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ORIGINAL RESEARCH

Pretreatment platelet count as a prognostic factor in patients with pancreatic cancer: a systematic review and meta-analysis

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Background: The relationship between platelet counts and pancreatic cancer as a prognostic factor has been reported in many studies. We aimed to evaluate the prognostic value of platelet counts in predicting the prognosis of pancreatic cancer patients.

Methods: We searched PubMed, Medline, EMBASE, and Google Scholar for eligible studies up to May 2017. Information about the characteristics of the study and relevant outcomes was extracted. A meta-analysis was performed to analyze the prognostic value of platelet counts using the hazard ratio (HR) and 95% confidence intervals (CIs).

Results: A total of 1,756 patients in 13 retrospective studies were included. The pooled HR of 1.51 (95% CI: 1.20–1.90, P<0.001) showed that patients with elevated platelet counts were expected to have poor overall survival after treatment. Subgroup analysis showed that prognostic value of platelet levels was stronger in patients who received surgical resection (HR =1.60, 95% CI: 1.09–2.34, P=0.02), followed by patients who received palliative therapy (HR =1.46, 95% CI: 1.03–2.06, P=0.03).

Conclusion: Platelet counts could be a useful prognostic marker for pancreatic cancer. Patients with high platelet counts are expected to have poor survival.

Keywords: pancreatic adenocarcinoma, blood parameters, platelet, prognosis, evidence-based medicine

Introduction

Thrombocytosis and coagulopathy are associated with the prognosis of patients with digestive malignant tumors, including pancreatic cancer.¹ Platelets are the smallest circulating blood cells that are involved in multiple links of the process of thrombus formation and development.^{2,3} Studies revealed that activated platelets can participate in tumor progression via multiple mechanisms.⁴ Platelets themselves can promote primary cancer growth by releasing different types of growth factors, chemokines, proangiogenic regulatory proteins, and proteolytic enzymes that promote tumor cell growth and invasion;⁵ platelets can also enhance the dissemination and implantation of cancer by assisting in the formation of platelet–tumor cell heteroaggregates.⁶

The platelet count is a widely used marker and can be easily acquired from complete blood cell parameters. Prognostic factors are essential for the stratification of cancer risk, medical treatment, and clinical research. Although emerging evidence supports a role for platelets in pancreatic cancer, the value of platelet counts in predicting pancreatic cancer prognosis is still controversial, and the association between platelet counts and survival in pancreatic cancer patients has not been previously reviewed. Hence, we aimed to conduct

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a systematic review and meta-analysis to reveal the value of platelet counts in predicting pancreatic cancer prognosis.

Methods

Literature search and filtration strategy

A literature search was performed in May 2017 without restriction to regions and publication types. We systematically searched PubMed, Medline, EMBASE, and Google Scholar to retrieve possible articles relevant to the topic of interest. The following MeSH terms or text words and their combinations were searched to retrieve potentially eligible studies: "blood", "hematological", "hemogram", "platelet", "pancreas", "pancreatic", "cancer", "carcinoma", "adenocarcinoma", "tumor", "neoplasm", and "malignancy". The initial selection was performed to exclude obviously irrelevant articles and to retain potentially relevant articles about platelet or pancreatic cancer prognostic risk factors by reviewing the title and abstract by two independent investigators (SC and NN). Thereafter, the full text of a study was further reviewed for eligibility if either reviewer found that the articles' title and abstract met the following criteria: 1) prospective or retrospective studies that researched the prognostic role of platelets in pancreatic cancer patients; and 2) the hazard ratios (HRs) of overall survival (OS), cancer-specific survival (CCS), or progression-free survival, along with their 95% confidence intervals (CIs) or P-values were available. When more than one study described a same population, the most recent or complete study was involved. Studies with the following criteria were excluded: basic research, duplicated literature, overlapping patients, studies without enough available data to obtain the HR values of outcomes, case reports, and letters. Meeting abstracts were not included due to lack of sufficient detail regarding methodology data and low quality.

