

A novel gene-pair signature for relapse-free survival prediction in colon cancer

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Background: Colon cancer (CC) patients with early relapse usually have a poor prognosis. In this study, we aimed to identify a novel signature to improve the prediction of relapse-free survival (RFS) in CC.

Methods: Four microarray datasets were merged into a training set (n=1,045), and one RNA-sequencing dataset was used as a validation set (n=384). In the training set, microarray meta-analysis screened out 596 common RFS-related genes across datasets, which were used to construct 177,310 gene pairs. Then, the LASSO penalized generalized linear model identified 16 RFS-related gene pairs, and a risk score was calculated for each sample according to the model coefficients.

Results: The risk score demonstrated a good ability in predicting RFS (area under the curve [AUC] at 5 years: 0.724; concordance index [C-index]: 0.642, 95% CI: 0.615–0.669). High-risk patients showed a poorer prognosis than low-risk patients (HR: 3.519, 95% CI: 2.870–4.314). Subgroup analysis reached consistent results when considering multiple confounders. In the validation set, the risk score had a similar performance (AUC at 5 years: 0.697; C-index: 0.696, 95% CI: 0.627–0.766; HR: 2.926, 95% CI: 1.892–4.527). When compared with a 13-gene signature, a 15-gene signature, and TNM stage, the score showed a better performance ($P<0.0001$; $P=0.0004$; $P=0.0125$), especially for the patients with a longer follow-up ($R^2=0.988$, $P<0.0001$). When the follow-up was >5 years (n=314), the score demonstrated an excellent performance (C-index: 0.869, 95% CI: 0.816–0.922; HR: 13.55, 95% CI: 7.409–24.78).

Conclusion: Our study identified a novel gene-pair signature for prediction of RFS in CC.

Keywords: colon cancer, relapse-free survival, gene pair, prognosis

Introduction

Colorectal cancer (CRC) is one of the most common cancers worldwide, with an estimated 1.4 million new cases and 693,900 deaths per year.¹ Despite great improvements in curative resection and adjuvant chemotherapy, the relapse rate remains high. Approximately 40–50% of patients relapse within the first year after initial resection.² The TNM staging system is widely used in relapse prediction.³ A higher TNM stage is usually associated with a higher relapse rate and shorter interval. However, this is controversial among a proportion of patients. Stage IIb patients tend to have a higher relapse rate than those with stage IIIa.⁴ Thus, in clinical practice, there is a critical need to improve the prediction of relapse in colon cancer (CC).

With the development of high-throughput gene detection technology, gene expression profiles became available to identify more reliable molecular subtypes,

and diagnostic and prognostic molecular signatures in multiple diseases. In 2013, Marisa et al classified CC into six molecular subtypes (C1–C6) by clustering gene expression data of 433 samples and found that the subtypes of C4 and C6 were associated with shorter relapse-free survival (RFS).⁵ In 2015, Guinney et al coalesced six CRC subtyping algorithms (including those from Marisa et al⁵) into four consensus molecular subtypes (CMSs) with distinguishing features: CMS1 (microsatellite instability immune), CMS2 (canonical), CMS3 (metabolic), and CMS4 (mesenchymal).⁶ CMS4 tumors displayed worse overall survival and RFS. Alderdice et al compared the CRC transcriptional signatures of CMSs and colorectal cancer intrinsic subtypes (CRISs), and found that CRISs provided a better classification of molecular subtypes than CMSs.⁷ More reliable molecular subtyping system will be detected in large-scale prospective designed studies.

Several studies have also proposed gene-expression-based prognostic signatures in CRC.^{8–11} However, those signatures are limited in sample size and sufficient cross-validation. In this study, we adopted a method based on relative ranking of gene expression levels, which could integrate samples from different datasets without the need for eliminating batch effects.^{12–14} Thus, we reached a large sample size and developed and validated a novel gene-pair signature for RFS prediction in CC.

Materials and methods

Data collection

The Gene Expression Omnibus (GEO) and European Bioinformatics Institute (EMBL-EBI) databases were searched for related microarray datasets from inception to April 1, 2018. A dataset was included if it fulfilled the following criteria: 1) it focused on gene expression profiles of CC tissues; 2) it was based on the chip platform of Affymetrix Human Genome U133 Plus 2.0 Array (GPL570); 3) the follow-up duration and RFS status (RFS was defined as the time from surgery to the first confirmed relapse) were available; and 4) the sample size was >50. Normalized gene expression profiles from the GEO database were downloaded as the training set, while RNA-sequencing data from The Cancer Genome Atlas (TCGA) database (TCGA Colon Cancer [COAD]) was selected as the validation set.

