Men’s preferences for the treatment of lower urinary tract symptoms associated with benign prostatic hyperplasia: a discrete choice experiment

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Objective: To explore and quantify men’s preferences and willingness to pay (WTP) for attributes of medications for lower urinary tract symptoms associated with benign prostatic hyperplasia using a discrete choice experiment.

Subjects and methods: Men in the UK aged ≥45 years with moderate-to-severe lower urinary tract symptoms/benign prostatic hyperplasia (based on self-reported International Prostate Symptom Score ≥8) were recruited. An online discrete choice experiment survey was administered. Eligible men were asked to consider different medication scenarios and select their preferred medication according to seven attributes: daytime and nighttime (nocturia) urinary frequency, urinary urgency, sexual and nonsexual side effects, number of tablets/day, and cost/month. A mixed-logit model was used to estimate preferences and WTP for medication attributes.

Results: In all, 247 men completed the survey. Men were willing to trade-off symptom improvements and treatment side effects. Men preferred medications that reduced urinary urgency and reduced day- and nighttime urinary frequency. Men preferred medications without side effects (base-case level), but did not care about the number of tablets per day. WTP for symptomatic improvement was £25.33/month for reduced urgency (urge incontinence to mild urgency), and £6.65/month and £1.39/month for each unit reduction in night- and daytime urination frequency, respectively. The sexual and nonsexual side effects reduced WTP by up to £30.07/month. There was significant heterogeneity in preferences for most attributes, except for reduced urinary urgency from urge incontinence to mild urgency and no fluid during ejaculation (dry orgasm).

Conclusion: To compensate for side effects, a medicine for lower urinary tract symptoms/benign prostatic hyperplasia must provide a combination of benefits, such as reduced urgency of urination plus reduced nighttime and/or reduced daytime urination.

Keywords: benign prostatic hyperplasia, discrete choice experiment, erectile dysfunction, lower urinary tract symptoms, storage symptoms, urge incontinence

Introduction

Lower urinary tract symptoms (LUTS) are prevalent in aging men, with ~90% of men aged 50 to 80 years suffering from potentially bothersome symptoms.1 Presenting symptoms include storage symptoms (ie, daytime urinary frequency, nocturia, urgency, urinary incontinence), voiding symptoms (ie, slow stream, splitting or spraying, intermittency, hesitancy, straining, terminal dribble), and/or post-micturition symptoms (sensation of incomplete emptying, post-micturition dribble).1 All three symptoms are found in ~one-third of men with LUTS,2 but storage symptoms have been reported to be the most bothersome.1,3

The mechanisms leading to the development of LUTS are multifactorial.3 Benign prostatic hyperplasia (BPH) is reported to be a common cause of LUTS6,7 and 25% to
50% of men with BPH have LUTS.1,4 LUTS associated with BPH (LUTS/BPH) is reported to negatively affect quality of life, daily activities, and general health.9,10 LUTS/BPH has also been linked to sexual dysfunction (eg, erectile dysfunction)11,12 and psychological dysfunction (eg, depression/anxiety)9,13 in men.

Drug treatment is recommended for men with moderate-to-severe LUTS, that is, those with an International Prostate Symptom Score (IPSS) ≥8, to provide symptomatic relief when conservative management strategies have been unsuccessful or are not appropriate.1,5 Drug treatment options include α1-blockers, 5α-reductase inhibitors (5-ARIs), antimuscarinics, and phosphodiesterase type 5 inhibitors.1,5 Combination therapy should be considered for men who have unresolved symptoms with monotherapy, for example, an α1-blocker plus an antimuscarinic or a 5-ARI.1,5

Patient preferences are an important consideration in the treatment of LUTS/BPH.5,14 Physicians should involve patients in the decision-making process to limit potential subsequent regret by the patient regarding some aspect(s) of the initiated treatment strategy.1,5 especially as patients’ and physicians’ preferences may not always be aligned.16 Discrete choice experiments (DCEs) are a common method to elicit patients’ preferences to understand what aspects of symptoms and treatment are important to patients and the relative importance of these factors.17 Treatment choices require trade-offs between treatment benefits and side effects; DCEs are designed to capture and quantify these trade-offs. DCEs are being increasingly used to assist with health care decision making during the regulatory process. Evaluation of patient preferences is recommended by the US Food and Drug Administration18 and the German Institute for Quality and Efficiency in Health Care; the latter explicitly encourages the use of DCEs.19 The European Medicines Agency also advocates consideration of patients’ assessments of the risks and benefits of medications through the use of quantitative methods.20 DCEs have been used to assess men’s preferences for treatments for LUTS/BPH,21 BPH,22 and overactive bladder.23 The previous study in LUTS/BPH used community dwelling men aged ≥40 years in the UK and, therefore, the present study is the first DCE that has focused specifically on men with symptoms of LUTS/BPH.

