Vitamin D deficiency and the risk of tuberculosis: a meta-analysis

Background and aim: To conduct meta-analyses of all published studies on various aspects of association between vitamin D and tuberculosis (TB).

Methods: PubMed and Web of Knowledge were searched for all properly controlled studies on vitamin D and TB. Pooled odds ratio, mean difference, or standardized mean difference, and its corresponding 95% confidence interval were calculated with the Cochrane Review Manager 5.3.

Results: A significantly lower vitamin D level was found in TB patients vs controls; vitamin D deficiency (VDD) was associated with an increased risk of TB, although such an association was lacking in the African population and in the human immunodeficiency virus-infected African population. A significantly lower vitamin D level was found in human immunodeficiency virus-TB-coinfected African patients receiving antiretroviral treatment who developed TB-associated immune reconstitution inflammatory syndrome vs those who did not develop TB-associated immune reconstitution inflammatory syndrome. VDD was associated with an increased risk of developing active TB in those subjects with latent TB infection and with an increased risk of tuberculin skin test conversion/ TB infection conversion, and the trend toward a lower vitamin D level in active TB patients vs latent TB infection subjects did not reach statistical significance, indicating that VDD was more likely a risk factor than a consequence of TB. This concept was further strengthened by our result that anti-TB treatment did not affect vitamin D level in TB patients receiving the treatment.

Conclusion: Our analyses revealed an association between vitamin D and TB. VDD is more likely a risk factor for TB than its consequence. More studies are needed to determine whether vitamin D supplementation is beneficial to TB prevention and treatment.

Keywords: vitamin D, vitamin D deficiency, tuberculosis, 25-hydroxyvitamin D, meta-analysis

Introduction

Tuberculosis (TB) remains a major public health problem globally. It was estimated that there were 8.7 million reported new TB cases and that 1.4 million people died from TB in 2011.1 Various factors that could possibly affect the incidence and progression of TB have been reported, one of them is vitamin D deficiency (VDD).1–6 Exposure to sunlight is the main source of vitamin D for human and induces the conversion of 7-dehydrocholesterol to vitamin D3 via previtamin D3 in the skin.7 Vitamin D3 is then converted to 25-hydroxyvitamin D (25(OH)D) in the liver and is further converted to the bioactive form of vitamin D, 1,25-dihydroxyvitamin D3 (1,25(OH)2D3) in the kidney.7 Studies have found that 1,25(OH)2D3 binds to vitamin D receptor (VDR), activates VDR signaling, and induces a series of antimicrobial responses such as induction of autophagy, phagolysosomal fusion, release and activation of the antimicrobial peptide cathelicidin, and killing of intracellular Mycobacterium tuberculosis.7–9

Shao-Jun Huang1
Xian-Hua Wang2
Zhi-Dong Liu1
Wen-Li Cao3
Yi Han1
Ai-Guo Ma2
Shao-Fa Xu1

1Department of Thoracic Surgery, Beijing Chest Hospital, Capital Medical University, Beijing, People’s Republic of China; 2Department of Nutrition and Food Hygiene, Institute of Human Nutrition, Medical College of Qingdao University, Qingdao, People’s Republic of China; 3Department of Pulmonary Medicine, Beijing Geriatric Hospital, Beijing, People’s Republic of China
Numerous studies have been conducted to study whether VDD was associated with TB; however, they produced inconsistent and varying results.¹⁻⁶,¹⁰⁻¹² There has been one meta-analysis on association between VDD and TB, and it concluded that a lower serum vitamin D level was associated with a higher risk of active TB.¹³ However, it was conducted in 2008 and did not include numerous later studies, and further it did not perform a subgroup analysis based on factors such as ethnicity.

In this study, we performed meta-analyses comparing serum vitamin D levels in TB cases vs controls and on association between VDD and risk of TB with subgroup analysis based on the ethnicity of the study population in order to further elucidate the role of vitamin D level in TB pathogenesis in a larger population pool. Second, assuming VDD is indeed associated with TB, it is still undetermined whether VDD is one of the causes of TB or manifestation of the malnutrition status generally seen in TB patients.¹ In order to explore this question, we further examined the association between vitamin D level/VDD and risk of developing active TB in subjects with latent TB infection (LTBI)/household contacts of TB patients and also the association between VDD and tuberculin skin test (TST) conversion/TB infection conversion (TBIC). An additional analysis was also performed on the effect of anti-TB treatment on vitamin D level in TB subjects for the same purpose. Third, since human immunodeficiency virus (HIV)-infected subjects are more susceptible to TB and have greater mortality and morbidity caused by TB,¹⁴ to further clarify whether vitamin D played any role in occurrence and deterioration of TB in HIV-infected subjects, we also analyzed the association between vitamin D level/VDD and TB in HIV-infected subjects and the association between vitamin D level and risk of developing TB-associated immune reconstitution inflammatory syndrome (TB-IRIS) in HIV-TB-coinfected patients receiving antiretroviral therapy (ART). Finally, although serum 25(OH)D is the commonly used measure of vitamin D status in a subject,¹ ¹ ² 5(OH)2D3 is the actual bioactive form of vitamin D that induces antimicrobial response,⁷ and so we also performed a meta-analysis on the association between 1,25(OH)2D3 and TB in order to determine whether there was any change of 1,25(OH)2D3 level in TB subjects compared to control without TB.

