

Positive association between *CD44* gene rs13347 C>T polymorphism and risk of cancer in Asians: a systemic review and meta-analysis

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Background: Cluster of differentiation 44 (CD44) is an important surface marker of cancer stem cells in a variety of tumors. A number of previous studies have been conducted to investigate the association between *CD44* gene rs13347 C>T polymorphism and cancer risk in humans; nevertheless, the results remain controversial. We therefore performed this meta-analysis to confirm the role of this polymorphism in susceptibility to human cancer.

Materials and methods: The studies published up to December 2015 were searched in PubMed, Web of Science, and China National Knowledge Infrastructure databases. Twelve eligible case-control studies were identified, involving a total of 6,982 cases and 7,430 controls. Pooled odds ratio (OR) and corresponding 95% confidence interval (CI) were calculated using a fixed or random-effect model to estimate the strength of the association.

Results: The results of the overall analyses indicated that *CD44* gene rs13347 polymorphism was significantly associated with cancer risk in Asians (CT vs CC: OR = 1.35, 95% CI = 1.12–1.62; TT vs CC: OR = 1.99, 95% CI = 1.52–2.60; TT + CT vs CC: OR = 1.41, 95% CI = 1.16–1.71; and TT vs CC + CT: OR = 1.74, 95% CI = 1.41–2.14), especially in Chinese population (CT vs CC: OR = 1.42, 95% CI = 1.16–1.75; TT vs CC: OR = 2.13, 95% CI = 1.58–2.86; TT + CT vs CC: OR = 1.50, 95% CI = 1.21–1.87; and TT vs CC + CT: OR = 1.80, 95% CI = 1.43–2.26). In stratified analyses by cancer types, there was evidence for an association between this polymorphism and nasopharyngeal cancer and breast cancer, respectively.

Conclusion: The results of this meta-analysis suggest that the *CD44* gene rs13347 C>T polymorphism is associated with elevated risk of human cancer in Asians, especially in Chinese population. Further well-designed studies on a larger population covering other ethnicities should be carried out to validate our results.

Keywords: cancer, CD44, polymorphism, meta-analysis

Background

Cancer is currently a serious public health burden in the world, which results from interactions between accumulations of genetic mutation and environmental risk factors.^{1,2} In recent years, a small subgroup of cancer cells, called cancer stem cells (CSCs), have been proved to be responsible for tumor initiation, progression, metastasis, recurrence, and drug resistance and cause cancer heterogeneity.³ Thus, eliminating CSCs is considered to be efficient and critical in cancer therapy.⁴ Several cell surface markers have been found to identify CSCs, and overexpression of these markers, such as cluster of differentiation 44 (CD44), CD24, CD133, CD166, and ALDH1A1, indicates severe clinical features and poor prognosis in a number of cancers.^{5–10} Among these cell surface markers, CD44 is one of the most frequently

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reported markers, which was utilized to isolate CSCs from other cancer cells.^{11–13}

CD44, a transmembrane glycoprotein, is ubiquitously expressed in many cell types.¹⁴ As a cell surface receptor for hyaluronate and osteopontin, CD44 is involved in many biological and physiological processes including cell migration, hematopoiesis, lymphocyte homing, embryonal development, and apoptosis.^{14,15} Besides its regulation of cellular processes, CD44 plays a critical role in tumor cell proliferation, differentiation, invasion, and migration, which contributes to the progression and metastasis of tumors.^{16–19} Therefore, overexpression of CD44 leads to the development of tumors and poor prognosis of several human malignancies.^{20–22} In addition, it has been reported that CD44⁺ cancer cells represent enhanced resistance to chemotherapy in the nude mice-engrafted tumor model.^{23,24} Recent studies have revealed that the genetic variants of *CD44* gene could influence tumor cell growth and migration,^{25,26} which were associated with the risk prediction and prognosis of various human cancers.^{27–29}

The gene encoding CD44 is located on chromosome 11p13.³⁰ A growing number of studies have been carried out to investigate the effects of several *CD44* gene single nucleotide polymorphisms (SNPs) on cancer risk. Among them, the rs13347 C>T polymorphism, located on the 3'-untranslated region (3'-UTR) of *CD44* gene, is the most frequently studied.^{20,31,32} This SNP might influence the *CD44* gene expression, since it has been reported that subjects carrying TT and CT genotypes had remarkably higher levels of CD44 protein than those carrying CC genotype in breast cancer, nasopharyngeal carcinoma, and acute myeloid leukemia.^{28,31,33}

To date, a number of case-control studies have been conducted to evaluate the role of *CD44* gene rs13347 polymorphism in predisposition to several human cancers.^{28,29,32–35} Nevertheless, the results from different articles remain controversial. To clarify the association between *CD44* gene rs13347 polymorphism and risk of cancer, we performed this meta-analysis by integrating data from eligible published studies.

