Incidence of febrile neutropenia during chemotherapy among patients with nonmyeloid cancer receiving filgrastim vs a filgrastim biosimilar

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Background: Filgrastim and other granulocyte colony-stimulating factors are recommended to decrease febrile neutropenia (FN) incidence among patients with nonmyeloid cancers undergoing chemotherapy. Data comparing biosimilar filgrastim-sndz with reference filgrastim (filgrastim-ref) are limited outside of clinical trials in the US.

Objective: To compare the incidence of FN across chemotherapy cycles 1–6 between patients treated with filgrastim-sndz vs filgrastim-ref.

Materials and methods: This was a retrospective claims analysis of patients with nonmyeloid cancer enrolled in commercial or Medicare Advantage plans from March 2015 to June 2016 and receiving filgrastim-sndz or filgrastim-ref during ≥1 completed chemotherapy cycle. Patients undergoing hematopoietic stem cell transplantation, pregnant patients, and those with missing data were excluded. FN was identified using the diagnosis codes for neutropenia + infection, neutropenia + bacterial/fungal infection, and neutropenia + infection + fever. Equivalence testing for FN incidence at the cycle level across chemotherapy cycles 1–6 was conducted for filgrastim-sndz vs filgrastim-ref after adjusting for baseline characteristics using inverse probability of treatment weighting. Results were considered equivalent if the 90% CIs for between-cohort differences were within ±6.0%.

Results: The analysis included 3,459 patients (162 filgrastim-sndz and 3,297 filgrastim-ref). Before weighting, the filgrastim-sndz cohort was younger than filgrastim-ref and had a higher proportion of men, a higher proportion with commercial insurance, and lower proportions undergoing hematopoietic stem cell transplantation, pregnant patients, and those with missing data were excluded. FN was identified using the diagnosis codes for neutropenia + fever, neutropenia + bacterial/fungal infection, and neutropenia + infection + fever. Equivalence testing for FN incidence at the cycle level across chemotherapy cycles 1–6 was conducted for filgrastim-sndz vs filgrastim-ref after adjusting for baseline characteristics using inverse probability of treatment weighting. Results were considered equivalent if the 90% CIs for between-cohort differences were within ±6.0%.

Conclusion: Filgrastim-sndz and filgrastim-ref are statistically equivalent for preventing FN across chemotherapy cycles 1–6 among patients with nonmyeloid cancer.

Keywords: biosimilar pharmaceuticals, febrile neutropenia, filgrastim, granulocyte colony-stimulating factor, retrospective studies

Introduction

Patients with cancer who receive myelosuppressive chemotherapy are susceptible to febrile neutropenia (FN),1 a potentially life-threatening side effect that increases
infection risk and frequently requires hospitalization and treatment with broad-spectrum antibiotics. Individuals who develop FN during chemotherapy treatment not only have increased rates of morbidity and mortality but are also more likely to experience treatment delay, reduction in chemotherapy dose intensity, or even treatment discontinuation, all of which can lead to reduced treatment response and lower survival.

Prophylaxis with granulocyte colony-stimulating factors (G-CSFs) such as filgrastim (Neupogen®, Amgen Inc., Thousand Oaks, CA, USA), which was approved in the US in 1991, has been shown to reduce the risk of these negative outcomes and improve chemotherapy-related quality of life among patients with solid tumors or nonmyeloid malignancies. Accordingly, current practice guidelines in the US and Europe recommend G-CSF prophylaxis for patients whose chemotherapy regimens are associated with high (≥20%) risk of FN, as well as for some patients on chemotherapy regimens with intermediate FN risk (10%–20%), depending on patient characteristics such as age, comorbidities, disease stage, and previous history of FN. However, only about 17% of patients who meet the high-risk criteria actually receive prophylaxis, and there is evidence that the high cost of G-CSFs—like many biologics—are expensive drugs that may constitute a treatment barrier for some patients.