Methods
Search strategy, inclusion and exclusion criteria, and data extraction
PubMed/Medline and Web of Science (Web of Knowledge) databases were searched to identify all studies relating to vitamin D and TB using the following search terms: “vitamin D”, “VDD”, “hypovitaminosis D”, “tuberculosis”, “TB”, “25-hydroxyvitamin D”, and “1,25-dihydroxyvitamin D”.

All properly controlled studies published in English on vitamin D and TB were potentially eligible for inclusion in our meta-analysis. We excluded reviews, meeting abstracts, and case-only studies.

Relevant data from each included study were extracted. Among the data extracted, relevant group meanvalues (standard deviations, SDs) were extracted from each included study whenever possible. When relevant values were presented only as median (interquartile range), mean (SD) was approximated as median (interquartile range/1.35);¹⁵ when relevant values were presented only as median (range), mean (SD) was approximated as median (range/4);¹⁶ when relevant values were presented only as mean values (standard errors, SEs), SD was extracted as SE × square root of sample size;¹⁷ and when relevant values were presented only as meanvalues (confidence intervals, CIs), SD was extracted as square root of sample size × (upper limit – lower limit)/(2 × Y), where Y is the t value from a size equal to sample size minus 1.¹⁷

Quality of the studies
Each study’s adequacy in three key areas (methodological, clinical, and statistical) was evaluated first. Then each study was screened with the Newcastle–Ottawa Scale.¹⁸

Statistical analysis
Cochrane Review Manager (RevMan 5.3, Copenhagen, Denmark, The Nordic Cochrane Centre) was used to perform all statistical analysis. Continuous outcomes were analyzed using mean difference (MD) or standardized MD (SMD), while dichotomous outcome was analyzed using pooled odds ratio (OR). Pooled OR and MD/SMD with their corresponding 95% CI were calculated with a random-effect model because this model assumes a genuine diversity in the results caused by between-studies heterogeneity and incorporates a between-studies variance into the calculation accordingly.¹⁹ The Z test was used to assess the statistical significance of the pooled OR and MD/SMD with their corresponding 95% CI. Between-studies heterogeneity was assessed with the Cochran Q statistic-based chi-square (χ²) test,²⁰ and F index was the tool used to assess the heterogeneity, where a F value around 25%, 50%, and 75% indicated a low, moderate, and large heterogeneity, respectively.²¹ Finally, if no heterogeneity was found using the random-effect model (F=0), data were analyzed again using a fixed-effect model. Statistical significance was considered with a P-value of <0.05, except for the Q statistic where a P-value of <0.10 indicated statistically significance.
Subpopulation analysis was performed according to the ethnicity of the subjects. In addition, funnel plots were used to evaluate the publication bias in each meta-analysis.

**Results**

**Eligible studies and characteristics of the studies**

Figure 1 illustrates our search process and results. From a total of 723 potentially eligible articles, 685 were excluded because they were irrelevant to our topic here, reviews, meeting abstracts, or case-only studies, or contained no detailed data needed for our meta-analysis. A total of 38 articles were included in our meta-analysis. Vitamin D level was significantly lower in TB patients vs controls, and VDD was associated with an increased risk of TB. A total of 25 studies with 3,599 TB cases and 3,063 control subjects were included in our analysis of serum/plasma 25(OH)D level in TB patients vs control (Figure 2A).1,4,5,11,12,14,22–28,30,35–44. Our analysis with a random-effect model showed a significantly lower serum/plasma 25(OH)D level in TB patients vs controls (MD = -13.05; 95% CI = [-19.02, -7.08]; P < 0.0001; I² = 96) (Figure 2A). Subgroup analysis showed similar result for the Asian population (MD = -14.64; 95% CI = [-20.15, -9.13]; P < 0.00001; I² = 89); however, there was no such significant

**Figure 1 Flow diagram of the publication selection process.**
Figure 2 (A) Forest plot of comparison of vitamin D level (serum 25-hydroxyvitamin D [25(OH)D]) in TB patients vs control: overall effect for continuous outcome using a random-effects model. (B) Forest plot of association between VDD and risk of TB: overall effect for dichotomous outcome using a random-effect model. The diamonds stand for pooled effect.

Abbreviations: TB, tuberculosis; vi, vitamin; VDD, vitamin D deficiency; SD, standard deviation; CI, confidence interval; df, degrees of freedom; IV, independent variable.
difference for the African population (MD = −6.05; 95% CI = [−20.54, −8.43]; P = 0.41; F = 98) (Figure 2A). We found a similar result when the same analysis was performed excluding studies where cases and controls were subjects with other diseases such as HIV and diabetes, or where control subjects had LTBI (data not shown).

