Prognostic value of FOXQ1 in patients with malignant solid tumors: a meta-analysis

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Background: Forkhead box Q1 (FOXQ1, also known as HFH1), a member of the forkhead transcription factor family, has been demonstrated to be overexpressed in multiple tumors and is thought to be an indicator of poor clinical outcomes.

Methods: A meta-analysis using qualified relevant literature was performed to evaluate the prognostic significance of FOXQ1 in various malignant solid tumors. A search of electronic databases was conducted in MEDLINE, Embase, and the Cochrane Library to identify relevant studies published from 1966 to July 30, 2016, and the studies were identified by further evaluation. The pooled hazard ratios (HRs) with 95% confidence intervals (CIs) for analyses were assessed to investigate the association between FOXQ1 expression and overall survival (OS) of patients with malignant solid tumors.

Results: A total of 1,520 patients from six studies (seven cohorts) with multiple malignant solid tumors were included. For OS, high FOXQ1 expression could significantly predict worse outcome with the pooled HR of 1.38 (95% CI: 1.17–1.59; P<0.001). The subgroup analysis suggested that the elevated levels of FOXQ1 appear to be associated with worse OS in hepatocellular carcinoma (HR =1.34; 95% CI: 1.11–1.57; P<0.001) and other cancers (HR =1.62; 95% CI: 1.09–2.14; P<0.001).

Conclusion: This meta-analysis indicated that the high expression of FOXQ1 is associated with an adverse OS in malignant solid tumors, suggesting that FOXQ1 may be a predictor of poor prognosis for the development of malignant solid tumors.

Keywords: FOXQ1, meta-analysis, cancer, prognosis, overall survival

Introduction
Forkhead box Q1 (FOXQ1, also known as HFH1) is a member of the forkhead transcription factor family. FOXQ1 is located at human chromosome 6p25.3 and consists of 2,319 bp. It is encoded by an open reading frame of 1,029 bp and produces a functional protein with 403 amino acids. It is expressed in different tissues, including stomach, salivary gland, duodenum, prostate, fetal liver tissue, bladder, and trachea, and plays important roles in development, metabolism, aging, and cancer. Initially, it has been shown to be involved in hair follicle differentiation, gastrulation, and mucin production in mice. Several recent studies demonstrated that FOXQ1 expression increased in many human cancers, including hepatocellular carcinoma (HCC), colorectal cancer, esophageal cancer, breast cancer, and non-small-cell lung cancer, and its expression correlated with metastasis and poor prognosis.

Therefore, the prognostic value of FOXQ1 overexpression across different malignant solid tumors is still unclear. Consequently, we initiated a meta-analysis to investigate the significance of increased FOXQ1 expression in the prediction of clinical outcomes of patients with solid tumors.
Methods

Search strategy

A search of electronic databases was conducted in MEDLINE, Embase, and the Cochrane Library to identify relevant studies published from 1966 to July 30, 2016, using the random combination of the following three sets of keywords: “forkhead box Q1 OR FOXQ1 OR HFH1”, “carcinoma OR cancer OR malignancy OR tumor”, and “prognosis OR survival OR outcome”. In addition, we manually searched references of the selected articles for the potential articles that have been missed at the first search.

Inclusion and exclusion criteria

The following criteria were applied to include the eligible studies: 1) human-based investigations, 2) must be a pathologically confirmed certain type of solid tumor, 3) articles published in English with full texts, 4) usage of immunohistochemistry (IHC) method to investigate the FOXQ1 expression in the primary tumor tissues, 5) reporting the correlation between FOXQ1 expression and overall survival (OS), and 6) availability of hazard ratios (HRs) and their 95% confidence intervals (CIs) directly or indirectly by calculating from the original data. The following ineligible studies were excluded: 1) studies conducted using cell lines and animals, 2) the types of case reports, letters, reviews, or meta-analyses, 3) studies published in non-English, 4) studies lacking necessary data, and 5) overlapped studies or duplicated data.

Data extraction

Data were extracted from each study in duplicate by two independent researchers (Xiaohai Cui and Jia Zhang) based on the inclusion and exclusion criteria described earlier. Disagreement was settled by consulting with a third investigator (Jing Zhang). Extracted data from eligible studies included the first author’s surname, year of publication, country, histology, number of patients, assay method, positive percentage, the sources of HR (95% CI), and survival results. HRs and 95% CIs were obtained from the original article in all cases.

Quality assessment

Newcastle–Ottawa scale (NOS) was used to perform the quality assessment of each study. Based on three items, including election, comparability, and outcome, the total score of NOS ranged from 0 to 9. A score of >6 is considered as high quality.