In the present study, a DCE was conducted in men with LUTS/BPH with the following objectives: 1) to explore men’s preferences for attributes of medications for LUTS/BPH; 2) to estimate their willingness to pay (WTP), a monetary measure of benefit for attributes of medications; 3) to determine the trade-offs they are prepared to make between the benefits and side effects of medication; and 4) to explore variations in preferences for attributes of medications.

Subjects and methods
DCEs are based on the assumption that medications are valued by patients because of their characteristics or attributes.24 In our DCE, medications (A, B, or no medicine) were described by a set of attributes (eg, symptom improvements, side effects, and cost) that differed in their attribute levels (eg, variations in extent of symptom improvement, presence or absence of side effects), thereby creating different scenarios associated with LUTS/BPH. Men were asked to hypothetically choose the medication they preferred in a series of such choice tasks. From these choices, preferences for medication were inferred, allowing the estimation of: relative importance of the attributes; trade-offs made between attributes; and WTP for attributes.

Attributes and levels
Following best practice recommendations,25 qualitative methods were used to select attributes and levels for the DCE. The LUTS/BPH medications in our study were described by seven attributes and corresponding levels (Table 1). These attributes and levels were identified from an earlier qualitative research phase of the study (ie, literature review

Table 1 Attributes and levels in the DCE

<table>
<thead>
<tr>
<th>Attribute</th>
<th>Level</th>
</tr>
</thead>
<tbody>
<tr>
<td>Daytime frequency of urination</td>
<td>Eight, ten, 12, or 14 times</td>
</tr>
<tr>
<td>Nighttime frequency of urination</td>
<td>One, two, three, or four times</td>
</tr>
<tr>
<td>Urgency of urination – when you need to urinate you usually…</td>
<td>Have to rush to the toilet and leak before you get there (urge incontinence)</td>
</tr>
<tr>
<td>Sexual side effects</td>
<td>No sexual side effects, No fluid during ejaculation (dry orgasm), Decreased sexual desire, Erectile dysfunction</td>
</tr>
<tr>
<td>Nonsexual side effects</td>
<td>No nonsexual side effects, Dry mouth, Headaches, Dizziness</td>
</tr>
<tr>
<td>Number of tablets</td>
<td>One or two tablets per day</td>
</tr>
<tr>
<td>Cost per month</td>
<td>£5, £10, £20, or £40</td>
</tr>
</tbody>
</table>

Notes: 1Frequency attributes enter the statistical analysis as deviation from the no-treatment alternative; 2base-case level, includes prescription fees, travel, and absenteeism from work. Note that for urgency of urination, the extended text is what the men read in the survey, while the short labels in brackets are used in the paper for ease of reading.

Abbreviation: DCE, discrete choice experiment.
and online discussion groups). Briefly, a structured literature review was conducted to find background information on LUTS/BPH and ensure that the subsequent study stages were based on existing evidence. Medline and Cochrane Database of Systematic Reviews were searched for articles in English that were published in 1991–2013. Combinations of the following search terms were used: “benign prostatic hyperplasia”, “prostatic hyperplasia”, “choice behavior”, “patient preference”, “decision making”, “patient participation”, and “attitude to health”. The search was augmented by searching the reference lists of the identified literature. The abstracts (where available) of all publications were read for relevance and the content of documents collected. A total of 362 unique papers were identified (after excluding 72 duplicates).

Five moderated asynchronous online discussion groups were used to collect qualitative data about men’s experiences of LUTS/BPH. Each discussion group lasted for 5 days and was hosted online on a secure and password protected bulletin board. Men residing in the UK with an IPSS score ≥8 were invited to take part in the discussion groups. The men (n=48) registered to a bulletin board with an anonymous username. Each day, a question about LUTS/BPH and its treatment was posed to and discussed by the men (day 1: LUTS/BPH and its effect on quality of life; day 2: management of symptoms; day 3: side effects of treatment; day 4: consultation with a doctor; and day 5: concluding comments). These questions were informed by the literature review. The discussions were downloaded and analyzed independently by three researchers using an inductive thematic analysis.

