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## ORIGINAL RESEARCH

Memantine treatment for Japanese patients with moderate to severe Alzheimer's disease: a meta-analysis of double-blind, randomized, placebo-controlled trials

## Taro Kishi<sup>1,\*</sup> Shinji Matsunaga<sup>1,2,\*</sup> Nakao Iwata<sup>1</sup>

<sup>1</sup>Department of Psychiatry, Fujita Health University School of Medicine, Toyoake, Aichi 470-1192, Japan; <sup>2</sup>Department of Geriatrics and Cognitive Disorders, Fujita Health University School of Medicine, Toyoake, Aichi 470-1192, Japan

\*These authors contributed equally to this work

Correspondence: Taro Kishi Department of Psychiatry, Fujita Health University School of Medicine, I-98 Dengakugakubo, Kutsukake-cho, Toyoake, Aichi 470-1192, Japan Tel +81 562 93 9250 Fax +81 562 93 1831 Email tarok@fujita-hu.ac.jp



**Purpose:** Although previous meta-analyses of randomized trials in the world literature have provided strong evidence that supports the efficacy and safety of memantine for the treatment of patients with Alzheimer's disease (AD), it is unclear whether the drug is beneficial in the treatment of Japanese patients with moderate to severe AD because of differences in the formulation and regimen of memantine and the cholinesterase inhibitor (ChEI) used in combination with memantine between the drugs made in Japan and those made in other countries. To address this issue, we conducted a meta-analysis on the efficacy and safety of memantine using data from only double-blind, randomized, placebo-controlled trials (DBRPCTs) in Japan on Japanese patients with moderate to severe AD.

**Patients and methods:** Our primary analysis was conducted using data from both memantine monotherapy (memantine vs placebo) and memantine combination therapy (memantine+ChEI vs ChEI+placebo) studies. The primary outcomes measured were cognitive function and behavioral disturbances. The secondary outcomes measured were the subscale scores of Behavioral Pathology in Alzheimer's Disease (Behave-AD), discontinuation rate, and individual adverse events.

**Results:** Four DBRPCTs (n=1,328) were detected. Memantine was superior to the control in cognitive functions (standardized mean difference [SMD]=-0.31, 95% CI=-0.53, -0.10) and behavioral disturbances (SMD=-0.16, 95% CI=-0.28, -0.05). Only memantine monotherapy was superior in both outcomes. It was also superior to the control in delusions, aggression, and diurnal rhythm disturbances based on the Behave-AD subscale scores. Although memantine was associated with a lower incidence of AD progression than that of the control, the incidence of somnolence was higher with memantine. There were no significant differences in other safety outcomes, including all-cause discontinuation, between the groups.

**Conclusion:** Our results suggest that memantine is useful for the treatment of patients in Japan with moderate to severe AD even though our meta-analysis comprised only four DBRPCTs. **Keywords:** Japanese patients, Alzheimer's disease, memantine, cognitive functions, behavioral disturbances, meta-analysis

## Introduction

Our previous meta-analysis demonstrated that memantine improves cognitive functions and behavioral disturbances in patients with Alzheimer's disease (AD) more efficiently than the controls, both as a monotherapy and in combination with donepezil.<sup>1</sup> We demonstrated that memantine improves positive symptoms of behavioral

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disturbances, such as agitation/aggression, delusions, disinhibition, and nighttime disturbance/diurnal rhythm disturbances in patients with AD.<sup>2</sup> The medication was also shown to be well tolerated based on no difference in all-cause discontinuation between those treated with memantine and the controls.<sup>1</sup> Moreover, although the meta-analysis comprised more than 200 safety outcomes, there were no significant differences in these outcomes between the groups with respect to safety except for differences in somnolence and dizziness.<sup>1</sup>