Data abstraction

Two investigators (SC and NN) independently reviewed each eligible study using a standardized form. Information about the characteristics of the study population, details of platelet values, and relevant outcomes was recorded. For studies showing only survival curves, the study authors were contacted to obtain the original data; otherwise, the HR values were estimated by the methods described by Tierney et al.⁷ Disagreements between the investigators were resolved through discussion. When necessary, the third investigator (ZJ) helped to reach a consensus.

Quality assessments

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Currently, there is no standard quality assessment tool for prognostic studies in systematic reviews and meta-analyses.

The quality assessment of included studies was evaluated using the "Newcastle–Ottawa Scale" for cohort studies, which consisted of three major events with eight items, for a total score of 0–9; studies achieving scores of \geq 6 were deemed as being of good quality.³⁶ The quality of each involved study was evaluated by two reviewers (SC and NN) independently. Disagreements between the investigators were resolved through discussion. When necessary, the third investigator (ZJ) helped reach a consensus.

Statistical analysis

All meta-analyses were performed using Review Manager 5.3 (The Cochrane Collaboration). We assessed heterogeneity among individual studies by the Cochran Q test and I^2 analysis. Any pooled estimates with $I^2 \ge 50\%$ were considered to have considerable heterogeneity. Pooled HRs and their 95% CIs were calculated using a random-effects model if the heterogeneity was considerable; otherwise, a fixed-effects model was performed. A *P*-value of <0.05 was considered as a statistically significant difference. Additionally, a funnel plot was used to estimate potential publication bias.

Results

Selection and characteristics of included studies

As shown in the flow chart of the literature search (Figure 1), the search strategy generated 221 studies, of which 65 studies were considered of potential value, and the full texts were retrieved for detailed evaluation. In further screening, studies were excluded for the following reasons: lack of sufficient data for estimating the HR, review articles, meeting abstracts without detailed data, letters, and comments. Finally, 13 studies published between 2005 and 2017 were included in the final meta-analysis.8-20 The characteristics of the included studies are shown in Table 1. The sample size of each study ranged from 46 to 211, and a total of 1,756 participants were involved. The studies came from the People's Republic of China (n=5), the USA (n=3), Japan (n=2), the UK (n=1), and Australia (n=1). All studies reported OS. Most of studies chose 300×10⁹/L as the cutoff value between high and low platelet levels. Nine of the 13 studies used multivariate analysis. The score for the quality assessment of included studies ranged from 6 to 7.

Main findings

All studies reported the effects of pretreatment platelet values on pancreatic cancer prognosis. The synthesized HR value

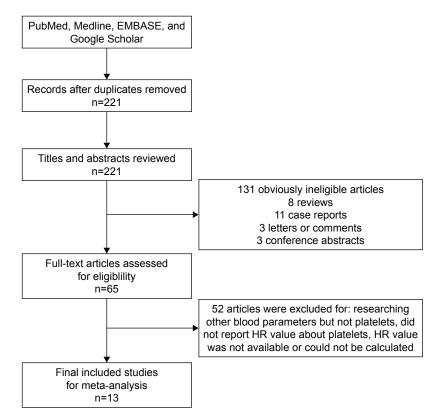


Figure I Literature screening flowchart. **Abbreviation:** HR, hazard ratio.

showed the significant superiority of a low platelet count on OS (pooled HR, 1.51, 95% CI: 1.20–1.90, P<0.001, I^2 =79%) (Figure 2). Six of the included studies defined a high platelet count as >300×10⁹/L;^{9,11,12,15,19,20} a pooled analysis of these six studies revealed that patients with pretreatment platelet

counts >300×10⁹/L might have poorer OS than those with platelet counts <300×10⁹/L (HR =1.62, 95% CI: 1.22–2.16, P<0.001). In addition, the synthesized HR for studies that used multivariate analysis was 1.78 (95% CI: 1.39–2.30, P<0.001).^{8–13,15,16,20}

