Quetiapine monotherapy versus placebo in the treatment of children and adolescents with bipolar depression: a systematic review and meta-analysis

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Background: Some studies have indicated the efficacy of quetiapine in the treatment of bipolar depression in adult patients. However, its efficacy has been not shown in child and adolescent patients.

Objective: This systematic review purposefully determined the efficacy and acceptability of quetiapine in the treatment of children and adolescents with bipolar depression.

Data sources: A database search of EMBASE, PubMed, CINAHIL, and Cochrane Controlled Trials Register was carried out in March 2016. All randomized controlled trials (RCTs) of bipolar depression in children and adolescents were considered for inclusion in this review.

Study eligibility criteria, participants, and interventions: RCTs of quetiapine in the treatment of child and adolescent patients with bipolar depression with end point outcomes were included in this study. Languages were not limited.

Study appraisal and synthesis methods: The full-text versions of relevant clinical studies were thoroughly examined and extracted. The primary efficacy of outcome was measured by using the pooled mean-changed scores of the rating scales for bipolar depression. However, the response and remission rates were also measured.

Results: A total of 251 randomized patients in the three RCTs of quetiapine versus placebo in the treatment of bipolar depression for children and adolescents were eligible in this review. The pooled mean-changed score of the quetiapine-treated group was not greater than that of the placebo-treated group. Similarly, the pooled response and remission rates were not different between the two groups. The pooled overall discontinuation rate and the discontinuation rate due to adverse events were not different between the two groups.

Limitations: Limited studies were eligible in this review.

Conclusion: According to the findings in this review, quetiapine may not be efficacious in the treatment of bipolar depression in children and adolescents. Its acceptability, however, was comparable to a placebo. Therefore, the use of quetiapine in children and adolescents with bipolar depression is not recommended. Further well-defined clinical studies should be performed to confirm these outcomes.

Keywords: monotherapy, depressive episodes, pharmacological treatment, Children’s Depression Rating Scale—Revised, CDRS-R

Introduction

Bipolar depression is considered as a major depressive episode of bipolar I or II disorders. Patients with bipolar I or II disorders often experience depressive symptoms, including time to clinical remission longer than either hypomanic or manic episodes.1–3 Depressive episodes in bipolar disorders are associated with significant impairment of various occupational, academic, and social functioning, negative effects on the quality
of life, and mortality of children and adolescents.\(^5\) However, a limited number of clinically controlled studies indicated the efficacy of pharmacological treatment for children and adolescents diagnosed with bipolar disorder.\(^6\)–\(^8\)

Quetiapine antagonizing to serotonergic (5-HT\(_{1A}\) and 5-HT\(_{2A}\)), dopaminergic (D\(_1\) and D\(_2\)), and histamine (H\(_1\)) receptors and adrenergic (\(\alpha_1\) and \(\alpha_2\)) receptors has efficacy in the treatment of schizophrenia,\(^9\) major depressive disorder (MDD),\(^10\) generalized anxiety disorder,\(^11\) and bipolar disorder, including its depressive episode.\(^12\) A previous study suggests that an adjunctive quetiapine treatment can alter the sleep architecture in patients with MDD or bipolar depression.\(^13\) In addition, it also has an active metabolite, norquetiapine (N-desalkyl quetiapine), which has a high affinity for norepinephrine transporters and partially agonizes to serotonin 5-HT\(_{1A}\) receptors.\(^14\) Hence, the efficacy of quetiapine in ameliorating depressive symptoms in adults with unipolar and bipolar depression may be associated with antagonizing at 5-HT\(_{2A}\), \(\alpha_2\)h adrenergic, and \(D_2\) receptors and agonizing at a 5-HT\(_{1A}\) receptor and its effect on sleep architecture.\(^15\)\(^15\)

Although some randomized controlled trials (RCTs) have shown the efficacy of quetiapine in the treatment of bipolar depression in children and adolescents,\(^6\)–\(^8\) each study had limited sample sizes. A previous review of quetiapine in bipolar depression was performed without separate synthesis of outcomes of the child and adolescent population, and one of the three RCTs of quetiapine versus placebo in the treatment of bipolar depression for children and adolescents was not included in such a review.\(^12\) Consequently, a statistical method for possibly determining the true effect size, meta-analysis, may examine quetiapine in the treatment of bipolar depression in children and adolescents in terms of efficacy and acceptability.

