REVIEW

NEDD9 overexpression predicts poor prognosis in solid cancers: a meta-analysis

This article was published in the following Dove Press journal: OncoTargets and Therapy

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Background: The oncogenicity of neural precursor cell-expressed developmentally downregulated 9 (NEDD9) has been demonstrated in multiple cancer types. However, the prognostic value of NEDD9 in some solid cancers remains controversial. Thus, this meta-analysis was conducted to evaluate the relationship between NEDD9 expression survival rates in solid tumors.

Method: Our meta-analysis included studies searched from various search engines with specific inclusion criteria and exclusion criteria. Combined HRs for overall survival (OS) and disease-free survival (DFS) or progression-free survival (PFS) or recurrence-free survival (RFS) or cancer-specific survival (CSS) were assessed using fixed-effects and random-effects models. The source of heterogeneity was identified by subgroup analysis. Additionally, publication bias was assessed using funnel plot and Egger's regression asymmetry test.

Result: Eighteen studies with a total of 2,476 patients were retrieved for analysis. Pooled HRs and 95% CIs were calculated. Both OS (HR=1.82; 95% CI: 1.43–2.31) and DFS/PFS/RFS/CSS (HR=2.54; 95% CI: 1.93–3.33) indicated that NEDD9 overexpression is associated with poor OS in cancer patients with solid tumors.

Conclusion: NEDD9 overexpression might be a potential marker to predict prognosis in solid cancer patients.

Keywords: NEDD9, solid cancer, prognosis, meta-analysis

Introduction

Epidemiological data show that cancers have become one of the world's leading causes of mortality.¹ From the latest data, solid cancers accounted for over 90% of all types of cancers and the top five of new cases and deaths are lung cancer, breast cancer, prostate cancer, colon cancer and nonmelanoma skin cancer.¹ Solid cancers are characterized by malignant tumors that form a discrete tumor mass. By contrast, lymphoproliferative malignancies diffusely infiltrate tissue without forming a mass. Although diagnosis of number of molecular markers and targeted therapies progressed with the efforts of many researchers and clinicians, the outcomes including death rates and overall survival (OS) of the majority of patients remain poor. Thus, more predictive molecular markers of solid cancers must be identified for prevention and individualized cancer treatment.

Neural precursor cell-expressed developmentally down-regulated 9 (NEDD9), also known as human enhancer of filamentation 1 (HEF1) and Crk-associated substrate lymphocyte type (Cas-L), belongs to the Crk-associated substrate family. NEDD9 coordinates the focal adhesion kinase and SRC signaling cascades that are involved in integrin-dependent adhesion and migration, invasion, cell apoptosis

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OncoTargets and Therapy 2019:12 4213-4222

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Methods

The meta-analysis was performed according to the PRISMA guideline and the statement for reporting systematic reviews and meta-analyses.¹⁵ Previously published studies were summarized and analyzed in this study (ethics board approval was not necessary).

Search strategy and study eligibility

We retrieved studies that measured NEDD9 expression and survival in solid cancer patients between 1995 and January 2019 from PubMed, EMBASE, Web of Science, and Google Scholar. The search terms included "NEDD9" or "neural precursor cell-expressed developmentally down-regulated 9" or "HEF1" or "human enhancer of filamentation 1"or "Cas-L" or "Crk-associated substrate lymphocyte type" and "neoplasms" or "cancer" or "tumor" or "malignancy" or "carcinoma" and "prognosis" or "survival". Non-English language studies were excluded. Results were restricted to human studies of solid cancer and 402 entries were found.

Inclusion criteria consisted of an evaluation of NEDD9 overexpression associated with OS (date of surgery to date of death as a result of any cause), disease-free survival (DFS) or progression-free survival (PFS) or recurrence-free survival (RFS) or cancerspecific survival (CSS), and immunohistochemistry (IHC) or ELISA analysis of NEDD9 expression. Tumors were classified by NEDD9 expression levels using the cutoffs defined in the studies (Table 1). All references of the included studies were scanned and studies of potential interest were reviewed for further analysis. Reviews, clinical endpoints other than OS or DFS/PFS/RFS/CSS, studies that enrolled less than 15 solid cancer patients, and studies without data that could be used to calculate the HR and 95% CI were excluded from our meta-analysis. Any disagreement was resolved by discussion among all investigators until a final consensus was reached.

