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ORIGINAL RESEARCH

Relationship between bone mineral density and the risk of breast cancer: a systematic review and dose–response meta-analysis of ten cohort studies

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Purpose: The evidence from recent epidemiological studies investigating the relationship between bone mineral density (BMD) and the risk of breast cancer (BC) remains inconsistent. Materials and methods: The PubMed, EMBASE, and Web of Science databases were comprehensively searched by two independent authors to identify related cohort studies from the inception of the databases through January 31, 2018. Similarly, two researchers separately extracted the data from the selected studies, and any differences were resolved by discussion. Summarized relative risks (RRs) and 95% CIs were summarized via inverse variance weighted random-effects meta-analysis. Heterogeneity among studies was assessed with the P statistic. **Results:** Ten studies with 1,522 BC patients among 81,902 participants were included in this meta-analysis. Compared to the participants with the lowest BMD at the lumbar spine, those with the highest BMD had a significantly lower RR for BC (RR =0.75; 95% CI =0.60-0.93; P=23.0%). In the subgroup analyses, although the directions of the results were consistent with those of the main findings, not all showed statistical significance. We failed to detect an association between BMD at the femoral neck or total hip and the risk of BC (RR =0.94; 95% CI = 0.66 - 1.33; $l^2 = 72.5\%$). Furthermore, the results of the dose–response analysis did not show a significant association between BMD at the lumbar spine, femoral neck, or total hip and the risk of BC. Funnel plot and statistical analyses showed no evidence of publication bias.

Conclusion: There is no relationship between BMD and the risk of BC. More prospective cohort studies are warranted to further investigate this issue.

Keywords: bone mineral density, breast cancer, cohort studies, meta-analysis, risk

Introduction

Breast cancer (BC) is the most common cancer (excluding non-melanoma skin cancer) among women worldwide.¹ Unsurprisingly, in 2018, BC was expected to account for 30% of all new cancer diagnoses in women in the USA.² Recently, estrogen has been proven to play an important role in the development and progression of this disease in numerous studies. Several risk factors for BC have been identified, many of which (eg, age at menarche, age at menopause, breastfeeding, and hormone replacement therapy [HRT]) are related to prolonged estrogen exposure.³ Furthermore, it is noteworthy that more than three-quarters of all BC and 85% of BC-related deaths occur in postmenopausal women.

Bone mineral density (BMD) is an essential component of the assessment of bone quality and is utilized to assess the osteoporotic status of the bone for the prevention of osteoporotic fractures. The standard method of assessing BMD is dual-energy X-ray

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absorptiometry (DXA) of the lumbar spine (LS), femoral neck (FN), or total hip (TH). Moreover, because estrogen regulates bone turnover by inhibiting bone resorption and upregulating hormones that enhance bone formation,⁴ high BMD is regarded as a marker of prolonged cumulative lifetime exposure to estrogen,^{5,6} and high BMD serves as an intermediate marker of BC risk. Although the number of prospective cohort studies exploring the role of BMD in the risk of BC has been increasing,7-16 the findings have been inconsistent. This inconsistency might be attributed to different study designs, menopausal status of participants, and BMD measurement methods and sites. Although Nagel et al⁷ performed a meta-analysis of this topic, the aforementioned differences still existed in their study. Herein, to further clarify the association between BMD and the risk of BC, we performed a systematic review and meta-analysis using currently available evidence from cohort studies.

Materials and methods Search strategy

We carried out a systematic review of English language articles published from the inception of the databases to the end of January 2018 in PubMed, EMBASE, and Web of Science that reported the association between BMD and the risk of BC with the following search algorithm: "(bone mineral density OR bone density OR BMD OR osteoporosis) AND (breast) AND (cancer OR neoplasm OR carcinoma OR tumour OR malignancy OR malignancies)". Figure 1 shows the flowchart detailing the process of identifying eligible studies; 12,040 abstracts were reviewed (by J-HC and QY). We followed the guidelines of the PRISMA.¹⁷

Study selection and exclusion

Paired reviewers (J-HC and QY) who have been trained in research methods both independently screened the titles/ abstracts and full texts to identify eligible articles, assessed the risk of bias, and extracted data from each eligible study using standardized pilot-tested forms with detailed instructions. Reviewers resolved discrepancies through discussion or, if necessary, arbitration by a third reviewer (S-HZ).