For our analysis on VDD and risk of TB, a total of 23 studies with 3,491 TB cases and 3,259 control subjects were included (Figure 2B). Our analysis showed that VDD was significantly associated with an increased risk of TB (OR = 2.57; 95% CI = [1.74, 3.80]; P < 0.00001; F = 83) (Figure 2B). Subgroup analysis revealed similar result for the Asian population (OR = 2.62; 95% CI = [1.63, 4.23]; P < 0.0001; F = 71); however, VDD was not associated with an increased risk of TB in the African population (OR = 1.89; 95% CI = [0.82, 4.33]; P = 0.13; F = 91) (Figure 2B). Results were similar when the same analysis was performed excluding studies where cases and controls were subjects with other diseases such as HIV and diabetes, or where control subjects had LTBI (data not shown).

VDD was more likely a risk factor for TB than its consequence

Four studies with 331 active TB cases and 326 LTBI subjects/household contacts of active TB patients were included in our meta-analysis (Figure 3A). We included household contacts of active TB patients as part of our control in this analysis because they were at a high risk of becoming infected and developing TB, and a meta-analysis by Fox et al reported a high prevalence of LTBI in contacts of active TB patients (51.5%; 95% CI = [47.1%, 55.8%]; F = 98.9%). Our analysis revealed a trend toward a lower serum/plasma 25(OH)D level in active TB patients than its consequence (data not shown).

Figure 3 (A) Forest plot of comparison of serum 25(OH)D level in active TB patients vs LTBI subjects/household contacts of active TB patients; overall effect for dichotomous outcome using a random-effect model. The diamonds stand for pooled effect. Abbreviations: TB, tuberculosis; LTBI, latent TB infection; vit, vitamin; VDD, vitamin D deficiency; TST, tuberculin skin test; TBIC, TB infection conversion; SD, standard deviation; CI, confidence interval; df, degrees of freedom; IV, independent variable.
significance (MD = −8.92; 95% CI = [−18.62, 0.79]; P = 0.07; F = 90) (Figure 3A).

VDD was significantly associated with an increased risk of developing active TB in LTBI subjects/household contacts of TB patients

Five studies with 372 active TB patients and 407 LTBI subjects/household contacts of active TB patients were included in our analysis (Figure 3B). 2,4,35,36 VDD was positively and significantly associated with an increased risk of developing active TB in LTBI subjects/household contacts of active TB patients (OR = 4.26; 95% CI = [2.48, 7.30]; P < 0.00001; F = 48%) (Figure 3B).

VDD was significantly associated with an increased risk of TST conversion/TBIC

Our meta-analysis here included three studies with 45 cases with TST conversion/TBIC and 366 control subjects without TST conversion/TBIC (Figure 3C). 49–50 Our analysis using a fixed-effect model revealed significant association between VDD and an increased risk of TST conversion/TBIC (OR = 3.99; 95% CI = [1.88, 8.45]; P = 0.0003; F = 0) (Figure 3C).

Vitamin D, TB, and HIV

No significant vitamin D level difference in HIV-TB-coinfected African patients vs African HIV patients without active TB, and VDD was not associated with an increased risk of TB in African HIV-infected patients

Our meta-analysis comparing vitamin D level in HIV-TB-coinfected patients vs HIV patients without active TB included four studies on the African population with 260 HIV-TB-coinfected cases and 164 control HIV subjects without active TB (Figure 4A). 11,12,14,35 Serum/plasma 25(OH)D level in HIV-TB-coinfected African patients was not

---

**Figure 4** (A) Forest plot of comparison of serum 25(OH)D level in HIV-TB-coinfected patients vs HIV patients without active TB: overall effect for continuous outcome using a random-effect model. (B) Forest plot of association between VDD and risk of TB in HIV-infected patients: overall effect for dichotomous outcome using a random-effect model. (C) Forest plot comparing 25(OH)D level in HIV-TB-coinfected patients receiving ART who developed TB-IRIS vs HIV-TB-coinfected patients receiving ART who did not develop TB-IRIS: overall effect for continuous outcome using a fixed-effect model. The diamonds stand for pooled effect.

**Abbreviations:** HIV, human immunodeficiency virus; TB, tuberculosis; ART, antiretroviral therapy; TB-IRIS, TB-associated immune reconstitution inflammatory syndrome; vit, vitamin; VDD, vitamin D deficiency; SD, standard deviation; CI, confidence interval; df, degrees of freedom; IV, independent variable.

---

**Table A**

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>TB</th>
<th>Mean (SD)</th>
<th>Control</th>
<th>Mean (SD)</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Odds ratio M-H, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>African</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Nansera et al 26</td>
<td>60</td>
<td>27.5</td>
<td>50</td>
<td>27.5</td>
<td>50</td>
<td>24.3</td>
<td>2011</td>
<td>1.00</td>
</tr>
<tr>
<td>Martinez et al 11</td>
<td>35</td>
<td>19.1</td>
<td>90</td>
<td>54.7</td>
<td>74</td>
<td>28.6</td>
<td>2011</td>
<td>0.02</td>
</tr>
<tr>
<td>Steinhoff et al 31</td>
<td>35</td>
<td>11.9</td>
<td>19</td>
<td>38.9</td>
<td>15</td>
<td>5.6</td>
<td>2012</td>
<td>0.02</td>
</tr>
<tr>
<td>Conesa-Botella et al 32</td>
<td>81</td>
<td>32.5</td>
<td>92</td>
<td>26.8</td>
<td>20</td>
<td>21.8</td>
<td>2012</td>
<td>0.02</td>
</tr>
<tr>
<td>Subtotal (95% CI)</td>
<td>164</td>
<td>100</td>
<td>260</td>
<td>93.6</td>
<td>100</td>
<td>1.89</td>
<td>2012</td>
<td>0.02</td>
</tr>
<tr>
<td>Heterogeneity: χ2 = 151.71; τ2 = 24.71, df = 3 (P &gt; 0.0001); I2 = 88% Test for overall effect: Z = 1.48 (P = 0.14)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