Data quality control

Microarray quality was assessed by sample clustering according to the distance between different samples in Pearson's correlation matrices. A height cut-off of 0.2 was chosen to

identify potential microarray outliers, which were removed from the subsequent analysis.¹⁵

RFS-related gene screening

To screen out the common RFS-related genes between datasets, microarray meta-analysis was conducted on the included datasets using R software and the MetaOmics package. In survival analysis, we took the log-rank test and permutation method ($n=300$) for an individual dataset, and then both Fisher and maxP methods were adopted to combine the P -values between datasets. A P -value <0.05 by both methods was considered statistically significant for the RFS-related probes, which were subsequently converted to gene symbols using the Affymetrix annotation files.

Development of gene-pair signature

The expression level of RFS-related genes in a specific sample underwent pairwise comparison to generate a score for each gene pair. If the first gene of one gene pair had a higher expression level than the second one, a gene pair score of 1 was assigned; otherwise, the gene pair score was 0. The profile of gene-pair scores was used to build the signature. To minimize the risk of over-fitting, we used a Cox proportional hazards regression model combined with the least absolute shrinkage and selection criteria operator (R glmnet package). The penalty parameter was estimated by 10-fold cross-validation at 1 standard error beyond the minimum partial likelihood deviance.

Signature evaluation and validation

The coefficients of significant gene pairs in the model were extracted to calculate a risk score for each sample. The concordance index (C-index) was used to assess the performance of the risk scores in RFS prediction. Then, the nearest neighbor estimation method was adopted to estimate the time-dependent receiver operating characteristics (ROC) curve of the risk scores (R survcomp package). The area under the ROC curve (AUC) was also used to assess the predictive ability of the signature risk score. The score corresponding to the shortest distance between the ROC curve and the point of 100% true positive and 0% false positive was selected as the cut-off value. The patients were divided into high- and low-risk groups according to the cut-off. Survival difference between the groups was assessed by the Kaplan–Meier estimate and compared using the log-rank test (R survival package). To validate the signature, we used the same methods in the RNA-sequencing dataset. Furthermore, univariate and multivariate Cox regression survival analyses were conducted on the risk score.

Functional enrichment analysis

To identify potential biological pathways related to the signature, gene set enrichment analysis (GSEA) was performed to detect whether a series of a priori defined biological pathways was enriched in the gene rank derived from differentially expressed genes between high- and low-risk groups. In addition, the gene-pair signature consisted of single genes, and gene ontology (GO) analysis was conducted on these genes to identify potential mechanism associated with the signature. A false discovery rate of <0.05 was chosen as the cut-off criterion.

Statistical analysis

All statistical analyses were performed using R 3.3.2 software. For use with GSEA software, the collection of annotated gene sets of c2.cp.kegg.v5.2.symbols.gmt in the Molecular Signatures Database was chosen as the reference gene sets. GO analysis was conducted with the online tool ToppGene (<https://toppgene.cchmc.org/>). A two-sided P -value <0.05 was considered statistically significant.

Results

Characteristics of included datasets

Three datasets (GSE17536, GSE17537, and GSE39582) were excluded owing to duplication, while one dataset (GSE33113) was excluded owing to incomplete expression data. Thirty cases of rectal cancer were excluded in GSE14333. Finally, four datasets (GSE14333 [n=196], GSE17538 [n=200], GSE38832 [n=92], and GSE40967 [n=557]) were included, with a total number of 1,045 patients. Those patients had a mean follow-up of 46.2 (0–201) months and a relapse rate of 37.7% (394/1,045). The validation set had a total of 384 patients, with a mean follow-up of 29.7 (0.2–150.1) months and a relapse rate of 20.1% (81/384). No samples were detected as outliers in the datasets (Figure 1).

Signature construction, evaluation, and validation

In microarray meta-analysis, 647 probes were significant ($P<0.05$) in both Fisher and maxP methods (Figure 2). These 647 probes could be mapped to 596 genes and constructed 177,310 gene pairs. The LASSO penalized generalized linear model identified 16 significant gene pairs (consisting of 24 genes) (Table 1). Then, a risk score was calculated for each sample according to the model coefficients as follows:

$$\begin{aligned} \text{Risk score} = & \text{Score}_{\text{CYP1B1 LIPE}} * 0.016 + \text{Score}_{\text{CYP1B1 GOLGA2P10}} * \\ & 0.027 + \text{Score}_{\text{SCG2 FOX P3}} * 0.283 + \text{Score}_{\text{SCG2 MED26}} * (-0.208) \\ & + \text{Score}_{\text{SCG2 LYPD5}} * 0.234 + \text{Score}_{\text{ACADSB SDHAF1}} * 0.350 + \\ & \text{Score}_{\text{ACADSB SLC25A19}} * 0.042 + \text{Score}_{\text{MATN3 R3HDM1}} * (-0.209) \\ & + \text{Score}_{\text{RGL2 SLC25A28}} * 0.223 + \text{Score}_{\text{RGL2 TMEM201}} * 0.409 + \\ & \text{Score}_{\text{GRB10 GOLGA2P10}} * 0.280 + \text{Score}_{\text{MAP1B SLC25A22}} * 0.334 \\ & + \text{Score}_{\text{ETS1 MED26}} * 0.507 + \text{Score}_{\text{WFDC1 HOOK3}} * (-0.418) + \\ & \text{Score}_{\text{GOLGA2P10 ANGPT2}} * 0.088 + \text{Score}_{\text{TCF21 ISM1}} * (-0.059) \end{aligned}$$

The time-dependent ROC curve analysis indicated an optimal cut-off value of 1.168 for the classification of high-risk and low-risk groups. The high-risk group had an obviously higher incidence of relapse in the follow-up (Figure 3).

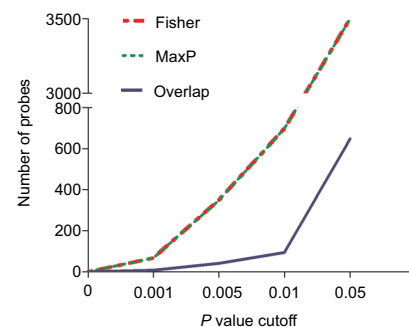


Figure 2 Number distribution of significant probes by different methods.

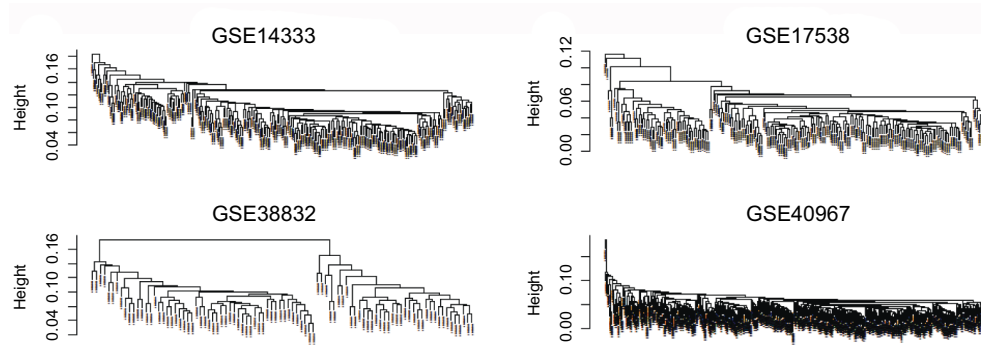


Figure 1 Sample clustering to detect outliers.

Table 1 Gene-pair signature information

Gene-pair 1	Probe	Gene-pair 2	Probe	Coefficient
<i>CYP1B1</i> (cytochrome P450 family 1 subfamily B member 1)	202435_s_at	<i>LIPE</i> (lipase E, hormone-sensitive type)	208186_s_at	0.016
<i>CYP1B1</i> (cytochrome P450 family 1 subfamily B member 1)	202435_s_at	<i>GOLGA2P10</i> (golgin A2 pseudogene 10)	223327_x_at	0.027
<i>SCG2</i> (secretogranin II)	204035_at	<i>FOXP3</i> (forkhead box P3)	224211_at	0.283
<i>SCG2</i> (secretogranin II)	204035_at	<i>MED26</i> (mediator complex subunit 26)	231724_at	-0.208
<i>SCG2</i> (secretogranin II)	204035_at	<i>LYPD5</i> (LY6/PLAUR domain containing 5)	236039_at	0.234
<i>ACADSB</i> (acyl-CoA dehydrogenase, short/branched chain)	205355_at	<i>SDHAF1</i> (succinate dehydrogenase complex assembly factor 1)	221855_at	0.350
<i>ACADSB</i> (acyl-CoA dehydrogenase, short/branched chain)	205355_at	<i>SLC25A19</i> (solute carrier family 25 member 19)	223222_at	0.042
<i>MATN3</i> (matrilin 3)	206091_at	<i>R3HDML</i> (R3H domain containing like)	234774_at	-0.209
<i>RGL2</i> (ral guanine nucleotide dissociation stimulator like 2)	209110_s_at	<i>SLC25A28</i> (solute carrier family 25 member 28)	221432_s_at	0.223
<i>RGL2</i> (ral guanine nucleotide dissociation stimulator like 2)	209110_s_at	<i>TMEM201</i> (transmembrane protein 201)	227617_at	0.409
<i>GRB10</i> (growth factor receptor bound protein 10)	210999_s_at	<i>GOLGA2P10</i> (golgin A2 pseudogene 10)	223327_x_at	0.280
<i>MAP1B</i> (microtubule associated protein 1B)	212233_at	<i>SLC25A22</i> (solute carrier family 25 member 22)	218725_at	0.334
<i>ETS1</i> (ETS proto-oncogene 1, transcription factor)	214447_at	<i>MED26</i> (mediator complex subunit 26)	231724_at	0.507
<i>WFDC1</i> (WVAP four-disulfide core domain 1)	219478_at	<i>HOOK3</i> (hook microtubule tethering protein 3)	226395_at	-0.418
<i>GOLGA2P10</i> (golgin A2 pseudogene 10)	223327_x_at	<i>ANGPT2</i> (angiopoietin 2)	236034_at	0.088
<i>TCF21</i> (transcription factor 21)	229529_at	<i>ISMI</i> (isthmin 1)	235182_at	-0.059