Statistical analysis

Chi-square test-based Cochran’s Q and \( I^2 \) statistics were used to calculate the heterogeneity of the individual HR. For Q statistics, \( P<0.05 \) was considered as statistically significant. For \( I^2 \) statistics, \( I^2 <25\% \) means no heterogeneity, \( 25\% <I^2 <50\% \) means moderate heterogeneity, and \( I^2 >50\% \) means strong heterogeneity. If no obvious heterogeneity was found among studies, the fixed effects model was used to combine the individual HR estimates; otherwise, the random effects model was applied. For FOXQ1 overexpression groups, HR >1 indicated worse survival. If there was no overlap between 95% CI and HR =1, it would be considered that the impact of FOXQ1 on survival was statistically significant. Sensitivity analysis was carried out to validate the credibility of the meta-analysis outcomes by sequentially excluding one study at a time. If the results did not obviously change when one study was removed, the sensitivity is low, and the results are robust. Publication bias was assessed through Begg’s test and Egger’s test, and \( P<0.05 \) was recognized as statistically significant. All statistical analyses were conducted using the STATA12.0 software (Stata Corporation, College Station, TX, USA). Ethical approval was not required as this study does not report on or involve the use of any animal or human data or tissue.

Results

Study selection and characteristics

The initial search identified 95 relevant studies. Based on the screening of the titles and abstracts of all studies, 85 studies were retrieved according to the inclusion criteria, leaving 10 articles available for full-text review. After carefully reviewing and assessing, four studies were removed due to various reasons. Therefore, six studies containing seven cohorts were finally included in this meta-analysis. Figure 1 shows the process of study selection.

![Figure 1: Selection flowchart of the meta-analysis.](https://www.dovepress.com/)

95 potentially relevant studies identified on the basis of key searching terms

85 studies were excluded by title or abstract

10 full-length articles assessed for eligibility

4 studies were excluded: 1 insufficiency information 1 data overlapping 2 without outcome of interest

6 studies included for final analysis

For Q statistics, \( P<0.05 \) was considered as statistically significant. For \( I^2 \) statistics, \( I^2 <25\% \) means no heterogeneity, \( 25\% <I^2 <50\% \) means moderate heterogeneity, and \( I^2 >50\% \) means strong heterogeneity. If no obvious heterogeneity was found among studies, the fixed effects model was used to combine the individual HR estimates; otherwise, the random effects model was applied. For FOXQ1 overexpression groups, HR >1 indicated worse survival. If there was no overlap between 95% CI and HR =1, it would be considered that the impact of FOXQ1 on survival was statistically significant. Sensitivity analysis was carried out to validate the credibility of the meta-analysis outcomes by sequentially excluding one study at a time. If the results did not obviously change when one study was removed, the sensitivity is low, and the results are robust. Publication bias was assessed through Begg’s test and Egger’s test, and \( P<0.05 \) was recognized as statistically significant. All statistical analyses were conducted using the STATA12.0 software (Stata Corporation, College Station, TX, USA). Ethical approval was not required as this study does not report on or involve the use of any animal or human data or tissue.
selection for this meta-analysis. The main characteristics of the six included studies containing seven cohorts are extracted and summarized in Table 1. Eventually, a total of 1,520 patients with malignant solid tumors were retained for subsequent meta-analysis. These six studies of good quality were published between 2012 and 2016. All the studies were conducted in China. All the studies examined the prognostic value of FOXQ1 overexpression in patients with solid tumors. Five of the six studies reported HRs (multivariate analysis) explicitly. An HR on OS could be extracted for our original data. All studies identified FOXQ1 overexpression as an indicator of poor OS.

### Table 1 Main characteristics of all studies included in the meta-analysis

<table>
<thead>
<tr>
<th>First author (year)</th>
<th>Patient source</th>
<th>Histology</th>
<th>Number of patients</th>
<th>Method</th>
<th>Positive (%)</th>
<th>HR estimation</th>
<th>Survival results</th>
</tr>
</thead>
<tbody>
<tr>
<td>Xia et al16 Cohort I</td>
<td>China</td>
<td>HCC</td>
<td>690</td>
<td>IHC</td>
<td>54.5</td>
<td>HR and 95% CI: 1.27 (1.03–1.56)</td>
<td>Poor</td>
</tr>
<tr>
<td>Xia et al16 Cohort II</td>
<td>China</td>
<td>HCC</td>
<td>312</td>
<td>IHC</td>
<td>53.8</td>
<td>HR and 95% CI: 1.48 (1.09–2.01)</td>
<td>Poor</td>
</tr>
<tr>
<td>Zhan et al11</td>
<td>China</td>
<td>PC</td>
<td>83</td>
<td>IHC</td>
<td>41.0</td>
<td>HR and 95% CI: 1.856 (1.063–3.234)</td>
<td>Poor</td>
</tr>
<tr>
<td>Feng et al13</td>
<td>China</td>
<td>NSCLC</td>
<td>103</td>
<td>IHC</td>
<td>79.6</td>
<td>HR and 95% CI: 2.091 (1.009–4.332)</td>
<td>Poor</td>
</tr>
<tr>
<td>Liang et al18</td>
<td>China</td>
<td>GC</td>
<td>158</td>
<td>IHC</td>
<td>53.8</td>
<td>HR and 95% CI: 1.286 (0.746–2.248)</td>
<td>Poor</td>
</tr>
<tr>
<td>Zhang et al22</td>
<td>China</td>
<td>GC</td>
<td>60</td>
<td>IHC</td>
<td>63.3</td>
<td>HR and 95% CI: 1.942 (1.057–3.567)</td>
<td>Poor</td>
</tr>
<tr>
<td>Wang et al15</td>
<td>China</td>
<td>HCC</td>
<td>114</td>
<td>IHC</td>
<td>68.4</td>
<td>HR and 95% CI: 2.691 (1.364–5.309)</td>
<td>Poor</td>
</tr>
</tbody>
</table>

Abbreviations: CI, confidence interval; GC, gastric cancer; HR, hazard ratio; HCC, hepatocellular carcinoma; IHC, immunohistochemistry; NSCLC, non-small-cell lung cancer; PC, pancreatic cancer.