Data collection and quality assessment

The included studies that were assessed independently by two investigators (J.L and Y.G). Disagreements were resolved by consensus. Data retrieved from the studies included the following: author, country, year of publication, patient number, gender, follow-up time, OS, DFS/PFS/ RFS/CSS, cutoff value for over/normal NEDD9 expression, histological type, and HR estimation. Some HR data were extracted from tables or Kaplan-Meier curves of NEDD9 overexpression and normal expression groups. We used the Newcastle-Ottawa Non-Randomized scale (NOS) as recommended by the Cochrane Non-Randomized Studies Methods Working Group for a quality assessment of the cohort studies.^{16,17} The NOS score ranged from 0 to 9 based on three board perspectives: study group selection, study group comparability, and ascertainment of the outcome of interest. A consensus NOS score for each study was obtained by a discussion between all investigators. Studies with a score ≥ 5 had a high-quality methodology.

Statistical analysis

Data were extracted from the primary publications and analyzed with Stata 12 (Stata Corporation, College Station, TX, USA). The HR and corresponding 95% CI estimates were calculated and pooled to determine the association of NEDD9 overexpression with OS and DFS/PFS/RFS/CSS. An HR>1 indicated a poor prognosis in patients with NEDD9 overexpression. The random-effects model and the fixed-effects model were conducted according to DerSimonian and Laird's method and Mantel-Haenszel's method, respectively.^{18,19} Statistical intrastudy heterogeneity was evaluated by the I^2 value to quantify the proportion of the total variation. The I² values of 25%, 50%, and 75% were the cutoff points of low, moderate, and high heterogeneity, respectively.²⁰ We used the fixed-effects model to pool the results if relatively low or moderated heterogeneity existed ($I^2 < 50\%$). The random-effects model was used when the I² value was \geq 50%. If high or moderate heterogeneity existed, a subgroup analysis of the cancer characteristics was conducted to determine possible causes.²¹ Differences between the subgroups were calculated from the Cochrane Handbook for Systematic Reviews of Interventions. Outcome credibility was validated by sensitivity analysis