This study purposefully and systematically reviewed the efficacy, acceptability, and tolerability of quetiapine monotherapy for acute depressive episodes in children and adolescents with bipolar I and II disorders based on the International Classification of Diseases (ICD) or the Diagnostic and Statistical Manual of Mental Disorders (DSM). In this review, efficacy was measured by using the pooled mean-changed scores of standardized rating scale for depressive symptoms, response rate, and remission rate. In addition, acceptability and tolerability were determined by the overall discontinuation rate and the discontinuation rate due to adverse events, respectively. Only such RCTs were eligible in this review.

**Methods**

**Types of included trials**

Any relevant RCTs that met the inclusion criteria were eligible.

**Types of participants**

All children and adolescents, up to 18 years of age and diagnosed with bipolar I or II disorder with depressive episodes utilizing either the ICD or DSM criteria, were included.

**Types of interventions**

Quetiapine as monotherapy was compared with a placebo. The doses, forms, and frequency of therapy were not limited.

**Types of outcome measures**

**Primary outcome measures**

The primary outcome measure was the mean-changed score of a standardized depressive rating scale (Children’s Depression Rating Scale–Revised, CDRS-R).\(^16\)\(^17\)

**Secondary outcome measures**

The secondary outcome measures included:

1. Response and remission rate defined by each trial
2. Clinical Global Impression–Bipolar Version Severity (CGI-BP-S)
3. Discontinuation rates
   a. Overall discontinuation rate
   b. Discontinuation rate due to adverse events

**Information sources**

The databases, including EMBASE, PubMed, CINAHL, and Cochrane Controlled Trials Register databases, were searched in March 2016. As the first publication of quetiapine administration commenced in 1991 in the PubMed, searching the publications was planned to start from January 1991 to March 2016. All searches were restricted to trials conducted in humans. In addition, the databases of ClinicalTrials.gov, EU Clinical Trials Register (EU-CTR), and AstraZeneca Clinical Trials were also searched. The relevant references of any article provided by any method were assessed. All relevant RCTs were considered. However, languages in each study were not limited.

**Searches**

To sensitize the optimal identification of the RCTs, the searching method for PubMed was limited to the following words and phrases: ([quetiapine] OR [seroquel]) AND ([Bipolar depression] OR [BD]). This similar searching strategy was used in the rest of the databases.

**Study selection**

To decide whether each study conformed to the eligibility criteria described earlier, inspection of all abstracts derived...
from those electronic databases was individually conducted by two reviewers (NM and BM). Then, the full-text versions of the relevant articles were collected and individually examined by the reviewers. If disputes occurred between the two reviewers, a conclusion was achieved by consensus.

**Data collection process**

Initially, data of the full-version articles were extracted and entered into the extraction form developed by the first reviewer (NM). Subsequently, the second reviewer (BM) meticulously rechecked the extracted data. Again, all disputes between two reviewers were also resolved by consensus. If those disagreements between two reviewers were still unresolved, the third reviewer (MS) would make the decision.

**Data items**

The extraction data gathered from the clinical studies comprised the following: 1) important information used for evaluating the study quality; 2) basic characteristic data including population, diagnostic criteria, study designs, and eligible/ineligible criteria; 3) forms, doses, and treatment duration of quetiapine vs placebo; 4) interesting outcomes; and 5) intention-to-treat outcomes.

**Risk of bias in individual studies**

The internal validity (quality) assessments of the included clinical trials were carried out by two reviewers (NM and BM). According to the quality assessment of Cochrane Collaboration, the risk of bias was examined, including 1) sequence generation (randomization); 2) allocation concealment; 3) blinding of participants, personnel, and outcomes; 4) incomplete outcome data; 5) selective outcome reporting; and 6) other biases. In addition, outcome quality was evaluated by using the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach.

**Summary measures**

The interesting outcomes were efficacious, acceptable, and tolerable. The end point or the mean-changed scores rated on the standardized depressive scale and the response rate defined by any set of criteria were considered as a treatment efficacy. Although acceptability and tolerability have been interchangeably applied, each term has its specific definition. Like the previous meta-analysis, the overall discontinuation rate, including rate of adverse events, and the discontinuation rate due to adverse events, related to only side effects of the medications, were considered as an acceptability and tolerability, respectively.