Table I Characteristics of all identified studies

Study	Year	Country	Sample size	Treatment	Sample of distant metastasis	Platelet cutoff (10º/L)	Outcome measured	HR estimation	NOS score
Miyamoto et al ⁸	2017	Japan	95	Surgical resection	None	278	OS	Multivariate	7
Wang et al ⁹	2015	China	163	Chemotherapy and radiation therapy	None	300	OS	Multivariate	7
Qi et al ¹⁰	2015	China	211	Chemotherapy	134, 37.0%	170	OS	Multivariate	7
Liu et al''	2015	China	168	Radiation therapy or radio-chemotherapy	None	300	OS	Multivariate	7
Chadha et al ¹²	2015	USA	199	Chemoradiation	None	300	OS	Multivariate	7
Asaoka et al ¹³	2016	Japan	46	Surgical resection	None	235	OS	Multivariate	6
Martin et al ¹⁴	2014	Australia	124	Chemotherapy or radio-chemotherapy	84, 67.7%	400	OS	Univariate	6
Wang et al ¹⁵	2014	China	125	Mixed	73, 58.4%	300	OS	Multivariate	7
Miura et al ¹⁶	2014	Japan	50	Surgical resection	2, 4.0%	150	OS	Multivariate	7
Wang et al ¹⁷	2012	China	177	Mixed	32, 18.1%	400	OS	Univariate	6
Bhatti et al ¹⁸	2010	UK	84	Surgical resection	None	Continuous	OS	Univariate	6
Dominguez et al ¹⁹	2008	USA	205	Surgical resection	None	300	OS	Survival curve	6
Brown et al ²⁰	2005	USA	109	Surgical resection	None	300	OS	Multivariate	7

Abbreviations: HR, hazard ratio; NOS, Newcastle–Ottawa Scale; OS, overall survival.

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Study or subgroup Log[HR] SE		SE	Weight (%)	HR IV, random, 95% Cl	Year	HR IV, random, 95% Cl
Brown et al ²⁰	0.7315	0.2718	7.5	2.08 (1.22, 3.54)	2005	
Dominguez et al19	0.063	0.1634	10.1	1.07 (0.77, 1.47)	2008	
Bhatti et al18	0	0.0026	12.5	1.00 (0.99, 1.01)	2010	+
Wang et al ¹⁷	0.6831	0.3889	5.3	1.98 (0.92, 4.24)	2012	
Miura et al16	2.0059	0.7025	2.3	7.43 (1.88, 29.45)	2014	
Wang et al ¹⁵	0.2677	0.3035	6.8	1.31 (0.72, 2.37)	2014	_ _
Martin et al14	-0.0943	0.2387	8.3	0.91 (0.57, 1.45)	2014	_ _
Qi et al ¹⁰	0.0917	0.1638	10.1	1.10 (0.80, 1.51)	2015	_ _
Wang et al ⁹	0.392	0.2538	7.9	1.48 (0.90, 2.43)	2015	_ _
Chadha et al12	0.7793	0.2223	8.7	2.18 (1.41, 3.37)	2015	
Liu et al11	0.802	0.2753	7.4	2.23 (1.30, 3.83)	2015	
Asaoka et al13	0.7227	0.4638	4.2	2.06 (0.83, 5.11)	2016	
Miyamoto et al ⁸	0.7006	0.2129	8.9	2.01 (1.33, 3.06)	2017	
Total (95% CI)			100	1.51 (1.20, 1.90)		▲
Heterogeneity: $\tau^2=0$.	11; γ^2 =56.26,	df=12 (P<0.	00001); / ² =79	,		
Test for overall effect: Z=3.52 (P=0.0004)						0.02 0.1 1 10 50
						Favors high platelet Favors low platelet

Figure 2 Forest plot and meta-analysis for the relationship between pretreatment platelet counts and OS in patients with pancreatic cancer.

Note: The results are presented as individual and combined HRs and 95% Cl.

Abbreviations: CI, confidence interval; HRs, hazard ratios; OS, overall survival; IV, inverse variance.