Table I Main cha	racteristics	of all studies in	cluded in t	he meta-anal)	/sis							
Study	Country	Tumor type	Case number	Age (years)	Gender (M/F)	TNM stage (I/II/III/ IV)	Follow-up (months)	Detected method	Survival analysis	Cut-off value	Multivariate analysis	뚜
Afsar 2018 ³⁸	Turkey	Pancrentic cancer	32	Media 61	20/12	NR	Max 27	ELISA	OS/PFS	NR	٩	SC
El-Babouly 2018 ³⁶	Egypt	Bladder cancer	105	Mean 57	69/36	32/73(I-II/II-IV)	Mean 25	IHC	PFS	Hs ^a >4	No	S
Erturk 2018 ³⁵	Turkey	Melanoma	112	Mean 52	NR	NR	Mean 20.8	ELISA	so	NR	No	Report
Harb 2017 ³¹	Egypt	Bladder cancer	50	Mean 52	35/15	24/26 (Ta-T1/T2-T3) (T stage)	Max 35	IHC	so	Hs ^a >4	Yes	Report
Karabulut 2015 ³²	Turkey	Gastric cancer	68	Media 60	49/19	NR	Mean 8	ELISA	OS/PFS	NR	No	sC
Kondo 2012 ²³	Japan	Lung cancer	60	Mean 62.5	26/34	31/5/23/1(////////)	06-0	HC	OS/RFS	>30% of nuclear Stained	Yes	Report
Li 2014 ²⁴	China	Colorectal cancer	92	Mean 62	52/40	15/30/37/10(1/11/11/V) (UICC stage)	5-60	IHC	so	Hs ^a >4	Yse	Report
Li 2016 ²⁵	China	Breast cancer	226	Median 56	0/226	79/83/53/11((I/II/IV)	Median 27	IHC	OS/DFS	Hs ^b ≥4	Yes	Report
Liu 2014 ³⁹	China	Gastric cancer	187	R	104/83	77/110(1-11/11-1V)	0-60	НС	S	>25% of nuclear stained	°Z	SC
Lu 2015 ²⁶	China	Hepatocellular cancer	164	Median 49	133/31	32/73(I-II/III-IV) (UICC Stage)	0-60	НС	S	Hs ^b ≥4	Yes	Report
Miao 2013 ²⁷	China	Lung cancer	105	Mean 60.4	63/42	46/59(I-II/III-IV) (UICC stage)	0-60	IHC	so	Postive cell score ≥3	٥	SC
Ostojic 2018 ³³	Croatia	Lung cancer	12	Median 49.2	112/53	55/110(I-II/III-IV) (UICC stage)	0-125	IHC	so	Postive cell score ≥3	°N	SC
Shi 2014 ²⁸	China	Gastric cancer	125	NR	61/64	73/52(I-II/III-IV)	0-60	IHC	so	Hs ^a ≥4	Yes	Report
Wang 2014 ⁸	China	Ovarian cancer	129	NR	0/129	73/52(I-II/II-IV) (FIGO stage)	0-60	IHC	OS/PFS	Nuclear staining	Yes	Report
Wang 2017 ³⁷	China	Renal cancer	68	NR	69/39	53/15(1-11/111-1V)	0-60	IHC	CSS	Hs ^c ≥5	No	sC
Xue 2013 ²⁹	China	Pancrentic cancer	106	NR	65/41	54/52(I-II/III-IV)	0-60	IHC	so	Hs ^b ≥4	No	Report
											(Cor	itinued)

Ħ	Report	Report	tensity score
Multivariate analysis	Yes	Yes	ge score [0–3] × in core of 7).
Cut-off value	Hs ^b ≥4	Hs ^b ≥4	scores = percentag 0 and a maximum s
Survival analysis	so	so	ic score (overall imum score of
Detected method	IHC	НС	(2); HS ^b , histolog 4], with the min
Follow-up (months)	0-60	090	naximum score of I intensity score [0–
TNM stage (I/II/II/ IV)	61/114(Ta-T1/T2-T3)	127/135/231/108(1/11/ 111/1V)	minimum score of 0 and a m = percentage score [0–3] +
Gender (M/F)	117/58	428/173	0–3], with the overall scores
Age (years)	Median 61.5	Median 64	intensity score [istologic score (
Case number	175	109	score [0–4] × e of 9); HS ^c , I
Tumor type	Bladder cancer	Gastric cancer	scores = percentage nd a maximum scor
Country	China	China	score (overall s n score of 0 a
Study	Zhang 2014 ³⁴	Zhang 2014 ³⁰	Notes: HS^{a} , histologic s [0–3], with the minimun

Abbreviations: OS, overall survival; RFS, recurrence-free survival; PFS, progression-free survival; DFS, disease-free survival; CSS, cancer-specific survival; HR, hazard ratio; UICC, union for international cancer control; FIGO, international federation of gynecology and obstetrics; NR, not report; SC, survival curve; IHC, immunohistochemistry: ELISA, enzyme-linked immunosorbent assay **Dove**press

and the publication bias was assessed by visual inspection of the funnel plots. Funnel plot asymmetry was determined by Egger's linear regression test.²² A two-sided P < 0.05 was considered statistically significant.

Results

Search results and study characteristics

Eighteen studies were eligible for the pooled meta-analysis. Eligible studies were published between 2012 and February 2019 with a total of 2476 patients.^{8,23–38} A flowchart depicting the selection of literature is shown in Figure 1.