The attributes and levels included in the DCE were those most relevant to men with LUTS/BPH and were identified in two separate researcher workshops based on findings from the literature review and the qualitative analysis. “Daytime frequency of urination” (or daytime urinary frequency) levels ranged from eight to 14 times based on evidence that the median urinary frequency was seven times in 24 hours (range 2 to 21) in asymptomatic men, with 95% voiding fewer than 12 times daily. The “nighttime frequency of urination” (nocturia) range was based on the IPSS instrument (range 0 to 5) and a study that reported data from 24-hour voiding diaries (range 0 to 4). The levels of “urgency of urination” (urinary urgency) were adapted from the validated Patient Perception of Intensity of Urgency Scale. The attributes and levels for the sexual side effects (hypoactive sexual desire, as well as erectile and ejaculatory dysfunction) and nonsexual side effects (headaches, dizziness, and dry mouth), were based on the reported side effects of current LUTS/BPH medications, including α₁-blockers, 5-ARIs, antimuscarinics, and phosphodiesterase type 5 inhibitors.

Two of the attributes were not identified from the literature review and online discussion groups: number of daily tablets and cost of medication. Men may receive combination therapy for persistent LUTS/BPH, and therefore have preferences for the number of tablets to improve symptoms. The cost attribute was included to allow estimation of WTP for marginal changes in attribute levels. This attribute was framed as the total cost for men, including any prescription fees, travel expenses and absenteeism. The range of levels for the cost attribute was derived from earlier related studies and adjusted for inflation using the consumer price index. This range was tested in the pilot study in two ways: by examining the proportion of men never selecting a treatment in the DCE choice sets and including a payment card contingent valuation instrument.

Experimental design and survey structure
The attributes and levels resulted in 8,192 (4⁵×2⁶) different possible combinations of medication. To reduce this number for presentation to men in a survey, a D-efficient design was generated. In total, 30 choice sets (grouped into three sets of ten choice sets) were selected to reduce the burden of the survey for men. Each choice set included two different medicine options and no medicine option (ie, continue with the baseline scenario described). Figure 1 presents an example choice set.

The questionnaire included four sections. Section 1: men were provided with information (text and pictures) about LUTS/BPH and then asked three multiple-choice questions to determine the level of bother from their urinary symptoms. These were measured on an 11-point scale (range, 0—not at all bothered to 10—very bothered; a “don’t know” option was also included). Men were also asked if they had ever visited a doctor about LUTS/BPH; were they currently taking medication to treat LUTS/BPH; and what treatments were received. Section 2: men considered a set of LUTS/BPH symptoms and were presented with questions about each of the seven attributes, one by one, to familiarize themselves with the attributes used in the DCE. Section 3: the DCE choice sets followed by debriefing questions about how difficult realistic they found the choice tasks and whether they considered all of the medication attributes when making their choices. Section 4: men were asked questions about their socioeconomic characteristics and to complete the validated Beliefs about Medication Questionnaire. At the end of the survey, men were invited to provide any comments they had about the questionnaire.

Ethical approval for this study was granted by the College of Life Science and Medicine, University of Aberdeen, College Ethics Research Board (CERB), November 1, 2013 – CERB/2013/8/942. The full survey is available from the corresponding author on request.
Participants and recruitment

The DCE was a self-administered online survey and participation was voluntary; the DCE survey was conducted during March 19–26, 2014. The target population was 300 men living in the UK, recruited from the online panel of a market research company. Men invited to take part in the survey (identified from screening questions) were aged ≥45 years and had moderate-to-severe LUTS/BPH, that is, self-reported IPSS ≥8. These inclusion criteria were selected based on a reported increased prevalence of storage symptoms in men aged ≥45 years and the IPSS categorization of LUTS as moderate or severe (scores of 8–19 or 20–35, respectively). The invited men received a detailed information sheet about the survey with a hyperlink.

We assumed that by clicking on the hyperlink men consented to taking part in the study. Men were free to withdraw at any point without having to give any reason(s). Men received a reward if they completed the 15-minute survey (£1.50).