Subgroup analysis

Figure 3 shows the subgroup analysis based on stages of pancreatic cancer. Studies were divided into three subgroups according to the patient characteristics they enrolled: 1) studies

enrolled only patients with resectable lesions;^{8,13,16,18–20} 2) studies enrolled only patients with unresectable lesions who were locally advanced and/or metastatic and subsequently received palliative therapy (chemotherapy, radiotherapy,

Study or subgroup	Log[HR]	SE	Weight (%)	HR IV, random, 95% Cl	Year	HR IV, 1 95% CI	random,	
Surgical resection								
Brown et al ²⁰	0.7315	0.2718	7.5	2.08 (1.22, 3.54)	2005			
Dominguez et al19	0.063	0.1634	10.1	1.07 (0.77, 1.47)	2008	-	-	
Bhatti et al ¹⁸	0	0.0026	12.5	1.00 (0.99, 1.01)	2010		•	
Miura et al ¹⁶	2.0059	0.7025	2.3	7.43 (1.88, 29.45)	2014			
Asaoka et al ¹³	0.7227	0.4638	4.2	2.06 (0.83, 5.11)	2016	-	· · · ·	
Miyamoto et al ⁸	0.7006	0.2129	8.9	2.01 (1.33, 3.06)	2017			
Subtotal (95% CI)			45.5	1.60 (1.09, 2.34)			•	
Heterogeneity: $r^2 = 0.1$	5; $\gamma^2 = 28.79$, di	f=5 (P<0.0	001); <i>P</i> =83					
Test for overall effect:	Z=2.41 (P=0.0	02)						
Palliative therapy								
Martin et al ¹⁴	-0.0943	0.2387	8.3	0.91 (0.57, 1.45)	2014			
Qi et al ¹⁰	0.0917	0.1638	10.1	1.10 (0.80, 1.51)	2015	-	•	
Chadha et al12	0.7793	0.2223	8.7	2.18 (1.41, 3.37)	2015		_ _	
Liu et al11	0.802	0.2753	7.4	2.23 (1.30, 3.83)	2015			
Wang et al ⁹	0.392	0.2538	7.9	1.48 (0.90, 2.43)	2015			
Subtotal (95% CI)			42.3	1.46 (1.03, 2.06)			◆	
Heterogeneity: $\tau^2=0.1$			1); <i>l</i> ²=68%					
Test for overall effect:	Z=2.12 (P=0.0	03)						
Mixed								
Wang et al ¹⁷	0.6831	0.3889	5.3	1.98 (0.92, 4.24)	2012		•	
Wang et al ¹⁵	0.2677	0.3035	6.8	1.31 (0.72, 2.37)	2014		•	
Subtotal (95% CI)			12.1	1.53 (0.96, 2.44)			◆	
Heterogeneity: $\tau^2 = 0.0$); <i>I</i> ²=0%					
Test for overall effect:	Z=1.78 (P=0.0	08)						
Total (95% CI)			100	1.51 (1.20, 1.90)			•	
Heterogeneity: $\tau^2=0.1$	1; χ^2 =56.26, d	f=12 (<i>P</i> <0.	00001); <i>l</i> ²=	=79%	H			
Test for overall effect:			,,		0.02	0.1	1 10	50
Test for subgroup diffe	erences: $\chi^2 = 0.7$	13, df=2 (F	P=0.94); /2=	:0%	Eave	ors high platelet	Favors low plate	lot
U		`			Favo	ns nign platelet	ravors low plate	Het

Figure 3 Subgroup analysis of the HRs of OS by types of cancer population and treatment.

Notes: Surgical resection: studies involved only cases with resectable lesions who received surgery; palliative therapy: studies involved unresectable cases who received chemotherapy and/or radiotherapy only; mixed: studies involved both resectable cases and unresectable cases.

Abbreviations: CI, confidence interval; HRs, hazard ratios; OS, overall survival; SE, standard error; IV, inverse variance.