Cohort studies that contained extractable information on BMD and BC outcomes were eligible. Furthermore, studies were included if they reported the relative risks (RRs) with 95% CIs or reported sufficient data to allow the calculation of those risk estimates. We excluded studies that met the following criteria: 1) case–control studies, randomized controlled trials, case reports, editorials, ecological studies, and reviews without original data and 2) studies that failed to provide risk estimates and 95% CIs. If multiple studies had the same participant cohort, only the study with the largest sample size was included for any given outcome.

Data extraction and quality assessment

Data were extracted in duplicate by two independent authors (J-HC and QY) using standardized forms. Disagreements were resolved by consensus. A single investigator (J-HC) extracted the following information from each included study: first author, publication year, geographic location, number of BC patients and controls, BMDs of BC patients and controls, site of BMD measurement, exposure category, and study-specific adjusted RRs with 95% CIs. For risk estimates, if both univariate and multivariate analyses were provided, data from multivariate analysis were used. The extracted data were entered into a standardized Excel spreadsheet (Microsoft Corporation, Redmond, WA, USA). The quality of each study was assessed according to the Newcastle–Ottawa Scale (NOS)¹⁸ by two authors (J-HC and QY) independently.

Statistical analysis

The RRs and 95% CIs are presented as summaries of the risk estimates; the RRs were calculated with a random-effects model to investigate the association between BMD and the risk of BC. For studies that only provided the results of a dose–response analysis, we used the scaling method proposed by Danesh et al¹⁹ to convert the reported risk estimates into a standard scale of effect to compare persons with levels of exposure in the top tertile to persons with exposure levels in the bottom tertile. More details about this method can be obtained elsewhere. For studies that did not use the lowest category of BMD as the reference,^{8–10,12,14} we used the effective counts method proposed by Hamling et al²⁰ to recalculate the RRs and 95% CIs.

The method described by Greenland and Longnecker²¹ was used for the dose–response analysis, and study-specific slopes (linear trends) and 95% CIs were computed from the natural logs of the RRs and CIs across categories of BMD values. This method requires that the distribution of cases and person-years or non-cases and RRs with variance estimates are known for at least three quantitative exposure categories. We assigned the median or mean level of BMD in each category to the corresponding RR for each study. For studies that reported ranges of BMD values, we estimated the midpoint in each category by calculating the average of the lower and upper bounds. When the highest category was open ended, we assumed the length of the open-ended interval to



Figure I Selection of studies for inclusion in this meta-analysis.

be the same as that of the adjacent interval. When the lowest category was open ended, we set the lower boundary to zero. The dose–response results in the forest plots are presented for 0.1 g/cm^2 increments in BMD.

*I*² statistic was used to assess the heterogeneity between studies,²² which was the amount of total variation that is explained by the variation between studies.²² Heterogeneity between subgroups was evaluated by meta-regression. Post hoc subgroup analyses were conducted according to the geographic location (North America, Europe, and others),

median number of BC cases (\geq 50 vs <50), exposure unit (g/cm² vs others), study quality (low risk of bias vs high risk of bias), and menopausal status of participants (post-menopausal vs not postmenopausal), and adjustments were made for potential confounders (including body mass index, menopausal status, HRT use, and any reproductive factors).

We generated a funnel plot and applied Egger's test²³ and Begg's test²⁴ to examine publication biases (eg, publication bias), with P<0.10 indicating the presence of bias. In addition, we visually explored the funnel plots for asymmetry.

We carried out sensitivity analyses by removing one study at a time to examine the effect of the data from each study on the overall estimate. The sequential exclusion strategy proposed by Patsopoulos et al^{25} was used to determine whether the overall estimates were influenced by the substantial observed heterogeneity. Studies that accounted for the largest proportions of the heterogeneity were sequentially and cumulatively excluded until I^2 was <50%. Then, further examinations were conducted to determine whether the risk estimates were consistent before and after the exclusion of those studies.²⁶ All statistical analyses were performed using the Stata statistical software package (version 12.0; StataCorp LP, College Station, TX, USA).

