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ORIGINAL RESEARCH

Prognostic value of high IMP3 expression in solid tumors: a meta-analysis

Luyao Chen^{1,2,*} Yongpeng Xie^{3,*} Xintao Li¹ Liangyou Gu¹ Yu Gao¹ Lu Tang¹ Jianwen Chen¹ Xu Zhang¹

¹Department of Urology, Chinese PLA General Hospital, Beijing, ²Department of Urology, First Affiliated Hospital of Nanchang University, Nanchang; ³School of Medicine, Nankai University, Tianjin, People's Republic of China

*These authors contributed equally to this work

Correspondence: Xu Zhang Department of Urology, Chinese PLA General Hospital, No 28 Fuxing Road, Haidian, Beijing 100853, People's Republic of China Tel +86 10 6693 8008 Fax +86 10 6822 3575 Email xzhang@tjh.tjmu.edu.cn



Background: Accumulated studies have investigated the prognostic role of insulin-like growth factor II mRNA-binding protein 3 (IMP3) in various cancers, but inconsistent and controversial results were obtained. Therefore, we performed a systematic review and meta-analysis to investigate the potential value of IMP3 in the prognostic prediction of human solid tumors.

Materials and methods: A systematic literature search in the electronic databases PubMed, Embase, Web of Science, and Cochrane library (updated to April 2016) was conducted to identify eligible studies. Pooled hazard ratios (HRs) with 95% confidence intervals (CIs) for survival outcomes were calculated and gathered using STATA 12.0 software.

Results: A total of 53 studies containing 8,937 patients with solid tumors were included in this meta-analysis. High IMP3 expression was significantly associated with worse overall survival (OS) of solid tumors (HR =2.08, 95% CI: 1.80–2.42, P<0.001). Similar results were observed in cancer-specific survival (CSS), disease-free survival (DFS), recurrence-free survival (RFS), progression-free survival (PFS), and metastasis-free survival (MFS). Further subgroup analysis stratified by tumor type showed that elevated IMP3 expression was associated with poor OS in renal cell carcinoma (RCC), lung cancer, oral cancer, urothelial carcinoma, hepatocellular carcinoma (HCC), colorectal cancer, pancreatic cancer, gastric cancer, and intrahepatic cholangiocarcinoma (ICC).

Conclusion: The current evidence suggests that high IMP3 expression is associated with poor prognosis in most solid tumors. IMP3 is a potential valuable prognostic factor and might serve as a promising biomarker to guide clinical decisions in human solid tumors.

Keywords: IMP3, prognosis, solid tumor, biomarker, meta-analysis

Introduction

Insulin-like growth factor II mRNA-binding protein 3 (IMP3 or IGF2BP3) is a member of the RNA-binding protein family, which plays an important role in RNA trafficking and stabilization, cell growth, and cell migration during the early stages of embryogenesis.¹ IMP3 was proposed to control the translation or turnover of various candidate target genes, including IGF2, CD44, HMGA2, and MMP9.^{2–5} This oncofetal protein has been reported to promote tumor cell survival, proliferation, chemoresis-tance, and tumor cell invasiveness in vitro. In recent years, accumulating studies have shown that IMP3 is specifically expressed in malignant tumors and acts as an important cancer-specific gene involved in many aggressive and advanced cancers.^{6,7}

Numerous studies have reported that upregulated IMP3 expression in tumor tissues is correlated with poor patient survival and can be used as a prognostic factor to guide clinical decisions and distinguish different prognoses in various solid tumors, such as renal cell carcinoma (RCC), lung cancer, oral cancer, bladder cancer, gastrointestinal tumors, and gynecological tumors.^{8–13} However, some other studies have reported

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Commercial use of this work, is published and licensed by Dove Medical Press Limited. The full terms of this license are available at https://www.dovepress.com/terms.php and incorporate the Creative Commons Attribution — Non Commercial (unported, v3.0) License (http://creativecommons.org/license2/b-nd/3.0/). By accessing the work you hereby accept the Terms. Non-commercial uses of the work are permitted without any further permission from Dove Medical Press Limited, provided the work is properly attributed. For permission for commercial use of this work, please see paragraphs 4.2 and 5 of our Terms (https://www.dovepress.com/terms.php). the absence of association between IMP3 expression and cancer prognosis.^{14,15} Some investigators have also replayed completely opposite results in ovarian cancer. For instance, Kobel et al¹⁶ proposed that IMP3 expression is a marker of unfavorable prognosis, whereas Noske et al¹⁷ asserted that IMP3 expression is associated with improved survival. Hence, the prognostic role of IMP3 expression in solid tumors remains unclear and controversial.

Therefore, we conducted a systematic review of published studies, with a standard meta-analysis combining available evidence, to evaluate the prognostic value of IMP3 expression in solid tumors.

Materials and methods

This meta-analysis was conducted according to the guideline of Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA)¹⁸ (Table S1). Because the data included in this study were retrieved from published articles, ethical approval from ethics committees was not needed.

Literature search

A comprehensive literature search was performed in PubMed, Embase, Web of Science, and Cochrane Library to identify studies evaluating IMP3 expression and clinical prognosis in solid tumors up to April 2016. The search strategy included the following terms through MeSH headings, keywords, and text words: "IMP3" or "Insulin-like growth factor 2 mRNA binding protein 3" or "IGF2BP3" combined with "cancer" or "carcinoma" or "neoplasm". The references cited in the identified articles were also screened for possible inclusions. The database search and preliminary evaluation of identified studies were performed independently by two investigators (LC and YX). No language limitation existed in the process.

Study selection

The inclusion criteria for selecting articles in our analysis are listed as follows: 1) studies that reported IMP3 expression in cancer tissues, 2) studies analyzing the relationship between IMP3 expression level and clinical cancer outcomes, 3) studies that directly reported survival outcomes with hazard ratio (HR) and corresponding 95% confidence interval (CI) or studies that provided sufficient data for estimating HR and 95% CI by using the methods described by Tierney et al,¹⁹ and 4) studies with a median follow-up of at least 6 months. Studies were excluded if they were 1) case reports, letters, conference abstracts, or reviews, 2) non-human research, 3) investigations on the diagnostic role, but not the

prognostic role, of IMP3, and 4) studies with insufficient data for calculating the HR and 95% CI. If duplicate publications by the same authors were retrieved, we included only the most informative and recent study. Two independent reviewers (LC and YX) evaluated the full articles for study eligibility, and any disagreement was resolved by consensus.