Study or subgroup	Log[HR]	SE	Weight (%)	HR IV, random, 95% Cl	Year	HR IV, 95% CI	random,	
Surgical resection								
Brown et al ²⁰	0.7315	0.2718	11.6	2.08 (1.22, 3.54)	2005			
Miura et al16	2.0059	0.7025	2.9	7.43 (1.88, 29.45)	2014			
Asaoka et al13	0.7227	0.4638	5.8	2.06 (0.83, 5.11)	2016	-		
Miyamoto et al8	0.7006	0.2129	14.4	2.01 (1.33, 3.06)	2017			
Subtotal (95% CI)			34.8	2.19 (1.59, 3.02)			•	
Heterogeneity: $\tau^2=0.0$	01; χ ² =3.23, di	f=3 (P=0.36)	; /2=7%					
Test for overall effect	: Z=4.76 (P<0	.00001)						
Palliative therapy								
Chadha et al12	0.7793	0.2223	13.9	2.18 (1.41, 3.37)	2015			
Liu et al11	0.802	0.2753	11.4	2.23 (1.30, 3.83)	2015			
Qi et al ¹⁰	0.0917	0.1638	17.2	1.10 (0.80, 1.51)	2015	-	-	
Wang et al ⁹	0.392	0.2538	12.4	1.48 (0.90, 2.43)	2015			
Subtotal (95% CI)			55.0	1.63 (1.12, 2.36)			•	
Heterogeneity: $\tau^2=0.$. ,,	l²=65%					
Test for overall effect	:: Z=2.57 (P=0	.01)						
Mixed								
Wang et al ¹⁵	0.2677	0.3035	10.3	1.31 (0.72, 2.37)	2014	_	•	
Subtotal (95% CI)			10.3	1.31 (0.72, 2.37)		-		
Heterogeneity: not a								
Test for overall effect	t: Z=0.88 (P=0	.38)						
Total (95% CI)			100	1.78 (1.39, 2.30)			•	
Heterogeneity: r ² =0.	07; χ²=16.18, α	df=8 (P=0.04); <i>I</i> ²=51%	-				
Test for overall effect: Z=4.50 (P<0.00001)					0.02	0.1	1 10	50
Test for subgroup dif	ference: $\chi^2 = 2.8$	80, <i>df</i> =2 (<i>P</i> =	0.25); /²=28.5	Fa	vors high platelet	Favors low plate	let	

Figure 4 Subgroup analysis by types of cancer population and treatment pooling studies used multivariate analysis.

Notes: Surgical resection: studies involved only cases with resectable lesions that received surgery; palliative therapy: studies involved unresectable cases that received chemotherapy and/or radiotherapy only; mixed: studies involved both resectable cases and unresectable cases. The results are presented as individual and combined HRs for OS and the 95% Cl.

Abbreviations: CI, confidence interval; HRs, hazard ratios; OS, overall survival; SE, standard error; IV, inverse variance.

or chemoradiotherapy);^{9–12,14} and 3) studies included both resectable cases and unresectable cases.^{15,17} The results showed that the prognostic value of platelet count was best for resectable cases (HR =1.60, 95% CI: 1.09–2.34, *P*=0.02), followed by unresectable patients (HR =1.46, 95% CI: 1.03–2.06, *P*=0.03).

The same subgroup analysis was also performed in studies that used multivariate analysis (Figure 4), and we found that the HR from four studies that only included resectable cases was 2.19 (95% CI: 1.59–3.02, P<0.001, I^2 =7%); the low I^2 value revealed that only minor heterogeneity existed across the studies in this subgroup. The HR synthesized from data from studies only involving advanced cases was 1.63 (95% CI: 1.12–2.36, P=0.01, I^2 =65%).

Publication bias

A funnel plot (Figure 5) was constructed, and it revealed a potential publication bias in this meta-analysis.