Approval of the filgrastim biosimilar filgrastim-sndz (US trade name Zarzio®, Sandoz Inc., Princeton, NJ, USA), which in 2015 became the first biosimilar available in the US via the abbreviated pathway established by the Biologics Price Competition and Innovation Act, was an important step in potentially expanding access to G-CSF treatment. Filgrastim-sndz was approved on the basis of a substantial body of analytical, preclinical, and clinical data demonstrating that there were no clinically meaningful differences between the biosimilar and the reference drug (filgrastim-ref) for preventing FN among patients receiving myelosuppressive chemotherapy. A subsequent multicenter prospective observational study (MONITOR-GCSF) demonstrated that safety and efficacy outcomes among patients treated with biosimilar filgrastim (EU trade name Zarzio®, Sandoz Inc.) were similar to those reported historically for filgrastim-ref, and studies conducted in the US and Europe have since indicated that biosimilar filgrastim is the most cost-efficient approach for preventing FN among patients undergoing myelosuppressive chemotherapy in routine clinical practice.

Until recently, real-world comparative data regarding clinical outcomes among patients treated with biosimilar filgrastim had been lacking. We conducted one of the first retrospective observational studies to compare outcomes between filgrastim-sndz–treated vs filgrastim-ref–treated patients undergoing chemotherapy treatment for nonmyeloid malignancies in US clinical practice, demonstrating that the incidence of FN was statistically equivalent between these groups during the first chemotherapy cycle. A contemporaneous analysis comparing FN incidence between filgrastim-sndz–treated vs filgrastim-ref–treated patients in the first 14 days after G-CSF use had similar results. While neutropenic complications are most common in the first cycle of chemotherapy, examination of subsequent cycles is warranted to confirm equivalence over a longer period. We conducted the present study to expand on earlier findings by extending the equivalency analysis up through the first six cycles of chemotherapy.

Materials and methods
Study design and data source
This was a retrospective claims analysis conducted from 01 September 2014 through 31 July 2016 (study period) using data from the Optum Research Database (ORD). The ORD contains fully de-identified medical and pharmacy claims and enrollment information for commercial and Medicare Advantage health plan members, is geographically diverse across the US, and, for individuals 18 years and older, is linked to dates of death (month and year) from the Social Security Administration death master file. The information included in medical claims and outpatient pharmacy claims within the ORD was described in detail in our previous report. Because no identifiable protected health information was extracted or accessed for the purpose of this study, institutional review board approval or waiver of authorization was not required.

Study sample
The study population comprised adult patients (aged 18 years or older) undergoing chemotherapy for nonmyeloid cancer who were treated with filgrastim-sndz or filgrastim-ref in at least one completed cycle of a chemotherapy treatment regimen from 01 March 2015 through 30 June 2016 (patient identification period). Patient selection and cohort assignment (filgrastim-sndz or filgrastim-ref) are illustrated in Figure 1 and described in detail in our previous report. The end of the follow-up period was defined as the end of continuous health plan enrollment, death, or 31 July 2016 (whichever came earliest).

Chemotherapy lines of therapy and cycle definitions
This analysis used up through the first six observed chemotherapy cycles of the first observed course of adjuvant, neoad-
Filgrastim vs biosimilar for preventing febrile neutropenia

Figure 1 Sample selection and attrition flow diagram.

juvant, or metastatic chemotherapy (line of therapy [LOT]) for each patient, which by definition began on the index date. The methodology for determining LOT end dates was described previously.23 The start of the first observed chemotherapy cycle began on the first day of the first observed LOT and included all chemotherapy regimens received during the first 6 days of the cycle. The end of the first observed chemotherapy cycle was defined as the earlier of the next fill/infusion of cancer
therapy occurring on or after day 8 of the LOT, or the end of
LOT. Cancer therapy fills/infusions that occurred on or after
day 8 of the LOT were defined as subsequent chemotherapy
cycles. Chemotherapy cycles that ended due to censoring (ie,
disenrollment or study end) were considered to be incomplete
and excluded from the analysis.