Materials and methods

Search strategy

Relevant reports were retrieved by searching the electronic databases: PubMed, Web of Science, and China National Knowledge Infrastructure (CNKI; from inception to December 20, 2015), using the following keywords: ("CD44") and ("tumor" or "cancer" or "carcinoma" or "neoplasm" or "malignancy") and ("polymorphism" or "polymorphisms" or "SNP" or "variant" or "variation"). The

search was filtered to English-language journals in PubMed. Besides, we also performed a manual search among the references of the relevant publications and related articles. The studies with overlapping data by the same investigators or based on the same population were checked prudently, and the most recent articles covering the largest numbers of cases and controls would be included.

Inclusion and exclusion criteria

The eligible studies in this meta-analysis were required to strictly follow the predetermined criteria: 1) use a case-control study design, 2) evaluate the association between *CD44* rs13347 C>T polymorphism and risk of cancer, and 3) report an estimation of odds ratio (OR) and 95% confidence interval (CI), or sufficient data to allow calculation of these two statistics. The main exclusion criteria were studies 1) that did not use a case-control design (eg, case reports, letters, animal studies, reviews, and editorials), 2) that are duplicate of previous publications, 3) that involve inherited cancers, 4) with sample size of cases or controls <100, and 5) in which genotype distribution of controls is not in agreement with the Hardy-Weinberg equilibrium.

Data extraction

All the eligible studies were reviewed by two authors independently to extract useful data. The following information were collected: the name of the first author, the year of publication, the ethnicity of study population, the country of origin, sample size of cases and controls, source of controls (hospital based or population based), genotyping method, and genotype distributions of cases and controls. The disagreements in this step were solved by rechecking the original data to reach a consensus.

Statistical analysis

Hardy-Weinberg equilibrium was estimated by the goodness-of-fit test based on the chi-square test in the control group of each study.³⁶ Pooled analysis was conducted to estimate the strength of the association between *CD44* rs13347 C>T polymorphism and cancer risk, using an OR with a corresponding 95% CI. The pooled ORs were calculated by comparisons with a codominant model (CT vs CC and TT vs CC), a dominant model (TT + CT vs CC), and a recessive model (TT vs CC + CT). The values of the pooled ORs were tested by the Z-test.³⁷ Stratified analyses were further performed based on country (People's Republic of China or India), specific cancer types (nasopharyngeal cancer, gallbladder cancer, or breast cancer), source of controls

(hospital based or population based), and genotyping method (TaqMan or matrix-assisted laser desorption/ionization time-of-flight). Heterogeneity among the included studies was evaluated by the chi-square-based Q -test.³⁸ Pooled ORs were calculated using a fixed (Mantel–Haenszel method³⁹) or a random (DerSimonian–Laird method⁴⁰) effective model, according to the absence ($P>0.10$ and $I^2<50\%$) or presence ($P<0.10$ or $I^2>50\%$) of heterogeneity. Sensitivity analyses were performed by omitting one study each time to evaluate the stability of the results. The potential publication bias of the included studies was assessed by Begg’s funnel plots graphically and Egger’s test quantitatively.⁴¹

All the statistical calculations were carried out with Stata/SE software Version 12.0 (StataCorp LP, College Station, TX, USA), using two-sided P -values, and $P<0.05$ was considered to be significant.

Results

Literature search and characteristics of eligible studies

A total of 890 potential relevant records were retrieved through the search strategy described previously. Eight hundred and seventy records were excluded after title or

abstract scanning. After full-text reviewing, eight studies were excluded. Finally, 12 studies were eligible for pooled analysis. The flow process of detailed literature search and study selection is shown in Figure 1.

There were 12 case–control studies included in this meta-analysis,^{28,29,31–35,42–46} involving a total of 6,982 cases and 7,430 controls. All the studies were conducted on an Asian population. Among them, nine originated from the People’s Republic of China and the other three from India. There was no study conducted on Caucasians or Africans. Cancer types included bladder cancer, non-small-cell lung cancer, gallbladder cancer, colorectal cancer, nasopharyngeal cancer, hepatocellular carcinoma, oral cancer, breast cancer, and acute myeloid leukemia. The distribution of genotypes in all the control groups was in agreement with Hardy–Weinberg equilibrium. The characteristics of each eligible study are listed in Table 1.