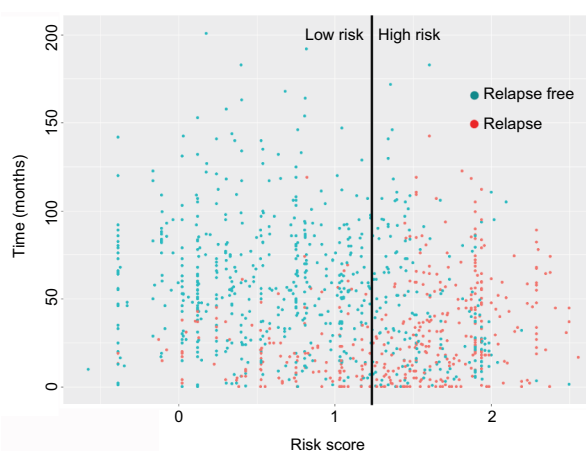
High-risk patients showed a poorer prognosis than low-risk patients (HR: 3.519, 95% CI: 2.870–4.314; $P < 0.0001$) (Figure 4). In the validation dataset, high-risk patients also had a poorer prognosis (HR: 2.926, 95% CI: 1.892–4.527; $P < 0.0001$).

In subgroup analysis, we evaluated the effects of age, gender, tumor location, TNM stage, genetic mutation (TP53, KRAS, and BRAF), molecular subtype (C1–C6), and adjuvant chemotherapy (Table 2). The results were consistent except for stage I and C1 subtype. We think that this may contribute to the excellent prognosis after curative surgery in these two subtypes.

As a continuous variable, the risk score showed a comparative ability in predicting RFS (AUC at 5 years: 0.724; C-index: 0.642, 95% CI: 0.615–0.669) (Figure 5). In the validation dataset, it had a similar performance (AUC at 5 years: 0.697; C-index: 0.696, 95% CI: 0.627–0.766). In multivariate Cox regression analysis, the risk score also showed a significant association with RFS (HR: 2.416, 95% CI: 1.824–3.198; $P < 0.001$) (Table 3).

Comparison with other signatures

To further investigate the predictive power of the risk score, we compared it with a 13-gene signature, a 15-gene signature, and TNM stage. The risk score in the 13-gene signature and 15-gene signature was calculated according to the coefficients provided by the primary studies.^{8,9} In ROC curve analysis, our signature showed a better predictive ability (AUC at 5 years: 0.724) than the 13-gene signature, 15-gene signature, and TNM stage (AUC at 5 years: 0.575, 0.509

**Figure 3** Distribution of relapse-free survival status and risk scores of the gene-pair signature.

and 0.703, respectively) (Figure 5). Then, we calculated the C-indices according to minimum follow-up duration. We found that in the patients with a longer follow-up, the risk score showed a better performance in RFS prediction ($R^2 = 0.988$, $P < 0.0001$) (Figure 6). In those with a follow-up of > 5 years, the C-index was up to 0.869 (95% CI: 0.816–0.922), and high-risk patients had a significantly poorer prognosis (HR: 13.55, 95% CI: 7.409–24.78) (Figure 7). In comparison, our signature had a higher C-index than the 13-gene signature (P for paired t -test < 0.0001), 15-gene signature ($P = 0.0004$), and TNM stage (P for paired t -test $= 0.0125$). Graphically, our signature also had better astringency and stability.

