Age-based model for metacarpophalangeal joint proprioception in elderly

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Abstract: Neurological injuries such as stroke can lead to proprioceptive impairment. For an informed diagnosis, prognosis, and treatment planning, it is essential to be able to distinguish between healthy performance and deficits following the neurological injury. Since there is some evidence that proprioception declines with age and stroke occurs predominantly in the elderly population, it is important to create a healthy reference model in this specific age group. However, most studies investigate age effects by comparing young and elderly subjects and do not provide a model within a target age range. Moreover, despite the functional relevance of the hand in activities of daily living, age-based models of distal proprioception are scarce. Here, we present a proprioception model based on the assessment of the metacarpophalangeal joint angle difference threshold in 30 healthy elderly subjects, aged 55–80 years (median: 63, interquartile range: 58–66), using a robotic tool to apply passive flexion–extension movements to the index finger. A two-alternative forced-choice paradigm combined with an adaptive algorithm to define stimulus magnitude was used. The mixed-effects model analysis revealed that aging has a significant, increasing effect on the difference threshold at the metacarpophalangeal joint, whereas other predictors (eg, tested hand or sex) did not show a significant effect. The adaptive algorithm allowed reaching an average assessment duration <15 minutes, making its clinical applicability realistic. This study provides further evidence for an age-related decline in proprioception at the level of the hand. The established age-based model of proprioception in elderly may serve as a reference model for the proprioceptive performance of stroke patients, or of any other patient group with central or peripheral proprioceptive impairments. Furthermore, it demonstrates the potential of such automated robotic tools as a rapid and quantitative assessment to be used in research and clinical settings.

Keywords: aging, difference threshold, hand function, joint position sense, MCP, robotic assessment, presbypropria, somatosensation

Introduction

Proprioception consists of limb position sense (sense of stationary position) and kinesthesia (sense of limb movement).1 Proprioceptive information originates from muscle spindle afferents, mechanoreceptors in joint capsules and cutaneous tactile receptors.2,3 The perception and correct interpretation of proprioceptive inputs is an essential prerequisite for many activities of daily living (ADL),4 such as haptic exploration of objects and grasping.5 Proprioceptive function can be impaired, for example, as a result of peripheral or central neurological injuries, as in about half of stroke patients.5,6 There is some evidence that proprioceptive impairments reduce the probability of functional recovery,4 which motivates an increased attention to assessing and treating proprioceptive deficits. To be able to quantify these impairments, normative data from the healthy population are required. There is some literature showing that...
proprioception in healthy subjects declines with increased age (also referred to as presbypropria,\textsuperscript{10} for review, Goble et al\textsuperscript{11}). Although the effect of age-related proprioceptive changes in the lower limbs on postural control in the elderly has been studied widely, relatively few studies have focused on the upper limbs, despite their importance for ADL.\textsuperscript{4-6} The decline in proprioceptive performance in more proximal joints (ie, shoulder, elbow, and wrist) could be demonstrated in a set of different experiments\textsuperscript{12-18} using mostly matching paradigms. For more distal joints (ie, in the hand) there exist some inconsistent results.\textsuperscript{19-22} The age-related effect on proprioception is usually investigated by comparing a young group of healthy subjects with an elderly group of healthy subjects. However, the incidence of neurological injuries may vary depending on the age, as it is the case for stroke, where incidence by age doubles each decade after the age of 55.\textsuperscript{23} Thus, stroke predominantly affects the elderly population. Therefore, to differentiate, for example, between healthy aging and proprioceptive deficits following a stroke, it would be of higher clinical utility to have a detailed model of how proprioception is affected by age within the target age group. To create a valid and accurate reference model, it is essential to base it on outcome measures from reliable and quantitative assessments.