However, even with our results, our question of whether memantine is beneficial in the treatment of Japanese patients with moderate to severe AD remained unanswered.<sup>1,2</sup> Although our previous meta-analysis included all randomized controlled trials on memantine for patients with AD worldwide, there are differences in the formulation and regimen of the medication and of the cholinesterase inhibitor (ChEI) used in combination with memantine among countries. First, there are two formulations of memantine: immediate-release and extended-release.<sup>3–7</sup> For example, the U.S. Food and Drug Administration (FDA) approves 28 mg/d extended-release memantine for the treatment of moderate to severe AD in the US;4 however, in Japan, only 20 mg/d immediate-release memantine was approved for treatment of the same condition.7 Other countries approve 10 mg BID immediate-release memantine for the treatment of moderate to severe AD.5,6 Second, although we included data from studies on low-dose memantine (<20 mg/d), 20 mg/d immediate-release memantine (or 28 mg/d extended-release memantine) was approved worldwide for the treatment of moderate to severe AD. Third, there is a difference in the approved dosage of donepezil among the countries. FDA approves 5 or 10 mg/d donepezil for mild to moderate AD, and 10 or 23 mg/d for moderate to severe AD. In Japan, 5 mg/d donepezil is approved for mild to moderate AD, and 10 mg/d donepezil is approved for severe AD. Recent metaanalyses have shown that there were significant differences in the efficacy and safety between high-dose donepezil and 5 mg donepezil.8 Fourth, although there are three formulations of rivastigmine (capsules, liquid solution, and patches) worldwide, only the patches are approved for the treatment of AD in Japan. The current meta-analysis of rivastigmine showed that the transdermal patch might have fewer side effects than the capsules.9

For direct evidence to answer our clinical question, we performed a meta-analysis of the efficacy and safety of memantine immediate-release using data from only doubleblind, randomized, placebo-controlled trials (DBRPCTs) in Japan on Japanese patients with moderate to severe AD. Although our previous meta-analysis included the treatment arms for both 10 and 20 mg/d memantine,<sup>1,2</sup> the current metaanalysis used data from only the treatment arm with 20 mg/d memantine to determine efficacy in Japan.<sup>7</sup> Our primary analysis was conducted using data from both memantine monotherapy and memantine combination therapy studies because there were only four DBRPCTs included in the current meta-analysis; however, we included data from both 10 and 20 mg/d memantine to conduct our current meta-analysis of specific safety characteristics. Given that only 5 mg/d memantine is first administered and the dose is increased by 5 mg/d weekly up to 20 mg/d,<sup>7</sup> we considered that the safety profile of lower dose memantine (<20 mg/d) must be evaluated.

# **Materials and methods** Search strategy and inclusion criteria

To identify relevant studies, two study authors (TK and SM) independently searched MEDLINE, Cochrane Library, Scopus, and PsycINFO without language restrictions from the inception of these databases to September 5, 2018, using the following search strategy: ("Alzheimer Disease" [Mesh] OR "Alzheimer disease" OR "Alzheimer's disease") AND ("Memantine" [Mesh] OR "memantine") AND ("randomized" OR "random" OR "randomly") AND ("Japanese" OR "Japan"). The authors also searched ClinicalTrials.gov (http://clinicaltrials.gov/), UMIN (http://www.umin.ac.jp/), and the International Clinical Trials Registry Platform (http:// www.who.int/ictrp/en/) to include DBRPCTs as comprehensively as possible and to minimize the possibility of publication bias. Only DBRPCTs of memantine treatment lasting >12 weeks in Japanese patients with moderate to severe AD were included. Two study authors (TK and SM) independently assessed the inclusion/exclusion criteria and selected the studies. The references of the included articles and review articles were also searched for citations on additional relevant published and unpublished studies, including conference abstracts.

## Data synthesis and outcome measures

Three primary outcomes were assessed based on two efficacy measures as follows: improvement in cognitive functions and behavioral disturbances. Cognitive function scores were derived from the Alzheimer's Disease Assessment Scale, cognitive subscale<sup>10</sup> score (which IE2201 study<sup>11</sup> used), and Severe Impairment Battery (SIB)<sup>12</sup> score (which Kitamura 2011 study,<sup>13</sup> Nakamura 2011 study,<sup>14</sup> and Nakamura 2016 study<sup>15</sup> used). The behavioral disturbances score comprised

the Neuropsychiatric Inventory<sup>16</sup> score (which Kitamura 2011 study<sup>13</sup> used) and the Behavioral Pathology in Alzheimer's Disease Rating Scale<sup>17</sup> score (which Nakamura 2011 study<sup>14</sup> and Nakamura 2016 study<sup>15</sup> used). The secondary outcome measures were the improvement of Behave-AD subscale scores, all-cause discontinuation, discontinuation because of adverse events, and the incidence of individual adverse events.

## Data extraction

Two study authors (TK and SM) independently extracted data from the included studies. Where possible, we used only an intention-to-treat or a full analysis set (FAS) population. When the data required for a meta-analysis were missing, we contacted the investigators (or the industries) of the relevant study and requested unpublished data.

