Children’s and adolescents’ relationship to pain during cancer treatment: a preliminary validation of the Pain Flexibility Scale for Children

Jenny Thorsell Cederberg
Sandra Weineland
Strandskov
JoAnne Dahl
Gustaf Ljungman

1Department of Women’s and Children’s Health, Pediatric Oncology, Uppsala University, Uppsala, Sweden; 2Närhälsan, Research and Development Center, Primary Health Care, Södra Älvsborg, Borås, Sweden; 3Department of Psychology, Uppsala University, Uppsala, Sweden

Objectives: Children with cancer often suffer from pain. Pain is associated with psychological distress, which may amplify the pain experience. In chronic pain, it has been shown that psychological acceptance is helpful for both adults and children. For experimentally induced pain, interventions fostering psychological acceptance have been shown to predict increases in pain tolerance and reductions in pain intensity and discomfort of pain. A single subject study aiming to nurture psychological acceptance for children with cancer experiencing pain has shown promising results. No instruments measuring psychological acceptance in acute pain are yet available. The aim of the current study was to develop and preliminarily evaluate an instrument to measure psychological acceptance in children experiencing pain during cancer treatment.

Methods: A test version of the Pain Flexibility Scale for Children was sent to all children aged 7–18 years undergoing cancer treatment in Sweden at the time of the study. Exploratory factor analysis was used. Internal consistency, test–retest reliability, and convergent validity were examined.

Results: Sixty-one children participated in the study. A two-factor solution with Promax rotation was found to best represent the data. Internal consistency was good to excellent (α =0.87–0.91). The total scale and the subscales demonstrated temporal stability (Intraclass correlation coefficient =0.56–0.61) and satisfactory convergent validity (r =−0.27 to −0.68).

Discussion: The Pain Flexibility Scale for Children measuring psychological acceptance in children experiencing pain is now available for use. This enables the evaluation of acceptance as a mediator for treatment change in the context of acute pain in children with cancer, which in turn is a step forward in the development of psychological treatments to help children cope with the pain during these difficult circumstances. The scale shows good psychometric properties but needs further validation, particularly considering the small sample size.

Keywords: acute pain, children, acceptance, psychological flexibility, factor analysis

Introduction

Children with cancer suffer from a number of symptoms throughout the cancer trajectory, of which pain is one of the most frequently reported and burdensome ones.1 The children experience pain most often as a result of the disease itself, side effects of the cancer treatment, and/or procedures pertaining to the medical management.2,3 The fact that pain is anxiety-provoking is well known,4 and children suffering from cancer are no exception.5,6 Anxiety, in turn, amplifies the pain experience.7,8 For persons with chronic pain, psychological acceptance has been shown to be helpful.9–11 An acceptance-based psychological treatment, acceptance and commitment therapy, has been shown to improve psychosocial as well as physical functioning for both adults and children with chronic pain.12,13 In acceptance and commitment therapy, the goal
is to help people to engage in their lives in the presence of
difficulties instead of being occupied with avoiding unpleas-
ant stimuli. This is enabled by fostering psychological flex-
ibility.14 Psychological acceptance is one of the key aspects
of psychological flexibility. The definition of psychological
acceptance of chronic pain is “living with pain without
reacting to, judging or attempting to reduce or avoid it.”215 It
entails actively engaging in meaningful life activities in the
presence of pain, in order to continue to live life instead of
putting it on hold, waiting for the pain to pass. Engaging in
meaningful activities in the presence of aversive stimuli has
been shown to precede a reduction in suffering from symp-
toms.16 For experimentally induced pain, acceptance-based
interventions have been shown to predict an increase in pain
tolerance and reductions in pain intensity and discomfort
of pain.17–21 In acceptance-based interventions, an attentive,
nonreactive stance toward unpleasant stimuli is cultivated.
The aim is to merely observe ongoing experiences without
further mental evaluation.22,23 This stance attenuates the pain
experience and at the same time helps the person in pain to
better choose his/her actions instead of rigidly reacting to
internal and/or external events. A single subject study of an
acceptance-based intervention for children reporting acute
pain during cancer treatment has recently been undertaken
at the Children’s University Hospital in Uppsala, Sweden
(Cederberg, unpublished data, 2017). The aim of the inter-
tervention of the study was to help the children to practice a
nonreactive stance toward painful stimuli with the purpose of
helping them to cope with the pain and the emotional distress
that the pain infers, thus giving them a means to continue to
engage in their daily life activities in the presence of pain.
All five participants reported decreased discomfort of pain
postmeasurement. Psychological acceptance was the main
treatment component, and hence the proposed mediator of
the intervention. It is essential to evaluate and understand
mechanisms of change in order to optimize treatments.24 In
the chronic pain area, several instruments measuring psy-
chological acceptance have been reported.25–27 In contrast,
no instrument for measuring psychological acceptance for
persons experiencing acute pain has yet been reported, least
of all for children. With such an instrument at hand, the
evaluation of psychological acceptance as a mediator for
change in psychological interventions in the context of acute
pain would be possible. This, in turn, would contribute to
the development of acceptance-based psychological inter-
ventions that may help children with cancer experiencing
pain to cope better with this challenging situation. The aim
of this study was to develop and preliminarily evaluate an
instrument measuring acceptance in the context of acute pain
in children with cancer.