Results

Search results, study characteristics, and quality assessment

In total, 12,040 articles on BMD and BC risk were screened for inclusion (Figure 1). After the removal of duplicates, 9,937 articles were screened via title and abstract to determine those eligible for full-text review. In total, 36 articles were selected for full-text review. Sixteen articles were excluded for various reasons. Finally, ten articles were selected for inclusion in this systematic review and meta-analysis.

The main characteristics of the ten included cohort studies are presented in Table 1. Most of the included studies were conducted in Europe (n=5) and North America (n=4), and one was conducted in Asia. All included studies included data on DXA scanning at the baseline but at different positions, including the LS, FN, and TH. Nine, eight, seven, and five studies adjusted for age, body mass index, menopausal status, and HRT use, respectively, while fewer studies adjusted for cigarette smoking (n=3), physical activity (n=3), alcohol consumption (n=2), and parity (n=2). None of the included studies adjusted for race (Table 2).

Table 3 provides the details of the study quality assessment as reflected by the NOS scores.

BMD and BC risk (highest vs lowest category)

We found that compared with participants with the lowest BMD at the LS, participants with the highest BMD had ~0.75-fold the risk of BC (95% CI =0.60–0.93; l^2 =23.0%; *P* for heterogeneity =0.254; n=7), as shown in Figure 2. In

	1		1	1	1	1
Author, year,	No. of cases	BMD of	No. of controls/	BMD of	Position	Exposure
country	(mean age,	cases (mean,	cohort (mean	controls/cohort	(assessment)	category
	years)	g/cm²)	age, years)	(mean, g/cm²)		
Nagel et al, ⁷ 2017,	52 (55.1)	LS (0.92)	1,380 (55.5)	LS (0.96)	LS (DXA)	Quartile per I z
Germany						score increase
Fraenkel et al, ⁸	86 (68.8)	LS (1.04)	15,268 (65.1)	LS (1.00)	LS/FN/TH (DXA)	Tertile
2013, Israel		FN (0.84)		FN (0.81)		
		TH (0.91)		TH (0.88)		
Grenier et al, ⁹ 2011,	794 (64.7)	LS (1.05)	37,860 (65.0)	LS (1.03)	LS/FN (DXA)	Quartile
Canada		FN (0.83)		FN (0.83)		
Burshell et al, ¹⁰	58 (N/A)	N/A	2,576 (66.5)	N/A	LS/FN (DXA)	Two groups
2008, USA						
Trémollieres et al, ¹¹	98 (52.6)	LS (1.03)	2,137 (53.2)	LS (1.05)	LS/FN (DXA)	Per I SD
2008, France		FN (0.82)		FN (0.85)		increase
Stewart et al, ¹²	87 (48.3)	LS (1.05)	3,013 (48.6)	LS (1.06)	LS/FN (DXA)	Per I SD
2005, UK		FN (0.86)		FN (0.88)		increase
Ganry et al, ¹³ 2004,	45 (79.4)	FN (0.75)	1,504 (78.8)	FN (0.70)	FN/Trochanter/Ward's	Tertile
France					triangle (DXA)	
van der Klift	74 (65.5)	LS (1.07)	3,107 (68.1)	LS (1.03)	LS/FN (DXA)	Tertile per I SD
et al, ¹⁴ 2003, the		FN (0.82)		FN (0.81)		increase
Netherlands						
Buist et al, ¹⁵ 2001,	131 (N/A)	N/A	8,203 (68.2)	TH (0.76)	TH (DXA)	Quartile
USA						
Cauley et al, ¹⁶ 1996,	97 (71.5)	LS (0.90)	6,854 (71.9)	LS (0.84)	LS/TH/distal radius/	Quartile per I
USA		TH (0.81)		TH (0.75)	proximal radius/calcaneus	SD increase
					(DXA)	

Table I Characteristics of the included cohort studies

Abbreviations: BMD, bone mineral density; DXA, dual-energy X-ray absorptiometry; FN, femoral neck; LS, lumbar spine; N/A, not available; TH, total hip.