Data extraction and quality assessment

Two authors (LC and YX) independently extracted data from each eligible study by using predefined item forms. The following information, if available, was recorded: first author's name, year of publication, study country or region, type of cancer, cancer stage, number of patients, detected method, cutoff definition, percentage of high IMP3 expression, follow-up period, and survival outcomes with their HRs and corresponding 95% CIs. If univariate and multivariate analyses were reported to obtain the HRs, the results of multivariate analysis were preferentially selected. If HRs and 95% CIs were not provided directly, we attempted to estimate these points with Kaplan-Meier curve or other required data in the original study by using Tierney et al's methods.¹⁹ Study quality was scored by two investigators (LC and YX) using the Newcastle–Ottawa Scale, which involves three main categories: selection, comparability, and outcome ascertainment. We defined studies with scores no less than 6 as gualified to be included in the meta-analysis. Discrepancies between investigators were resolved through discussion.

Statistical analysis

Pooled HRs and corresponding 95% CIs were calculated to evaluate the prognostic role of high IMP3 expression in the clinical outcomes of solid tumors. An observed HR greater than 1 implied a worse prognosis in patients with high IMP3 expression, and an HR less than 1 indicated a better prognosis. Statistical heterogeneity of combined HR was assessed using Cochrane Q-test and Higgins I^2 metrics. $I^2 > 50\%$ was considered a measure of obvious heterogeneity.²⁰ If no evident heterogeneity existed, the fixed-effect model (Mantel-Haenszel method) was used to pool the results.²¹ Otherwise, the randomeffect model (DerSimonian and Laird method) was selected.22 The potential sources for heterogeneity, if significant, were further explored using a predefined subgroup analysis and meta-regression analysis (based on cancer type, ethnicity, case number, cutoff, cancer stage, HR obtained method, and analysis method). To assess the stability of the pooled results, sensitivity analysis was performed by sequential omission of each single study. Publication bias was also estimated by

Prognostic value of high IMP3 expression in solid tumors

visually assessing the asymmetry of the funnel plot and then quantitatively evaluated by Begg's and Egger's tests.^{23,24} All the abovementioned analyses were performed using STATA version 12.0 (Stata Corporation, College Station, TX, USA). All statistical tests were two sided, and statistical significance was defined as a *P*-value less than 0.05.

Results

Search results and study characteristics

The flowchart of the literature search is shown in Figure 1. A total of 420 potentially relevant studies were retrieved from the initial literature search in the aforementioned electronic databases. A total of 144 duplicated records were excluded by a literature manager software. After carefully screening titles and abstracts of the remaining 120 records, 46 studies were excluded and 74 studies were selected for full-text assessment. Given the inclusion and exclusion criteria, 21 studies that belonged to duplicate publication or failed to offer sufficient prognostic information were excluded. Finally, 53 studies satisfied our eligibility criteria and were included in this meta-analysis.

The characteristics of these enrolled studies are summarized in Table 1. The 53 studies involved 8,937 patients with different cancer types, including 6 studies of RCC, 8,25-29 6 lung cancer,^{9,30-34} 4 oral cancer,^{10,35-37} 4 urothelial carcinoma,³⁸⁻⁴¹ 4 ovarian cancer,^{16,17,42,43} 3 hepatocellular carcinoma (HCC),⁴⁴⁻⁴⁶ 4 colorectal cancer,^{12,47-49} 3 prostate cancer,^{14,15,50} 3 pancreatic cancer,^{51–53} 2 gastric cancer,^{11,54} 2 intrahepatic cholangiocarcinoma (ICC),^{55,56} and one study each of tongue cancer,57 thyroid carcinoma,58 sacral chordoma,59 pilocytic astrocytoma and pilomyxoid astrocytoma (PA/ PMA),60 neuroblastoma,61 meningioma,62 melanoma,63 breast cancer,64 giant cell tumor,65 bile duct carcinoma,66 esophageal carcinoma.⁶⁷ and cervical cancer.¹³ A total of 25 studies involved Caucasians and 28 involved Asians. The survival outcomes in these studies, including overall survival (OS), cancer-specific survival (CSS), disease-free survival (DFS), recurrence-free survival (RFS), progression-free survival (PFS), and metastasis-free survival (MFS), were investigated in 40, 10, 8, 7, 4, and 5 studies, respectively. HRs were reported directly in most of these studies (43/53) and were estimated indirectly in the 10 other studies. Multivariate Cox

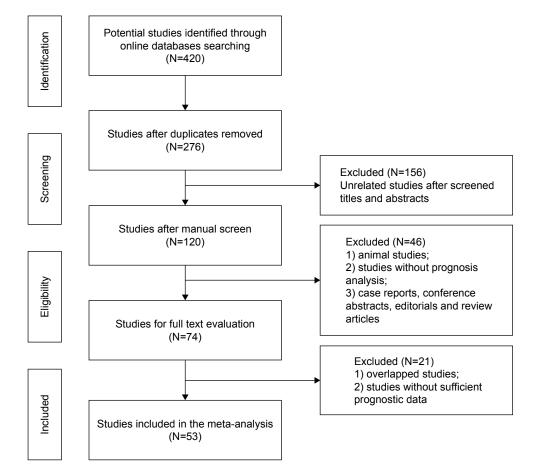


Figure I Flowchart of the study selection process.