Methods
Participants and procedures
All children aged 7–18 years being treated for cancer in
Sweden at the time of the study were invited to participate in
the study. Two hundred thirty-three patients were identified
by the Swedish Childhood Cancer Registry in November
2015. For one child, complete patient information was lack-
ing, and he was therefore excluded. The research nurses at
the six pediatric oncology centers in Sweden were consulted
double checked that the children had not gone into pal-
liation and was therefore excluded. Two hundred thirty-one
children were contacted in December 2015 via mail at their registered
addresses. The study material consisted of information about
the study, the test version of the scale, evaluation questions,
and two measures for validation. The children were offered
inclusion in a lottery of ten movie tickets on participation in
the study. For the children, consent was given through par-
ticipation in the study. In addition, a written parental consent
was required for children under 15 years of age. Two weeks
after the first dispatch, a reminder was sent out. One month
after collection of the first measurement, the test material
was sent out again for test–retest analysis. The study material
contained no patient information, but was coded. The code
key was kept in a locked space that could only be accessed by
one of the researchers. Three dispatches were returned by the
Postal Service. Sixty-two children (27%) participated in the
study, of whom 39 participated at both measurements and 23
participated at one measurement, and one was excluded due
to insufficient completion of the scales. Ten children declined
participation. One hundred fifty-six children did not respond.
The study was closed in May 2016. Table 1 provides a demo-
graphic overview. The study was approved by the Regional
Ethical Review Board in Uppsala, Sweden [Dnr 2014/375].

Background information
Background information included age, gender, type and date
of diagnosis, and date of end of treatment (if applicable).
Descriptive pain information included current level of pain
and discomfort; highest, lowest, and average level of pain
during the past week; average level of discomfort of pain during
the past week; and type of pain. Pain and level of discomfort
was rated on a scale from 0 = “No pain/discomfort at all” to
10 = “Unbearably lot of pain/discomfort.”28
Table 1  Gender, age, and diagnosis group of children

<table>
<thead>
<tr>
<th>Sample</th>
<th>Male (%)</th>
<th>Female (%)</th>
<th>Mean age (SD), years</th>
<th>Age range, years</th>
<th>Diagnosis</th>
</tr>
</thead>
<tbody>
<tr>
<td>Children (n=61)</td>
<td>33 (54.1)</td>
<td>28 (45.9)</td>
<td>12.7 (3.4)</td>
<td>7–18</td>
<td>Leukemia 23</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>Brain tumor 13</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>Solid tumor 25</td>
</tr>
</tbody>
</table>

Abbreviation: SD, standard deviation.