Quantitative synthesis

The pooled results of the present meta-analysis are shown in Table 2 and Figure 2. There was an association between *CD44* gene rs13347 polymorphism and risk of cancer in the overall analyses. Significantly elevated cancer risk was revealed in the codominant genetic model (CT vs CC: OR =1.35, 95%

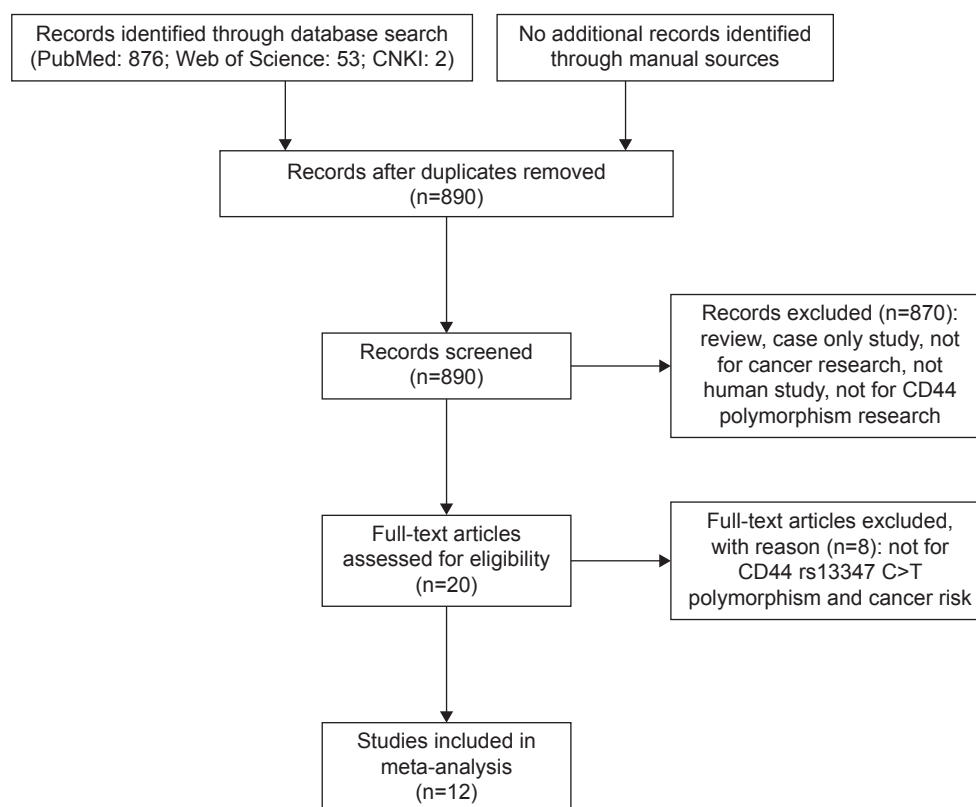


Figure 1 Flow diagram of the study identification process.

Table 1 Characteristics of eligible studies in the meta-analysis

ID	Reference	Year	Country	Source of control	Cancer type	Genotyping method	Sample size case/control	Genotype (case/control)			HWE
								CC	CT	TT	
1	Weng et al ⁴⁴	2015	People's Republic of China	HB	Bladder	TaqMan	275/275	138/143	111/117	26/15	0.153
2	Liu et al ³²	2015	People's Republic of China	PB	NSCLC	TaqMan	234/468	179/337	51/121	4/10	0.823
3	Yadav et al ⁴⁶	2016	India	PB	Gallbladder	TaqMan	610/250	378/162	201/80	31/8	0.620
4	Wu et al ⁴⁵	2015	People's Republic of China	HB	Colorectal	MALDI-TOF	946/989	416/578	441/348	89/63	0.279
5	Lou et al ⁴²	2014	People's Republic of China	HB	Nasopharyngeal	TaqMan	272/489	104/288	126/174	42/27	0.915
6	Chou et al ³⁴	2014	People's Republic of China	HB	Hepatocellular	TaqMan	203/561	110/295	72/223	21/43	0.924
7	Chou et al ³⁵	2014	People's Republic of China	HB	Oral	TaqMan	599/561	287/295	262/223	50/43	0.924
8	Sharma et al ⁴³	2014	India	NA	Gallbladder	TaqMan	405/200	293/154	104/42	8/4	0.572
9	Wu et al ³¹	2015	People's Republic of China	PB	AML	MALDI-TOF	421/461	163/254	196/171	62/36	0.340
10	Tulsyan et al ²⁹	2013	India	HB	Breast	TaqMan	258/241	191/178	60/57	7/6	0.577
11	Xiao et al ³³	2013	People's Republic of China	PB	Nasopharyngeal	MALDI-TOF	906/943	386/606	418/297	102/40	0.637
12	Jiang et al ²⁸	2013	People's Republic of China	PB	Breast	MALDI-TOF	1,853/1,992	812/1,146	850/727	190/119	0.795