**Table B**

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>TB Events</th>
<th>Mean (SD)</th>
<th>Control Events</th>
<th>Mean (SD)</th>
<th>Weight (%)</th>
<th>Odds ratio M-H, random, 95% CI</th>
<th>Year</th>
<th>Odds ratio M-H, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>African</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Martinez et al 33</td>
<td>65</td>
<td>39</td>
<td>75</td>
<td>75</td>
<td>32.6</td>
<td>5.60</td>
<td>2011</td>
<td>0.02</td>
</tr>
<tr>
<td>Nansera et al 26</td>
<td>6</td>
<td>50</td>
<td>50</td>
<td>50</td>
<td>24.3</td>
<td>1.00</td>
<td>2011</td>
<td>0.02</td>
</tr>
<tr>
<td>Conesa-Botella et al 32</td>
<td>15</td>
<td>19</td>
<td>15</td>
<td>19</td>
<td>20.1</td>
<td>0.70</td>
<td>2012</td>
<td>0.02</td>
</tr>
<tr>
<td>Subtotal (95% CI)</td>
<td>260</td>
<td>100</td>
<td>260</td>
<td>100</td>
<td>1.89</td>
<td>0.69</td>
<td>2012</td>
<td>0.02</td>
</tr>
<tr>
<td>Heterogeneity: χ2 = 67.6, τ2 = 67.87, df = 3 (P = 0.03); I2 = 65% Test for overall effect: Z = 1.24 (P = 0.22)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

**Table C**

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>TB-IRIS Mean (SD)</th>
<th>Control Mean (SD)</th>
<th>Weight (%)</th>
<th>Mean difference IV, fixed, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, fixed, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>African</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Nansera et al 26</td>
<td>60</td>
<td>30</td>
<td>15</td>
<td>17.5</td>
<td>5.1</td>
<td>−5.00 (−21.25, 11.25)</td>
</tr>
<tr>
<td>Conesa-Botella et al 32</td>
<td>22.5</td>
<td>7.26</td>
<td>23</td>
<td>7.19</td>
<td>85</td>
<td>−5.50 (−9.46, −1.54)</td>
</tr>
<tr>
<td>Subtotal (95% CI)</td>
<td>77</td>
<td>51</td>
<td>100</td>
<td>−5.67 (−9.33, −2.01)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

*Figure 4 A* Forest plot of comparison of serum 25(OH)D level in HIV-TB-coinfected patients vs HIV patients without active TB: overall effect for continuous outcome using a random-effect model.

*Figure 4 B* Forest plot of association between VDD and risk of TB in HIV-infected patients: overall effect for dichotomous outcome using a random-effect model.

*Figure 4 C* Forest plot comparing 25(OH)D level in HIV-TB-coinfected patients receiving ART who developed TB-IRIS vs HIV-TB-coinfected patients receiving ART who did not develop TB-IRIS: overall effect for continuous outcome using a fixed-effect model. The diamonds stand for pooled effect.
significantly different from HIV-infected African patients without active TB (MD = −9.83; 95% CI = [−22.87, 3.22]; P = 0.14; F = 88) (Figure 4A).

The same four studies on African population with 260 HIV-TB-coinfected cases and 164 control HIV subjects without active TB were included in our analysis on the association between VDD and risk of TB in HIV-infected patients (Figure 4B).

There was no significant association between VDD and risk of TB in African HIV-infected patients (OR = 1.89; 95% CI = [0.69, 5.14]; P = 0.22; F = 65) (Figure 4B).

Significantly lower vitamin D level in HIV-TB-coinfected African patients receiving ART who developed TB-IRIS vs those who did not develop TB-IRIS

Three studies with 77 HIV-TB-coinfected African cases receiving ART who developed TB-IRIS and 117 African control patients who did not develop TB-IRIS were included in this analysis (Figure 4C). HIV-TB-coinfected African cases receiving ART who developed TB-IRIS had a significantly lower vitamin D level vs those who did not develop TB-IRIS (MD = −5.67; 95% CI = [−9.33, −2.01]; P = 0.002; F = 0) (Figure 4C).