The DCE was pilot tested among a subgroup of 63 men to assess the clarity of wording used in the survey, as well as to check the appropriateness of the attributes and levels. The mean WTP for treatment using a payment card was £15, and more than a third of men always chose the no medicine option. Based on these responses, the range of the cost attribute was lowered from £15, £30, £50, and £75 to £5, £10, £20, and £40 in the final DCE.

Data analysis

In the DCE, each man (i) was presented with a choice between three alternative actions (m), that is, taking medicine A, taking medicine B, or taking no medicine. Men were assumed to choose the alternative (ie, their most preferred option) that provided them with the highest utility (ie, overall benefit) in terms of symptom improvement, side effects, number of tablets, and cost (ie, the attributes included in the DCE). The utility ($u_{im}$) that a man receives from an alternative is a function of the characteristics included in the DCE and an error term ($e_{im}$):

$$
    u_{im} = \beta_0 + (\beta_1 + \sigma_{i1}) \text{daytime frequency}_{im} + (\beta_2 + \sigma_{i2}) \text{nighttime frequency}_{im} + (\beta_3 + \sigma_{i3}) \text{mild urgency}_{im} + (\beta_4 + \sigma_{i4}) \text{moderate urgency}_{im} + (\beta_5 + \sigma_{i5}) \text{severe urgency}_{im} + (\beta_6 + \sigma_{i6}) \text{no fluid during ejaculation sexual side effect}_{im} + (\beta_7 + \sigma_{i7}) \text{decreased desire sexual side effect}_{im} + (\beta_8 + \sigma_{i8}) \text{erectile dysfunction sexual side effect}_{im} + (\beta_9 + \sigma_{i9}) \text{dry mouth nonssexual side effect}_{im} + (\beta_{10} + \sigma_{i10}) \text{headache nonssexual side effect}_{im} + (\beta_{11} + \sigma_{i11}) \text{dizziness nonssexual side effect}_{im} + (\beta_{12} + \sigma_{i12}) \text{number of tablets}_{im} + \beta_1 \text{cost per month}_{im} + e_{im}
$$

$\beta_0$ is a constant term that denotes men’s general preference for medication versus no medication. Given dummy variables are used to analyze the categorical variables (later); the constant term also values these reference categories.

In the estimated model, each attribute can be associated with two coefficients, one representing the average utility of the attribute ($\beta_1$, $\beta_2$, $\ldots$), and another representing individual-specific preference variation for the attributes ($\sigma_{i1}$ to $\sigma_{i12}$). These individual-specific preference variations allow

Please imagine this situation:

- You have to urinate 14 times during the day. When you need to urinate you usually have to rush to the toilet and leak before you get there. During the night you wake up four times to go to the bathroom and urinate.
- You do not take any medicine to treat your symptoms. However, your general practitioner has told you that there are medicines available that can improve the symptoms described above. These medicines, however, may have side effects.

Please compare the medicines and tick which, if any, you would take:

<table>
<thead>
<tr>
<th>Description</th>
<th>Medicine A</th>
<th>Medicine B</th>
<th>No medicine</th>
</tr>
</thead>
<tbody>
<tr>
<td>Daytime frequency of urination</td>
<td>Eight times</td>
<td>14 times</td>
<td>14 times</td>
</tr>
<tr>
<td>Nighttime frequency of urination</td>
<td>Three times</td>
<td>One time</td>
<td>Four times</td>
</tr>
<tr>
<td>Urgency – when you need to urinate you usually…</td>
<td>Have to rush to the toilet and leak before you get there</td>
<td>Cannot postpone and have to rush to the toilet in order not to wet yourself</td>
<td>Have to rush to the toilet and leak before you get there</td>
</tr>
<tr>
<td>Sexual side effects of medicine</td>
<td>No sexual side effects</td>
<td>Decreased sexual desire</td>
<td>You do not have any side effects</td>
</tr>
<tr>
<td>Nonsexual side effects of medicine</td>
<td>Dry mouth</td>
<td>Headaches</td>
<td>You do not have any side effects</td>
</tr>
<tr>
<td>Number of tablets per day</td>
<td>Two tablets</td>
<td>One tablet</td>
<td>You do not take any medicine</td>
</tr>
<tr>
<td>Cost per month</td>
<td>£5</td>
<td>£20</td>
<td>£0</td>
</tr>
</tbody>
</table>

Please select your answer here: [ ] [ ] [ ]

Figure 1 Example DCE choice task.