## Meta-analysis methods

The meta-analysis was conducted using Review Manager software.<sup>18</sup> The random-effects model was selected for this meta-analysis because of the potential heterogeneity across studies. Dichotomous outcomes were presented as risk ratios (RRs) with 95% CIs. When the random-effects model showed significant differences among groups, the number needed to harm (NNH) was calculated. NNH values were then derived from the risk difference (RD) using the formula NNH=1/ RD. Continuous outcomes were analyzed using the mean difference (MD) or, when different studies used different scales, the standardized mean difference (SMD). Lower SIB scores indicate more impairment or more severe symptoms; hence, we reversed the algebraic sign of the numerical scores for these scales. We assessed the quality of the trial methods according to the Cochrane risk-of-bias criteria.<sup>19</sup> Most commonly discussed in the Cochrane Handbook for Systematic Reviews of Interventions is the SMD, also known as Cohen's d,<sup>19</sup> which is the MD divided by the pooled standard deviation of the two groups. For the SMD, a negligible effect is considered to range from 0.0 to 0.2, a small effect is between 0.2 and 0.5, a medium effect ranges from 0.5 to 0.8, and a large effect is considered to be > 0.8.<sup>19</sup>

Study heterogeneity was tested using the  $I^2$  statistic, considering  $I^2 \ge 50\%$  to reflect considerable heterogeneity.<sup>19</sup> In addition, we performed a sensitivity/subgroup analysis by dividing the primary analysis of the primary outcomes into memantine monotherapy and combination therapy studies. Because there was difference in percentage of males among the DBRPCTs included in the current meta-analysis (Table 1), a meta-regression analysis was performed to evaluate the association between meta-analysis results for cognitive functions or behavioral disturbances and percentage of males using Comprehensive Meta-Analysis software version 2 (Biostat Inc., Englewood, NJ, USA). Egger's regression test was used to detect publication bias in the meta-analyses using the same software.

# **Results** Study characteristics

Of the 35 results obtained from our literature search, 21 were excluded because they were duplicates, 11 after a review of the abstract or title, and two after a review of the full text (two review articles).<sup>1,2</sup> In addition, three studies<sup>11,14,15</sup> were retrieved by searching through the review articles.<sup>1,2</sup> We did not detect any new DBRPCTs from the clinical trial registries. Four DBRPCTs (memantine monotherapy vs placebo: two DBRPCTs, n=747; combination therapy with memantine and ChEIs vs ChEI monotherapy: two DBRPCTs, n=581) were identified.<sup>11,13–15</sup> All DBRPCTs were published in Japanese and sponsored by a pharmaceutical company (Table 1). The study design for all DBRPCTs was the same (ie, 24 weeks, DBRPCTs) (Table 1). The researchers used a FAS population. Evaluations regarding the quality of the study methods that included DBRPCTs were conducted according to the Cochrane risk-of-bias criteria, as shown in our previous article.1

## Efficacy outcomes

Memantine was shown to significantly improve two primary efficacy outcomes compared with the control – cognitive function (SMD=-0.31, 95% CIs=-0.53 to -0.10, P=0.004,  $I^2=61\%$ ; N=4, n=1,192) and behavioral disturbances (SMD=-0.16, 95% CIs=-0.28 to -0.05, P=0.006,  $I^2=0\%$ ; N=3, n=1,166) (Figure 1). The data on cognitive function and behavioral disturbances scores in each treatment group were simulated with no publication bias (Egger's test *P*-value: cognitive function scores=0.463, behavioral disturbances score=0.509).

The sensitivity/subgroup analyses showed that only memantine monotherapy was superior to the controls in both primary outcomes (cognitive function scores of memantine monotherapy: SMD=-0.41, 95% CIs=-0.57 to -0.26, P < 0.00001,  $I^2=0\%$ , behavioral disturbances score of memantine monotherapy: SMD=-0.21, 95% CIs=-0.37 to -0.06, P=0.007,  $I^2=0\%$ ). Meta-regression analysis detected no associations between the effect size of memantine treatments with respect to cognitive functions or behavioral disturbances and percentage of males.

