higher in septic non-survivors than in patients who survived. The pooled NLR was 0.25, suggesting that the possibility of a negative result was only a chance of one-fourth. Further, the AUC of 0.77 indicated an acceptable prognostic accuracy that presepsin could be a new and promising biomarker for predicting sepsis mortality.

We further compared the results with PCT, a useful but also controversial biomarker in predicting mortality for sepsis. Some suggested that the levels of PCT showed no significant difference between survivors and decedents and PCT could also rise transiently in patients with non-septic situations such as trauma, surgery or systemic inflammatory response syndrome (SIRS). In addition, studies suggested that the peak of PCT level was reached later than 24 hrs, which might not be reasonably early enough as biomarkers for sepsis diagnosis and stratification of the sepsis severity. However, to date, PCT is extensively accepted for mirroring the poor outcome in infectious diseases and is also recognized for diagnosis, guided therapy and prognostication of patients with sepsis. In a recent meta-analysis, Prkno et al found that under a PCT-guided medicine treatment, the use of antibiotic therapy decreased remarkably from 8 days in the control group to 6 days. The lesser duration of the antibiotic consumption appeared to be associated with a reduction in the mortality, morbidity, length of hospitalization and health care costs. More importantly, recent studies show that PCT might confer positive prediction value in mortality. In Liu’s meta-analysis exploring the effect of PCT on mortality, the pooled SEN was 0.72 (95% CI, 0.58–0.82) compared to 0.76 (95% CI, 0.55–0.89) in our result and SPE was 0.77 (95% CI, 0.55–0.90) compared to our data of 0.74 (95% CI, 0.63–0.84).
The AUC in Liu’s publication was 0.70 (95% CI, 0.75–0.83) and that in our study is 0.81 (95% CI, 0.78–0.84). In spite of the fact that publication bias existed for the limited number of articles about PCT included in the present study, through the comparison with the published meta-analysis, our results demonstrated that PCT is a stable and reliable biomarker in the prognostic value on mortality. Therefore, it is credible for us to make the contrast of the predicting value in these two biomarkers. Though the pooled SEN for presepsin is higher while the SPE is inferior than that of PCT, there showed no statistically significant difference in both groups, with a p-value of 0.39 and 0.71. Hence, presepsin could be another promising predictor, which appears to be comparable to PCT in predicting mortality in septic patients.

Figure 6 Fagan’s nomogram for likelihood ratios of presepsin (A) and PCT (B).

Presepsin was a newly identified marker for sepsis in 2004, and the use of presepsin to make an assessment of severity and prognosis was first reported in 2011. The prominent advantage of this indicator was the early secretion after infection, while the peak concentration of its counterpart was observed later than 24 hrs after infection. This could mean that when presepsin is secreted at its maximum, the level of PCT has not increased completely or reached its summit. This earlier increase in the concentration of presepsin may also partially explain why presepsin expresses a slightly higher sensitivity than PCT in our present study. In 2013 Guidelines of the Surviving Sepsis Campaign, it was recommended that confirming the potential infectious site within 6 hrs of presentation was critical to the outcomes in sepsis patients and that...
broad-spectrum antibiotic intervention must be treated within 1 hr after sepsis identification. Presepsin was also considered as more of an independent mortality predictor than PCT. However, there is some inherent limitation for this novel serum biological marker. Presepsin is detected while being releasing into the circulation after activating host cells by recognition of bacterial lipopolysaccharide (LPS), which is more of a characteristic of gram-negative bacteria. Thus, in general, the levels of presepsin were often examined elevated most in the gram-negative organisms, followed by fungi and lesser in gram-positive bacteria. In addition, in the same way that PCT could be detected in other critical conditions such as multiple trauma, pancreatitis and extensive burns, several studies revealed that high levels of presepsin were also found in patients with cardiac surgery. Particularly, presepsin was found to be elevated in patients with renal or liver dysfunction such as severe chronic kidney disease or liver cirrhosis without infection. This could be explained by the fact that presepsin was detained in the serum when the function of kidneys had failed and the degradation ability of liver was impaired. Therefore, more studies including sepsis patients with or without kidney and liver dysfunction should be conducted to fully exploit the diagnostic and prognostic value of presepsin. Considering the merits and demerits of both biomarkers, the combination use of presepsin and PCT appears to provide greater clinical value.

Significant publication bias was found after being visually tested by the Deek’s funnel plot. The exclusion of the following modalities, including review, letters, case reports, abstracts of conference and negative studies, etc., are the probable sources attributed to the publication bias in this review. In addition, language bias could be a factor for the reason that we include articles written in English.

This systematic review and meta-analysis has several limitations as well. First, the study presented some heterogeneity which is inherent in the present available data. The cut-off value revealed a large gap varying from 556 to 2455 pg/mL among different studies despite the same
presepsin assay method. This threshold effect caused by diverse cut-off values could be one of the reasons and thus the optimal cut-off could not be defined. In addition, the reference standard was not uniform, ranging from ACCP/SCCM 1992 criteria, 2001 version (Sepsis 1.0 or 2.0), International Guidelines, to the newly defined sepsis 3.0. Based on the new guideline (sepsis 3.0), SIRS with infection is no longer defined as sepsis which might explain the highest presepsin level of 2,455 pg/mL in Kim’s article. And a further reason for this discrepancy could be from the study designs, with five in PR, two in RR, and one being unknown. With a limited number and high heterogeneity between selected studies, these results should be carefully and cautiously interpreted.

**Conclusion**

According to the results of our study, both presepsin and PCT are prospective predictors for contrasting decedents and survivors with sepsis. However, with the limitation of both biomarkers, the combination of presepsin and PCT might provide a more accurate prognostic value. Finally, continuous re-evaluation should be conducted to determine a promising mortality predictor with high sensitivity and specificity in the clinical practice and define an optimized cut-off value.

**Data sharing statement**

The data used to manage and analyses are archived and be available from the PubMed, Embase and Cochrane databases.

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**Author contributions**

All authors contributed to data analysis, drafting or revising the article, gave final approval of the version to be published, and agree to be accountable for all aspects of the work.
Disclosure
The authors have declared that no competing interests exist in this work.

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