Proprioception is commonly assessed with clinical tests, such as the up-down test in which the finger or toe is moved passively and the patient has to report the direction of movement,\textsuperscript{1,24} or a recent extension to the latter, named dual joint position test, where two digits are simultaneously moved in the same or reverse direction, which was shown to be superior.\textsuperscript{25} However, these tests are administered manually and suffer from poor inter-rater reliability and sensitivity.\textsuperscript{26} Due to the use of ordinal scales, these tests do not provide a precise quantification of proprioceptive function, and the provided outcome measures cannot serve as a basis for a reference model.

With the development of novel methods combined with simple tools, it has become possible to create fine-graded scales.\textsuperscript{21,27,28} Since with robotic technology it is possible to reduce or prevent manual intervention of the experimenter (eg, repositioning of the limb) and to take advantage of the control and sensing capabilities of robotic devices, stimuli can be presented in a more reproducible and well-controlled manner. This is suggested to increase reliability as well as sensitivity\textsuperscript{29} or may reduce flooring and ceiling effects by using continuous outcome measures. As a result, many different studies quantifying proprioception with the help of robotic tools in combination with various assessment paradigms have emerged.\textsuperscript{12,22,30-38} These kinds of assessments would allow creating reference models of healthy performance as a function of age. However, so far there exists only a very limited number of models, namely, for arm position matching\textsuperscript{12} and grasp aperture discrimination using spherical objects.\textsuperscript{21} Hence, since different assessment paradigms may target particular aspects of proprioception, it is necessary to create models, specific to the proprioceptive task, describing healthy performance of a target age group.

The purpose of this study was to create a model of proprioceptive function at the metacarpophalangeal (MCP) joint of the index finger in healthy elderly subjects to be used as a reference model for the prospective evaluation of proprioceptive deficits following peripheral and central neurological injuries, such as stroke. The index finger is essential for most grasp types used in ADL,\textsuperscript{39-41} and flexion of the MCP is a strong contributor to a major synergy for natural grasp patterns.\textsuperscript{42} The MCP joint angle difference threshold or limen (DL) was assessed with an automated robotic tool using an adaptive procedure named Parameter Estimation by Sequential Testing (PEST)\textsuperscript{43} in a two-alternative forced-choice (2AFC)\textsuperscript{44} paradigm. The influence of age, sex, dominance of the tested hand, finger length, measurement order, and number of trials on the DL was examined. We hypothesized that the model would reveal a major influence of age on the DL, thus, demonstrating the sensitivity of the proposed assessment method and its suitability for research on proprioception as well as clinical settings for a more informed diagnosis, prognosis, and treatment planning after stroke or other neurological injuries.

**Methods**

**Subjects**

Thirty healthy elderly subjects (S01–S30, 62.8±6.4 years, range 55–80 years, 14 males and 16 females, 29 right and 1 left handed) completed the study. Average finger length (measured from the MCP joint to the tip of the index finger) was 97.3±7.2 mm (across both hands). Handedness was assessed with the Edinburgh Handedness Inventory.\textsuperscript{45} Subjects were excluded from the study if they had somatosensory or motor deficits affecting hand function, or any history of neurological (central or peripheral) or hand injury. Before participating in the experiment, all subjects provided written informed consent. Ethical approval was obtained from the institutional ethics committee of ETH Zurich and the University of Konstanz.

**Robotic apparatus**

The assessment of MCP joint proprioception was performed with the Robotic Sensory Trainer (Figure 1) previously used...
This device can provide well-controlled and reproducible passive finger movements (flexion and extension) around the MCP joint through an actuated remote center of motion (RCM) mechanism. The index finger is attached by means of two Velcro<sup>®</sup> straps to a sliding finger carriage mounted on the RCM mechanism. The hand and forearm supports can be adjusted with 6 degrees of freedom to allow for a comfortable posture. The tested hand of the subject is occluded from vision by a touchscreen to avoid visual cues on finger position. The subject can provide feedback on perceived stimuli directly on the touch-screen placed above the tested hand by using the nontested hand. Based on LabVIEW (National Instruments, Austin, TX, USA), the assessment runs autonomously and does not require intervention of the experimenter once initiated.