Author	Year	Country or region	Cancer type	Case number	Method	Cutoff	High expression	Follow-up	Outcomes	Analysis	HR obtained	NOS
liang et a ^{l8}	2006	USA	RCC	371	HC	Positive vs negative*	71 (19.1%)	Median 63 months	OS MES	Multi	Report	6
Pei et al ²⁶	2015	NSA	RCC	346	HC	Positive vs negative	73 (21.1%)	>10 years	OS RFS	Multi	Report	œ
Hoffmann et al ²⁵	2008	NSA	RCC	716	HC	Positive vs negative	213 (29.7%)	9.5 years	CSS MFS	Multi	Report	œ
Park et al ²⁷	2014	Korea	RCC	148	HC	>5% of cells stained	43 (29.1%)	Median 55.5 months	CSS	Multi	Report	7
Jiang et al ²⁸	2008	NSA	RCC	317	HC	Positive vs negative	40 (12.6%)	8.8 years	OS MFS	Multi	Report	6
Tantravahi et al ²⁹	2015	NSA	RCC	27	НС	>20% of cells stained	14 (51.9%)	>2 years	SO	Multi	Report	9
Del Gobbo et al ³⁴	2014	ltaly	Lung cancer	74	HC	Positive vs negative	24 (32.4%)	Mean 65.6 months	OS DFS	Uni	Report	7
Sun et al ³²	2015	China	Lung cancer	196	IHC	H-score >100 (0-300)	83 (42.3%)	Range (16.5–69.0) months	OS DFS	Multi	Report	œ
Yan et al ⁹	2016	China	Lung cancer	95	НС	>25% of cells stained	39 (41.1%)	>5 years	SO	Multi	Report	7
Zhang et al ³³	2015	China	Lung cancer	186	IHC	>5% of cells stained	139 (74.7%)	>5 years	SO	Multi	Report	8
Lin et al ³⁰	2015	China	Lung cancer	92	HC	Positive vs negative	62 (67.4%)	>5 years	SO	Multi	Report	8
Beljan Perak et al ³¹	2012	Croatia	Lung cancer	60	HC	>10% of cells stained	61 (67.8%)	>5 years	SO	Uni	SC	9
Clauditz et al ³⁵	2013	Germany	Oral cancer	145	НС	>10% of cells stained	79 (54.5%)	Mean 41.3 months	SO	Multi	Report	8
Lin et al ³⁷	2011	Taiwan	Oral cancer	93	IHC	>25% of cells stained	51 (54.8%)	Mean 44.8 months	SO	Multi	Report	6
Li et al ³⁶	2010	Korea	Oral cancer	96	IHC	Positive vs negative	65 (67.7%)	Median 73 months	SO	Multi	Report	6
Kim and Cha ¹⁰	2011	Korea	Oral cancer	95	HC	Positive vs negative	67 (70.5%)	>5 years	SO	Multi	Report	7
Szarvas et al ⁴⁰	2012	Germany	Urothelial carcinoma	106	HC	Staining index >7 (0–9)	17 (16.0%)	Median 15 months	OS CSS MFS	Multi	Report	7
Sitnikova et al ³⁹	2008	NSA	Urothelial carcinoma	214	IHC	Positive vs negative	42 (19.6%)	Median 35 months	PFS DFS	Multi	Report	8
Lee et al ⁴¹	2013	Multicenter	Urothelial carcinoma	622	НC	Positive vs negative	76 (12.2%)	Median 27 months	OS CSS RFS	Multi	Report	6
Niedworok et al ³⁸	2015	Germany	Urothelial carcinoma	26	HC	H-score >100 (0–300)	7 (26.9%)	Median 50 months	OS PFS	Uni	Report	7
Bi et al ⁴³	2016	China	Ovarian cancer	73	НС	>10% of cells stained	46 (63.0%)	>5 years	SO	Uni	SC	7
Kobel et al ⁱ⁶	2009	British and	Ovarian cancer	278	IHC	>5% of cells stained	147 (52.9%)	>4.6 years	CSS	Multi	Report	8
		North America										
Hus et al ⁴²	2015	Taiwan	Ovarian cancer	140	HC	The median value (IRS: 0–9)	NR	Median 39 months	PFS	Multi	Report	9
Noske et al ¹⁷	2009	Germany	Ovarian cancer	68	IHC	IRS >6	32 (47.1%)	Median 37 months	SO	Uni	SC	7
Hu et al ⁴⁴	2014	China	HCC	160	HC	Staining score (2–7 vs 0–1)	97 (60.6%)	Median 36 months	OS RFS	Uni	SC	8
Wachter et al ⁴⁵	2011	Germany	HCC	365	HC	Staining group (2–3 vs 0–1)	67 (18.4%)	Mean 23.3 months	SO	Multi	Report	7
Chen et al ⁴⁶	2013	China	HCC	92	ΗC	Positive vs negative	65 (70.7%)	>3 years	SO	Multi	Report	7
Yuan et al ⁴⁸	2009	Taiwan	Colorectal cancer	186	НС	>50% of cells stained	66 (35.5%)	Median >5 years	SO	Multi	Report	8
Li et al ⁴⁹	2009	China	Colorectal cancer	203	IHC	Staining score (2–7 vs 0–1)	132 (65.0%)	Median 61 months	OS DFS	Multi	Report	6
Lochhead et al ¹²	2012	USA	Colorectal cancer	671	НС	Intense or moderate vs	234 (34.9%)	Median 160 months	OS CSS	Multi	Report	œ
l in at 2 ¹³⁰	2012	China	Coloroctol concor	196			(%) 42/ 27/		Š	M. H.	P on ort	7
lkanhara at al ¹⁵	0100	Switzerland	Prostate cancer	475		Positive vs negative	354 (83 3%)	Madian 63 months	D C C C C C C C C C C C C C C C C C C C		Report	. σ
Chromecki et al ^{i₄}	2011	USA	Prostate cancer	232		> 10% of cells stained	42 (18.1%)	Median 69.8 months	RFS	Multi	Report	6
Szarvas et al ⁵⁰	2014	Germany	Prostate cancer	124	HC	>10% of cells stained	30 (24.2%)	Median 155 months	OS CSS	Uni	Report	~ ~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~
Wang et al ⁵²	2014	China	Pancreatic cancer	50	aPCR	Cutoff value based on the	30 (60.0%)	>2 vears	SO	Multi	Report	-
	-)		2	5	ROC curve	(22.2.2.) 2.2)	5		
Schaeffer et al ⁵¹	2010	Canada	Pancreatic cancer	127	HC	IHC score >5	80 (63.0%)	Mean 13 months	SO	Multi	Report	œ