Anti-TB treatment and vitamin D level

Anti-TB treatment did not affect vitamin D level in TB patients receiving the treatment

Three studies with 224 TB patients receiving 1–4 months of anti-TB treatments and four studies with 391 TB patients receiving full course of anti-TB treatment were included in our analysis on whether anti-TB treatment affected vitamin D level in TB patients receiving the treatment (Figure 5A and B). Neither 1–4 months of anti-TB treatment (MD = 1.47; 95% CI = [−2.79, 5.73]; P = 0.50; F = 57) (Figure 5A) nor a full course of anti-TB treatment (MD = 5.03; 95% CI = [−3.20, 13.25]; P = 0.23; F = 86) (Figure 5B) affected 25(OH)D level significantly.

TB patients still had a significantly lower vitamin D level after completion of anti-TB treatment than controls without TB

Four studies with 327 TB cases after completion of anti-TB treatment and 592 control subjects without TB were included (Figure 5C). Vitamin D level was still significantly lower in TB patients after completion of their anti-TB treatment vs control subjects without TB (MD = −8.05; 95% CI = [−13.56, −2.54]; P = 0.004; F = 68) (Figure 5C), especially for the Asian population (MD = −9.36; 95% CI = [−15.23, −3.48]; P = 0.002; F = 80) (Figure 5C).

A trend of higher 1,25(OH)2D3 level in TB patients vs controls without TB

Although serum 25(OH)D is the commonly used measure of vitamin D status in a subject, 1,25(OH)2D3 is the actual bioactive form of vitamin D that induces antimicrobial response, and therefore we performed a meta-analysis comparing 1,25(OH)2D3 level in TB subjects vs controls without TB. Four studies with 160 TB cases and 338 control subjects without TB were included (Figure 6). There was a trend toward higher 1,25(OH)2D3 level in TB patients vs controls; however, this trend did not reach the level of statistical significance (SMD = 1.02; 95% CI = [−0.09, 2.14]; P = 0.07; F = 95) (Figure 6).

Between-studies heterogeneity and publication bias

Between-studies heterogeneity varied from none to high for our meta-analyses with F values ranging from 0 to 96 (Figures 2A and B, 3A–C, 4A–C, 5A–C, and 6). Furthermore, funnel plots for all of our meta-analyses were symmetrical, suggesting the presence of none or very little publication bias in our analyses (data not shown).

Discussion

Our meta-analysis on the association between vitamin D and TB showed a significantly lower level of vitamin D in TB patients vs controls and that VDD was positively associated with an increased risk of TB, which is consistent with the results of the meta-analysis by Nnoaham et al. In the subgroup analysis, although we found similar results for the Asian population, such an association was lacking in the African population even when we excluded studies where cases and controls had other diseases such as HIV and diabetes, or where the control subjects had LTBI. This suggested the possibility of ethnic difference in the role of vitamin D in TB that warrants further investigation. As sunlight and diet are two major sources of vitamin D for humans, adequate exposure to sunlight and/or adjustment of diet to maintain proper vitamin D level in our bodies is deemed a necessity.

As HIV-infected subjects are more susceptible to TB and have greater mortality and morbidity caused by TB, in order to further clarify whether vitamin D played any role in occurrence and deterioration of TB in HIV-infected subjects, we also analyzed the association between vitamin D level/VDD and TB in HIV-infected subjects and the association...
### Table A

<table>
<thead>
<tr>
<th>Study group</th>
<th>After treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Before treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Davies et al’s patient group Oct to Mar</td>
<td>21.75</td>
<td>8.88</td>
<td>14</td>
<td>22</td>
<td>14</td>
<td>14</td>
<td>15.9</td>
<td>-0.25 (-0.93, 0.43)</td>
<td>1987</td>
<td>-0.25 (-0.93, 0.43)</td>
</tr>
<tr>
<td>Davies et al’s patient group Apr–Sep</td>
<td>16</td>
<td>3.06</td>
<td>13</td>
<td>14.25</td>
<td>1.19</td>
<td>13</td>
<td>43.1</td>
<td>1.75 (-0.03, 3.53)</td>
<td>1987</td>
<td>1.75 (-0.03, 3.53)</td>
</tr>
<tr>
<td>Treumann et al’s</td>
<td>10.11</td>
<td>33.7</td>
<td>81</td>
<td>91</td>
<td>20.37</td>
<td>81</td>
<td>16.1</td>
<td>10.10 (1.52, 18.68)</td>
<td>2010</td>
<td>10.10 (1.52, 18.68)</td>
</tr>
<tr>
<td>Koo et al’s</td>
<td>31.28</td>
<td>21.04</td>
<td>116</td>
<td>34.75</td>
<td>24.07</td>
<td>116</td>
<td>24.9</td>
<td>-3.50 (-9.32, 2.32)</td>
<td>2012</td>
<td>-3.50 (-9.32, 2.32)</td>
</tr>
<tr>
<td><strong>Total (95% CI)</strong></td>
<td></td>
<td></td>
<td></td>
<td>224</td>
<td>224</td>
<td>100</td>
<td></td>
<td>1.47 (-2.79, 5.73)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: $I^2 = 10.13$; $r^2 = 6.93$, $df=3$ ($P=0.07$); $I^2 = 87\%$ Test for overall effect: $Z=0.68$ ($P=0.50$)