### Experimental protocol

Subjects were seated in front of the assessment apparatus, and forearm, hand and index finger were strapped to the device after adjusting the supports. The MCP joint was carefully aligned with the RCM indicated by the extension of the black arrow mounted on the device (Figure 1). The MCP joint position DL was assessed for the index fingers of both hands in randomized order within one experimental session.

Every trial consisted of two successive passive finger movements to different flexion angles of the MCP joint (two-interval design, Figure 1). Passive movements were induced by the robotic apparatus, always starting from the same resting position (with all finger segments aligned, indicated by the dashed line in Figure 1). Flexion movements and movements back to resting position lasted 1 s each, whereas the MCP flexion angle was maintained for 1.5 s. According to the 2AFC paradigm, subjects were asked after each trial to indicate on the touchscreen, using the nontested hand, which of the two presented angular displacements was larger. No feedback on correctness of the answer was provided. The difference between the two angles was defined as positive and centered around the flexion reference of 20°. The difference was adjusted from trial to trial using the adaptive PEST algorithm in order to converge toward the DL. PEST is based on a set of heuristic rules taking the subject’s responses to past stimuli into account, leading to smaller stimulus differences in case of high proportion of correct responses, and larger difference in case of low proportion of correct responses. The range of angular differences was limited to (0°, 40°) due to the mechanical limitations of the device. The same starting parameters (first presented angular difference of 5.5°, and first decreasing or increasing step of 2°), termination conditions (minimum step of ±0.1°, or 20 consecutive trials at the same level), and logarithmic mapping as reported for the previous pilot study were used. A maximum of 120 trials was permitted in case of nonconvergence, in order to keep the assessment duration short. These parameter values were selected based on our prior experimental knowledge and experience.

### Data analysis

In order to estimate the DL from the experimental data, the proportion of correct responses at stimulus levels x (ie, at differences between the consecutively presented angles) was fitted with the psychometric function \( \psi(x) \) in (1) using a Maximum Likelihood criterion:

\[
\psi(x; \alpha, \beta, \gamma, \lambda) = \gamma + (1 - \gamma - \lambda) F(x; \alpha, \beta)
\]

\( F(x) \) corresponds to a cumulative normal function with the inflection point at \( \alpha \) and the slope \( \beta \) at this point. The guessing rate \( \gamma \) was set to 0.5, according to the 2AFC paradigm, and the lapse rate \( \lambda \) (taking into account stimulus-independent errors, or “lapses”) was allowed to vary between 0 and 0.1, to reduce estimation bias. The DL is defined at \( x = \psi^{-1} \) (0.75). The Weber fraction \( K \) (DL divided by the reference angle, here 20°) is reported together with the group average DL in degrees. Such a hybrid procedure combining adaptive sampling procedures and fitting of parametrized functions, as described by Hall, allows estimating the DL even when the adaptive sequence (here PEST sequence) does not converge and terminate within a maximum number of trials.

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**Figure 1** Side view on the RCM mechanism of the robotic device used to induce passive movements around the MCP joint.

**Note:** The sequence of pictures shows one trial, during which two different flexion angles (A1 and A2) are presented.

**Abbreviations:** MCP, metacarpophalangeal; RCM, remote center of motion.
A linear mixed-effects model analysis to describe the DL was performed. As fixed effects, age (in years), sex (male versus female), tested hand (dominant versus nondominant), index finger length (in millimeters), measurement order (first versus second assessment within the session), and number of trials were entered into the model. Furthermore, the interaction between the factors tested hand and measurement order was included. As random effects, intercepts for the subjects were added. Handedness (right versus left) was not included into the model, as the right and left handed groups were not balanced (29/1). A log_{10} transform was applied to the DL before fitting the model, since the DL is only supported in the positive, semi-infinite interval (0,∞). Visual inspection of residual plots did not reveal any obvious deviations from homoscedasticity or normality when using the log_{10} transform. Furthermore, the predictors were tested for collinearity. In order to assess the significance of fixed effects, the P-values were obtained by simulated (n=1,000) likelihood ratio tests (LRTs, MATLAB function compare) of the full model H_A with the effect in question against the model H_0 without the effect in question. This method generates a reference distribution of the LRT statistic X^2 under H_0 and compares it to the observed X^2. This is computationally more intensive, but could be more accurate than comparing the test statistic X^2 to a χ^2-distribution using an analysis of variance, which is not always a very good approximation and tends to be anticonservative. To test whether clinical utility could be improved by reducing the maximum number of trials (120 to 60) to shorten the assessment duration, the same mixed-effects model analysis was conducted with truncated PEST sequences.