### Meta-analysis

The results of the meta-analysis regarding the correlation between FOXQ1 overexpression and OS are shown in Figure 2. Overall, the combined HR for all six eligible studies (seven cohorts) evaluating FOXQ1 expression and OS indicated that high FOXQ1 expression was associated with poor prognosis of solid tumors for OS (HR = 1.38; 95% CI: 1.17–1.59; P < 0.001; random effects). No significant heterogeneity was present among all included cohorts ($I^2 = 0.0\%$, $P = 0.568$).

In subgroup analyses (Figure 2), three cohorts related to HCC were divided into the same group, with the combined...
HR = 1.34 (95% CI: 1.11–1.57). There was a statistical significance between the expression of FOXQ1 and the prognosis of HCC \((P < 0.001)\). There was no heterogeneity among various studies \(\left(I^2 = 17.6\%, \ P = 0.297\right)\). Four cohorts related to other solid tumors were divided into others, with the combined HR = 1.62 (95% CI: 1.09–2.14). There was a significant statistical significance between the expression of FOXQ1 and the prognosis of others \(\left(P < 0.001\right)\). There was also no heterogeneity among various studies \(\left(I^2 = 0.00\%, \ P = 0.682\right)\).

### Sensitivity analysis

To test the robustness of the HR estimates, sensitivity analysis was performed to evaluate the impact of individually excluding studies on the final result of meta-analysis. As shown in Figure 3, the corresponding pooled HRs were not significantly altered, suggesting the robustness of our results.

### Publication bias

The interpretability of publication bias assessed by Begg’s and Egger’s tests was limited when only six studies were included in this meta-analysis.

### Discussion

As a transcription factor, forkhead family protein plays important roles in cell cycle regulation, embryonic development, lipid and carbohydrate metabolism, aging, immune regulation, and other biological processes.\(^4,23,24\) As a member of forkhead family protein, in mammals, FOXQ1 was initially found to take part in hair follicle development.\(^5\) Thereafter, it was found to be overexpressed in various cancers.\(^9–19\) Its upregulation enhances tumor growth and invasion.\(^10,12,14,16\) Moreover, increases in FOXQ1 are predictive of poor prognosis in patients with various malignant solid tumors.\(^9–19\) However, thus far, no meta-analyses have been carried to evaluate the prognostic value of increased FOXQ1 in patients with solid tumors. To the best of our knowledge, this is the first comprehensive meta-analysis of the effects of increased FOXQ1 expression on the survival of malignant solid tumors.

In this meta-analysis, we performed a meta-analysis to evaluate the prognostic value of FOXQ1 by combining six eligible studies including 1,520 malignant solid tumor patients. The data were organized according to OS. Combined HRs indicated that high FOXQ1 expression was associated with poor prognosis of OS \((HR = 1.38; 95\% CI: 1.17–1.59)\) in patients with solid tumors. The subgroup analysis revealed that the elevated levels of FOXQ1 appear to be worse OS in HCC \((HR = 1.34; 95\% CI: 1.11–1.57; P < 0.001)\) and other cancers \((HR = 1.62; 95\% CI: 1.09–2.14; P < 0.001)\).

Publication bias was a problem to be analyzed in all of the meta-analysis. In this study, the interpretability of publication bias assessed by Begg’s and Egger’s tests was limited when only six studies were included in this meta-analysis;

### Figure 3

Results of sensitivity analysis by omitting each study at a time.

**Notes:** Rows represent the results of meta-analysis of all studies except the omitted study named in that row. Omission of any study did not affect the whole estimate results significantly.

**Abbreviation:** CI, confidence interval.
so, we did not perform Begg’s and Egger’s tests. Sensitivity analyses further revealed the robustness of the meta-analysis findings.

Several limitations should be taken into consideration when interpreting the results of our meta-analysis. 1) Only six relevant articles of seven cohorts were only involved with 1,520 patients and all designed in China. 2) The low number of included cohorts limited our evaluation of a potential publication bias. Therefore, publication bias may exist. 3) The method used for the assessment of FOXQ1 expression is IHC, and experimental IHC protocols could have confounded the results, and the cutoff definition of FOXQ1 overexpression may led to between-study heterogeneity.

**Conclusion**

Our study indicated that overexpression of FOXQ1 may lead to poor prognosis. With the gradual deepening of the study, FOXQ1 may become an important target for cancer therapy in the future.

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**Disclosure**

The authors report no conflicts of interest in this work.

**References**