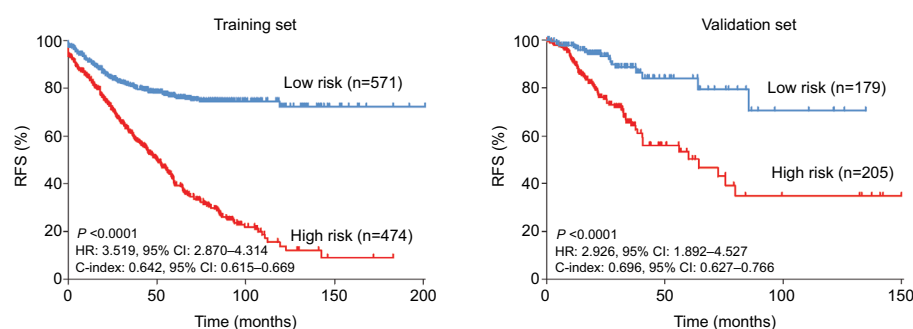


Figure 4 Kaplan–Meier curves of RFS in the training and validation sets.

Abbreviation: RFS, relapse-free survival.

Table 2 Subgroup analysis for the HRs between high- and low-risk groups divided by the gene-pair signature

Subgroup	Number of patients		HR (95% CI)	Log-rank <i>P</i> -value
	High risk	Low risk		
All	474	571	3.519 (2.870–4.314)	<0.0001
Age (years)				
<65	172	217	3.856 (2.804–5.302)	<0.0001
≥65	228	335	4.349 (3.288–5.752)	<0.0001
Gender				
Male	217	293	4.250 (3.198–5.648)	<0.0001
Female	183	260	3.984 (2.919–5.439)	<0.0001
Tumor location				
Right	163	156	3.514 (2.519–4.903)	<0.0001
Left	186	246	3.796 (2.802–5.144)	<0.0001
TNM stage				
I	17	65	0.289 (0.009–9.594)	0.4871
II	106	259	1.959 (1.156–3.322)	0.0125
III	112	203	2.579 (1.708–3.895)	<0.0001
IV	46	41	3.993 (2.032–7.845)	<0.0001
TP53 mutation				
Yes	57	131	5.809 (3.387–9.962)	<0.0001
No	45	111	3.814 (1.960–7.422)	<0.0001
KRAS mutation				
Yes	65	149	2.831 (1.666–4.811)	<0.0001
No	87	235	5.412 (3.265–8.968)	<0.0001
BRAF mutation				
Yes	10	39	6.132 (1.339–28.09)	0.0205
No	129	325	3.837 (2.561–5.749)	<0.0001
Molecular subtype				
C1	10	104	42.84 (9.319–196.9)	<0.0001
C2	13	88	0.499 (0.131–1.891)	0.3061
C3	10	61	4.728 (1.116–20.04)	0.0350
C4	42	17	2.622 (1.206–5.697)	0.0149
C5	57	95	4.078 (2.173–7.654)	<0.0001
C6	26	34	2.353 (1.055–5.247)	0.0365
Adjuvant chemotherapy				
Yes	148	164	3.409 (2.423–4.796)	<0.0001
No	191	234	3.933 (2.892–5.347)	<0.0001

Biological pathways associated with the gene-pair signature

In the validation set, 374 patients were divided into high- and low-risk groups according to the risk score. Three KEGG pathway-related gene sets were enriched in the low-risk group,

namely “peroxisome”, “pyruvate metabolism”, and “citrate cycle (TCA cycle)” (where TCA is tricarboxylic acid) (Figure 8). In GO analysis of the 24 genes comprising the signature, these genes were associated with multiple biological processes, such as angiogenesis, endothelial cell migration, and T-cell toleration.