**Statistical analysis and synthesis of results**

Basically, a weighted mean difference (WMD) or a standardized mean difference (SMD) with 95% confidence interval (CI) is estimated by the mean difference between the comparison groups divided by an estimate of the within-group standard deviation (SD). Mean difference, with 95% CI, was applied for synthesis of the continuous outcomes. When the same outcomes are measured by various rating scales across studies, the direct comparison or combination of such study results was implausible. Hence an SMD, a statistic summary method that can be used when the same outcome is measured in a variety of ways, is a reasonable measurement of comparison or combination for those findings. When a similar rating scale was used, a WMD, directly comparing, or combining the study results, was applied. Since all included clinical trials used a similar rating scale to assess the same outcomes, the WMD was applied for estimating the efficacy of quetiapine by using all of the continuous data. If the SD of the changed mean scores was not available, it was estimated by using any of the statistical analyses or by direct substitution. The statistical method for combining the results of multiple studies by an inverse variance for the estimated measure effect, which weighed the influence of each study, was used to calculate the pooled mean-changed scores with 95% CIs.

Synthesis of dichotomous outcomes used the relative risk (RR), with the 95% CI. If the RR was exactly 1, it indicated that the differences in the outcomes did not take place between the intervention and control groups. When the RR was > or <1, it suggested that the intervention, respectively, increased or decreased the risk of the outcomes. In this review, the RRs were applied to compare the response rates, the remission rates, the overall discontinuation rates, and the discontinuation rates due to adverse events between the two groups. The Mantel–Haenszel technique was applied to calculate all pooled RRs of dichotomous data with 95% CIs.

In a systematic review, the outcomes are synthesized by using either the fixed- or random-effects model. When relying on the fixed-effects model, all of the eligible studies expected that the true effect size was the same in all trials, and the summary effect was the estimation of the common effect size. Hence, after the individual study was weighted, the outcomes of smaller studies could be ignored due to the better outcomes regarding the same effect size in the larger studies provided. In this event, a fixed-effects model may be used. In fact, the assumption of one true effect size is hardly plausible. Even if all the included studies were relatively homogenous, it was not possible to conclude that they were
absolutely identical. By this reason, a random-effects model, assuming that the true effect size varies across the studies, was applied to synthesize all the outcomes in this review. The RevMan 5.1 was applied through synthesis to all of the results in this systematic review.

**Risk of bias across studies**
If possible, a funnel plot is used for evaluating the reporting bias. A funnel plot is a simple scatter plot of the treatment effect estimated from each study against a measure of each study’s size. If bias does not occur, the plot should resemble a symmetrical inverted funnel.

**Test of heterogeneity**
A test of heterogeneity can examine the similarities of clinical outcomes. When the test was performed in this systematic review, we hypothesized that the effect size had differences because of the differences in the quality of the methodology in each study. The findings of all studies were determined to whether they were higher and different from the anticipated outcomes by chance alone. To evaluate those findings, we reviewed them by showing them as graphs and also used the test of heterogeneity. In the case of an $I^2$ of 50% or more, those findings were noted as having significance of heterogeneity.

**Results**

**Study selection**
According to searching such databases, there was a total of 160 citations (EMBASE = 35, PubMed = 34, CINAHL = 15, Cochrane Controlled Trials Register = 68, ClinicalTrials.gov = 5, and EU-CTR = 3; Figure 1). After the duplicates were removed, 129 citations persisted. When their titles and abstracts were assessed, seven citations still met the eligibility criteria. Therefore, full citations of seven citations were examined. Of the seven citations, four were excluded from

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![Diagram](https://example.com/diagram.png)

**Figure 1** Flow diagram of the study.

**Abbreviation:** EU-CTR, EU Clinical Trials Register.
this review since they were studied in an adult population.\textsuperscript{23–26} Consequently, three clinical trials were included in this meta-analysis.\textsuperscript{27–36} However, a relevant or unpublished study fitting the included criteria was not detected.

### Study characteristics

All eligible studies included patients with bipolar I or II associated with depressive episode with a CDRS-R score of $\geq 40–45$. The duration of included clinical trials was 8 weeks. Each randomized patient was assigned to receive either quetiapine or placebo treatment. The response criterion was similar in each of the eligible studies. The dose of quetiapine was 150–600 mg/day (Table 1). The demographic and basic characteristics of the quetiapine-treated group vs the placebo-treated group were generally well matched across the three studies.\textsuperscript{6–8}

A total of 251 randomized participants were enrolled in this review. All the included patients met a depressive episode of either bipolar I or II disorder based on the criteria of the DSM, Fourth Edition, Text Revision (DSM-IV-TR). The mean (SD) ages of the quetiapine-treated group and the placebo-treated group were 14.5 (2.3) years and 14.2 (2.1) years, respectively. The basic characteristics of three eligible trials have been presented in Table 1.