Development of the pain flexibility scale for children

Three psychologists theoretically and clinically familiar with the concept of acceptance were involved in the development of the Pain Flexibility Scale for Children (PFS-C). First, a draft of different potential dimensions of acceptance was elaborated. Second, the Chronic Pain Acceptance Questionnaire (CPAQ) was used as a basis for the new scale. The CPAQ is designed to measure acceptance in patients with chronic pain and contains 20 items divided into two subscales. The Activity Engagement subscale measures engagement in meaningful activities in spite of the presence of pain, and the Pain Willingness subscale measures the degree to which the respondent tries to avoid or control pain. A higher score indicates a higher level of acceptance. Internal consistency has been shown to be α =0.78–0.82. The CPAQ correlates negatively with measures of physical and psychosocial disability. The same response format was used; a seven-point Likert scale. The scale ranged from 0 = “Completely disagree (Never true)” to 6 = “Entirely agree (Always true).” Eleven items from the CPAQ that were clearly chronic pain oriented (#2, #4, #5, #6, #9, #10, #12, #13, #14, #18, and #19), such as “My life is going well, even though I have chronic pain,” were deleted. Nine items from the CPAQ (#1, #3, #7, #8, #11, #15, #16, #17, and #20) were retained. These were reframed to suit the process of acceptance in the context of acute pain in children. For example, Item 11 “My thoughts and feelings about pain must change before I can take important steps in life” was reframed to “The pain needs to pass before I can focus on anything else.” Third, 29 new items were generated in accordance with the draft of potential dimensions of acceptance in the context of acute pain. The language was adapted to suit children. Two children, aged 8 and 10 years, filled in the test scale to assess the appropriateness of the level of language. No adjustments were called upon based on their feedback. The final test version contained 38 items. In order to synchronize the direction of the scale, twenty-three items reflecting the opposite pole of the dimension, such as “Being in pain is too difficult for me” and “I need to focus on getting rid of the pain,” were reversed before performing the statistical analyses.

Measures used for validation

Two measures were used to evaluate convergent validity. The Pain Catastrophizing Scale for Children (PCS-C) is designed to measure catastrophicizing thoughts in children in pain. The scale consists of 13 statements with which the children rate their agreement on a five-point Likert scale. Examples of statements are: “When I have pain, I worry all the time about whether the pain will end” and “When I have pain, I get scared that the pain will get worse.” The score range is 0–52, and a higher score indicates a higher level of catastrophizing. The PCS-C correlates with measures of depressed mood and trait anxiety. Internal consistency has been shown to be good (α =0.87). The Avoidance and Fusion Questionnaire for Youth (AFQ-Y) is designed to measure psychological inflexibility in youths. Respondents rate to what extent they agree with statements targeting experiential avoidance and cognitive fusion such as “My life won’t be good until I feel happy” and “I am afraid of my feelings.” The response format is a five-point Likert scale. The score range is 0–32, and a higher score indicates a higher level of psychological inflexibility. The short version of eight items was used, which correlates positively with child-reported anxiety, physical symptoms, and problem behavior and negatively with general quality of life. Internal consistency has been shown to be good (α =0.83).

Statistical analyses

Initial analyses of the test version of the scale were carried out to assess the suitability of factor analysis. Internal consistency was calculated, frequency distributions were examined, and inter-item and item–total correlations were inspected. Preliminary factor analysis was performed whereby eigenvalue, scree plot, and pattern matrices were evaluated to select the number of factors to retain for final factor analysis. Principal component analysis was used. Internal consistency and test–retest reliability was calculated for the final total scale and the subscales. The intraclass correlation coefficient was used to calculate test–retest reliability where a two-way random effects model using an absolute agreement definition was applied. The Single Measures value was assessed. An ICC of less than 0.40 indicates poor agreement, between 0.40 and
0.59 fair, between 0.60 and 0.74 good, and more than 0.75 excellent. Correlations with other measures were carried out in order to assess convergent validity. The data on all scales was normally distributed, and Pearson correlation was used. Correlation coefficients were interpreted according to the guidelines recommended by Cohen (r = 0.10–0.29 small, 0.30–0.49 medium, and 0.5–1.0 large). Level of statistical significance was set at p < 0.05. All statistical analyses were performed in IBM SPSS Statistics, version 24 (Armonk, NY, USA).