Author, year	Adjustment for potential confounders in the primary analysis of each study								
	Age	Race	BMI	Smoking	Alcohol consumption	PA	Parity	Menopause	HRT use
Nagel et al, ⁷ 2017	\checkmark	×	\checkmark	\checkmark	×	\checkmark	×	\checkmark	\checkmark
Fraenkel et al, ⁸ 2013	\checkmark	×	\checkmark	×	×	×	×	×	×
Grenier et al, ⁹ 201 l	\checkmark	×	\checkmark	×	×	×	×	×	\checkmark
Burshell et al, ¹⁰ 2008	\checkmark	×	×	×	×	×	×	×	×
Trémollieres et al, ¹¹ 2008	×	×	×	×	×	×	×	\checkmark	\checkmark
Stewart et al, ¹² 2005	\checkmark	×	\checkmark	×	×	×	×	\checkmark	\checkmark
Ganry et al, ¹³ 2004	\checkmark	×	\checkmark	\checkmark	\checkmark	\checkmark	\checkmark	\checkmark	\checkmark
van der Klift et al, ¹⁴ 2003	\checkmark	×	\checkmark	×	×	×	×	\checkmark	×
Buist et al, ¹⁵ 2001		×	\checkmark	×	X	×	×	\checkmark	×
Cauley et al, ¹⁶ 1996	\checkmark	×		\checkmark	\checkmark		\checkmark	\checkmark	×

Table 2 Adjustment for potential confounders in the included cohort studies

Abbreviations: BMI, body mass index; HRT, hormone replacement therapy; PA, physical activity.

the analysis of BMD at the FN or TH, we failed to detect an association between BMD and BC risk (summarized RR =0.94; 95% CI =0.66–1.33; P=72.5%; P for heterogeneity <0.001; n=8), as shown in Figure 3. The funnel plot and statistical analyses showed no evidence of publication bias (Figures S1 and S2).

In subgroup analyses, although the directions of the results were consistent with those of the main findings, not all subgroups showed statistical significance. For example, there was a significant association between BMD at the LS and the risk of BC in the subgroups of studies that used exposure units such as T-score or z score. Furthermore, similar significant results for BMD at the LS and the risk of BC were observed in studies with a high risk of bias and studies with postmenopausal women. However, in metaregression analyses, there was no evidence of heterogeneity between the subgroups stratified by the study characteristics or those adjusted for confounding factors (Table 4), except in the analysis of BMD at the FN or TH. We obtained significant results for differences in the quality of studies in meta-regression analyses. Sensitivity analyses using an alternative statistical model showed robust results (data not shown). Additionally, the sensitivity analysis presented the summarized RR of BMD at the LS for the risk of BC ranged from 0.68 (95% CI =0.58–0.81; *I*²=0%; exclusion of Cauley et al¹⁶) to 0.79 (95% CI =0.58–1.07; *I*²=31.5%; exclusion of Grenier et al⁹). Moreover, the sensitivity analysis presented the summarized RR of BMD at the FN or TH for the risk of BC ranged from 0.85 (95% CI =0.61-1.19; P=69.5%; exclusion of Ganry et al¹³) to 1.13 (95% CI =0.82–1.56; *I*²=55.8%; exclusion of Fraenkel et al8). For BMD at the FN or TH, when the studies that contributed the largest amount to betweenstudy heterogeneity were sequentially excluded until I^2 was

<50%, the summarized HR was 1.03 (95% CI =0.78–1.37; I^2 =43.4%), which was similar to the original estimate.