Gao et al ³⁶	2014	China		77		Positive vs negative	(%61.9%)	Median 14.9 months	So	Multi	Keport	20
Gao et al ³⁶	2014	China		77	ЦЦ	Positive vs negative	59 (81.9%)	Median 14.9 months	SO	Multi	Report	œ
Li et al ⁵⁷	2011	China	Tongue carcinoma	65	НC	Positive vs negative	50 (76.9%)	Median 36 months	CSS	Uni	SC	8
Asioli et al ⁵⁸	2010	USA	Thyroid carcinoma	103	HC	Final score >2 (0–6)	61 (59.2%)	>5 years	OS DFS MFS	Multi	Report	6
Zhou et al ⁵⁹	2014	China	Sacral chordoma	32	HC	Staining score (2–7 vs 0–1)	20 (62.5%)	Median 110 months	DFS	Uni	S	8
Barton et al ⁶⁰	2013	USA	PA/PMA	77	HC	Three groups (1–2 vs 0)	24 (31.2%)	Mean 8.8 years	PFS	Uni	Report	7
Chen et al ⁶¹	2011	Taiwan	Neuroblastoma	06	HC	>10% of cells stained	52 (57.8%)	Median 39.5 months	SO	Multi	Report	8
Hao et al ⁶²	2011	USA	Meningioma	107	HC	Positive vs negative	7 (6.5%)	Median 53 months	OS RFS	Multi	Report	7
Sheen et al ⁶³	2014	Taiwan	Melanoma	97	НС	>10% of cells stained	72 (74.2%)	Median 5.2 years	SO	Multi	Report	7
Walter et al ⁶⁴	2009	NSA	Breast cancer	138	HC	> 10% of cells stained	45 (32.6%)	Median 71.5 months	SO	Multi	Report	7
Zhang et al ³³	2015	China	Giant cell tumor	38	HC	Staining score (3–7 vs 0–2)	13 (34.2%)	Median 88.0 months	RFS	Uni	S	9
Riener et al ⁶⁶	2009	Switzerland	Bile duct carcinoma	115	HC	Intense or moderate vs	67 (58.3%)	Median 9 months	CSS	Multi	Report	œ
						weak or absent						
Takata et al ⁶⁷	2014	2014 Japan	Esophageal carcinoma 191	161	HC	> 10% of cells stained	113 (59.2%)	Mean 41 months	SO	Multi	Report	6
Wei et al ¹³	2014	2014 China	Cervical carcinoma	96	HC	> 10% of cells stained	54 (56.3%)	Median 58.1 months	SO	Multi	Report	8
Note: *Positive vs no Abbreviations: CSS	egative: tu	imor cells with any pecific survival; DF	Note: *Positive vs negative: tumor cells with any detectable staining were considered positive. Abbreviations: CSS, cancer-specific survival: DFS, disease-free survival; HCC, hepatocellular c	nsidered pr , hepatoce	ositive. Ilular carcinor	Note: *Positive vs negative: tumor cells with any detectable staining were considered positive. Abbreviations: CSS, cancer-specific survival: PCS, disease-free survival: HCC, hepatocellular carcinoma; HR, hazard ratio; ICC, intrahepatic cholangiocarcinoma; IHC, immunohistochemistry; IRS, immunoreactivity score; MFS, metastasis-free	ic cholangiocarc	inoma; IHC, immunohistocherr	nistry; IRS, immunc	oreactivity sc	ore; MFS, meta	stasis-free
survival; NOS, Newc	astle-Ott	awa Scale; NR, not	: reported; OS, overall surviva	ıl; PA/PM∕	A, pilocytic ast	survival; NOS, Newcastle-Ottawa Scale; NR, not reported; OS, overall survival; PA/PMA, pilocytic astrocytoma and pilomyxoid astrocytoma; PFS, progression-free survival; qPCR, quantitative polymerase chain reaction; RFS, recurrence-free	ma; PFS, progres	sion-free survival; qPCR, quan	titative polymeras	e chain react	ion; RFS, recur	rence-free

OncoTargets and Therapy 2017:10

SC, survival curve

renal cell carcinoma;

survival; RCC,

analysis was performed to evaluate the prognostic role of IMP3 in 38 studies; and univariate analysis was conducted in the other 15 studies. Immunohistochemistry (IHC) staining and quantitative polymerase chain reaction (qPCR) were used to test the IMP3 expression in cancer tissues. Notably, the definition and cutoff of high IMP3 expression were heterogeneous among these studies. The majority of included studies used the percentage of positive staining cells (0%, 10%, 25%, or 50%) as the criteria, whereas in some other studies, staining scores with the percentage and intensity score were obtained as cutoff values for high IMP3 expression. The percentage of high expression in the cohort population varied in different cancer types and ranged from 6.5% to 83.3%. Quality score assessment suggested that the scores of enrolled studies ranged from 6 to 9, which were considered adequate for quantitative meta-analysis.

Association of IMP3 with OS

The association of IMP3 expression and OS was investigated in 40 studies containing 6,425 patients with different cancer types. A random-effect model was selected because of the evident interstudy heterogeneity ($I^2=59.1\%$, P=0.005). Combined analysis revealed that high IMP3 expression was associated with the worse OS of solid tumors (HR =2.08, 95% CI: 1.80–2.42, P<0.001, Figure 2). The effect of IMP3 expression on OS was further analyzed by tumor types, and the results are presented in Figure 3A. High IMP3 expression was significantly associated with poor OS in RCC (HR =2.80, 95% CI: 1.59-4.93, P<0.001), lung cancer (HR =1.87, 95% CI: 1.22–2.84, P=0.004), oral cancer (HR =1.66, 95% CI: 1.27-2.18, P<0.001), urothelial carcinoma (HR =1.92, 95% CI: 1.42–2.59, P<0.001), HCC (HR =2.25, 95% CI: 1.65-3.06, P<0.001), colorectal cancer (HR =1.52, 95% CI: 1.23-1.90, P<0.001), pancreatic cancer (HR =3.54, 95% CI: 2.06-6.09, P<0.001), gastric cancer (HR =2.67, 95% CI: 1.38-5.17, P=0.003), and ICC (HR =2.10, 95% CI: 1.52-2.92, P<0.001) but not in ovarian cancer (HR =1.05, 95% CI: 0.18-6.15, P=0.957). To explore the source of heterogeneity, subgroup analysis and meta-regression were performed by the following stratification: patient ethnicity, study number, cutoff value, cancer stage, HR obtained method, and analysis style (Table 2). The results indicated that the combined HR estimates for OS in Caucasians and Asians were 2.08 (95% CI: 1.54-2.81, P<0.001) and 1.96 (95% CI: 1.73–2.22, P<0.001), respectively. Differences in the case number, cutoff value, cancer stage, HR obtained method, and analysis method did not influence the effect of IMP3 expression on the OS of solid tumors. Further meta-regression analysis revealed that cancer