All included studies were retrospectively analyzed. The sample size ranged from 32 to 601 (median 137). NEDD9 overexpression was reported in 1309 (52.9%) of the 2476 included patients. Amongst all study cohorts, 3 evaluated lung cancer, 23,27,33 4 assessed gastric cancer, 28,30,32,39 3 undertook urinary bladder carcinoma, 31,34,36 2 illustrated pancreatic carcinoma,^{29,38} and single studies focused on breast cancer,²⁵ colorectal cancer,²⁴ hepatocellular carcinoma,²⁶ renal carcinoma,³⁷ melanoma,³⁵ and ovarian cancer.⁸ Studies were conducted in China (11), Japan (1), Turkey (3), Croatia (1), and Egypt (2). The Newcastle-Ottawa Quality Assessment Scale indicated that the quality scores ranged from 6 to 8 (median 7) with a relatively high study quality. The main characteristics of the 18 eligible publications are reported in Table 1. Seven studies^{8,23,25,32,36–38} reported that the HR of DFS/PFS/RFS/CSS associated with NEDD9 expression. Sixteen studies^{8,23–35,38,39} included OS and a follow-up time of 8-125 months. The HR and 95% CI were directly recorded in 11 studies.^{8,23-26,28-31,34,35} HR was inferred from the remaining studies using the Tierney method to estimate HR from Kaplan-Meier survival curves.

Meta-analysis results

Moderate intrastudy heterogeneity was observed ($I^2=58.4\%$) in 16 studies where OS was determined by univariate analysis or multivariate analysis. The random-effects model was used to pool the results. The combined HR estimate of OS was 1.82 (95% CI: 1.43–2.31; *P*<0.01). NEDD9 overexpression was associated with poor OS (Figure 2).

Studies were divided into five groups by cancer location for subgroup analysis: (1) digestive system cancer (n=8);^{24,26,28–30,32,38,39} (2) respiratory system cancer (n=3);^{23,27,33} (3) breast and ovarian cancer (n=2);^{8,25} (4) urinary system cancer (n=2);^{31,34} and (5) melanoma (n=1).³⁵ The

Table I (Continued)



Figure I Flow diagram of the study selection process.

Study D		HR (95% CI)	% Weight
Digestive system cancers			
Afsar(2018) -		1.30 (0.60, 2.79)	5.31
Karabulut(2015)	- i -	1.10 (0.56, 2.19)	6.01
Li (2014)		2.19 (1.02, 4.71)	5.34
Liu (2014)		1.52 (0.85, 2.73)	6.92
Lu(2015)	1.	2.42 (1.59, 3.69)	8.62
Shi(2014)		2.02 (1.05, 3.89)	6.25
Xue(2013)		1.45 (0.75, 2.62)	6.52
Zhang(2014)	•	1.36 (1.02, 1.81)	10.06
Subtotal (I-squared = 8.4%, P=0.365)	\diamond	1.61 (1.32, 1.95)	55.02
Respiratory system cancers			
Kondo(2012)		5.36 (1.98, 14.50)	3.87
Miao(2013)		2.15 (1.03, 4.48)	5.57
Ostojic(2018)	•	0.81 (0.50, 1.29)	8.04
Subtotal (I-squared = 85.0%, P=0.001)		1.96 (0.68, 5.70)	17.49
Breast and gynecological cancer			
Li(2016)	*	5.06 (1.21, 21.14)	2.26
Wang(2014)		3.79 (2.06, 6.97)	6.67
Subtotal (I-squared = 0.0%, P=0.716)	\sim	3.96 (2.26, 6.94)	8.93
Jrinary system cancers			
Harb(2017)		2.00 (0.95, 4.22)	5.49
Zhang(2014)	-	2.75 (1.53, 4.94)	6.89
Subtotal (I-squared = 0.0%, P=0.510)		2.44 (1.54, 3.86)	12.38
Melanoma			
Enturk/2018)	100	1 14 (0 59 2 22)	6 19
Subtotal (Leonared = $\% P = $)		1 14 (0.59, 2.22)	6 19
Subtotal (I-Squaled = .70, P =.)		1.14 (0.58, 2.21)	0.10
Overall (I-squared = 58.4%, P = 0.002)	💠	1.82 (1.43, 2.31)	100.00
NOTE: Weights are from random effects analysis			
	1	1	

Figure 2 Forest plot of NEDD9 overexpression with overall survival.

combined HR estimate for OS in the digestive system cancer group was 1.61 (95% CI: 1.32–1.95; I^2 =8.4%). The HR estimate for OS in the respiratory system cancer group was

1.96 (95% CI: 0.68–5.70; I^2 =85%), 3.96 (95% CI: 2.26–6.94; I^2 =0%) for breast and ovarian cancer groups, and 2.44 (95% CI: 1.54–3.86; I^2 =0%) for urinary system groups.