BMD and BC risk (dose-response analysis)

Four studies were included in the dose–response analysis of BMD at the LS, with 234 BC patients among 11,237 participants. The summary RR was 0.94 (95% CI =0.82–1.07) for each 0.1 g/cm² incremental change in BMD at the LS, with moderate heterogeneity (l^2 =59.6%; *P* for heterogeneity =0.059), as shown in Figure 4. Furthermore, four studies were included in the dose–response analysis of BMD at the FN or TH, with 227 BC patients among 6,668 participants. The summary RR was 0.94 (95% CI =0.73–1.22) for each 0.1 g/cm² incremental change in BMD at the FN or TH, with significant heterogeneity (l^2 =71.6%; *P* for heterogeneity =0.014), as shown in Figure 5.

Discussion

On the basis of the ten included cohort studies, we found a significant association between BMD at the LS and the risk of BC in a categorical meta-analysis. However, no significant association was observed between BMD at the FN or TH and the risk of BC. Similar null findings were also observed in dose–response meta-analyses.

The present meta-analysis has several strengths. Since we carried out the analyses on the basis of prospective cohort studies, we have effectively avoided recall bias and reduced the possibility of selection bias. Compared with previous meta-analyses, we conducted more detailed subgroup analyses and dose–response analyses. This meta-analysis may also have several limitations that must be taken into consideration. BMD may be associated with other factors, including body

Author, year	Selection				Comparability	Outcome		
	Representativeness	Selection of	Ascertainment	Outcome of interest	Controlled for	Assessment	Follow-up long	Adequacy
	of the exposed	the unexposed	of exposure	not present at the	important factors or	of outcome	enough for	of follow-up
	cohort	cohort		start of study	additional factors ^a		outcomes to occur ^b	of cohorts ^c
Nagel et al, ⁷ 2017	*	*	*	*	**	*	*	*
Fraenkel et al, ⁸ 2013	1	*	*	*	*	*	1	*
Grenier et al, ⁹ 2011	*	*	*	*	*	*	1	*
Burshell et al, ¹⁰ 2008	1	*	*	*	*	*	*	*
Trémollieres et al, ^{II} 2008	1	*	*	*	*	*	*	*
Stewart et al, ¹² 2005	*	*	*	*	*	*	*	*
Ganry et al, ¹³ 2004	*	*	*	*	**	*	*	*
van der Klift et al, ¹⁴ 2003	*	*	*	*	**	*	*	*
Buist et al, ¹⁵ 2001	*	*	*	*	**	*	1	*
Cauley et al, ¹⁶ 1996	*	*	*	*	**	*	1	*
Notes: A study could be awarded	a maximum of one star for ea	ch item except for the	item "controlled for in	portant factors or additional f	actors". ^a A maximum of two st of star ^b A cohort study with a	ars could be award	ded for this item. Studies that	controlled for age

study with a follow-up rate >75% was assigned one star

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mass index, physical activity, alcohol consumption, smoking, reproductive factors, and the use of HRT, which could possibly confound the associations we observed. The direction of results for BMD at the LS persisted in studies that adjusted for these potential confounders; however, statistical significance was only observed in these studies adjusting for body mass index. We could not easily interpret these results because of the limited number of studies in some of these subgroup analyses. There was no evidence of heterogeneity between these subgroups with and without adjustment for these potentially confounding factors. For BMD at the FN or TH, although null results were observed throughout the main subgroup analyses, the direction of the results was not consistent. Further studies are needed to fully adjust for these potential confounders in the future. Second, a high degree of heterogeneity was observed in the analysis of BMD at the FN or TH. In the subgroup analyses, according to study characteristics with adjustment for potential confounders, many of the I^2 estimates were judged to be moderate or high. Although we explored the potential sources of heterogeneity, because of the scarcity of the data, most causes of heterogeneity could not be identified. It is noteworthy that we found only study quality to be a source of heterogeneity (P=0.042). These issues may reduce the strength of the conclusions that can be drawn from this meta-analysis. Furthermore, after excluding the study performed by Cauley et al¹⁶ in 1996, we observed significant results in the categorical and doseresponse analyses. Notably, in that study, they measured BMD twice by different methods, which may explain the significant results we obtained after the exclusion of that study. Third, a limited number of included studies provided information that could be used in the dose-response analyses. Therefore, we could only conduct the dose-response analysis with a limited number of studies.^{7,11–14,16} Since the BMD measurement units varied among the included studies, a future meta-analysis should perform another dose-response analysis to clarify the association between BMD and the risk of BC. Fourth, the quality of the included studies varied. Interestingly, we only generated significant results after summarizing the high-risk studies, although this phenomenon might be attributed to the limited number of studies. Lastly, most studies investigating BMD and BC risk combined both estrogen receptor (ER)-positive and ER-negative tumors as a single outcome, although estrogen is primarily associated with the development of ER-positive cancers. Furthermore, an ER mutation in the ER isoform has been described by Fugua et al²⁷ that is associated with the increased proliferation of BC cells and that has been found in a high percentage of patients with