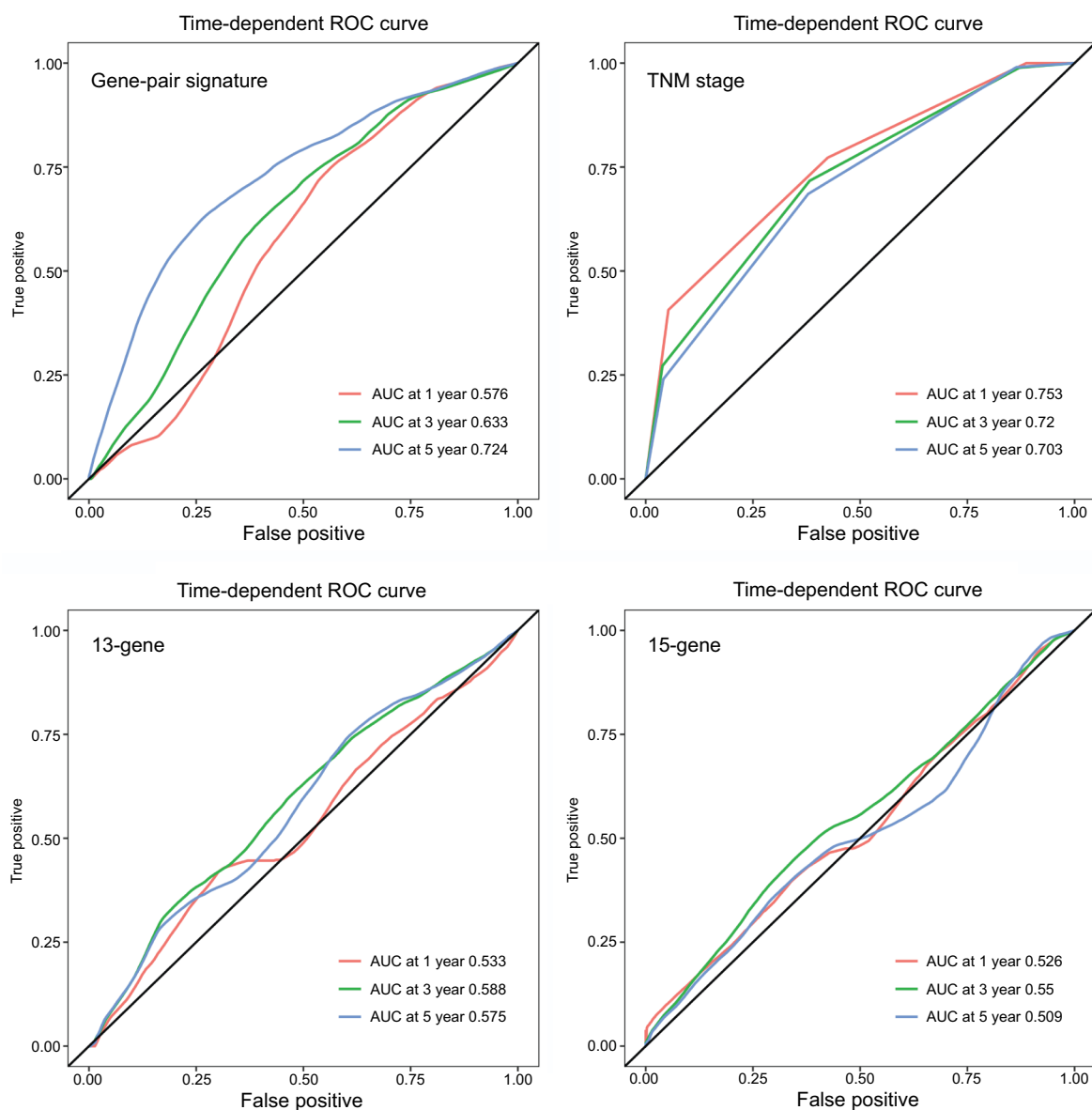


Figure 5 Time-dependent ROC curve analysis.

Abbreviations: ROC, receiver operating characteristics; AUC, area under the curve.

Table 3 Univariate and multivariate Cox regression analyses in colon cancer

Variable	Univariate analysis		Multivariate analysis	
	HR (95% CI)	P-value	HR (95% CI)	P-value
Age	1.001 (0.993–1.009)	0.854	1.006 (0.992–1.019)	0.411
Gender	1.066 (0.872–1.303)	0.530	1.412 (1.022–1.951)	0.036
Tumor location	1.083 (0.870–1.347)	0.477	0.763 (0.547–1.062)	0.109
TNM stage	2.847 (2.378–3.408)	0.000	2.093 (1.623–2.701)	0.000
Adjuvant chemotherapy	1.220 (0.978–1.523)	0.079	1.057 (0.738–1.513)	0.762
Risk score	2.719 (2.320–3.185)	0.000	2.416 (1.824–3.198)	0.000

Discussion

Tumor relapse within the first year after initial resection occurred in almost 50% patients, which was closely related to survival outcomes. Currently, the TNM staging system

is widely used to predict RFS. However, CRC patients with the same stage and treatment varied considerably in the RFS prediction, which may be caused by the genetic differences between individuals.¹⁶ Thus, genomic biomarkers for relapse