Study measures
Baseline characteristics
Patient demographic and clinical characteristics were
assessed during the 6 months prior to the index date (baseline
period). Comorbidity burden was evaluated using the Quan–
Charlson comorbidity index.26 Because FN rarely occurs
within the first 5 days of chemotherapy, use of filgrastim-
sndz or filgrastim-ref that was initiated on or before day 5
within the first 5 days of chemotherapy, use of filgrastim-
CSF doses, length of chemotherapy cycle 1, use of G-CSF
insurance type, Quan–Charlson index score, number of
completed cycles, age, sex, geographic region,
assignment. The covariates used in propensity scoring were
are used to create a synthetic sample in which the distri-
bution of baseline covariates is independent of treatment
assignment. The covariates used in propensity scoring were
number of completed cycles, age, sex, geographic region,
insurance type, Quan–Charlson index score, number of
CSF doses, length of chemotherapy cycle 1, use of G-CSF
prophylaxis, cancer type, metastatic cancer in general and
specific to bone, infection, radiation, surgery, renal or
hepatic dysfunction, and prior use of CSF agents. As in our
previous analysis of cycle 1, only complete cycles were used
and patients whose first cycle was censored were dropped
from the study.27 For the present multiple-cycle analysis,
the loss of some cycles 2–6 due to censoring was adjusted
for by applying additional inverse probability of censoring
weights to the non-censored (ie, complete) cycles. These
additional weights depended only on the cycle number and the
treatment cohort based on cycle 1 for the patient. The
probability of censoring used in the weights was calculated
using separate Kaplan–Meier analyses for each cycle number
2 through 6, stratified by treatment cohort.30 Baseline
characteristics were analyzed descriptively on the cycle
level and presented by unweighted and weighted treatment
cohorts; numbers and percentages were provided for cat-
egorical variables, and means were provided for continuous
variables. Standardized differences were calculated to com-
pare characteristics between filgrastim-sndz vs filgrastim-
ref cohorts (unweighted and weighted); values <10% are
by convention considered to indicate high similarity.29 All
statistical analyses were performed using SAS 9.4 (SAS
Institute, Cary, NC, USA).

Statistical equivalence testing
FN incidence was calculated among the weighted cohorts
for all three definitions of FN. Bootstrap methods were
used to calculate 90% CIs of the difference in FN incidence
between groups.31 CIs within ±6% were considered to denote
statistical equivalence between cohorts on the basis of the
0%–11% range of neutropenic fever incidence reported in
the PIONEER trial.32

Results
Study sample
A total of 3,542 patients (172 filgrastim-sndz and 3,370
filgrastim-ref) met the study criteria and were included in
the unweighted treatment cohorts (Figure 1). There were 162
patients in the weighted filgrastim-sndz cohort and 3,297 in
the weighted filgrastim-ref cohort after IPTW, including a
total of 349 chemotherapy cycles for filgrastim-sndz–treated
patients and 6,546 chemotherapy cycles for filgrastim-ref–
treated patients (excluding censored cycles and patients;
Figure 1). Before weighting, the filgrastim-sndz cohort was
younger than the filgrastim-ref cohort (mean age 60.9 vs 64.8
years) and had a higher proportion of men (42.1% vs 35.6%),
a higher proportion of patients with commercial insurance
(61.0% vs 44.1%), a lower proportion of patients with G-CSF prophylaxis (45.6% vs 51.6%), and a lower proportion of patients with metastatic cancer (39.0% vs 46.6%), as shown in Table 1. The unweighted mean numbers of G-CSF doses received for filgrastim-sndz and filgrastim-ref were 2.4 and 2.2, respectively, with prophylaxis occurring during 45.6% of chemotherapy cycles among filgrastim-sndz patients and 51.6% of chemotherapy cycles among filgrastim-ref patients. After weighting, baseline characteristics were similar between cohorts and standardized differences were acceptably low (Table 1).