Study ID	HR (95% CI)	% weigh
Jiang et al (2006) ⁸	5.84 (3.60–9.49)	3.30
Pei et al (2015) ²⁶	◆ 2.22 (1.41–3.50)	3.45
Jiang et al (2008) ²⁸	1.95 (1.15–3.31)	3.10
Tantravahi et al (2015) ²⁹	♦ 2.17 (0.83–6.25)	1.51
Del Gobbo et al (2014) ³⁴	II.30 (3.13−41.45)	1.04
Sun et al (2015) ³²	1.14 (0.50–2.60)	1.98
Yan et al (2016) ⁹	2 .31 (1.19–4.48)	2.53
Zhang et al (2015) ³³	- 1.61 (1.13–2.28)	3.98
Lin et al (2015) ³⁰	2.29 (1.30–4.05)	2.92
Beljan Perak et al (2012) ³¹	1.09 (0.68–1.75)	3.36
Clauditz et al (2013) ³⁵	1.67 (1.02–2.75)	3.25
Lin et al (2011) ³⁷	1.42 (1.13–1.78)	4.56
Li et al (2010) ³⁶	2.35 (1.16–4.75)	2.37
Kim and Cha (2011) ¹⁰	3.01 (1.17–7.75)	1.66
Szarvas et al (2012) ⁴⁰	1.82 (1.00–3.30)	2.81
Lee et al (2013) ⁴¹	2.07 (1.45–2.95)	3.95
Niedworok et al (2015) ³⁸	0.70 (0.16–3.18)	0.81
Bi et al (2016)43	2.55 (1.23–5.28)	2.28
Noske et al (2009) ¹⁷	0.42 (0.18–0.99)	1.90
Hu et al (2014) ⁴⁴	2.43 (1.57–3.75)	3.55
Wachter et al (2011) ⁴⁵	1.89 (1.02–3.53)	2.69
Chen et al (2013)46	• 2.28 (1.24–4.20)	2.73
Yuan et al (2009)48	1.83 (1.06–3.15)	3.02
Li et al (2009) ⁴⁹	2 .37 (1.20–4.70)	2.45
Lochhead et al (2012) ¹²	1.32 (1.05–1.66)	4.56
Lin et al (2013)30	1.62 (1.03–2.54)	3.47
Szarvas et al (2014) ⁵⁰	1.09 (0.45–2.62)	1.83
Wang et al (2014) ⁵²	3.74 (1.74–8.03)	2.16
Schaeffer et al (2010) ⁵¹	→ 3.34 (1.92–8.96)	2.14
Wang et al (2010)54	2.68 (1.24–5.80)	2.14
Okada et al (2011) ¹¹	2.66 (0.90–11.40)	1.07
Chen et al (2013)46	2.65 (1.51–4.66)	2.94
Gao et al (2014) ⁵⁶	1.87 (1.25–2.80)	3.71
Asioli et al (2010)58	8.96 (3.11–25.84)	1.41
Chen et al (2011)61	2.48 (1.05–5.86)	1.88
Hao et al (2011) ⁶²	→ 10.34 (1.00–106.75)	0.37
Sheen et al (2014) ⁶³	2.36 (0.89–6.27)	1.59
Walter et al (2009)64	3.14 (1.00–9.85)	1.26
Takata et al (2014)67	1.84 (1.18–2.93)	3.45
Wei et al (2014) ¹³	• 8.16 (1.80–36.90)	0.81
Overall (/2=59.1%, P=0.000)	2.08 (1.80–2.42)	100
	<u> </u>	
0.1 1	10	

Figure 2 Forest plot of studies evaluating HR of high IMP3 expression in solid tumors for OS.

Notes: A pooled analysis showed that high IMP3 expression was associated with poor OS in solid tumors (HR =2.08, 95% CI: 1.80-2.42, P<0.001). Weights are from random-effects analysis.

Abbreviations: CI, confidence interval; HRs, hazard ratios; IMP3, insulin-like growth factor II mRNA-binding protein 3; OS, overall survival.

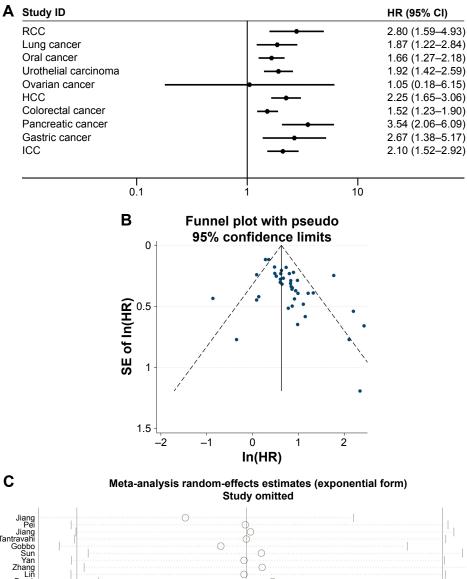
stage is a potential significant contributor to heterogeneity (P=0.017), unlike other factors (P>0.05).

To assess the credibility of the pooled outcomes, we performed a sensitivity analysis through the sequential omission of individual studies. The results were not obviously influenced by any single study (Figure 3C). The publication bias of all included studies was evaluated using a vertical funnel plot, Begg's, and Egger's tests. However, the funnel plot in Figure 3B appears asymmetrical, and the Begg's (P=0.015) and Egger's tests (P=0.002) revealed existing

evidence of publication bias, which may be attributed to only seven studies that reported negative results among all the enrolled studies.