Abbreviation: DCE, discrete choice experiment.
that not all men will have the same preferences for medication and provide a measure of the variation in preferences in the population. \( \sigma_1 \) to \( \sigma_{12} \) represents preference variation in the sample of men; if these coefficients are significant, this implies that preferences vary across the sample.

The interpretation of the average utility coefficient depends on the unit of measurement of the attribute. \( \beta_i \) and \( \beta_j \) represent the effect of a one-time reduction in day- and nighttime frequency of urination, respectively (and are modeled as continuous variables). The urgency attribute, side effects attributes, and number of tablets attributes were coded as dummy variables. In this case, the coefficients representing the average utility are interpreted relative to the reference categories – these are “urinary incontinence” for the urgency attribute, “none” for both the sexual and nonsexual side effects attributes, and one tablet per day for the number of tablets attribute. \( \beta_{13} \) indicates the effect of a £1 increase in the cost of treatment. The signs (+/−) of the coefficients, representing average utility, indicate whether a change in the attribute as described earlier has a positive or negative effect on medication utility. DCE responses were analyzed using a mixed-logit model;19 (Supplementary material).

From the mixed-logit results, we calculated WTP estimates, which show the monetary amounts that men are willing to pay per month for a unit change in each of the significant attributes. This placed all attributes on a common and meaningful metric (money), which allows the relative importance of each attribute as well as strength of preference and the trade-offs men make between attributes to be compared. WTP is calculated as the ratio of the average utility coefficient for the attribute and negative of the coefficient for the cost attribute. For example, \( (\beta_{i}/-\beta_{13}) \) is men’s WTP for one less daytime urination, while \( (\beta_{j}/-\beta_{13}) \) is men’s WTP to avoid the erectile dysfunction sexual side effect. Confidence intervals for WTP estimates were calculated using the Delta method.19 The WTP estimates were also used to assess the trade-offs that men were willing to make between symptom improvements and the side effects of medications.

Results

Men’s characteristics

A total of 5,212 individuals were invited to participate and 1,097 opened the online survey (ie, clicked the hyperlink). Of those who opened the DCE, 450 individuals did not meet the eligibility criteria, 136 met the criteria but were excluded because the quota was full, and 201 only partially completed the survey. A total of 310 men completed the survey (63 men in the pilot sample and 247 in the final sample); responses from the final sample of 247 men were used for analysis.

The mean age of the sample was 62.32 years, 51.01% of men had their day-to-day activities limited because of a health problem or disability, and 46.15% of men were retired (Table 2). The reported mean IPSS of the population was 15.69 (standard deviation 5.99). In addition, men were generally bothered by urinary symptoms, more during the night (mean score 5.86) than the day (mean score 4.92) and by urinary urgency (mean score 5.73). Approximately half of the men had spoken to a doctor (45.34%) and less than a third received medication for their urinary symptoms (26.32%); a summary of prior medications is shown in Table 3.

Estimation results

A total of 216 men (87.5%) stated that they considered all of the presented attributes. Across all of the choices in the DCE,
The regression coefficients were statistically significant at a 5% level for day- and nighttime urinary frequency, two of three attribute levels for urinary urgency, all sexual side effects, two of three attribute levels for nonsexual side effects, and the cost of medication (Table 4). The positive coefficients indicate that men preferred reductions in the frequency of urination both during the day and at night, as well as reductions in urgency of urination from urge incontinence (base case) to moderate and mild urgency. The negative coefficients for sexual and nonsexual side effects and cost per month showed that, on average, men preferred medication without these attributes.

Three regression coefficients were not statistically significant: men did not distinguish between urge incontinence and severe urgency, between taking one or two tablets per day, or viewed dry mouth as something to avoid. These results imply that, on average, these attributes did not influence men’s choice of medication.

There was a statistically significant preference variation in the population for several attributes (Table 4): reductions in day- or nighttime frequency; reductions in urinary urgency from urge incontinence to severe or moderate urgency; the sexual side effects of decreased sexual desire and erectile