Significance levels were set to α=0.05. Descriptive statistics are reported as mean ± standard deviation (SD). All statistical analyses were performed in MATLAB R2014a (MathWorks, Natick, MA, USA).

Results
At group level (both hands), the MCP joint angle position DL averaged at 1.81° ± 0.96° (K=9.0%±4.8%). On average, 65.3±27.3 trials (ranging from 24 to 120) were required, resulting in an average duration of 14.3±6.0 min per assessment. In 5 out of the 60 assessments, the PEST algorithm did not converge, leading to a convergence rate of 91.7%. The psychometric functions fitted to the proportion of correct responses and the evolution of the corresponding PEST sequences for both hands are shown for a representative subject (S22) in Figure 2. As visible in the top plots, stimulus levels are primarily sampled around the steeper part of the psychometric function, and the DL of the left, nondominant hand is smaller. In addition, only the assessment of the left, nondominant hand converged prior to 120 trials (bottom plots).

The mixed-effects model analysis revealed that age significantly affected the DL in increasing manner according to the simulated LRT (P_{age} =0.029). The effect of age on the DL is illustrated in Figure 3 using averaged parameters of this sample multiplied by the estimates of the corresponding predictors not in question. Adding other fixed effects, that is, sex, finger length, handedness, tested hand, measurement, and number of trials as well as the interaction between tested hand and measurement, did not significantly improve the model for estimating the DL. Despite handedness not significantly improving the model, in 19 of 30 subjects, the DL of the nondominant hand was lower compared to the DL of the dominant hand. The simplest model for estimating the DL in healthy elderly (age range: 55–80 years) containing only the fixed effect age and a constant accounting for the other parameters was DL =10^{(0.014 \cdot age – 0.703)} in degrees. The complete summary of the mixed-effects model is provided in Table 1. When truncating the assessment to a maximum of 60 trials, the number of trials was 51.5±11.9 on an average, with 33 of 60 assessments requiring all 60 trials. The changes in the model were minor (DL =10^{(0.015 \cdot age – 0.704)}, simulated LRT P_{age} =0.032, all other P-values >0.05).

Discussion
This study aimed to create a model of MCP joint proprioception in the healthy elderly population to serve as a reference for patients suffering from proprioceptive deficits. Proprioception was quantified by the joint angle DL and assessed with a robotic tool applying well-controlled flexion movements to the index finger. As hypothesized, the proposed assessment approach is sensitive enough to capture a declining effect of MCP joint proprioception with increasing age. According to the model, the index finger MCP joint angle DL increases by around 2° from age 55 to 80. Furthermore, age was the only fixed effect having a significant influence on the DL, and it is thus essential to include it in a model of proprioception.

Age-related decline of proprioception
The observed age effect is consistent with most of the literature on proprioception at more proximal joints12–18 and distal joints.20–22 This decline in proprioception could be a consequence of increased proprioceptive attentional demand in older adults10 due to age-related central19 or peripheral22 physiological changes (for review, Goble et al11). However, there are also studies showing no age-related decline,16,19 or some showing age-related deficits depending on the task.50
Figure 2 Resulting psychometric functions and PEST sequences of a representative subject.