Study ID	HR (95% CI)	% Weight
Digestive system cancers		
Afsar(2018)	1.28 (0.59, 2.77)	12.44
Karabulut(2015)	1.96 (1.16, 3.30)	27.21
Subtotal (I-squared = 0.0%, P=0.371)	1.71 (1.11, 2.64)	39.64
Respiratory system cancers		
Kondo(2012)	4.24 (1.93, 9.26)	12.09
Subtotal (I-squared = .%, P=.)	4.24 (1.94, 9.29)	12.09
Breast and gynecological cancer		
Li(2016)	2.93 (1.04, 8.24)	6.94
Wang(2014)	2.49 (1.40, 4.43)	22.41
Subtotal (I-squared = 0.0%, P=0.788)	2.59 (1.56, 4.28)	29.35
Urinary system cancers		
Wang(2017)	1.59 (0.36, 6.96)	3.39
El-Babouly(2018)	♦ 4.94 (2.47, 9.86)	15.52
Subtotal (I-squared = 45.9%, P=0.174)	4.03 (2.15, 7.55)	18.91
Heterogeneity between groups: P=0.075		
Overall (I-squared = 37.6%, P=0.142)	2.54 (1.93, 3.33)	100.00
.101	1 9.86	

Figure 3 Forest plot of NEDD9 overexpression with DFS/PFS/RFS/CSS.

Abbreviations: DFS, disease-free survival; PFS, progression-free survival; RFS, recurrence-free survival; CSS, cancer-specific survival.

There was no apparent intrastudy heterogeneity ($I^2=0\%$) among 7 studies where DFS/PFS/RFS/CSS was determined by univariate analysis. The fixed-effects model was used to pool the results. The combined HR estimate was 2.54 (95% CI: 1.93–3.33; P=0.142), which indicated that NEDD9 over-expression was associated with poor DFS/PFS/RFS/CSS (Figure 3). The studies were subdivided by cancer location: digestive system (n=2),^{32,38} respiratory system (n=1),²³ breast and ovarian cancer (n=2),^{8,25} and urinary system cancer (n=2).^{36,37}The combined HR estimate for DFS/PFS/RFS/CSS were 1.71 (95% CI: 1.11–2.64; $I^2=0\%$) in the group digestive system cancer and 2.59 (95% CI: 1.56–4.28; $I^2=0\%$) in the group breast and ovarian cancers. The HR estimate for DFS/PFS/RFS/CSS in urinary cancer was 4.03 (95% CI: 2.15–7.55; $I^2=45.9\%$).

Publication bias

Funnel plot and Egger's test were used to evaluate the publication bias. The shape of the funnel plots did not

show obvious asymmetry for OS determined by univariate analysis or multivariate analysis (Figure 4). Egger's test also failed to provide evidence of significant publication bias. Publication bias was also investigated for DFS/PFS/ RFS/CSS obtained by univariate analysis and the funnel plots showed basic symmetry (Figure 5). Our results suggested that publication bias was not evident.

Discussion

Regarding prognostic value of level of NEDD9 and survival rates of solid tumors patients, we use the tool of metaanalysis to demonstrate that the high expression of NEDD9 acts as a prediction of the poor outcome for patients with solid tumors.

NEDD9 contains multiple functional modules for protein interaction, yet lacks any known enzymatic function. Adult lung and kidney tissues, which are rich in immature lymphoid cells, as well as the fetal brain, contain the highest levels of NEDD9.^{40–42} NEDD9 is crucial in



Figure 4 Funnel plots for OS. **Abbreviation:** OS, overall survival.