Association of IMP3 with CSS, DFS, RFS, PFS, and MFS

Ten studies that involved a total of 2,877 patients provided sufficient data for CSS analysis. No heterogeneity was observed among these studies ($I^2=31.3\%$, P=0.158). Thus, a fixed model was applied to pool the results. The combined



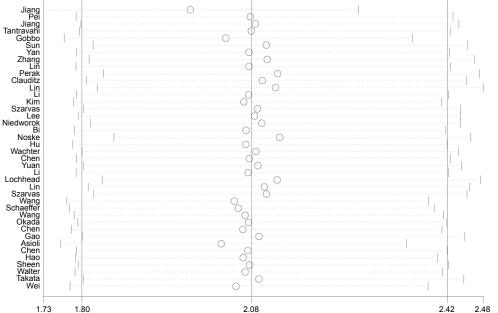


Figure 3 Subgroup analysis of OS stratified by tumor types, funnel plot of OS for publication bias, and sensitive analysis of OS.

Notes: (A) High IMP3 expression was significantly associated with poor OS in RCC, lung cancer, oral cancer, urothelial carcinoma, HCC, colorectal cancer, pancreatic cancer, gastric cancer, and ICC but not in ovarian cancer. (B) The funnel plot for OS was asymmetric, which indicated the probability of publication bias. (C) Sensitivity analysis by sequential omission of individual studies did not alter the significance, which confirmed the credibility of outcomes.

Abbreviations: CI, confidence interval; HCC, hepatocellular carcinoma; HR, hazard ratio; ICC, intrahepatic cholangiocarcinoma; In, natural logarithm; IMP3, insulin-like growth factor II mRNA-binding protein 3; OS, overall survival; RCC, renal cell carcinoma; SE, standard error.

С

Subgroups	Studies	Patients	Pooled HR and 95% CI	P-value	Heterogeneity (I²)	Meta-regressior P-value
Ethnicity						0.748
Caucasian	18	3,827	2.08 (1.54-2.81)	< 0.00 I	76.3%	
Asian	22	2,598	1.96 (1.73-2.22)	< 0.00 I	9.6%	
No of patients						0.659
>100	20	4,850	2.08 (1.71–2.53)	<0.001	62.3%	
<100	20	1,575	2.11 (1.66–2.67)	< 0.001	57.6%	
Cutoff						0.421
Positive vs negative	13	2,562	2.50 (1.96–3.19)	<0.001	53.9%	
>10% of cells stained	11	1,201	1.95 (1.50–2.53)	<0.001	29.7%	
>25% of cells stained	2	188	1.63 (1.06–2.52)	0.027	46.5%	
Others	14	2,474	1.87 (1.42-2.46)	< 0.001	65.6%	
Cancer stage						0.017
Nonmetastatic	14	2,918	2.01 (1.77-2.29)	<0.001	23.4%	
Mixed (metastatic and nonmetastatic)	26	3,507	1.77 (1.58–1.97)	<0.001	16.8%	
HR obtain method						0.326
Reported	34	5,881	2.14 (1.84–2.50)	<0.001	55.5%	
Extracted	6	544	1.70 (1.03–2.82)	0.040	76.2%	
Analysis			× /			0.319
, Univariable analysis	9	768	1.76 (1.09-2.85)	0.020	74.7%	
Multivariable analysis	31	5,657	2.14 (1.84-2.48)	< 0.00 I	52.9%	

Abbreviations: CI, confidence interval; HR, hazard ratio.

HR was 1.75 (95% CI, 1.50–2.05, P<0.001), indicating that high IMP3 expression was associated with worse CSS in the patients with solid tumors (Figure 4A). The subgroup analysis stratified by cancer types showed that high IMP3 expression significantly affected the RCC (HR =1.49, 95% CI: 1.11– 2.01, P=0.008) and urothelial carcinoma (HR =2.17, 95% CI: 1.54–3.07, P<0.001). Further sensitivity analysis did not alter the significance of combined HR, which validated the outcome credibility. Eight studies that involved 979 patients reported HRs for DFS, and the effect of high IMP3 expression is presented in Figure 4B. A combined analysis showed that high IMP3 expression was associated with poor DFS in solid tumors (HR =3.30, 95% CI: 1.82–5.99, P<0.001).

Seven studies with 1,930 patients investigated the prognostic role of IMP3 expression in the RFS of solid tumors. Pooled results demonstrated that high IMP3 adversely influenced the RFS in patients with solid tumors (HR =2.11, 95% CI: 1.43–3.12, P<0.001, Figure 5A). For PFS, four studies with 457 patients were included in the analysis. A forest plot of study-specific HRs for PFS is presented in Figure 5B. The combined results indicated that high IMP3 expression was significantly associated with worse PFS in solid tumors (HR =2.18, 95% CI: 1.11–4.29, P=0.023). In addition, five studies, including 1,613 patients, focused on the influence of IMP3 on solid tumor metastasis. Meta-analysis of these studies suggested that IMP3 expression was also associated with poor MFS (HR =4.91, 95% CI: 2.05–11.73, P<0.001, Figure 5C).

Discussion

Over the past decades, increasing correlative studies describe the elevated IMP3 expression in human cancers, and various functional in vitro or in vivo studies provide strong evidence indicating that this oncofetal protein serves an essential role in modulating tumor cell fate.⁶ As a molecular biomarker, IMP3 has attracted extensive attention and can be used to distinguish different prognoses, improve prediction accuracy, and better guide clinical decisions in different tumor types.⁷ Nevertheless, the relationship between IMP3 expression and oncological outcome remains controversial and requires a consensus. Consequently, we attempted to perform a systematic review of published relevant studies and conduct a meta-analysis to clarify the prognostic value of IMP3 expression in patients with solid tumors.