### Table 3 Previous treatments for urinary symptoms (n=65)

<table>
<thead>
<tr>
<th>Drug</th>
<th>Frequency, n (%)</th>
<th>Relative frequency (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>α₁-blockers</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Tamsulosin</td>
<td>29</td>
<td>44.6</td>
</tr>
<tr>
<td>Doxazosin</td>
<td>6</td>
<td>9.2</td>
</tr>
<tr>
<td>Alfuzosin</td>
<td>4</td>
<td>6.2</td>
</tr>
<tr>
<td>Terazosin</td>
<td>2</td>
<td>3.1</td>
</tr>
<tr>
<td>5α-reductase inhibitor</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Finasteride</td>
<td>12</td>
<td>18.5</td>
</tr>
<tr>
<td>Dutasteride</td>
<td>2</td>
<td>3.1</td>
</tr>
<tr>
<td>Antimuscarinic</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Oxybutinin</td>
<td>5</td>
<td>7.7</td>
</tr>
<tr>
<td>Solifenacitin</td>
<td>3</td>
<td>4.6</td>
</tr>
<tr>
<td>Tolterodine</td>
<td>2</td>
<td>3.1</td>
</tr>
<tr>
<td>Flavoxate</td>
<td>2</td>
<td>3.1</td>
</tr>
<tr>
<td>Darifenacitin</td>
<td>1</td>
<td>1.5</td>
</tr>
<tr>
<td>Propiverine</td>
<td>1</td>
<td>1.5</td>
</tr>
<tr>
<td>Combination therapy</td>
<td>Tamsulosin plus dutasteride</td>
<td>6</td>
</tr>
</tbody>
</table>

**Note:** Ten men had received two previous treatments.

25.71% of men chose medication A, 32.75% chose medication B, and 41.54% chose no medication. Furthermore, 12.55% of men always chose the “no medicine” alternative compared to 21.05% of men who always chose one of the medicines (either medicine A or medicine B).

### Table 4 Mixed-logit regression coefficients, willingness-to-pay estimates, and standard deviations of significant coefficients

<table>
<thead>
<tr>
<th>Attribute</th>
<th>Regression coefficient (β)</th>
<th>P-value</th>
<th>Standard deviation (σ)</th>
<th>P-value</th>
<th>WTP</th>
<th>95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Constant**</td>
<td>−0.377</td>
<td>0.212</td>
<td>2.771**</td>
<td>&lt;0.001</td>
<td>NS</td>
<td>NS</td>
</tr>
<tr>
<td>Daytime frequency**</td>
<td>0.080**</td>
<td>0.002</td>
<td>0.114**</td>
<td>&lt;0.001</td>
<td>£1.39</td>
<td>£0.55 to £2.23</td>
</tr>
<tr>
<td>Nighttime frequency**</td>
<td>0.384**</td>
<td>0.000</td>
<td>0.847**</td>
<td>0.001</td>
<td>£6.65</td>
<td>£4.36 to £8.94</td>
</tr>
<tr>
<td>Urinary urgency†</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Severe urgency</td>
<td>−0.164</td>
<td>0.376</td>
<td>1.255**</td>
<td>&lt;0.001</td>
<td>NS</td>
<td>NS</td>
</tr>
<tr>
<td>Moderate urgency</td>
<td>1.381†</td>
<td>0.000</td>
<td>0.863**</td>
<td>&lt;0.001</td>
<td>£23.93</td>
<td>£17.36 to £30.51</td>
</tr>
<tr>
<td>Mild urgency</td>
<td>1.462†</td>
<td>0.000</td>
<td>0.518</td>
<td>0.054</td>
<td>£25.33</td>
<td>£19.00 to £31.67</td>
</tr>
<tr>
<td>Sexual side effects†</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No fluid during ejaculation (dry orgasm)</td>
<td>−0.969‡‡</td>
<td>0.000</td>
<td>0.316</td>
<td>0.534</td>
<td>−£16.76</td>
<td>−£23.48 to −£10.11</td>
</tr>
<tr>
<td>Decreased sexual desire</td>
<td>−1.135‡‡</td>
<td>0.000</td>
<td>0.926‡‡</td>
<td>&lt;0.001</td>
<td>−£19.66</td>
<td>−£26.63 to −£12.68</td>
</tr>
<tr>
<td>Erectile dysfunction</td>
<td>−1.735‡‡</td>
<td>0.000</td>
<td>0.035‡‡</td>
<td>&lt;0.001</td>
<td>−£30.07</td>
<td>−£38.01 to −£22.13</td>
</tr>
<tr>
<td>Nonsexual side effects*</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Dry mouth</td>
<td>−0.042</td>
<td>0.748</td>
<td>0.135</td>
<td>0.662</td>
<td>NS</td>
<td>NS</td>
</tr>
<tr>
<td>Headaches</td>
<td>−0.939‡‡</td>
<td>0.000</td>
<td>0.785‡‡</td>
<td>0.002</td>
<td>−£16.26</td>
<td>−£22.00 to −£10.52</td>
</tr>
<tr>
<td>Dizziness</td>
<td>−0.955‡‡</td>
<td>0.000</td>
<td>0.666</td>
<td>0.016</td>
<td>−£16.54</td>
<td>−£22.25 to −£10.83</td>
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<td>Number of tablets per day</td>
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<td>0.425</td>
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<tr>
<td>Cost per month</td>
<td>−0.058‡‡</td>
<td>0.000</td>
<td>Fixed coefficient</td>
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<td>Log likelihood</td>
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<td>Akaike information criterion</td>
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<td>McFadden R²</td>
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<td>N (observations)</td>
<td>7,410</td>
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**Notes:** *P<0.05; †P<0.01; ‡ coefficient is normally distributed; ‡‡ coefficient is log-normally distributed.