Table 1 Baseline patient characteristics by weighted and unweighted cycles

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Unweighted</th>
<th>Filgrastim-sndz (n=349)</th>
<th>Filgrastim-ref (n=6,546)</th>
<th>Stand. diff. (%)</th>
<th>Weighteda</th>
<th>Filgrastim-sndz (n=349)</th>
<th>Filgrastim-ref (n=6,546)</th>
<th>Stand. diff. (%)</th>
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<tbody>
<tr>
<td>Observed cycle, n (%)</td>
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<tr>
<td>1</td>
<td>162 (46.4)</td>
<td>3,297 (50.4)</td>
<td>–7.9</td>
<td>174 (49.9)</td>
<td>3,115 (47.6)</td>
<td>4.7</td>
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<tr>
<td>2</td>
<td>61 (17.5)</td>
<td>1,116 (17.0)</td>
<td>1.1</td>
<td>56 (16.2)</td>
<td>1,100 (16.9)</td>
<td>–2.1</td>
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<tr>
<td>3</td>
<td>51 (14.6)</td>
<td>846 (12.9)</td>
<td>4.9</td>
<td>49 (14.0)</td>
<td>863 (13.2)</td>
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<td>4</td>
<td>33 (9.5)</td>
<td>590 (9.0)</td>
<td>1.5</td>
<td>34 (9.9)</td>
<td>640 (9.8)</td>
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<tr>
<td>5</td>
<td>26 (7.4)</td>
<td>414 (6.3)</td>
<td>4.5</td>
<td>24 (6.7)</td>
<td>481 (7.3)</td>
<td>–2.4</td>
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<tr>
<td>6</td>
<td>16 (4.6)</td>
<td>283 (4.3)</td>
<td>1.3</td>
<td>12 (3.3)</td>
<td>338 (5.2)</td>
<td>–9.2</td>
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<tr>
<td>Mean age, years</td>
<td>60.9</td>
<td>64.8</td>
<td>–28.6</td>
<td>65.6</td>
<td>64.6</td>
<td>7.4</td>
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<tr>
<td>Male, n (%)</td>
<td>147 (42.1)</td>
<td>2,333 (35.6)</td>
<td>13.3</td>
<td>97 (27.9)</td>
<td>2,356 (36.0)</td>
<td>–17.5</td>
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<tr>
<td>Geographic region, n (%)</td>
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<tr>
<td>Northeastd</td>
<td>10 (2.9)</td>
<td>980 (15.0)</td>
<td>–43.5</td>
<td>141 (40.4)</td>
<td>2,576 (39.3)</td>
<td>2.2</td>
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<tr>
<td>Midwestd</td>
<td>106 (30.4)</td>
<td>1,624 (24.8)</td>
<td>12.5</td>
<td>–</td>
<td>–</td>
<td>–</td>
<td></td>
<td></td>
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<tr>
<td>South</td>
<td>178 (51.0)</td>
<td>3,168 (48.4)</td>
<td>5.2</td>
<td>171 (48.9)</td>
<td>3,182 (48.6)</td>
<td>0.5</td>
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<tr>
<td>West</td>
<td>55 (15.8)</td>
<td>773 (11.8)</td>
<td>11.5</td>
<td>37 (10.7)</td>
<td>788 (12.0)</td>
<td>–4.2</td>
<td></td>
<td></td>
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<tr>
<td>Otherd</td>
<td>0 (0.0)</td>
<td>1 (0.0)</td>
<td>–1.7</td>
<td>–</td>
<td>–</td>
<td>–</td>
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<tr>
<td>Insurance type, n (%)</td>
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<tr>
<td>Commercial</td>
<td>213 (61.0)</td>
<td>2,890 (44.1)</td>
<td>34.3</td>
<td>151 (43.2)</td>
<td>2,948 (45.0)</td>
<td>–3.7</td>
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<tr>
<td>Medicare</td>
<td>136 (39.0)</td>
<td>3,656 (55.9)</td>
<td>–34.3</td>
<td>198 (56.8)</td>
<td>3,598 (55.0)</td>