Results
Descriptives
Sixty-one children participated in the study. Reports of level of pain and discomfort are presented in Table 2, and reports of type of pain are presented in Table 3.

Factor analysis
Cronbach’s α for the test version of the scale was 0.78, and hence internal consistency was acceptable. Frequency distributions showed that the data on some items were skewed, which was expected. The variability was however considered acceptable for all items. The data were normally distributed for the total test scale. Outliers were identified on items 1 and 2 and on the total scale. The outliers had very little effect on the mean and were retained in the analyses. Eleven items had corrected item–total correlations below zero and were eliminated from further analysis. They were as follows: Item 1, “I prepare to fight when I get pain”; Item 2, “Even though it is difficult for me to be in pain, I know that I can handle it”; Item 3, “I refuse to feel the pain”; Item 12, “If I think about something else I can handle being in pain”; Item 20, “The pain gets easier if I try to control it”; Item 22, “How I react when I get pain is different from one time to another”; Item 23, “If I grit my teeth I can stand being in pain”; Item 26, “Sometimes it is unavoidable to have pain”; Item 33, “If I try to feel what I really actually feel, it is easier,” ”; Item 36, “Sometimes I am actually curious about the pain”; and Item 37, “The pain gets worse if I try to control it.” After the elimination of these eleven items, five items had item–total correlations below 0.3 and were eliminated. They were as follows: Item 4, “Sometimes it feels OK to experience pain”; Item 14, “Even though it is difficult to be in pain I have learned that I can actually handle it”; Item 24, “I need to control the pain”; Item 32, “I try to help myself cope with the pain”; and Item 35, “I do things to flee from the pain.” Furthermore, after elimination of these five items, the item–total correlation for Item 19 “Sometimes I feel that I am greater than the pain” had sunk to below 0.3, and consequently Item 19 was eliminated. This, in turn, lowered the item–total correlation for Item 34 “If I try to feel what I really actually feel, it is more difficult” to below 0.3, and therefore Item 34 was also eliminated. Principal component analysis was performed on the remaining 20 items. Preliminary factor analyses showed no items loading independently of the others. Bartlett’s test of sphericity was significant and the Kaiser–Meyer–Olkin Index was 0.76. Interdependence between factors was indicated, and oblique rotation was used. From the preliminary factor analyses, six factors were extracted with eigenvalues above 1, while the scree plot indicated two factors to retain. The component and pattern matrices supported a two-factor solution, which was chosen. All items had factor loadings above 0.4 and communalities above 0.3. Twenty items were included in the final solution, and Promax was chosen as the rotation method. Variance explained by the factor solution was 54%; 37% by the first and 17% by the second factor. Table 4 provides the final