Since all eligible studies reported the scores of the CDRS-R as a measure of depressive severity, the WMDs of the mean-changed scores could be applied to estimate and synthesize the outcome. In addition, these studies also reported the remission, response, and discontinuation rates.

### Risk of bias within studies

The random consequence generation, allocation concealment, binding of participants, and personnel and binding of outcome assessment were not clearly explained in all eligible trials (Figures 2 and 3). Two studies used the intention-to-treat analysis,\textsuperscript{6,7} and one used the modified intention-to-treat analysis.\textsuperscript{8}

### Table 1 The basic characteristics of controlled trials of quetiapine vs placebo in bipolar disorder

<table>
<thead>
<tr>
<th>Study</th>
<th>Number of randomized patients</th>
<th>Age of subjects (years)</th>
<th>Study duration (weeks)</th>
<th>Drug/dose</th>
<th>Diagnostic criteria</th>
<th>Response criteria</th>
<th>Remission criteria</th>
<th>Outcome measures</th>
</tr>
</thead>
<tbody>
<tr>
<td>DelBello et al\textsuperscript{6}</td>
<td>32</td>
<td>12–18</td>
<td>8</td>
<td>Quetiapine/ 300–600 mg/day</td>
<td>DSM-IV-TR</td>
<td>$\geq$50% reduction in the CDRS-R total score</td>
<td>$\leq$28 in the CDRS-R scores and CGI-BP-I $\geq$2</td>
<td>CDRS-R, CGI-BP-I, CGI-BP-S, HAM-A, YMRS, CGi-BP-i, CGi-C</td>
</tr>
<tr>
<td>Chang et al\textsuperscript{7}</td>
<td>26</td>
<td>12–18</td>
<td>8</td>
<td>Quetiapine/ 300–600 mg/day</td>
<td>DSM-IV-TR</td>
<td>$\geq$50% reduction in the CDRS-R total score</td>
<td>$\leq$28 in the CDRS-R scores and CGI-BP-I $\geq$2</td>
<td>CDRS-R, CGI-BP-S, CGI-I, CGi-S</td>
</tr>
<tr>
<td>Findling et al\textsuperscript{8}</td>
<td>193</td>
<td>10–17</td>
<td>8</td>
<td>Quetiapine XR/ 150–300 mg/day</td>
<td>DSM-IV-TR</td>
<td>$\geq$50% reduction in the CDRS-R total score</td>
<td>$\leq$28 in the CDRS-R scores and CGI-BP-I $\geq$2</td>
<td>CDRS-R, CGI-BP-S, CGI-BP-C</td>
</tr>
</tbody>
</table>

analysis. Based on the GRADE assessment, all outcome qualities were moderate to high (Table 2).

**Synthesis of results**

**Efficacy**

The significance of heterogeneity was not found in the WMDs for the pooled mean-changed scores of the CDRS-R and the CGI-BP-S and response and remission rates. The pooled WMD for the mean-changed score of the CDRS-R total and the CGI-BP-S scores in the quetiapine- and placebo-treated groups had no significant differences with WMD (95% CI) of $-1.82 (-5.98, 2.34)$, $I^2=0\%$ and $-0.29 (-0.67, 0.09)$, $I^2=0\%$, respectively (Figures 4 and 5). Similarly, the overall pooled response and remission rate between the two groups were not also significantly different with RRs (95% CI) of $1.10 (0.89, 1.35)$, $I^2=0\%$ and $1.23 (0.90, 1.68)$, $I^2=0\%$, respectively (Figures 6 and 7).

**Overall discontinuation rate (acceptability)**

A significance of heterogeneity was not noted in the overall discontinuation rate since the pooled overall discontinuation rate was not significantly different between the two groups with RRs (95% CI) of $0.73 (0.36, 1.49)$, $I^2=47\%$, it indicated that acceptability of quetiapine was not better than that of placebo in the treatment of children and adolescents with bipolar depression.

**Discontinuation rate due to adverse events (tolerability)**

The significance of heterogeneity was not found in the discontinuation rate due to adverse events between the two groups. The pooled discontinuation rate due to adverse events of the quetiapine- and placebo-treated groups having no significant difference (RR [95% CI] of $0.33 [0.11, 1.01]$, $I^2=0\%$) indicated that the tolerability of quetiapine was comparable with placebo in the treatment of children and adolescents with bipolar depression.