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<tr>
<td>Mean Quan–Charlson score</td>
<td>5.6</td>
<td>5.8</td>
<td>–7.2</td>
<td>5.9</td>
<td>5.8</td>
<td>7.3</td>
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<td></td>
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<tr>
<td>Mean number of G-CSF doses</td>
<td>2.3</td>
<td>2.2</td>
<td>9.2</td>
<td>2.3</td>
<td>2.2</td>
<td>3.2</td>
<td></td>
<td></td>
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<tr>
<td>Mean length of cycle 1, days</td>
<td>20.9</td>
<td>21.5</td>
<td>–7.1</td>
<td>21.9</td>
<td>21.4</td>
<td>5.7</td>
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<tr>
<td>G-CSF prophylaxis, n (%)</td>
<td>159 (45.6)</td>
<td>3,379 (51.6)</td>
<td>–12.1</td>
<td>185 (53.0)</td>
<td>3,366 (51.4)</td>
<td>3.1</td>
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</tr>
<tr>
<td>Cancer</td>
<td></td>
<td></td>
<td></td>
<td></td>
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<td></td>
</tr>
<tr>
<td>None</td>
<td>4 (1.1)</td>
<td>52 (0.8)</td>
<td>3.6</td>
<td>1 (0.4)</td>
<td>53 (0.8)</td>
<td>–4.9</td>
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</tr>
<tr>
<td>Lymphoid</td>
<td>38 (10.9)</td>
<td>366 (5.6)</td>
<td>19.4</td>
<td>12 (3.5)</td>
<td>388 (5.9)</td>
<td>–11.2</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Solid</td>
<td>295 (84.5)</td>
<td>5,562 (85.0)</td>
<td>–1.2</td>
<td>304 (87.2)</td>
<td>5,563 (85.0)</td>
<td>6.5</td>
<td></td>
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</tr>
<tr>
<td>Both</td>
<td>12 (3.4)</td>
<td>566 (8.6)</td>
<td>22.0</td>
<td>31 (8.8)</td>
<td>542 (8.3)</td>
<td>1.9</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Metastatic cancer</td>
<td>136 (39.0)</td>
<td>3,053 (46.6)</td>
<td>–15.5</td>
<td>171 (49.0)</td>
<td>3,022 (46.3)</td>
<td>5.6</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Metastatic cancer to bone</td>
<td>45 (12.9)</td>
<td>883 (13.5)</td>
<td>–1.8</td>
<td>39 (11.3)</td>
<td>882 (13.5)</td>
<td>–6.7</td>
<td></td>
<td></td>
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<tr>
<td>Prior infection</td>
<td>142 (40.7)</td>
<td>3,164 (48.3)</td>
<td>–15.4</td>
<td>195 (56.0)</td>
<td>3,134 (47.9)</td>
<td>16.3</td>
<td></td>
<td></td>
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<tr>
<td>Prior radiation</td>
<td>30 (8.6)</td>
<td>1,003 (15.3)</td>
<td>–20.8</td>
<td>67 (19.2)</td>
<td>974 (14.9)</td>
<td>11.4</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Prior surgery</td>
<td>138 (39.5)</td>
<td>2,490 (38.0)</td>
<td>3.1</td>
<td>142 (40.6)</td>
<td>2,495 (38.1)</td>
<td>5.2</td>
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<td></td>
</tr>
<tr>
<td>Renal dysfunction</td>
<td>83 (23.8)</td>
<td>1,409 (21.5)</td>
<td>5.4</td>
<td>82 (23.5)</td>
<td>1,413 (21.6)</td>
<td>4.5</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hepatic dysfunction</td>
<td>81 (23.2)</td>
<td>1,953 (29.8)</td>
<td>–15.1</td>
<td>122 (34.9)</td>
<td>1,935 (29.6)</td>
<td>11.4</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Prior use of G-CSF agents</td>
<td>56 (16.0)</td>
<td>2,151 (32.9)</td>
<td>–39.9</td>
<td>111 (31.7)</td>
<td>2,089 (31.9)</td>
<td>–0.5</td>
<td></td>
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</tr>
</tbody>
</table>