#### Table B

<table>
<thead>
<tr>
<th>Study group</th>
<th>After treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Before treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Asian</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Koo et al’s</td>
<td>27.5</td>
<td>17.28</td>
<td>116</td>
<td>34.75</td>
<td>24.07</td>
<td>116</td>
<td>24.0</td>
<td>-7.25 (-12.64, -1.86)</td>
<td>2012</td>
<td>-7.25 (-12.64, -1.86)</td>
</tr>
<tr>
<td>Kim et al’s</td>
<td>31.5</td>
<td>19.03</td>
<td>165</td>
<td>33</td>
<td>21.58</td>
<td>165</td>
<td>24.8</td>
<td>-1.50 (-5.69, 2.89)</td>
<td>2013</td>
<td>-1.50 (-5.69, 2.89)</td>
</tr>
<tr>
<td>Hong et al’s</td>
<td>32.95</td>
<td>16.87</td>
<td>83</td>
<td>28.5</td>
<td>14.99</td>
<td>83</td>
<td>24.5</td>
<td>4.45 (-0.35, 9.20)</td>
<td>2014</td>
<td>4.45 (-0.35, 9.20)</td>
</tr>
<tr>
<td><strong>Subtotal (95% CI)</strong></td>
<td>364</td>
<td>364</td>
<td>73.3</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-1.34 (-7.65, 4.97)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: $r^2 = 24.92$; $r^2 = 10.17$, $df=2$ ($P=0.006$); $I^2 = 80\%$ Test for overall effect: $Z=0.42$ ($P=0.68$)

**Ethnicity unspecified**

<table>
<thead>
<tr>
<th>Study group</th>
<th>After treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Before treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Davies et al’s patient group Apr–Sep</td>
<td>30</td>
<td>31.56</td>
<td>13</td>
<td>14.25</td>
<td>1.19</td>
<td>13</td>
<td>12.3</td>
<td>15.75 (-1.42, 32.92)</td>
<td>1987</td>
<td>15.75 (-1.42, 32.92)</td>
</tr>
<tr>
<td>Davies et al’s patient group Oct to Mar</td>
<td>50.5</td>
<td>24.38</td>
<td>14</td>
<td>22</td>
<td>14</td>
<td>14</td>
<td>14.4</td>
<td>28.50 (13.77, 43.23)</td>
<td>1987</td>
<td>28.50 (13.77, 43.23)</td>
</tr>
<tr>
<td><strong>Subtotal (95% CI)</strong></td>
<td>27</td>
<td>27</td>
<td>26.7</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>22.92 (10.52, 35.32)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: $r^2 = 14.69$; $r^2 = 12.2$, $df=1$ ($P=0.27$); $I^2 = 18\%$ Test for overall effect: $Z=3.82$ ($P=0.003$)

**Total (95% CI)**

<table>
<thead>
<tr>
<th>Study group</th>
<th>After treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Before treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Asian</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Koo et al’s</td>
<td>27.5</td>
<td>17.28</td>
<td>116</td>
<td>33</td>
<td>17.96</td>
<td>86</td>
<td>28.5</td>
<td>-5.50 (-10.43, -0.57)</td>
<td>2012</td>
<td>-5.50 (-10.43, -0.57)</td>
</tr>
<tr>
<td>Kim et al’s</td>
<td>31.25</td>
<td>19.03</td>
<td>101</td>
<td>46.75</td>
<td>20.83</td>
<td>197</td>
<td>29.1</td>
<td>-15.60 (-20.22, -10.78)</td>
<td>2013</td>
<td>-15.60 (-20.22, -10.78)</td>
</tr>
<tr>
<td><strong>Subtotal (95% CI)</strong></td>
<td>300</td>
<td>565</td>
<td>88.5</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-9.36 (-15.23, -3.48)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: $r^2 = 21.58$; $r^2 = 10.04$, $df=2$ ($P=0.007$); $I^2 = 80\%$ Test for overall effect: $Z=3.12$ ($P=0.002$)

**Ethnicity unspecified**

<table>
<thead>
<tr>
<th>Study group</th>
<th>After treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Before treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Davies et al’s patient group Apr–Sep</td>
<td>30</td>
<td>31.56</td>
<td>13</td>
<td>33</td>
<td>37.5</td>
<td>13</td>
<td>3.8</td>
<td>-3.00 (-20.64, 23.64)</td>
<td>1987</td>
<td>-3.00 (-20.64, 23.64)</td>
</tr>
<tr>
<td>Davies et al’s patient group Oct to Mar</td>
<td>50.5</td>
<td>24.38</td>
<td>14</td>
<td>46</td>
<td>23.5</td>
<td>14</td>
<td>7.7</td>
<td>4.50 (-13.24, 22.24)</td>
<td>1987</td>
<td>4.50 (-13.24, 22.24)</td>
</tr>
<tr>
<td><strong>Subtotal (95% CI)</strong></td>
<td>27</td>
<td>27</td>
<td>11.5</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>2.29 (-12.57, 16.96)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: $r^2 = 0.00$; $r^2 = 0.21$, $df=1$ ($P=0.65$); $I^2 = 0\%$ Test for overall effect: $Z=0.29$ ($P=0.77$)