Notes: (Top) Psychometric functions (thick lines) for both hands of subject S22 (67-year-old male). The size of the black dots indicates the number of presentations of a stimulus at a certain stimulus level ($\Delta$ stimulus). (Bottom) Corresponding PEST sequences for the same subject. The thick line represents the stimulus level at each trial.

Abbreviations: DL, difference threshold or limen; PEST, Parameter Estimation by Sequential Testing.

At the level of the hand, the assessment of proximal interphalangeal joint position sense through position matching using a robotic tool that induced velocities below the movement detection threshold, reported a significant deterioration from the young to the elderly group, as well as a moderate, but significant, positive correlation ($r=0.466$) between age and the magnitude of the matching error. Similarly, comparisons between different age groups in two proprioceptive tasks (indicating overlap of fingers during passive crisscross movements and onset of passive finger movement) revealed

Figure 3 Age–DL relationship.

Notes: Aging has a significant increasing effect on the DL ($P<0.05$). The DL of the dominant hand is indicated by an upward-pointing triangle and the DL of the nondominant by a downward-pointing triangle. Both assessments are connected through a gray vertical line for each subject. Black lines show the age-based DL model (thick line) ±SE (dashed lines) obtained from the mixed-effects model analysis.

Abbreviations: DL, difference threshold or limen; SE, standard error.
Table 1 Summary of the mixed-effects model analysis with the estimates and their SE predicting the log\_DL

<table>
<thead>
<tr>
<th>Name</th>
<th>Estimate</th>
<th>SE</th>
<th>t-Value</th>
<th>DF</th>
<th>P-value</th>
<th>Confidence interval (95%)</th>
<th>Simulated LRT</th>
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<tr>
<td></td>
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<td></td>
<td></td>
<td>Lower</td>
<td>Upper</td>
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<tr>
<td>Fixed effects</td>
<td></td>
<td></td>
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<td></td>
<td></td>
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</tr>
<tr>
<td>(Intercept)</td>
<td>-0.857</td>
<td>0.678</td>
<td>-1.27</td>
<td>52</td>
<td>0.212</td>
<td>-2.217</td>
<td>0.503</td>
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<tr>
<td>Age(^a)</td>
<td>0.014</td>
<td>0.005</td>
<td>2.74</td>
<td>52</td>
<td>0.008(^a)</td>
<td>0.004</td>
<td>0.025</td>
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<td>Sex(^a)</td>
<td>0.142</td>
<td>0.082</td>
<td>1.72</td>
<td>52</td>
<td>0.091</td>
<td>-0.023</td>
<td>0.307</td>
</tr>
<tr>
<td>Finger length(^b)</td>
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<td>0.005</td>
<td>-0.09</td>
<td>52</td>
<td>0.928</td>
<td>-0.011</td>
<td>0.010</td>
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<td>Trials(^d)</td>
<td>0.002</td>
<td>0.001</td>
<td>1.96</td>
<td>52</td>
<td>0.055</td>
<td>-0.000</td>
<td>0.004</td>
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<tr>
<td>Tested hand(^d)</td>
<td>-0.054</td>
<td>0.085</td>
<td>-0.64</td>
<td>52</td>
<td>0.527</td>
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<tr>
<td>Measurement(^e)</td>
<td>-0.010</td>
<td>0.083</td>
<td>-0.12</td>
<td>52</td>
<td>0.903</td>
<td>-0.177</td>
<td>0.157</td>
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<tr>
<td>Measurement*tested hand(^d)</td>
<td>0.083</td>
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<td>0.331</td>
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Random effects

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<th>Group</th>
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<tr>
<td>Subject (Intercept)</td>
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<tr>
<td>Residual</td>
<td></td>
<td>0.213</td>
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</tbody>
</table>

Notes: The last column shows the P-values from the simulated LRT. The fixed effect age affected the DL significantly in increasing direction (*P*<0.05). *Baseline* = 0 years.