<table>
<thead>
<tr>
<th>Table 2 Reports of level of pain and discomfort</th>
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<tbody>
<tr>
<td>Variable</td>
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<tr>
<td>Metric</td>
</tr>
<tr>
<td>Mean (SD)</td>
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<tr>
<td>Minimum</td>
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<tr>
<td>Maximum</td>
</tr>
<tr>
<td>Mean (SD)</td>
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<tr>
<td>Minimum</td>
</tr>
<tr>
<td>Maximum</td>
</tr>
<tr>
<td>Current pain</td>
</tr>
<tr>
<td>1.1 (1.6)</td>
</tr>
<tr>
<td>0</td>
</tr>
<tr>
<td>8</td>
</tr>
<tr>
<td>0.8 (1.1)</td>
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<tr>
<td>0</td>
</tr>
<tr>
<td>5</td>
</tr>
<tr>
<td>Current discomfort</td>
</tr>
<tr>
<td>1.0 (1.7)</td>
</tr>
<tr>
<td>0</td>
</tr>
<tr>
<td>7</td>
</tr>
<tr>
<td>0.6 (0.9)</td>
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<tr>
<td>0</td>
</tr>
<tr>
<td>3</td>
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<tr>
<td>Most pain last week</td>
</tr>
<tr>
<td>2.5 (2.5)</td>
</tr>
<tr>
<td>0</td>
</tr>
<tr>
<td>9</td>
</tr>
<tr>
<td>2.1 (2.4)</td>
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<tr>
<td>0</td>
</tr>
<tr>
<td>10</td>
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<tr>
<td>Least pain last week</td>
</tr>
<tr>
<td>0.5 (1.3)</td>
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<tr>
<td>0</td>
</tr>
<tr>
<td>7</td>
</tr>
<tr>
<td>0.3 (0.7)</td>
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<tr>
<td>0</td>
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<tr>
<td>3</td>
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<tr>
<td>Average pain last week</td>
</tr>
<tr>
<td>1.4 (1.5)</td>
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<tr>
<td>0</td>
</tr>
<tr>
<td>6</td>
</tr>
<tr>
<td>1.1 (1.5)</td>
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<tr>
<td>0</td>
</tr>
<tr>
<td>7</td>
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<tr>
<td>Discomfort last week</td>
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<tr>
<td>1.5 (2.0)</td>
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<tr>
<td>0</td>
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<tr>
<td>8</td>
</tr>
<tr>
<td>1.2 (1.6)</td>
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<tr>
<td>0</td>
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<tr>
<td>7</td>
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</table>

<table>
<thead>
<tr>
<th>Table 3 Reports of type of pain</th>
</tr>
</thead>
<tbody>
<tr>
<td>Pain due to</td>
</tr>
<tr>
<td>Type of pain at measurement 1 (n=49)</td>
</tr>
<tr>
<td>Cancer disease</td>
</tr>
<tr>
<td>Side effects of treatment</td>
</tr>
<tr>
<td>Medical procedures</td>
</tr>
<tr>
<td>Several medical causes</td>
</tr>
<tr>
<td>Other causes of pain</td>
</tr>
</tbody>
</table>
factor solution. The theoretical analysis of the item content of the factors brought forth the following factor labels: 1) Valued Action and 2) Pain Resistance. The first factor, Valued Action, is about continuing to live in the presence of pain instead of being occupied with trying to control pain. It is also characterized by a nonevaluative perspective of pain in relation to one's ability to cope with it. The second factor, Pain Resistance, is about resisting and trying to control pain and/or the feelings that being in pain infers. It is also characterized by a kind of reactivity to pain as pain is seen as threatening and unmanageable. The score range is 0–120 for the total scale, 0–54 for the Valued Action subscale, and 0–66 for the Pain Resistance subscale.

Reliability and validity
Scale characteristics and reliability and validity coefficients are presented in Table 5. The test–retest correlation coefficients indicated good agreement for the total scale and fair agreement for the subscales. Controlling for change in level of pain had a negligible effect on these correlations. The correlations with the PCS-C were large for the total scale and the Valued Action subscale and medium for the Pain Resistance subscale. When controlling for level of pain, the correlations between the PCS-C and the total scale and the Valued Action subscale were unchanged, but the correlation between the PCS-C and Pain Resistance subscale changed from −0.43 to −0.41. The correlations with the AFQ-Y were medium for the total scale and the Pain Resistance subscale and small for the Valued Action subscale. When controlling for level of pain, the correlation between the AFQ-Y and the Valued Action subscale was unchanged, but the correlation between the AFQ-Y and the total scale changed from −0.36 to −0.33 and correlation with the Pain Resistance subscale from −0.32 to −0.28. Hence, for the Pain Resistance subscale, controlling for level of pain changed the effect from medium to small. Regarding all other effects of level of pain on the correlations, these did not change the interpretation of the strength of the correlation. All correlations were significant (p<0.05).