**Total (95% CI)**

<table>
<thead>
<tr>
<th>Study group</th>
<th>After treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Before treatment Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Mean difference IV, random, 95% CI</th>
<th>Year</th>
<th>Mean difference IV, random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Asian</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Koo et al’s</td>
<td>27.5</td>
<td>17.28</td>
<td>116</td>
<td>33</td>
<td>17.96</td>
<td>86</td>
<td>28.5</td>
<td>-8.05 (-13.56, -2.54)</td>
<td>2012</td>
<td>-8.05 (-13.56, -2.54)</td>
</tr>
</tbody>
</table>

Heterogeneity: $r^2 = 21.40$; $r^2 = 12.48$, $df=4$ ($P=0.01$); $I^2 = 68\%$ Test for overall effect: $Z=2.85$ ($P=0.045$)

**Figure 5 (A)** Forest plot of the effect of I–4 months of anti-TB treatment on vitamin D level (25(OH)D) in TB patients: overall effect for continuous outcome using a random-effect model. **(B)** Forest plot of the effect of a full course of anti-TB treatment on vitamin D level (25(OH)D) in TB patients: overall effect for continuous outcome using a random-effect model. **(C)** Forest plot of comparison of vitamin D level (25(OH)D) in TB patients after anti-TB treatment vs control without TB: overall effect for continuous outcome using a random-effect model. The diamonds stand for pooled effect.

**Abbreviations:** TB, tuberculosis; SD, standard deviation; CI, confidence interval; vit, vitamin; df, degrees of freedom; IV, independent variable.
between vitamin D level and risk of developing TB-IRIS in HIV-TB-coinfected patients receiving ART. Our analysis showed that there was no significant difference in vitamin D level in HIV-TB-coinfected African patients vs African HIV patients without active TB and that VDD was not associated with an increased risk of TB in African HIV-infected patients. However, a significantly lower vitamin D level was found in HIV-TB-coinfected African patients receiving ART who developed TB-IRIS vs those who did not. In view of the fact that our analysis showed lack of association between vitamin D and TB in the African population, such lack of association between vitamin D and TB in the HIV-infected African population should be interpreted with caution as to whether it represents a true lack of association in HIV-infected subjects; studies on HIV-infected subjects of other ethnicities are needed to confirm these results. However, if such a lack of association is indeed true, Steenholf et al.\(^7\) suggested that as HIV-infected subjects are already immunocompromised, it is possible that any immune response induced by vitamin D would also be compromised, thus masking any role vitamin D played in fighting TB; this, along their observation that HIV-infected subjects without TB were already plagued with VDD, could provide a possible explanation for such a lack of association between vitamin D and TB in HIV-infected subjects. On the other hand, our finding that HIV-TB-coinfected subjects who developed TB-IRIS (a paradoxical worsening of TB after initial improvement upon TB treatment after ART initiation or manifestation of TB clinically not apparent before ART initiation) had a lower vitamin D level that those who did not suggest that vitamin D was somehow still involved in TB occurrence and progression in HIV-infected patients. Conesa-Botella et al.\(^3\) reported an association between VDD and elevated proinflammatory cytokine and that corticosteroid therapy could modify patients’ inflammatory profile and possibly even reduce the severity of TB-IRIS symptom and suggested vitamin D supplement prior to ART initiation to determine whether it decreased TB-IRIS incidence. Meanwhile, as only a limited number of studies with limited sample sizes were included in our analyses, more studies with larger sample sizes are needed to further elucidate whether there is any association between VDD and TB/TB-IRIS.

Also, our analysis found a significant association between VDD and TB, but it did not tell us whether VDD was a risk factor or a consequence of TB; however, our analysis revealed that VDD was significantly associated with an increased risk of developing active TB in LTBI subjects/household contacts of TB patients, that VDD was significantly associated with an increased risk of TST conversion/TBIC, and that the trend toward a lower vitamin D level in active TB patients vs LTBI subjects/household contacts of active TB patients did not reach statistical significance. These results suggested that VDD was more likely a risk factor for TB rather than its consequence. Had VDD been a manifestation of malnutrition generally associated with TB, we would expect a more significant difference in vitamin D level between active TB patients vs LTBI subjects/household contacts of active TB patients and VDD would not have been significantly associated with an increased risk of LTBI developing into active TB or TST conversion/TBIC. On the contrary, our result suggested that vitamin D was possibly involved in

### Figure 6
Forest plot of comparison of the level of the vitamin D metabolite – serum/plasma 1,25-dihydroxyvitamin D3 (1,25(OH)\(_2\)D3) in TB patients vs control without TB: overall effect for continuous outcome using a random-effect model.

**Note:** The diamonds stand for pooled effect.

**Abbreviations:** TB, tuberculosis; SD, standard deviation; CI, confidence interval; df, degrees of freedom; IV, independent variable.