Abbreviations: DF, degrees of freedom; DL, difference threshold or limen; LRT, likelihood ratio test; SD, standard deviation; SE, standard error.

larger errors in the elderly. In addition, an increase of 0.4 decision errors/decade for a total of 21 decisions was reported for an assessment where subjects had to indicate whether they perceived a comparison polystyrene ball to be larger, smaller, or equal to the volume of reference polystyrene ball.\(^21\) In contrast, a study also assessing proprioception at the MCP joint, but by determining the detection threshold of sinusoidal movements, reported no major age-related decline in joint motion sensation.\(^19\) However, it is difficult to compare the different findings on age-related changes in proprioception quantitatively due to the diversity of outcome measures arising from the different experimental paradigms. Models reported in the literature are based on matching paradigms and provide matching errors,\(^12,20,21\) which cannot be directly related to different thresholds of proprioceptive perception.

One limitation of this study is the fact that the created model is limited to an age range of 55–80 years, and that the recruited subjects were not uniformly distributed across age. In particular, the number of subjects with age >70 years was small. However, additional verification of the model parameters showed that the model was robust to exclusion of the four most elderly subjects (>70 years), demonstrating that the effect of declining proprioception is supported by the entire data set. Furthermore, no comparison to a young group of healthy subjects was made. Yet, as expected, the group average of the DL in the elderly sample population of this study is slightly higher compared to the one reported for healthy young subjects (1.73, K=8.6%) in the pilot study using the same apparatus and practically identical protocol.\(^36\)

Furthermore, although the incidence by age doubles each decade after the age of 55,\(^23\) the age range of our study covers the major part of our target population. Although the comparison of young and elderly subjects can support the investigation of age-related changes, it has limited clinical value compared to an age-based reference model.

Influence of hand dominance and sex

Although the mixed-effects model did not show a significant effect of the tested hand (dominant versus nondominant) on the DL, about two-thirds of the subjects showed a better performance (ie, lower DL) with the nondominant limb. This trend is well in line with some literature, suggesting proprioceptive processing advantages of the nondominant limb in some conditions,\(^12,16,31–35\) while others suggest gain differences of sensory-motor loops as an explanation.\(^36\) In contrast to those studies, other groups did not identify any difference between dominant and nondominant limb.\(^18,22\) There are also studies showing smaller matching errors with the dominant limb.\(^28\) It has been suggested that long-term use-dependent superiority of the dominant hand may enhance proprioception.\(^57\) As a conclusion, proprioception may be superior in the nondominant limb, although the dominant hand is generally more dexterous. However, as the results from the different studies show, this may strongly depend on the assessment paradigm used, as motor function may be a strong confound in some assessments, as for example in active matching tasks.

As in most previous studies, we did not find an effect of sex on proprioception.\(^16,28,51,58\) There was one study that
identified sex-related differences, however, only in some of the outcome measures of a matching task including movements of the elbow and shoulder.12

Robustness to fatigue and learning, number of trials, and clinical utility

Based on the results from this study, the DL is robust with respect to the measurement order (first versus second of two consecutive assessments, also if hand dominance is taken into account) and the number of trials. Whether the dominant/nondominant hand was assessed first or second within the session had no influence on the DL. This suggests that there is no learning effect of the task, which is essential for an assessment, as the goal is to measure the capacity of the subject and not the improvement due to increased level of familiarization with the task. Despite both hands being assessed consecutively, no fatigue effect could be observed. It could be, though, that learning and fatigue effects cancel each other. However, the fact that a larger number of trials does not have a significant influence on the DL either, provide further evidence that there is no important fatigue effect.