Notes: *Includes adjustment for covariate imbalance using inverse probability of treatment weighting and an adjustment for censored cycles using Kaplan–Meier estimates.

*Numbers in header row refer to the number of chemotherapy cycles. By convention, stand. diff. <10% indicates high between-cohort similarity for that variable. Bold text indicates stand. diff. >10%.

*Northeast/midwest/other regions are combined for weighted cohorts.

Abbreviations: G-CSF, granulocyte colony-stimulating factor; stand. diff., standardized difference.

Incidence of FN

FN incidence in chemotherapy cycles 1–6 was statistically equivalent between the weighted treatment cohorts (90% CI of difference within ±6%; Table 2). Adjusted FN incidence in the filgrastim-sndz vs filgrastim-ref cohorts, respectively, was 0.81% vs 0.61% for neutropenia + fever (difference [90% CI] = 0.20 [−0.57 to 1.56]), 1.21% vs 1.33% for neutropenia + infection (difference [90% CI] = −0.12 [−1.17 to 2.28]), and 0.0% vs 0.14% for neutropenia + infection + fever (difference = −0.14; CI was not calculated for this definition because there were no FN events in the filgrastim-sndz cohort).
Table 2 Adjusted incidence of febrile neutropenia by weighted* cycle

<table>
<thead>
<tr>
<th>FN definition, n (%)</th>
<th>Filgrastim-sndz (n=349)*</th>
<th>Filgrastim-ref (n=6,546)*</th>
<th>Difference of percentages (90% CI)*</th>
</tr>
</thead>
<tbody>
<tr>
<td>Neutropenia + fever</td>
<td>3 (0.81)</td>
<td>40 (0.61)</td>
<td>0.20 (−0.57 to 1.56)</td>
</tr>
<tr>
<td>Neutropenia + infection</td>
<td>4 (1.21)</td>
<td>87 (1.33)</td>
<td>−0.12 (−1.17 to 2.28)</td>
</tr>
<tr>
<td>Neutropenia + infection + fever</td>
<td>0 (0.0)</td>
<td>9 (0.14)</td>
<td>−0.14</td>
</tr>
</tbody>
</table>

Notes: *Includes adjustment for covariate imbalance using inverse probability of treatment weighting and an adjustment for censored cycles using Kaplan–Meier estimates. *Numbers in header row refer to the number of chemotherapy cycles. *CIs are based on the bias-adjusted bootstrap method with 10,000 bootstrap replications. *CI is not displayed because the incidence of FN in the weighted filgrastim-sndz cohort is 0.

Abbreviation: FN, febrile neutropenia.

Discussion

In this real-world analysis, the incidence of FN among patients undergoing chemotherapy for nonmyeloid cancers was statistically equivalent between those treated with filgrastim-sndz vs filgrastim-ref up through six chemotherapy cycles. These findings build upon those of our previous first-cycle analysis by demonstrating equivalence over a longer period among the same study group, and bolster existing evidence from randomized controlled trials regarding the noninferiority of biosimilar filgrastim compared with the reference drug.

In the present study, the incidence of FN ranged from 0.0% to 1.2% of cycles in the filgrastim-sndz cohort and from 0.14% to 1.33% of cycles in the filgrastim-ref cohort (depending on the FN definition used), similar to the patient-level FN incidence ranges found in chemotherapy cycle 1 in our previous study (filgrastim-sndz, 0.0%–2.3%; filgrastim-ref, 0.3%–1.7%). Our findings are also comparable to those of the only other real-world, head-to-head comparison of these two drugs to date, a claims analysis conducted by Douglas et al in a similar patient group. In the Douglas et al study, the incidence of FN within the first 14 days after G-CSF administration was 2.0% for filgrastim-sndz and 1.1% for filgrastim-ref, when FN was defined as hospitalization plus diagnoses for both neutropenia and infection. Taken together, these contemporary results suggest a lower FN rate than the 1.3%–3.5% reported in previous claims analyses, which may be partly attributable to the fact that earlier studies defined FN as hospitalization with a diagnosis code for neutropenia. This broader definition has higher sensitivity, but a lower positive predictive value for FN than definitions that also require diagnoses of infection and/or fever, such as the ones used in the present study, as well as in clinical trials. In addition, earlier analyses were conducted prior to a 2006 modification of the American Society of Clinical Oncology guidelines that lowered the FN risk cutoff for G-CSF prophylaxis from 40% to the current 20%.