Discussion
The aim of the current study was to develop and preliminarily evaluate an instrument for measuring acceptance in
the context of acute pain in children. This would enable the investigation of acceptance as a mediator for change in acceptance-based interventions that may help children with cancer experiencing pain to cope better. Factor analysis was used, and a two-factor solution was chosen. The final scale, the PFS-C, consisted of 20 items. The two subscales were Valued Action and Pain Resistance. Regarding the name of the scale, the term “Flexibility” was chosen instead of “Acceptance” to indicate the theoretically slightly broader scope of the scale, including the Valued Action subscale. The total scale and the Valued Action subscale showed excellent internal consistency, while the Pain Resistance subscale showed good internal consistency. Furthermore, the PFS-C demonstrated temporal stability and satisfactory convergent validity.

The sample of the study was small, especially considering the statistical method used, ie, factor analysis. It is often a challenge to achieve large enough samples in clinical studies in general, and in pediatric clinical studies in particular. A consequence of this challenge is that research in pediatric clinical settings runs the risk of being overlooked and thus not being conducted. The population of 231 children was an in-built limitation. However, given the significance of the prevailing of pediatric clinical research in spite of the challenge of small populations and the importance of the development and evaluation of instruments enabling investigation of mediators for treatment change in order to optimize interventions for children in pain, the study was considered important despite this limitation. Almost a third of the children participated in the study. Considering the format of the study and the often intense situation that undergoing cancer treatment implies for these children, this response frequency must be deemed good enough under the circumstances. This should however be kept in mind when generalizing the results of the study. Some respondents communicated that the questions were difficult to understand. Given the nature of the questions, this was expected and considered inevitable to some extent. The respondents were evenly distributed across the whole age span, ranging from 7 to 18 years, showing that younger children participated to the same extent as older ones. Yet, the possibility of children not participating in the study due to perceived difficulty is, also, something to keep in mind when generalizing the results. All children aged 7–18 years undergoing cancer treatment at the time of the study were invited to participate in the study. Respondents were asked to rate their level of pain. Many respondents had previously experienced pain but were not in pain at the time of the measurement. For those children, the measurements were completed retroactively. This may be the reason why the reported level of pain is as low as it is. Even though retrospective measurements are not desirable, taking into consideration the likelihood that experiencing pain is a strongly unpleasant experience for a child and that the pain episode is likely to have occurred relatively recently, these ratings were considered to be valid. Information about previous experience of pain and its time frames was not collected. In the absence of such background questions, there was a risk of collecting data from children who had not experienced any pain during cancer treatment. Previous research1 and clinical experience suggest however that this would be unlikely. Furthermore, several respondents commented that they referred to a previous pain episode when filling in the scale. The children who explicitly declined participation in the study often stated that their pain had been very limited in time, for example as a side effect of surgery. The risk of including children who had not experienced any pain associated with their cancer or cancer treatment is therefore considered small.

The study is a preliminary validation of the PFS-C. Further validation is always important in the development of new scales, particularly if they are to be used for other populations. In this case, this is especially important, considering the small sample size of the study. A Swedish version of the scale has been developed and evaluated. To be used as an English version, it needs to be validated first. Sensitivity to change also needs to be assessed.

In summary, a scale for measuring acceptance of acute pain in children with cancer is now available for use, enabling the investigation of acceptance as a mechanism of treatment change in this context. This is a step forward in the development of acceptance-based psychological interventions that may help children and adolescents with cancer to cope better with the pain that is often associated with cancer treatment. Given the small sample size of the study, the results should be seen as tentative.

Acknowledgments

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to acknowledge the developers of the CPAQ, upon which the test version of the PFS-C was based.

Disclosure
The authors report no conflicts of interest in this work.

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