### Table

<table>
<thead>
<tr>
<th>Study or subgroup</th>
<th>TB Mean</th>
<th>SD</th>
<th>Control Mean</th>
<th>SD</th>
<th>Total</th>
<th>Weight (%)</th>
<th>Std mean difference N, random, 95% CI</th>
<th>Year</th>
</tr>
</thead>
<tbody>
<tr>
<td>Asian</td>
<td>Yonemura et al(^1)</td>
<td>45.9</td>
<td>31.5</td>
<td>6</td>
<td>11.8</td>
<td>7.5</td>
<td>110</td>
<td>22.7</td>
</tr>
<tr>
<td></td>
<td>Selvaraj et al(^2)</td>
<td>92.5</td>
<td>50.39</td>
<td>65</td>
<td>62.5</td>
<td>43.78</td>
<td>60</td>
<td>26.4</td>
</tr>
<tr>
<td></td>
<td>Gao et al(^3)</td>
<td>365.9</td>
<td>235.7</td>
<td>74</td>
<td>264.4</td>
<td>335.6</td>
<td>153</td>
<td>26.7</td>
</tr>
<tr>
<td>Subtotal (95% CI)</td>
<td>145</td>
<td>323</td>
<td>75.8</td>
<td>1.16 (-0.24, 2.55)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: χ\(^2\)=1.44; χ\(^2\)=64.58, df=2 (P=0.00001); I²=97%
Test for overall effect: Z=1.62 (P=0.11)

| Ethnicity unspecified | Davies et al\(^4\) | 35.7 | 10.48 | 15 | 28.7 | 9.75 | 15 | 24.2 | 0.67 (-0.07, 1.41) | 1985 |
| Subtotal (95% CI) | 15 | 15 | 24.2 | 0.67 (-0.07, 1.41) | |

Heterogeneity: not applicable
Test for overall effect: Z=1.79 (P=0.07)
Test for subgroup differences: χ\(^2\)=0.36, df=1 (P=0.55); I²=0%

Total (95% CI) | 160 | 338 | 100 | 1.02 (-0.09, 2.14) | |

Heterogeneity: χ\(^2\)=1.19; χ\(^2\)=65.94, df=3 (P=0.00001); I²=96%
Test for overall effect: Z=1.81 (P=0.07)
Test for subgroup differences: χ\(^2\)=0.36, df=1 (P=0.55); I²=0%
the early stage of TB infection and symptom manifestation; in addition, it has been shown that 1,25(OH)2D3 activates VDR signaling and induces antimicrobial responses such as induction of autophagy, phagolysosomal fusion, release and activation of the antimicrobial peptide cathelicidin, and killing of intracellular M. tuberculosis. All of these results pointed to the fact that VDD is more likely a risk factor for TB. This possibility was further strengthened by our results that anti-TB treatment did not significantly affect vitamin D level in TB patients receiving the treatment and that TB subjects after completion of anti-TB treatment still had a significantly lower vitamin D level than control subjects without TB. This was because if VDD had been a consequence of TB, we would expect a significantly elevated vitamin D level in TB subjects after effective anti-TB treatment. Overall, our analyses indicated that VDD was most likely a risk factor for TB.

Since our analysis indicated that VDD was most likely a risk factor for TB, this raises a question of whether vitamin D supplement would be beneficial to TB prevention and treatment. The meta-analysis by Xia et al in 2012 reported that vitamin D supplementation did not show a significant benefit to TB treatment; however, the meta-analysis did not address the question of whether it would be beneficial to TB prevention. Rigorously controlled studies will be needed to further determine whether vitamin D supplementation would be beneficial.

Because 1,25(OH)3D is the bioactive form of vitamin D, we also analyzed the level of 1,25(OH)2D3 in TB subjects vs controls and found a trend of higher 1,25(OH)2D3 level in TB patients vs controls without TB. Selvaraj et al postulated that this possible increase in 1,25(OH)2D3 could be because of CYP27B1 expression upregulation that led to increased conversion of 25(OH)D to 1,25(OH)2D3, which could then cause 25(OH)D deficiency because of its increased use. However, if this is the case, it would mean that VDD would be a consequence of TB rather than one of its causes and would be inconsistent with our previous conclusions. On the other hand, as only four studies with limited sizes were included in our analysis, this result is far from certain, and more studies are needed to confirm or refute this result.

There are certain strengths and limitations in our study. Its strength lies in the fact that it was a comprehensive analysis of multiple aspects of the relationship between vitamin D and TB and helped to answer the question of whether vitamin D was one of the causes or consequences of TB, and our analysis included all relevant available studies of reasonable quality to ensure that the resulting picture was as complete as possible. However, our study also has certain limitations. For some of our analyses, the number of relevant studies was limited and their sample sizes were small, which undoubtedly would affect the certainty of some of our results. Furthermore, because different studies used different standards for VDD, such inconsistency in VDD definition would also affect our results.

**Conclusion**

Our analyses revealed a significantly lower vitamin D level in TB subjects and that VDD was associated with more risk of TB. VDD is more likely a risk factor for TB than its consequence, and more studies are needed to determine whether vitamin D supplement is beneficial to TB prevention and treatment.

**Disclosure**

The authors report no conflicts of interest in this work.

**References**

Vitamin D and tuberculosis: a meta-analysis


27. Cochrane Handbook for Systematic Reviews of Interventions. Version 5.1.0 [updated March 2011], Section 7.7.3.4