Discussion

It is now widely regarded that outcomes in cancer patients are dependent on not only the tumor characteristics but also several host-related factors.²¹ Platelets serve various roles in physiological pathways, including coagulation and inflammation.^{22,23} Abnormal platelet levels have been found in various types of solid tumors, and some studies have suggested that elevated platelet levels are correlated with disease progression, metastasis, and poor patient prognosis.^{24,25} However, although pancreatic cancer is traditionally associated with hypercoagulable states, the prognostic significance of pretreatment platelet counts in predicting pancreatic

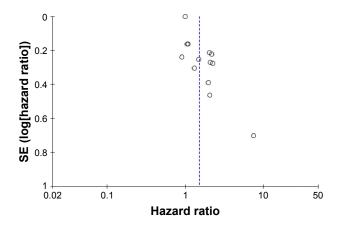


Figure 5 Funnel plots based on the value of platelets in predicting OS in patients with pancreatic cancer. Abbreviations: OS, overall survival; SE, standard error.

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cancer outcome remains controversial. Based on our metaanalysis, we found evidence that an elevated treatment platelet level is associated with poor outcome in patients with pancreatic cancer.

The mechanisms by which platelets contribute to cancer progression remain not completely answered. The tumor itself may stimulate thrombocytopoiesis by secreting cytokines.²⁶ Tumor cells can also induce platelet activation through a variety of pathways, such as enhancing the expression of surface integrins and selectins and the secretion of granules, which amplify platelet aggregation.^{27–29} Activated platelets, in turn, promote primary cancer cell growth and invasion by releasing a large number of growth factors, including angiogenic cytokines, mitogenic proteins, and growth factors.⁶ Platelets can adhere to tumor cells through adhesion receptors such as GPIIb/IIIa and P-selectin.^{30,31} The formation of heteroaggregates between tumor cells and platelets is enhanced by this type of platelet-tumor cell interaction.^{27,32,33} According to our present understanding, tumor cells in circulating blood stream face physical damage and immunological cytotoxicity that can potentially inhibit tumor metastasis and implantation, but heteroaggregates protect cancer cells from these potential damages.34,35

The majority of the included studies enrolled only patients with resectable pancreatic cancer, and the remaining studies included locally advanced and/or metastatic lesions that then received adjuvant therapy only. In addition to confirming the association between platelet counts and the prognosis of pancreatic cancer patients, our subgroup analysis revealed that the prognostic efficacy of platelet counts seems to be stronger in cases with resectable tumors than in those who could only receive adjuvant therapy. Studies have suggested that aggressive pancreatic cancer cells may disseminate to the bone marrow, and platelet counts may decrease in advanced pancreatic cancer patients; additionally, locally advanced cancer cases are usually accompanied by thrombocytopenia caused by splenic vein occlusion.9 Therefore, pretreatment platelet levels might be affected by many more factors in end-stage pancreatic cancer patients than in resectable cases. However, the real reason why platelet counts show better prognostic value in resectable cases remains unanswered.

There are several limitations in our study. First, all of the included studies were retrospective. In observational studies, selection bias is impossible to avoid, although a multivariate analysis can control the confounding factors to a certain extent. Second, the platelet counts could be affected by multiple factors, especially undetected diseases, such as chronic infection, hematological disease, and peripheral arterial disease. Most of the included studies did not report CCS, which is an essential outcome for cancer patients. Third, we noted that all studies used the dichotomous platelet counts to determine the prognostic value and most of them chose the normal hemogram platelet values as the cutoff value. Thus, we recommend using a continuous platelet counts variable rather than a categorical variable in future studies.

Conclusion

This meta-analysis is the first to evaluate the platelet count as a prognostic marker for patients with pancreatic cancer, and the pooled results suggest that the pretreatment platelet count is closely correlated with the prognosis of patients with pancreatic cancer. However, due to the limitations of the included studies, more multi-center prospective cohorts need to be conducted to validate the role and exact cutoff value for platelet counts in predicting the prognosis of pancreatic cancer.

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Disclosure

The authors report no conflicts of interest in this work.

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