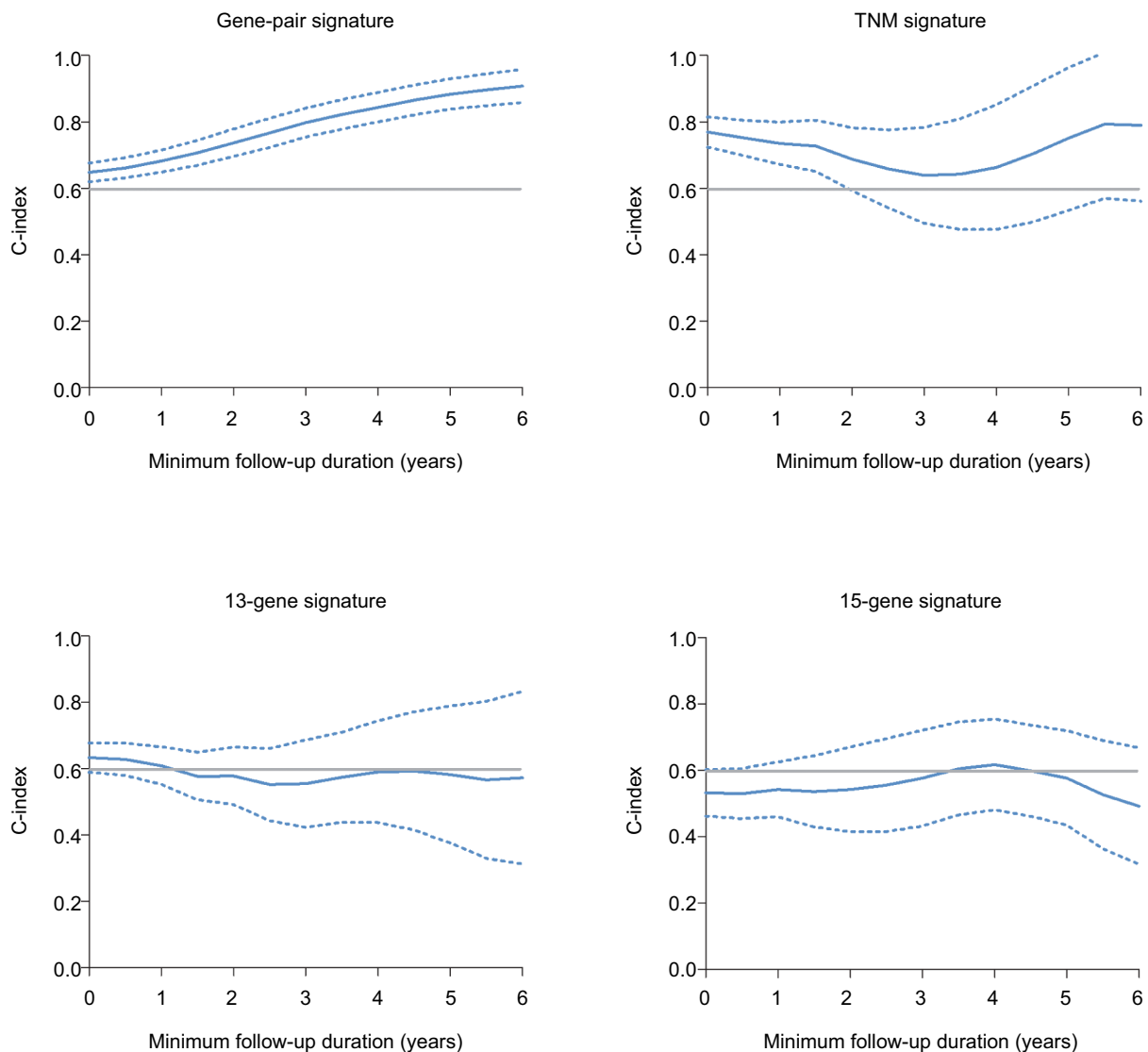


Figure 6 C-indices according to minimum follow-up duration.

Note: Dashed line: 95% CI.

Abbreviation: C-index, concordance index.

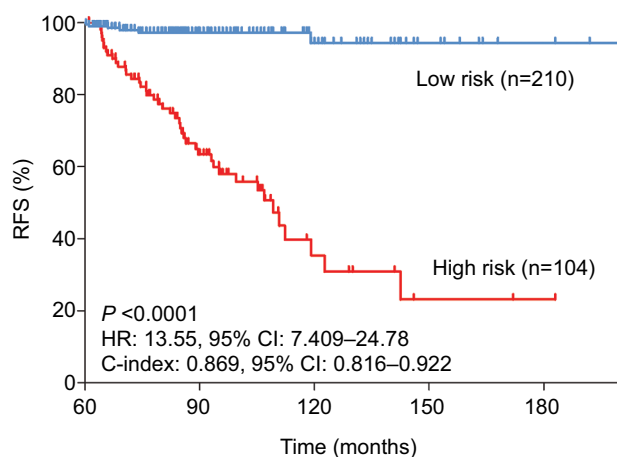


Figure 7 Kaplan-Meier curves of relapse-free survival in patients with a follow-up of >5 years.

Abbreviations: C-index, concordance index; RFS, relapse-free survival.

detection could make up for the deficiency of TNM classification, and thereby identify high-risk patients to improve personalized cancer care at an earlier stage.^{17,18}

With the development of gene microarray and sequencing technologies, we are faced with a large amount of biological data. Genome and transcriptome information is available in more and more samples, prompting a search for new biomarkers for the discrimination of high-risk individuals in disease diagnosis or prognosis. A genomic signature is a biomarker in which the genomic data are combined in a defined manner to provide either a continuous score or a categorical classifier for clinical decision-making.