Figure 5 Funnel plots for DFS/PFS/RFS/CSS.

Abbreviations: DFS, disease-free survival; PFS, progression-free survival; RFS, recurrence-free survival; CSS, cancer-specific survival.

integrin-dependent signaling cascades to activate cell migration and is involved in lateral communication to Ras signaling cascades.⁷ NEDD9 can drive mesenchymal-type movement by activating RAC2 GTPase and WAVE

in complex with GEF DOCK3 to inhibit GTPase Rho and amoeboid movement.⁴³ Invasion is accompanied by ECM proteolysis through activation of MMP14, MMP2, and MMP9 metalloproteinases.⁴⁴ Epithelial–mesenchymal

transition (EMT) is a transdifferentiation strategy where carcinoma cells acquire the ability to execute the steps of invasion-metastasis cascade.⁴⁵ Epithelial cells express mesenchymal markers and cytoskeletal changes that increase motility and invasiveness during EMT. NEDD9 also interacts with key proteins to coordinate the signaling cascades controlling cell cycle transitions, apoptosis, migration, and invasion.^{4,5,46,47} NEDD9 expression has been linked to tumorigenesis of various malignancies, cancers of the breast,²⁵ colon.^{25,48,49} including pancreas,²⁹ head and neck,⁵⁰ ovaries,⁸ gastric,^{24,30,32} lung,^{23,27,51} liver,⁵² kidney,⁵³ glioblastoma,⁵⁴ and neuroblastoma.55 However, some studies have the opposite point of view. One example is a study from Ostojić et al (2018) that is also included in our meta-analysis.³³ It is interesting that they concluded the relationship between NEDD9 expression and survival did not reach statistically significant difference and there was a trend toward the correlation of higher NEDD9 to longer survival. The explanation is that expression of NEDD9 over certain level inevitably leads to apoptosis.⁵⁶ Furthermore, it is also possible that elevated expression of NEDD9 may indicate better responses to therapy.²³ Although there are some different opinions, our findings through metaanalysis suggested that elevated NEDD9 expression was correlated with poor survival in most solid cancers. Further studies are needed to elucidate the mechanisms of NEDD9 in cancer pathogenesis.

The prognostic value of NEDD9 in solid cancer is controversial despite previous investigations of this marker in multiple cancers. This is the first meta-analysis to analyze NEDD9 overexpression in relation to OS or DFS/PFS/RFS/CSS of solid cancer patients. We systematically evaluated the survival data of 2393 solid cancer patients from 18 eligible studies. We demonstrated that NEDD9 overexpression was associated with poor survival in solid cancer patients. The detrimental effect of NEDD9 overexpression on survival was statistically significant in OS or DFS/PFS/RFS/CSS. NEDD9 could be a prognostic biomarker and novel therapeutic target of solid cancer.

Several limitations should be considered when interpreting the findings of our meta-analysis. Differences among dilution solubility, antibodies, and cutoff values influenced the assessment of NEDD9 overexpression. A large multicenter clinical study using consistent antibodies and cutoff values is needed to validate our results. Language bias may exist in our meta-analysis because the search strategy was limited to English. Most eligible studies included Asian populations, especially Chinese, leading to population homogeneity. Some studies did not report HR and 95% CI directly. Data extracted by Tierney's methods may introduce bias to the original data.

Conclusion

Here, we provide evidence that NEDD9 overexpression is associated with poor OS or DFS/PFS/RFS/CSS in solid cancer patients. Further investigation is required to elucidate the role of NEDD9 overexpression and provide new insights into tumor metastasis and progression for the development of target-specific therapies.

Author contributions

Conception/Design: Fei Zheng. Provision of study materials: Yang Gu, Jingjing Lu, Chen Chen. Collection and extract data: Yang Gu, Jingjing Lu. Data analysis and statistical guidance: Yang Gu. All authors contributed to data analysis, drafting and revising the article, gave final approval of the version to be published, and agree to be accountable for all aspects of the work.

Disclosure

The authors report no conflicts of interest in this work.

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