Study name	Total, n	Study methods	Patients	Age	Male	Baseline	Intervention (mean	Study results
(study design)		I. Duration of study	I. Diagnosis	(mean±SD)	(%)	cognitive	dose, mg/d)	1. Cognitive functions
		2. Analyzed population 3. Sponsorship	<ol> <li>Inclusion criteria</li> <li>Study defined disease severity</li> </ol>			tunction scales (mean ± SD)		2. Benavioral disturbances
Kitamura et al (2011) <sup>13</sup> (DBRPCT)	315	I. 24 weeks 2. FAS 3. Industry	<ol> <li>AD, DSM-IV, and NINCDS-ADRDA</li> <li>Age ≥50 years, MMSE 5–14, FAST 6a-7a</li> <li>Moderate to severe</li> </ol>	73.3±9.4	29.3	SIB: 71.1±17.8, MMSE: 10.1±3.0	MEM (20 or 10, fixed) vs PBO	I. SIB-J; МЕМ20>РВО 2. NPI: МЕМ20=РВО
Nakamura et al (2011) <sup>14</sup> (DBRPCT)	432	1. 24 weeks 2. FAS 3. Industry	<ol> <li>AD, DSM-IV, and NINCDS-ADRDA</li> <li>Age ≥50 years, MMSE 5–14, FAST 6a-7a</li> <li>Moderate to severe</li> </ol>	74.6±8.4	35.7	SIB: 71.0±17.9, MMSE: 9.9±3.0	MEM (20, fixed) vs PBO	I. SIB-J; MEM20>PBO 2. Behave-AD: MEM20=PBO
IE2201 <sup>a.11</sup> (unpublished study, DBRPCT)	35	1. 24 weeks 2. FAS 3. Industry	1. AD, NINCDS-ADRDA 2. Age ≥50 years, MMSE 10–20 3. Moderate	73.2±6.7	54.3	ADAS-cog: 25.7, MMSE: 15.6	MEM (20, fixed)+ChEls (NR, DON=83%) vs MEM 10 (fixed)+ChEls (NR, DON=91%) vs PBO+DON (NR, DON=100%)	1. NR 2. NR
Nakamura et al (2016) <sup>15</sup> (DBRPCT)	546	1. 24 weeks 2. FAS 3. Industry	<ol> <li>AD, DSM-IV-TR, and NINCDS-ADRDA</li> <li>Age ≥50 years, MMSE 1–14, SIB 30–85, patients receiving a ChEI for ≥36 weeks</li> <li>Moderate to severe</li> </ol>	78.5±7.8	27.2	SIB: 77.0±13.6, MMSE: 10.8±3.7	MEM (20, fixed)+DON (6.9) vs PBO+DON (6.9)	I. SIB-J: MEM20=PBO 2. Behave-AD: MEM20=PBO
Note: "This study is Abbreviations: AD diagnostic and statisti and Communicative [	presented in , Alzheimer , ical manual of Disorders an	the Clinical Study Report or appl disease; ADAS-cog, Alzheimer's d f mental disorders fourth edition ( d Stroke and the Alzheimer's Dis	roval application ( <u>https://www.pmda.go.jp/</u> ). <sup>11</sup> lisease assessment scale-cognitive subscale; ChEI, chc text revision); FAS, full analysis set; MEM, memantine ease and Related Disorders Association; NR, not rep	Ninesterase inhibito MMSE, mini ment iort; NPI, neuropsy	or; DBRP al state e: /chiatric ii	CT, double-blind rand camination; n, number iventory; PBO, placeb	omized, placebo-controlled trial; C of patients; NINCDS-ADRDA, Na o; SIB, severe impairment battery.	DON, donepezil; DSM-IV(-TR), ational Institute of Neurological

Table I Characteristics of included double-blind, randomized, placebo-controlled trials