In the present research, given the inclusion criteria, 53 studies involving 8,937 patients were eligible, and the HRs of cumulative survival rates were summarized quantitatively by standard meta-analysis techniques. Our results suggested that high IMP3 expression was associated with worse OS of the solid tumors. Further subgroup analysis stratified by tumor type presented detailed results as follows. The negative prognostic effects of IMP3 on OS were specifically observed

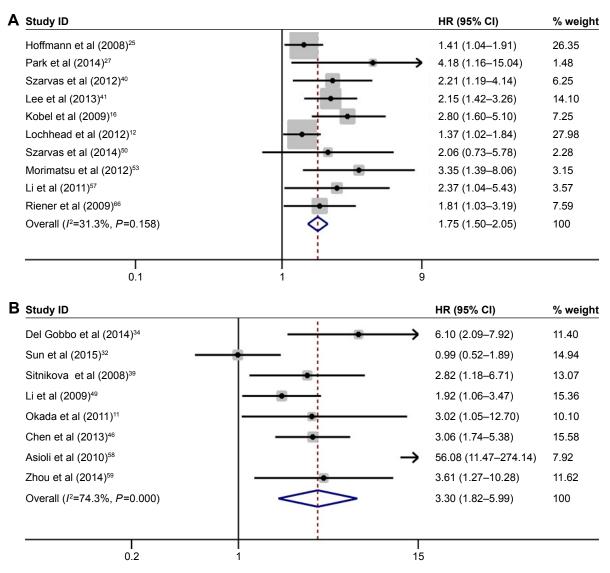


Figure 4 Forest plot of studies evaluating HRs of high IMP3 expression in solid tumors for CSS and DFS.

Notes: (A) High IMP3 expression was associated with poor CSS in solid tumors (HR =1.75, 95% CI: 1.50–2.05, P<0.001). (B) High IMP3 expression was associated with poor DFS in solid tumors (HR =3.30, 95% CI: 1.82–5.99, P<0.001). Weights are from random-effects analysis.

Abbreviations: CI, confidence interval; CSS, cancer-specific survival; DFS, disease-free survival; HRs, hazard ratios; IMP3, insulin-like growth factor II mRNA-binding protein 3; OS, overall survival.

in RCC, lung cancer, oral cancer, urothelial carcinoma, HCC, colorectal cancer, pancreatic cancer, gastric cancer, and ICC. Besides OS, we also investigated other frequently used survival outcomes, including CSS, DFS, RFS, PFS, and MFS. Similar influences were found for high IMP3 expression regarding the abovementioned end points, which provide a relatively comprehensive assessment of the value of IMP3 acting as a prognostic biomarker in solid tumors.

Accumulated literature suggests that IMP3 contributes to various aspects of cancer by promoting target genes expression by either preventing mRNA decay or stimulating mRNA translation. IMP3 knockdown in vitro can significantly inhibit the translation of IGF2 mRNA resulting in the marked inhibition of cell proliferation.² By using solid cancer transcriptome data, IMP3 was also found to be correlated with HMGA2 mRNA expression in a dose-dependent manner. Additional assay for elucidating the mechanism indicated that IMP3 may function as a cytoplasmic safe house and prevents miRNA-directed mRNA decay of HMGA2 during tumor progression.⁴ Another recent study identified IMP3 as capable of directly binding the mRNAs of cyclins D1, D3, and G1 in vivo and in vitro. The study also found that IMP3 can regulate the expression of these cyclins depending on their protein partner HNRNPM in six human cancer cell lines of different origins.⁶⁸ In addition, IMP3 promotes tumor cell invasion and migration by targeting the epithelial–mesenchymal

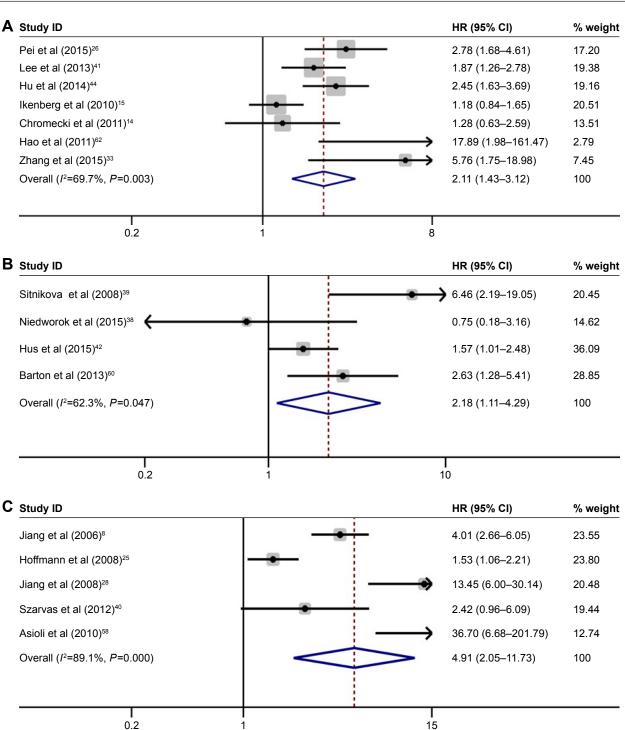


Figure 5 Forest plot of studies evaluating HRs of high IMP3 expression in solid tumors for RFS, PFS, and MFS.

Notes: (A) High IMP3 expression was associated with poor RFS in solid tumors (HR =2.11, 95% CI: 1.43–3.12, P<0.001). (B) High IMP3 expression was associated with poor PFS in solid tumors (HR =2.18, 95% CI: 1.11–4.29, P=0.023). (C) High IMP3 expression was associated with poor MFS in solid tumors (HR =4.91, 95% CI: 2.05–11.73, P<0.001). Weights are from random-effects analysis.