As previously shown in a pilot study with healthy young subjects, the adaptive sampling procedure PEST allows reducing the assessment time considerably compared to the widely used, but inefficient, method of constant stimuli,36 resulting in an average assessment duration of ~15 min. Truncating PEST sequences to a maximum of 60 trials resulted in an average assessment duration of 11 min and only in minor changes of the mixed-effects model, demonstrating robustness of the assessment. This is crucial because shortening the assessment duration can significantly increase the assessment’s clinical utility, especially for patient groups, where assessment time is expensive, and where time constraints are perceived as a barrier.59 Furthermore, reliability and validity of an assessment could benefit from short assessments, as the influence of confounding factors such as inattention and other cognitive factors would be decreased.60

Advantages and limitations of the assessment paradigm

This paradigm assesses proprioception in an isolated way, in contrast to assessments using ipsilateral and contralateral matching tasks, which most often require the subject to move actively to reproduce a presented position. Thus, this assessment can be used to investigate somatosensory deficits independently of motor deficits. This allows investigating their contribution to functional impairments and effect on recovery after neurological injuries, such as stroke.

Besides potential confounding factors such as learning and fatigue, the 2AFC method could also be affected by short-term memory. Previous studies have shown that absolute errors in position matching tasks are significantly influenced by several factors such as the type of position matching task or reference position establishment (eg, reference joint angle and how the limb was displaced to present this angle).51 A study with 10 healthy subjects demonstrated better proprioceptive performance in an ipsilateral matching task requiring short-term memory than in a contralateral (simultaneous) matching task requiring interhemispheric transfer, which in turn is better than in a contralateral remembered task requiring both, memory and interhemispheric transfer.61 Similar results have been presented in another study for these three task types revealing mostly disproportionate increases in matching errors for the contralateral remembered task.15 A parallel can be drawn to the assessments of difference thresholds using either two intervals (as was the case here) requiring short-term memory (the two stimuli have to be remembered and compared postpresentation) or two locations stimulated simultaneously (eg, on both index fingers) requiring interhemispheric transfer for comparison, inducing temporal or spatial errors, respectively. Hence, this suggests that the paradigm used in this study (two-interval 2AFC) should be less affected by factors besides healthy aging of proprioception and also be less error prone, as no interhemispheric transfer is required. Furthermore, many limb matching tests rely on the sensorimotor function of the ipsilesional “unimpaired” limb of the patient, which may also be affected by a cerebral lesion,59,62 and their outcomes might also be influenced by deficits in the central integration of proprioceptive information across the two limbs. These confounds are also fully addressed by the two-interval 2AFC approach. Moreover, compared to other psychophysical paradigms such as Yes–No, Remainder, and Same–Different, the 2AFC approach is more robust against decision criteria (ie, response bias), and thus more objective.43,63

Although in each trial of the 2AFC assessment paradigm movements with different amplitudes are presented, there is a trade-off between constant movement duration versus constant movement velocity. Thus, subjects could rely on one of these potential confounds besides the presented joint position angle. There have been approaches where a subthreshold movement velocity was used to resolve this trade-off.20 However, this approach is only usable in research, as it leads to an increased trial duration resulting in overly long
assessments, thus limiting their clinical applicability. For our assessment, a constant duration with varying velocity was chosen because perception of movement velocity (ie, kinesthesia) is a subpart of proprioception, and both position and velocity information is suggested to be incorporated within internal models. In contrast, discriminating time intervals, apart from detection of movement onset and cessation following discharge of muscle spindles, is not strictly part of the proprioceptive sense.

**Conclusion**

The results of this study demonstrate that taking age into account when creating a model of healthy performance of proprioception as a reference for neurological patients suffering from proprioceptive deficits is essential. All the more, as in some neurological injuries, such as stroke, the ipsilesional “unimpaired” limb of the patient can also be potentially affected and should thus not be used as a reference to quantify the performance of the impaired limb. Furthermore, this study supports the clinical utility of the objective and automated assessment approach using a robotic tool for quantifying the angular DL at the MCP joint through its sensitivity and rapid administration. This highlights its potential as an assessment tool to be used in combination with age-based reference models for an informed diagnosis, prognosis, and planning of clinical interventions.

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**Disclosure**

The authors report no conflicts of interest in this work.

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