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Figure 2 Forest plot (random-effects model) of bone mineral density at the lumbar spine and the risk of breast cancer.

Note: The squares indicate study-specific RRs (the size of the square reflects the study-specific statistical weight); the horizontal lines indicate 95% CIs; and the diamond indicates the summary RR estimate with its 95% CI.

Abbreviation: RR, relative risk.



Figure 3 Forest plot (random-effects model) of bone mineral density at the FN or TH and the risk of breast cancer.

Note: The squares indicate study-specific RRs (the size of the square reflects the study-specific statistical weight); the horizontal lines indicate 95% CIs; and the diamond indicates the summary RR estimate with its 95% CI.

Abbreviations: FN, femoral neck; RR, relative risk; TH, total hip.

hyperplastic breast tissue.²⁸ If such a mutation was present in both the bone and breast tissue, very low levels of estradiol could preserve BMD and still increase the risk of BC. However, only one included study¹² provided risk estimates stratified by ER status. In addition to the effects of different study designs, BMD measurement sites, BMD measurement methods, and characteristics of study participants, the inconsistent findings of these published studies might also be attributed to the different menopausal status of the participants in various studies.

No. of studies RR 0.75 95% CI 0.60-0.93 l² (%) 23.0 P _h ^a h P _h ^b h No. of studies HR 9 95% CI 0.94 l² (%) 0.66-1.33 P _h ^a h P _h ^b h Overall 7 0.75 0.60-0.93 23.0 0.254 9 0.94 0.66-1.33 72.5 <0.001 Subgroup analyses - - - - - - - 0.189
Overall 7 0.75 0.60–0.93 23.0 0.254 9 0.94 0.66–1.33 72.5 <0.001
Subgroup analyses Output Geographic location 0.842
Geographic location 0.842 0.189
North America 3 0.86 0.51–1.47 72.1 0.028 4 1.21 0.87–1.67 32.7 0.216
Europe 3 0.72 0.51-1.04 0 0.918 3 1.09 0.54-2.21 73.4 0.023
Others I 0.61 0.35–1.07 N/A N/A 2 0.48 0.33–0.70 0 0.915
No. of cases 0.881 0.289
≥50 5 0.75 0.56-1.01 48.1 0.103 6 0.82 0.54-1.25 77.1 0.001
<50 2 0.78 0.47-1.28 0 0.960 3 1.40 0.60-3.31 70.0 0.036
Exposure unit 0.225 0.150
g/cm ² 4 0.90 0.59–1.39 44.7 0.143 5 1.28 0.75–2.19 66.4 0.018
Others 3 0.67 0.55–0.82 0 0.899 4 0.70 0.41–1.19 81.4 0.001
Study quality 0.409 0.042
Low risk of bias 5 0.81 0.60–1.10 43.4 0.132 6 1.21 0.85–1.73 58.4 0.034
High risk of bias 2 0.62 0.41–0.93 0 0.939 3 0.55 0.40–0.76 0 0.434
Menopausal status 0.289 0.482
Postmenopausal 3 0.68 0.56-0.83 0 0.960 4 1.09 0.71-1.69 64.6 0.037
Not postmenopausal 4 0.89 0.56–1.40 50.0 0.112 5 0.82 0.48–1.41 73.0 0.005
Adjustment for potential confounders or risk
Tes 6 0.77 0.00-0.77 33.6 0.102 6 0.77 0.00-1.42 73.4 <0.001
NO I 0.63 0.34–1.14 N/A I 0.77 0.42–1.42 N/A N/A
Menopausal status 0.00
Tes 4 0.70 0.39–1.37 44.7 0.143 3 1.07 0.34–2.21 73.4 0.023
HKT use 0.4/1 0.844 Var 0.94 0.77 0.070 4 0.00 0.77 1.75 0.000
Tes 4 0.84 0.57-1.23 57.5 0.070 4 0.99 0.56-1.75 80.3 0.002
NO 3 0.66 0.47-0.73 0 0.854 5 0.70 0.55-1.47 66.7 0.017
Any reproductive
1es 5 0.01 0.00-1.10 43.4 0.132 / 1.0/ 0.75-1.38 /0.6 0.002 No 2 0.42 0.41 0.92 2 0.59 0.22 0.41 0.32