Several CRC prognostic signatures have been reported using the public microarray data from the GEO database. In

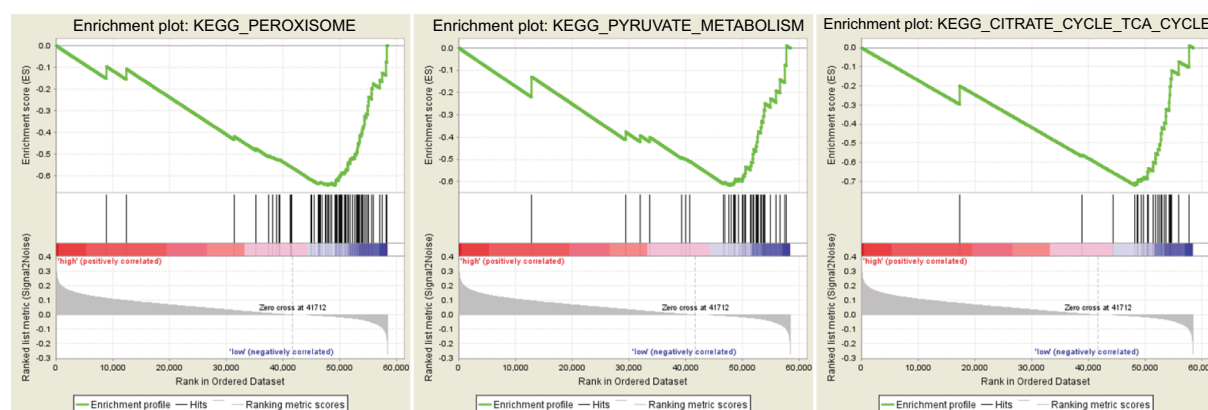


Figure 8 Gene set enrichment analysis.

the study by Dai et al, the dataset GSE39582 was used to identify a 15-mRNA signature as an early relapse classifier in CC, which was validated in the datasets of GSE14333, GSE17538, GSE33113, and GSE37892.⁸ Tian et al used GSE17536 to build a 13-mRNA signature for RFS prediction in CRC and validated it in GSE14333 and GSE33113.⁹ Among these three signatures, only one gene, *SCG2*, overlapped between the gene-pair signature and 13-gene signature, which has been reported in multiple cancers.^{19–21} This may have contributed to the limited sample size and subsequent selection bias, and a systematic method was needed to identify the prognostic genes. In addition, no signatures based on the microarray data were validated in the RNA-sequencing data, and this may affect their further application in clinical practice.

In this study, we constructed an RFS-related signature consisting of 16 gene pairs using a penalized Cox proportional hazards regression model. With specific coefficients, a risk score was calculated for each patient based on the gene expression profile, and this could be used without considering batch effects. Thus, we achieved a large sample size which also helped us to identify a robust signature, which showed a better performance than other single-gene signatures, especially among the patients with a longer follow-up. A study by Tauriello et al used a k-top scoring pair (k-TSP) approach to classify CRCs between CMS4 and CMS1–3. It also adopted the method of gene pairs and had a good performance.²² In our study, CMS information was unavailable, but C1–C6 subtyping information was available. High-risk individuals in our signature had a poor prognosis in C2–C6, indicating a signature stability of different molecular subtypes across patients.

To investigate the potential mechanism associated with the signature, GSEA and GO analyses were conducted. Three gene sets were significantly enriched in low-risk patients. The role of the peroxisome pathway in cancer is controversial. Its inactivation leads to metabolic stress, mTOR inhibition, and

lethality in liver cancer cells.²³ However, it protects lymphoma cells from histone deacetylase inhibitor-mediated apoptosis.²⁴ Fransen et al present potential therapeutic opportunities for targeting the TCA cycle in cancer cells.²⁵ Moreover, in GO analysis, 24 genes in the signature were associated with multiple biological processes, including angiogenesis, endothelial cell migration, and T-cell toleration, which have been reported in tumor carcinogenesis, progression, and metastasis.^{26–28}

Several limitations to our study should be acknowledged. First, this was a retrospective study, rather than a prospective cohort study. Second, in the training set, we considered the microarray data based only on GLP570 and ignored other unusual platforms, which may have led to selection bias. Third, we ignored batch effects caused by other confounders, such as genetic background and environmental factors, which may have affected the model coefficients and cut-off value.

Conclusion

Our study identified a novel gene-pair signature for RFS prediction in CC.

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Disclosure

The authors report no conflicts of interest in this work.

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