**Risk of bias across studies**

If a systematic review includes the RCTs of <10 studies, a funnel plot examining the publication bias may not have enough power to identify the chances of real asymmetry.
Quetiapine monotherapy vs placebo in the treatment of bipolar depression

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>Quetiapine Mean (SD)</th>
<th>Placebo Mean (SD)</th>
<th>Weight (%)</th>
<th>RR M–H, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>DelBello et al</td>
<td>–19 (14)</td>
<td>–20 (17)</td>
<td>14.6</td>
<td>1.00 (0.98, 1.18)</td>
</tr>
<tr>
<td>Finding et al</td>
<td>–29.6 (15.8)</td>
<td>–27.3 (16)</td>
<td>85.4</td>
<td>–2.30 (–6.80, 2.20)</td>
</tr>
<tr>
<td>Total (95% CI)</td>
<td>109</td>
<td>115</td>
<td>100</td>
<td>–1.82 (–5.98, 2.34)</td>
</tr>
</tbody>
</table>

Heterogeneity: $\chi^2=0.00, \chi^2=0.30, df=1 (P=0.58); I^2=0$

Test for overall effect: $Z=0.86 (P=0.39)$

Figure 4: The forest plot of comparison of the mean-changed scores from baseline of a CDRS-R scale (95% CI) of quetiapine vs placebo in children and adolescents with bipolar depression.

Abbreviations: CDRS-R, Children’s Depression Rating Scale–Revised; CI, confidence interval; df, degrees of freedom; IV, inverse variance; SD, standard deviation.

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>Quetiapine Mean (SD)</th>
<th>Placebo Mean (SD)</th>
<th>Weight (%)</th>
<th>RR M–H, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>DelBello et al</td>
<td>–1.8 (1.8)</td>
<td>–1.6 (1.6)</td>
<td>10.2</td>
<td>–0.20 (–1.38, 0.98)</td>
</tr>
<tr>
<td>Finding et al</td>
<td>–1.7 (1.4)</td>
<td>–1.4 (1.4)</td>
<td>89.8</td>
<td>–0.30 (–0.70, 0.10)</td>
</tr>
<tr>
<td>Total (95% CI)</td>
<td>109</td>
<td>115</td>
<td>100</td>
<td>–0.29 (–0.67, 0.09)</td>
</tr>
</tbody>
</table>

Heterogeneity: $\chi^2=0.00, \chi^2=0.02, df=1 (P=0.87); I^2=0$

Test for overall effect: $Z=1.51 (P=0.13)$

Figure 5: The forest plot of comparison of the mean changes from baseline of CGI-BP-S (95% CI) of quetiapine vs placebo in children and adolescents with bipolar depression.

Abbreviations: CGI-BP-S, Clinical Global Impression–Bipolar Version Severity; CI, confidence interval; IV, inverse variance; df, degrees of freedom; SD, standard deviation.

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>Quetiapine Events</th>
<th>Placebo Events</th>
<th>Weight (%)</th>
<th>RR M–H, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>DelBello et al</td>
<td>12</td>
<td>10</td>
<td>19.2</td>
<td>1.06 (0.66, 1.70)</td>
</tr>
<tr>
<td>Chang et al</td>
<td>5</td>
<td>5</td>
<td>4.7</td>
<td>0.63 (0.24, 1.62)</td>
</tr>
<tr>
<td>Finding et al</td>
<td>58</td>
<td>55</td>
<td>76.2</td>
<td>1.15 (0.90, 1.45)</td>
</tr>
<tr>
<td>Total (95% CI)</td>
<td>125</td>
<td>125</td>
<td>100</td>
<td>1.10 (0.89, 1.35)</td>
</tr>
</tbody>
</table>

Total events 75

Heterogeneity: $\chi^2=0.00, \chi^2=1.50, df=2 (P=0.47); I^2=0$

Test for overall effect: $Z=0.88 (P=0.38)$

Figure 6: The forest plot of comparison of RR M–H (95% CI) for clinical response rates of quetiapine vs placebo in children and adolescents with bipolar depression.

Note: Rounding for weight (%).

Abbreviations: CI, confidence interval; df, degrees of freedom; M–H, Mantel–Haenszel; RR, relative risk.