Study or	Memant	ine	1e		ol		Weight	Std mean difference	Std mean difference
subgroup	Mean	SD	Total	Mean	SD	Total	(%)	IV, random, 95% CI	IV, random, 95% Cl
Cognitive functions									
IE2201	-0.6	4.81	12	3.16	6.9	12	5.8	-0.61 (-1.43, 0.21)	
Kitamura 2011	-0.1	6.54	100	3.42	9.84	107	25.7	-0.42 (-0.69, -0.14)	<b>_</b> _
Nakamura 2011	0.42	9.87	218	4.87	11.66	206	33.1	-0.41 (-0.60, -0.22)	
Nakamura 2016	1.34	8.267	268	2.15	8.151	269	35.3	-0.10 (-0.27, 0.07)	
Subtotal (95% CI)			598			594	100	-0.31 (-0.53, -0.10)	$\bullet$
Heterogeneity: $\tau^2$ =0.03; Test for overall effect: Z	; χ²=7.79, « ζ=2.91 ( <i>P</i> =	df=3 (P= 0.004)	•0.05);	/²=61%					
Behavioral disturbance	ces								
Kitamura 2011	-1.73	8.9	100	-0.38	9.92	107	17.8	-0.14 (-0.42, 0.13)	
Nakamura 2011	-0.25	4.22	217	0.91	5.08	208	36.3	-0.25 (-0.44, -0.06)	
Nakamura 2016	-0.1	3.7	267	0.3	4.1	267	45.9	-0.10 (-0.27, 0.07)	
Subtotal (95% CI)			584			582	100	-0.16 (-0.28, -0.05)	$\bullet$
Heterogeneity: $\tau^2$ =0.00; Test for overall effect: Z	; χ²=1.28, 2=2.77 ( <i>P</i> =	df=2 (P= 0.006)	•0.53);	/²=0%					
									Favors memantine Favors control

Figure I Cognitive function and behavioral disturbances.

Memantine was also superior to the control in delusions, aggression, and diurnal rhythm disturbances based on Behave-AD subscale scores (Figure 2).

## Safety outcomes

Although memantine was associated with a lower incidence of the progression of AD than the control, the drug was associated with a higher incidence of somnolence than the control (Table 2). There were no significant differences in other safety outcomes, including all-cause discontinuation, between the groups (Table 2).

## Discussion

The current meta-analysis was conducted using data from only Japanese patients with moderate to severe AD based on DBRPCTs in Japan. Memantine was superior to the control in improving cognitive function and behavioral disturbances. The results of the current meta-analysis were similar to those of previous meta-analyses, including all randomized trials of memantine for the treatment of AD worldwide; however, although our previous meta-analysis exhibited that both monotherapy and combination therapy with donepezil were superior to treatment with placebo in improving cognitive function and behavioral disturbances scores,<sup>1</sup> the current meta-analysis showed that only memantine monotherapy was superior to the controls in those outcomes. We considered several reasons for these results. First, although there were two DBRPCTs using combination therapy, one of the two was a very small study (n=35). There was also only one DBRPCT using combination therapy

in the behavioral disturbances score; therefore, we might not be able to detect the superiority of memantine over the control in these efficacy outcomes because of insufficient statistical power. Second, the safety profile of ChEI (most of the patients included in the current meta-analysis received donepezil, Table 1) might be a confounding factor in the results of subgroup meta-analysis of combination therapy because a recent systematic review and meta-analysis of donepezil reported that the drug was associated with a higher incidence of anorexia (10 mg/d), diarrhea (5 and 10 mg/d), dizziness (10 mg/d), fatigue (10 mg/d), hallucinations (10 mg/d), insomnia (10 mg/d), muscle cramps (5 and 10 mg/d), nausea (10 mg/d), peripheral edema (10 mg/d), tremor (10 mg/d), vertigo (10 mg/d), vomiting (10 mg/d), and weight loss (10 mg/d) than those observed with placebo treatment. Further studies are necessary to investigate whether combination therapy is beneficial in the treatment of AD in Japanese patients.

Although memantine also improved positive symptoms of behavioral disturbances (delusions, aggression, and diurnal rhythm disturbances), it did not increase any negative symptoms such as mood. These results were similar to those of the previous meta-analysis.<sup>2</sup> Moreover, memantine was associated with a lower incidence of the progression of AD compared with the control. Although memantine was associated with a higher incidence of somnolence than the control, there were no significant differences in other safety outcomes, including all-cause discontinuation, between the groups. We considered that memantine was well tolerated by Japanese patients with moderate to severe AD.