It is well established that prophylactic administration of G-CSF among patients receiving myelosuppressive chemotherapy reduces the risk of FN, FN-related hospitalization, and death, while allowing higher chemotherapy regimen intensity. Filgrastim prophylaxis is recommended to be administered daily until neutrophil counts return to normal; in clinical trials, this has been shown to require 9–14 injections. However, while the proportion of chemotherapy cycles with G-CSF prophylaxis in our study sample ranged from 45.6% (filgrastim-sndz) to 51.6% (filgrastim-ref), the mean numbers of G-CSF injections administered per cycle for filgrastim-sndz and filgrastim-ref were only 2.4 and 2.2, respectively, similar to the patient-level mean found in our previous study. These findings suggest that many patients receive G-CSF dosing that is inconsistent with existing guidelines, a phenomenon that has been observed in a number of other real-world studies. Unfortunately, such practices may be putting patients at risk, as shorter courses of G-CSF have been associated with increased rates of FN, neutropenia-related hospitalization, and chemotherapy regimen disruptions. Potential reasons for G-CSF underutilization cannot be discerned from the present analysis, but may include factors such as physicians’ lack of familiarity with neutropenia guidelines, patients’ aversion to additional injections, or decreased level of monitoring after chemotherapy administration. In addition, there is evidence that the high cost of filgrastim precludes treatment for some patients. Given the lower cost of filgrastim-sndz compared with filgrastim-ref, increased use of the biosimilar has the potential to improve drug access and guideline compliance. In Europe, for example, not only have significant cost savings been achieved since the 2008 approval of the filgrastim biosimilar Zarzio, but also the use of G-CSF prophylaxis has become substantially more widespread. Recent US analyses have demonstrated that prophylaxis with filgrastim-sndz in routine clinical practice is cost-efficient compared with filgrastim-ref and pegfilgrastim; moreover, use of filgrastim-sndz in place of filgrastim-ref could potentially increase access to curative
cancer care, as the resulting cost savings could be applied to provide expensive curative treatments to more patients on a budget-neutral basis.44

Limitations
This study is subject to several limitations, which were described in detail in our previous report23 and are outlined briefly below. Our findings may not be generalizable to populations that were excluded from the study sample (eg. uninsured patients and those on fee-for-service health care plans). The analysis did not determine causality or severity of the study outcomes, as this information is not available in claims data. In addition, because it cannot be definitively determined from claims data whether G-CSF was given as treatment or prophylaxis, we did not attempt to exclude patients with prophylaxis from the study; however, prophylaxis was included as a propensity scoring covariate to avoid confounding the equivalence analysis. The number of patients using filgrastim-sndz was low because the drug was relatively new to the US market during our patient identification period; however, the effect of small sample size on study outcome was mitigated by adjustment with IPTW. Finally, possible misidentification of patients with the outcome of interest due to reliance on diagnosis codes from medical claims is an inherent limitation of claims analyses, and exists for this study also.

Conclusion
The findings of this study not only demonstrate statistical equivalence between filgrastim-sndz vs filgrastim-ref for preventing FN among patients with nonmyeloid cancer undergoing chemotherapy, but also suggest that G-CSF may be underutilized in this population. With burgeoning evidence for its safety and efficacy in clinical practice, biosimilar filgrastim has the potential to improve patient access to G-CSF treatment and increase physician compliance with existing prophylaxis guidelines.

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Disclosure
SB and KC are employees of Sandoz Inc., which is the manufacturer of the filgrastim biosimilars Zarzio® and Zarxio®. LSL, LB, CE, and SK are employees of Optum, which was contracted by Sandoz Inc. to conduct this study. LSS is a consultant for Sandoz Inc. The authors report no other conflicts of interest in this work.

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