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>Quetiapine Events</th>
<th>Placebo Events</th>
<th>Weight (%)</th>
<th>RR M–H, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>DelBello et al</td>
<td>6</td>
<td>6</td>
<td>12.1</td>
<td>0.88 (0.36, 2.16)</td>
</tr>
<tr>
<td>Chang et al</td>
<td>6</td>
<td>4</td>
<td>9.9</td>
<td>0.94 (0.35, 2.52)</td>
</tr>
<tr>
<td>Finding et al</td>
<td>42</td>
<td>34</td>
<td>77.9</td>
<td>1.34 (0.94, 1.91)</td>
</tr>
<tr>
<td>Total (95% CI)</td>
<td>125</td>
<td>125</td>
<td>100</td>
<td>1.23 (0.90, 1.68)</td>
</tr>
</tbody>
</table>

Total events 54

Heterogeneity: $\chi^2=0.00, \chi^2=1.06, df=2 (P=0.59); I^2=0$

Test for overall effect: $Z=1.31 (P=0.19)$

Figure 7: The forest plot of comparison of RR M–H (95% CI) for clinical remission rates of quetiapine vs placebo in children and adolescents with bipolar depression.

Note: Rounding for weight (%).

Abbreviations: CI, confidence interval; df, degrees of freedom; M–H, Mantel–Haenszel; RR, relative risk.
occuring. Hence, the test of funnel plot was not carried out in this review since this review included only three studies.

Discussion
The findings from this systematic review suggested that quetiapine monotherapy as measured by the mean-changed score and response rate was not effective in the acute treatment of bipolar depression in children and adolescents. Similarly, its remission rate was not higher than that of placebo. Quetiapine monotherapy could not reduce the severity of depressive symptoms measured by clinical global impression. In addition, its acceptability and tolerability were also not better than those of placebo.

The previous systematic review found that quetiapine monotherapy was an effective treatment in adult patients with bipolar depression, while our findings did not support those results. The different outcomes between adults and child and adolescent populations may be caused by either high comorbid psychopathology, such as attention-deficit hyperactivity disorder and conduct disorder, in children and adolescents with bipolar I or II disorder associated with poor response treatment or methodological limitations of the included studies.

Based on the previous evidence, antidepressant effect is associated with the norquetiapine/quetiapine ratio, quetiapine extended release may have a different antidepressant profile compared with quetiapine immediate release. In this review, two of three included studies administered immediate-release quetiapine which may be associated with the negative outcomes. Further studies should be carried out to confirm those results.

Recent evidence suggested that the acceptability and tolerability of quetiapine was not higher than placebo and selective serotonin reuptake inhibitors, which was compatible with our findings. The patients have experienced some adverse events such as headache, sedation, dizziness, somnolence, diarrhea, fatigue, and nausea. In addition, previous studies have supported that the development of obesity and metabolic syndrome is possibly associated with the use of second-generation antipsychotics, including quetiapine, while obesity is one of the risk factors for the development of rapid cycling in patients with bipolar disorder. For this reason, use of quetiapine should be not recommended in those patients.

This systematic review had some limitations. First, since there were a small number of included studies in this review, the pooled sample size may be affected. Further reviews including augmentation trials with quetiapine in such patients may be useful. Second, the eligible trials were funded by a patent-holding company for quetiapine. Interpretation of those results should be viewed with caution. Third, this review did not include patients with bipolar depression and disruptive mood dysregulation disorder diagnosed by using the DSM, Fifth Edition (DSM-5). Fourth, this review used narrow search terms as a previous review; therefore, the sensitivity of searching possibly decreased. Finally, selection, detection, and reporting biases of all the included studies were performed and the test of funnel plot to assess an asymmetry was unable to be carried out because of a limited number of eligible clinical studies. Hence, in this review, publication bias was impossible to be excluded.

Conclusion
According to the evidence from this review, quetiapine as monotherapy is inefficacious in the acute treatment of children and adolescents with bipolar depression. In addition, its acceptability and tolerability were not better than those of placebo. Hence, the use of quetiapine in such patients should be avoided. Further well-defined studies should be performed to validate these findings.

Acknowledgments
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Author contributions
All authors contributed toward data analysis, drafting and critically revising the paper and agree to be accountable for all aspects of the work.

Disclosure
BM has received honoraria and/or travel reimbursement from Lundbeck and Pfizer. NM has received travel reimbursement from Lundbeck and Pfizer. SS has received honoraria and/or research grants from Janssen-Cilag, Thai-Otsuka, Lundbeck, and AstraZeneca. CC received honoraria and/or travel reimbursement from Janssen-Cilag, GlaxoSmithKline, and Thai-Otsuka. MS has received honoraria, consultancy fees, research grants, and/or travel reimbursement from AstraZeneca, GlaxoSmithKline, Pfizer, Janssen-Cilag, Johnson & Johnson, Lundbeck, Thai-Otsuka, Sanofi-Aventis, and Servier. SP and PW report no conflicts of interest in this work.

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