Study or subgroup	Memar Mean	ntine SD	Total	Contro Mean	SD	Total	Weight (%)	Mean difference IV, random, 95% Cl	Mean difference IV, random, 95% Cl	
Delusions										
Nakamura 2011	-0.15	1.68	217	0.13	1.9	208	30.7	-0.28 (-0.62, 0.06)		
Nakamura 2016	-0.1	1.16	267	0.1	1.5	267	69.3	-0.20 (-0.43, 0.03)	<b>_</b>	
Subtotal (95% CI)			484			475	100	-0.22 (-0.41, -0.04)		
Heterogeneity: $\tau^2=0$	.00; $\chi^2 = 0$	).15, df:	=1 ( <i>P</i> =0	.70); /2=0	)%					
lest for overall effec	t: Z=2.33	3 (P=0.0	02)							
Hallucinations										
Nakamura 2011	0	0.93	217	0.05	1 19	208	29.2	-0.05 (-0.25, 0.15)		
Nakamura 2016	0.1	0.73	267	0.1	0.81	267	70.8	0.00 (-0.13, 0.13)		
Subtotal (95% CI)			484			475	100	-0.01 (-0.12, 0.10)		
Heterogeneity: $\tau^2=0$	.00; χ <sup>2</sup> =0	).16, <i>df</i> :	=1 ( <i>P</i> =0	.69); /²=0	)%					
Test for overall effect	t: Z=0.26	6 (P=0.7	79)							
Activity disturbanc	es									
Nakamura 2011	-0.03	1.25	217	0.31	134	208	0.0	-0.34 (-18.55, 17.87)		+
Nakamura 2016	0	1.23	267	0.1	1.33	267	100	-0.10 (-0.32, 0.12)		
Heterogeneity: $\tau^2=0$	$00. \gamma^2 = 0$	00 df:	404 =1 (P=0	98)· /2=0	1%	4/5	100	-0.10 (-0.32, 0.12)		
Test for overall effect	t <sup>.</sup> <b>7</b> =0.90	) (P=0 :	37)		//0					
		,, 0.	01)							
Aggression										
Nakamura 2011	-0.05	1.19	217	0.26	1.55	208	43.8	-0.31 (-0.57, -0.05)	<b>_</b>	
Nakamura 2016	-0.1	1.31	267	0.1	1.43	267	56.2	-0.20 (-0.43, 0.03)	<b>_</b>	
Subtotal (95% CI)			484			475	100	-0.25 (-0.42, -0.07)		
Heterogeneity: $\tau^2=0$ .	.00; $\chi^2 = 0$	).38, df:	=1 ( <i>P</i> =0	.54); /²=0	)%					
Test for overall effect	t: Z=2.79	) (P=0.0	005)							
Diurnal rhythm dia	turbana									
Nakamura 2011		0.53	217	0.06	0.61	208	45.4	-0.06 (-0.17, 0.05)		
Nakamura 2016	_0 1	0.57	267	0	0.6	267	54.6	-0.10 (-0.20, -0.00)		
Subtotal (95% CI)	0.1	0.07	484	Ū	0.0	475	100	-0.08 (-0.16, -0.01)	<b>.</b>	
Heterogeneity: $\tau^2=0$	.00; χ²=0	).28, df:	=1 ( <i>P</i> =0	.59); /²=0	)%				•	
Test for overall effect	t: Z=2.19	) (P=0.0	03)							
Mood										
Nakamura 2011	-0.02	0.69	217	-0.02	0.77	208	41.4	0.00 (-0.14, 0.14)		
Nakamura 2016	0	0.68	267	0	0.7	267	58.6	0.00 (-0.12, 0.12)		
Subtotal (95% CI)	$00. x^2 = 0$	00 df:	484 =1 (P=1	00): /2=0	10/2	475	100	0.00 (-0.09, 0.09)	-	
Test for overall effect	t <sup>.</sup> 7=0.00	) (P=1 (	- 1 (/ - 1 00)	.00), 7 –0	//0					
		, , , , ,	,							
Anxieties										
Nakamura 2011	-0.05	0.98	217	-0.05	1.06	208	54.0	0.00 (-0.19, 0.19)	<b>_</b>	
Nakamura 2016	0	1.21	267	-0.1	1.27	267	46.0	0.10 (-0.11, 0.31)		
Subtotal (95% CI)			484			475	100	0.05 (–0.10, 0.19)		
Heterogeneity: $\tau^2=0$	.00; $\chi^2 = 0$	).47, df:	=1 ( <i>P</i> =0	.49); /2=0	)%					
Test for overall effec	t: Z=0.63	8 (P=0.	53)							
										_
									Eavore momanting Eavore control	

Figure 2 Behavioral pathology in Alzheimer's Disease Rating Scale subscale.