Table 4 Risk estimate summary of the association between bone mineral density and the risk of breast cancer (highest category vs lowest category)

Notes: *P-value for heterogeneity within each subgroup. *P-value for heterogeneity between subgroups according to the meta-regression analysis. **Abbreviations:** BMI, body mass index; HRT, hormone replacement therapy; N/A, not available; RR, relative risk.

Significant results were observed in the analysis of BMD at the LS in postmenopausal participants (Table 4). Because BMD could be a marker of cumulative exposure to estrogen, this effect could be more apparent in postmenopausal women than in premenopausal women. Furthermore, the rates of HRT use were different among these studies. For example, Stewart et al¹² reported that over half the women had received HRT at some point by the end of the study in a population-based screening program for osteoporosis risk. Additionally, they found that those with low baseline levels of BMD were more likely to receive HRT. However, the Framingham²⁹ and Rotterdam¹⁴

studies included all subjects and had 17.7% and 10.2% of their participants who reported the use of HRT, respectively. Therefore, it may be that those in the lowest quartile of BMD were taking HRT, thereby increasing their risk of BC and while not increasing their BMD. Although we could not perform an analysis stratified by the use of HRT, interestingly, we found significant results in the subgroup analysis that did not adjust for HRT use. More studies should focus on this issue.

In summary, on the basis of the present meta-analysis, we were unable to demonstrate an association between BMD and the risk of BC.



Figure 4 Forest plot (random-effects model) of BMD at the lumbar spine and the risk of breast cancer in the dose-response analysis (per 0.1 g/cm² incremental change in BMD).

Note: The squares indicate study-specific relative risks (the size of the square reflects the study-specific statistical weight); the horizontal lines indicate 95% CIs; and the diamond indicates the summary RR estimate with its 95% CI.

Abbreviations: BMD, bone mineral density; RR, relative risk.



Figure 5 Forest plot (random-effects model) of bone mineral density at the femoral neck or total hip and the risk of breast cancer.

Note: The squares indicate study-specific relative risks (the size of the square reflects the study-specific statistical weight); the horizontal lines indicate 95% CIs; and the diamond indicates the summary RR estimate with its 95% CI.

Abbreviation: RR, relative risk.

Author contributions

J-HC, QY, Y-NM, and D-LW designed the research; J-HC and QY conducted the research; QY and S-HZ analyzed data; J-HC and QY wrote the draft; and all authors contributed to data analysis, drafting and revising the article, gave final approval of the version to be published, and agree to be accountable for all aspects of the work.

Disclosure

The authors report no conflicts of interest in this work.

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Supplementary materials



Figure SI Test for publication bias for bone mineral density at the lumbar spine through Begg's funnel plot. Notes: The circles represent real studies. The vertical lines represent the summary effect estimates, and the dashed lines represent pseudo-95% CI limits. Abbreviations: RR, relative risk; SE, standard error.



Figure S2 Test for publication bias for bone mineral density at the femoral neck or total hip through Begg's funnel plot. Notes: The circles represent real studies. The vertical lines represent the summary effect estimates, and the dashed lines represent pseudo-95% CI limits. Abbreviations: RR, relative risk; SE, standard error.

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