**Abbreviations:** CI, confidence interval; NS, not significant; WTP, willingness to pay.
dysfunction; and the nonsexual side effects of headache and dry mouth. However, there were no statistically significant variations in preference for an improvement from urge incontinence to mild urgency. This may reflect the variation in the sample about the degree of improvement represented by a move from urge incontinence to severe or moderate urgency, and the agreement in the sample about the benefit of the largest improvement from urge incontinence to mild urgency. The side effect of dry mouth also showed no significant preference variation, implying this attribute was not important to men in our sample.

**WTP**

The value of reducing day- or nighttime frequency by one time was small compared to the value for a reduction of urgency and side effects (Table 4). In terms of symptom improvements, men were willing to pay more to reduce the frequency urination by one toilet visit at night compared with the day (£6.65/month vs £1.39/month). Men were willing to pay more to reduce urgency of urination than to reduce frequency of urination, for example, reducing urgency of urination from urge incontinence to mild urgency was valued at £25.33/month; this urinary urgency reduction is equivalent to reducing nighttime frequency by four urination events (ie, £25.33/£6.65). All sexual side effects reduced the value of medication; erectile dysfunction was considered to be the worst (£30.07/month). The other sexual side effects, no fluid during ejaculation (dry orgasm), and decreased sexual desire, had a similar impact on medication value as the nonsexual side effects of headaches and dizziness. Furthermore, the value of reduced urgency (to moderate or mild urgency) would compensate any one of the side effects, except for erectile dysfunction.

**Discussion**

This study estimated men’s preferences for medication attributes of treatments for LUTS/BPH. In terms of symptom improvement, reducing urinary urgency was the most important benefit of medication, followed by a one-time reduction in night- and daytime urinary frequency. In terms of side effects, avoiding sexual and nonsexual side effects was also important to men; erectile dysfunction was perceived as the worst side effect. The results also showed preference variations in the importance of medication attributes for treatment choice in the sample.

Based on key findings from the earlier qualitative phase of this study, the DCE was designed to focus on the most bothersome symptoms (ie, storage symptom attributes of frequency and urgency) and men’s experience of their symptoms and treatment. As such, this DCE did not consider markers of treatment failure (eg, acute urinary retention or surgery) or treatment outcomes (eg, time to achieve improved symptoms or reduced prostate size), endpoints specifically associated with $\alpha_1$-blockers and 5-ARIs, which have been reported to be important to men with BPH\(^{20,41}\) and included in the previously published DCEs of BPH.\(^{21,35}\) The attribute levels for nonsexual side effects were expanded versus the previously published DCEs to include dry mouth, headache, and dizziness. Overall, these changes versus previous DCEs in BPH were designed to ensure our DCE focused on men’s experiences and make it more relevant to all currently recommended therapies.

Storage symptoms, which include urgency, increased daytime frequency, increased nighttime frequency, and urge incontinence, represent the most bothersome LUTS\(^{13,4}\) and are reported in 62% of men with LUTS.\(^{3}\) We found that the most valued attribute of medication was reduced urgency of urination from urge incontinence to moderate or mild levels. Men were willing to pay up to £24–25/month for the ability to postpone urination temporarily (moderate urgency) or as long as necessary, without fear of wetting themselves (mild urgency). In contrast, a change from urge incontinence to severe urgency (cannot postpone and have to rush to the toilet in order to not wet yourself) appeared to have little importance and men were not willing to pay for this level of change. Men also valued reductions in day- and nighttime frequency. These data are not unexpected given that storage symptoms are reported to impair quality of life\(^{42,43}\) and limit daily activities\(^{44}\) in men with LUTS.