Abbreviations: CI, confidence interval; HRs, hazard ratios; IMP3, insulin-like growth factor II mRNA-binding protein 3; MFS, metastasis-free survival; PFS, progression-free survival; RFS, recurrence-free survival.

transition-associated molecular makers, including E-cadherin, Slug, and vimentin.⁶⁹ Overall, IMP3 plays an essential and multifaceted role in human cancers. Hence, targeting IMP3 may serve as a potential strategy for anticancer therapy. To our knowledge, our study is the first meta-analysis that comprehensively evaluated the association between IMP3 expression and prognosis in patients with solid tumors. However, several limitations of our study must

be acknowledged. First, we only extracted summarized population-level data rather than individual subject data from published literature. Second, different cutoff values and definitions of high IMP3 expression were used in these included studies. Third, a marked study heterogeneity existed in some analyses. The subgroup analyses and meta-regression revealed that cancer stage might be a significant contributor to heterogeneity. Moreover, several potential factors such as cancer type, cutoff value, baseline characteristics (sample size, sex, age, and pathological subtype), and duration of follow-up may partially contribute to the heterogeneity. Among the enrolled studies, 10 works did not directly report the HRs. The calculated HRs, which were estimated using the methods of Tierney et al, might not be as dependable as those retrieved directly from the reported results. As such, the HRs inevitably introduced some statistical errors and may have influenced the pooled analysis. Furthermore, some studies only provided univariate analysis results, which may have introduced a bias toward overestimation of the prognostic value compared with multivariate analysis. The funnel plot and Egger's test suggested the probability of publication bias because of fewer studies reporting negative results. However, the greater difficulty in publishing studies with insignificant results than those with significant results may be unavoidable. Finally, despite the well-recognized advantages of systematic review and meta-analysis, the results were based on the quality of the included studies. Thus, further high-quality studies with larger samples and a unified detection method are entailed to achieve a consensus on this matter.

Conclusion

The current evidence suggests that high IMP3 expression in tumor tissues is associated with adverse survival in various cancers. Hence, IMP3 might be a potential and promising biomarker that can be used to improve prognosis stratification and guide decision making in the treatment of solid tumors. Further well-designed studies are needed to confirm our findings and obtain more precise evaluations of the prognostic value of IMP3 in cancers.

Acknowledgments

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Disclosure

The authors report no conflicts of interest in this work.

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Supplementary material

Table SI Checklist of PRISMA 2009

Section/topic	#	Checklist item	Reported
			on page #
Title			
Title	I	Identify the report as a systematic review, meta-analysis, or both.	I
Abstract			
Structured	2	Provide a structured summary including, as applicable: background;	2
summary		objectives; data sources; study eligibility criteria, participants, and	
		interventions; study appraisal and synthesis methods; results; limitations;	
		conclusions and implications of key findings; systematic review	
		registration number.	
Introduction			
Rationale	3	Describe the rationale for the review in the context of what is already known.	3
Objectives	4	Provide an explicit statement of questions being addressed with reference	3,4
Objectives	Т	to participants, interventions, comparisons, outcomes, and study design	5,4
		(PICOS).	
Methods			
Protocol and	5	Indicate if a review protocol exists, if and where it can be accessed (eg,	No
registration	-	Web address), and, if available, provide registration information including	-
0		registration number.	
Eligibility criteria	6	Specify study characteristics (eg, PICOS, length of follow-up) and report	4,5
. .		characteristics (eg, years considered, language, publication status) used as	
		criteria for eligibility, giving rationale.	
Information	7	Describe all information sources (eg, databases with dates of coverage,	4
sources		contact with study authors to identify additional studies) in the search and	
		date last searched.	
Search	8	Present full electronic search strategy for at least one database, including any	4
		limits used, such that it could be repeated.	
Study selection	9	State the process for selecting studies (ie, screening, eligibility, included in	5
		systematic review, and, if applicable, included in the meta-analysis).	
Data collection	10	Describe method of data extraction from reports (eg, piloted forms,	5
process		independently, in duplicate) and any processes for obtaining and confirming	
		data from investigators.	
Data items	11	List and define all variables for which data were sought (eg, PICOS, funding	5,6
B. I. (1).		sources) and any assumptions and simplifications made.	- /
Risk of bias in	12	Describe methods used for assessing risk of bias of individual studies	5,6
individual studies		(including specification of whether this was done at the study or outcome	
C	12	level), and how this information is to be used in any data synthesis.	F /
Summary measures Synthesis of results	13 14	State the principal summary measures (eg, risk ratio, difference in means). Describe the methods of handling data and combining results of studies, if	5,6 6
Synchesis of results	14	done, including measures of consistency (eg, l^2) for each meta-analysis.	0
Risk of bias across	15	Specify any assessment of risk of bias that may affect the cumulative evidence	6
studies	15	(eg, publication bias, selective reporting within studies).	0
Additional analyses	16	Describe methods of additional analyses (eg, sensitivity or subgroup analyses,	6
		meta-regression), if done, indicating which were pre-specified.	·
Results			
Study selection	17	Give numbers of studies screened, assessed for eligibility, and included in the	7
,		review, with reasons for exclusions at each stage, ideally with a flow diagram.	
Study	18	For each study, present characteristics for which data were extracted (eg,	7
characteristics		study size, PICOS, follow-up period) and provide the citations.	
Risk of bias within	19	Present data on risk of bias of each study and, if available, any outcome level	7–14
studies		assessment (see item 12).	
Results of	20	For all outcomes considered (benefits or harms), present, for each study: (a)	7–14
individual studies		simple summary data for each intervention group; (b) effect estimates and	
		confidence intervals, ideally with a forest plot.	

(Continued)

Table SI (Continued)

Section/topic	#	Checklist item	Reported
			on page #
Synthesis of results	21	Present results of each meta-analysis done, including confidence intervals and measures of consistency.	7–14
Risk of bias across studies	22	Present results of any assessment of risk of bias across studies (see item 15).	7–14
Additional analysis	23	Give results of additional analyses, if done (eg, sensitivity or subgroup analyses, meta-regression [see Item 16]).	7–14
Discussion			
Summary of evidence	24	Summarize the main findings including the strength of evidence for each main outcome; consider their relevance to key groups (eg, healthcare providers, users, and policy makers).	14,15
Limitations	25	Discuss limitations at study and outcome level (eg, risk of bias), and at review-level (eg, incomplete retrieval of identified research, reporting bias).	15,16
Conclusions	26	Provide a general interpretation of the results in the context of other evidence, and implications for future research.	17
Funding			
Funding	27	Describe sources of funding for the systematic review and other support (eg, supply of data); role of funders for the systematic review.	None

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Reference

 Moher D, Liberati A, Tetzlaff J, et al. The PRISMA Group (2009). Preferred Reporting Items for Systematic Reviews and Meta-Analyses: The PRISMA Statement. *PLoS Med.* 2009;6(7):e1000097.

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