### Table 2 Safety outcomes

	Ν	n	<b>1</b> <sup>2</sup>	RR (95% CIs)	P-value
All-causes discontinuation	4	1,328	31%	0.98 (0.71, 1.37)	0.93
Discontinuation due to adverse events	4	1,328	45%	0.88 (0.51, 1.53)	0.66
Discontinuation due to death	4	1,328	0%	1.70 (0.51, 5.72)	0.39
Abdominal pain	2	747	0%	1.07 (0.21, 5.36)	0.93
Abnormal behavior	4	1,328	0%	0.66 (0.22, 1.96)	0.45
Abnormal gait	3	782	0%	0.90 (0.29, 2.82)	0.86

(Continued)

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#### Table 2 (Continued)

	N	n	<b>1</b> <sup>2</sup>	RR (95% CIs)	P-value
Aggression	2	747	0%	0.90 (0.17, 4.62)	0.90
Agitation	2	747	0%	1.06 (0.30, 3.72)	0.93
AD	3	782	0%	0.45 (0.22, 0.91)	0.03 (NNH=33)
Anorexia/appetite decreased	3	1,293	0%	1.40 (0.68, 2.85)	0.36
Anxiety	2	747	0%	0.56 (0.23, 1.36)	0.20
Aphasia	2	747	0%	2.12 (0.22, 20.30)	0.51
At least one adverse event	4	1,328	0%	1.04 (0.98, 1.11)	0.22
At least one serious adverse event	3	1,293	0%	0.92 (0.65, 1.30)	0.63
Back pain	4	1,328	18%	0.92 (0.43, 1.99)	0.84
Cardiac failure	3	1,293	0%	1.57 (0.27, 9.03)	0.61
Cataract	3	782	0%	0.80 (0.28, 2.31)	0.68
Constipation	3	1,293	0%	1.20 (0.83, 1.73)	0.33
Dehydration	3	1,293	0%	1.23 (0.41, 3.70)	0.71
Delirium	2	747	0%	2.04 (0.27, 15.13)	0.49
Delusion/hallucination	3	782	0%	1.17 (0.31, 4.31)	0.82
Diabetes	3	782	0%	3.84 (0.68, 21.63)	0.13
Diarrhea	4	1,328	0%	0.78 (0.46, 1.32)	0.35
Dizziness/vertigo	4	1,328	0%	1.84 (0.92, 3.66)	0.08
Edema	3	782	6%	0.71 (0.33, 1.52)	0.38
Fall	4	1,328	0%	0.88 (0.59, 1.31)	0.53
Fatigue	2	747	0%	0.71 (0.07, 6.77)	0.76
Fever	3	782	0%	1.08 (0.51, 2.27)	0.84
Headache	2	747	0%	1.07 (0.40, 2.84)	0.89
Insomnia	3	1,293	0%	1.29 (0.76, 2.21)	0.35
Irritability	2	747	0%	3.54 (0.42, 30.11)	0.25
Muscle weakness	2	747	0%	1.18 (0.15, 9.56)	0.88
Nasopharyngitis	4	1,328	17%	0.94 (0.69, 1.29)	0.72
Nausea	3	782	26%	0.84 (0.15, 4.77)	0.84
Parkinsonism	2	467	0%	2.83 (0.32, 24.99)	0.35
Rash	2	747	34%	1.16 (0.14, 9.74)	0.89
Renal failure	2	747	0%	1.18 (0.15, 9.56)	0.88
Seizure	3	1,293	0%	1.14 (0.22, 5.91)	0.87
Somnolence	3	1,293	0%	2.87 (1.09, 7.55)	0.03 (NNH=ns)
Vomiting	3	782	7%	0.62 (0.27, 1.43)	0.26
Weight decreased	2	747	0%	1.08 (0.49, 2.39)	0.85
Weight increased	2	747	0%	1.78 (0.46, 6.86)	0.40

Notes: Bold face: memantine was associated with a lower incidence of the outcome than the control. Italic face: memantine was associated with a higher incidence of the outcome than the control.

Abbreviations: AD, Alzheimer's disease; N, number of studies; n, number of patients; NNH, number needed to harm; ns, not significant; RR, risk ratio.

# Conclusion

We concluded that memantine was highly useful as a treatment for Japanese patients with moderate to severe AD. Although this current meta-analysis did not compare memantine with other antidementia drugs in Japanese patients with AD, a network meta-analysis that includes all antidementia drugs and that uses data from only Japanese patients with AD will be necessary to evaluate their safety and efficacy.

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