The DCE showed that men prefer to avoid medications with sexual side effects and that erectile dysfunction has the strongest influence on choice, reducing the WTP by £30.07 per month. This is consistent with previous DCEs, which reported erectile dysfunction was the least desirable sexual side effect and confirmed the undesirability of decreased sexual desire and no fluid during ejaculation (dry orgasm).\(^{21,35}\) Our results suggest that restrictions on sexual activity, specifically described as an inability to be sexually active, have a large negative effect on men’s quality of life.

Men preferred treatments without headache and dizziness, which are reported nonsexual side effects of $\alpha_1$-blockers and phosphodiesterase type 5 inhibitors.\(^{33,45,46}\) In contrast, the prospect of dry mouth, associated with antimuscarinic use,\(^{33}\) had a minimal impact on men’s preferences. This could be due to the fact that headache/migraine and dizziness are reported to impact work- and/or social-related activities.\(^{57,48}\) It may also be because few men in our sample ($n=14, 5.67\%$) had previously received antimuscarinic agents.
Trade-offs are likely required when considering the efficacy and side effect profiles of pharmacological treatment options available to men with LUTS/BPH. Considering the net benefit of a medicine for men based on the estimated WTP, a medicine targeting a mix of symptoms with a particular focus on reducing urgency of urination combined with reducing night-/daytime frequency is likely to provide the highest benefit for men. Such a combination of benefits may compensate for the negative impact of side effects. For example, a therapy that reduced urgency to moderate urgency and provided one less nighttime urination (combined WTP estimate, £30.58/month) has the potential to compensate for the side effect of erectile dysfunction (WTP estimate, −£30.07/month). In contrast, the WTP estimates suggest that a therapy which only reduced nighttime frequency by one or two episodes per night (£6.65/month) would not compensate for decreased sexual desire (−£19.66/month).

A potential limitation of our study was that the DCE was limited to seven attributes that included urgency but excluded other LUTS symptoms (ie, voiding and/or post-micturition) and other potential treatment-related factors as described earlier. However, increasing the number of attributes increases the difficulty of completing the tasks, meaning that respondents may provide less precise answers or may not consider all of the attributes. In addition, a greater number of attributes increases the number of choices needed to ensure the impact of each attribute can be identified in the statistical model. We consider that the selection of these seven attributes balances comprehensiveness of the symptoms and medication description and practicality of the survey administration. Another potential limitation of this DCE was that all surveys were completed online thereby excluding men without access to the Internet. However, the men were spread geographically across the UK limiting any potential regional bias and Internet usage in the UK is high, even in people aged ≥65 years. The results of this study are not generalizable beyond the UK.

**Conclusion**

In conclusion, treatment choice and the efficacy and tolerability profiles of medication are important factors for men with moderate-to-severe LUTS/BPH. Receiving medicines that significantly reduce storage symptoms, such as urge incontinence and nighttime urinary frequency, is perceived as the most important benefit of medication. However, men would prefer to avoid medicines that are likely to elicit sexual and nonsexual side effects. This study also highlights that treatment decisions may be optimized by understanding what symptoms and treatment benefits/side effects are of the greatest concern to patients with LUTS/BPH.

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**Author contributions**

All authors contributed toward data analysis, drafting and critically revising the paper and agree to be accountable for all aspects of the work.

**Disclosure**

DI, SH, MR, and VW are employed by the University of Aberdeen and performed the research which was funded by Astellas. CM, CN and JN are employees of Astellas.

**References**


Supplementary material
Description of mixed-logit model used to analyze discrete choice experiment responses

This approach allows preferences for attributes to vary across individuals and requires the specification of a distribution for each coefficient hypothesized to vary across participants. The distribution of the coefficients of the day- and nighttime frequency attributes was assumed to be log normal, given that all respondents ought to prefer reductions in the frequency of urination.1,2 All other coefficients were assumed to have a normal distribution, allowing for both positive and negative effects on utility. The estimation was based on 3,000 Halton draws using Stata version 13 (StataCorp LP, College Station, TX, USA).3 A fixed parameter was assumed for the cost attribute, thus assuming all individuals had the same negative preference, and allowing easier estimation of willingness to pay for all other attributes.4

References