Abbreviations: PB, population based; HB, hospital based; NA, not available; HWE, Hardy–Weinberg equilibrium; AML, acute myeloid leukemia; NSCLC, non-small-cell lung cancer; MALDI-TOF, matrix-assisted laser desorption/ionization time-of-flight.

CI=1.12–1.62 and TT vs CC: OR=1.99, 95% CI=1.52–2.60), dominant model (TT+CT vs CC: OR=1.41, 95% CI=1.16–1.71), and recessive model (TT vs CC + CT: OR=1.74, 95% CI=1.41–2.14). In the subgroup analyses stratified by country, there was evidence in Chinese population for an association between this SNP and cancer risk (CT vs CC: OR=1.42, 95% CI=1.16–1.75; TT vs CC: OR=2.13, 95% CI=1.58–2.86; TT + CT vs CC: OR=1.50, 95% CI=1.21–1.87; and TT vs CC + CT: OR=1.80, 95% CI=1.43–2.26). According to the source of control, both population based and hospital based subgroups were linked to cancer risk. For specific cancer types, increased risk among studies of nasopharyngeal cancer and breast cancer was observed in several genetic models.

Test for heterogeneity and sensitivity analysis

There was significant heterogeneity in the overall comparisons under all the genetic models (CT vs CC: $P<0.001$ and $I^2=82.4\%$ for heterogeneity, TT vs CC: $P<0.001$ and $I^2=67.7\%$ for heterogeneity, TT + CT vs CC: $P<0.001$ and $I^2=85.8\%$ for heterogeneity, and TT vs CC + CT: $P=0.031$ and $I^2=48.3\%$ for heterogeneity; Table 2). The source of heterogeneity was investigated by covariates, such as country, cancer type, source of control, and genotype method. According to subgroup

analyses, country and cancer type might be the source of heterogeneity. To explore the main origin of heterogeneity, a meta-regression analysis was performed in the comparison TT vs CC, which indicated that cancer type contributed to the most proportion in heterogeneity. Moreover, two studies on nasopharyngeal cancer^{33,42} were under suspicion. After omitting these two studies, the pooled OR was not altered (TT vs CC: OR=1.74, 95% CI=1.39–2.16), whereas the heterogeneity remarkably decreased (TT vs CC: $P=0.110$ and $I^2=34.7\%$ for heterogeneity).

Sensitivity analysis was performed, and the pooled ORs were not influenced qualitatively in all the genetic models by removing any single study, which indicated that the pooled results of this meta-analysis were statistically stable (Figure 3).

Publication bias

The potential publication bias of the eligible studies was assessed by Begg's funnel plots graphically and Egger's test statistically. The shapes of funnel plots in all the genetic models did not indicate any evidence of an obvious asymmetry (Figure 4). Meanwhile, Egger's test revealed that there was no publication bias either (CT vs CC: $P=0.139$, TT vs CC: $P=0.755$, TT + CT vs CC: $P=0.186$, and TT vs CC + CT: $P=0.975$).

Table 2 Pooled analysis of association between CD44 rs13347 C>T polymorphism and cancer risk in Asians

Study	N	CT vs CC			TT vs CC			TT + CT vs CC			TT vs CC + CT		
		OR (95% CI)	P _h	I ² (%)	OR (95% CI)	P _h	I ² (%)	OR (95% CI)	P _h	I ² (%)	OR (95% CI)	P _h	I ² (%)
Overall	12	1.35 (1.12–1.62)	<0.001	82.4	1.99 (1.52–2.60)	<0.001	67.7	1.41 (1.16–1.71)	<0.001	85.8	1.74 (1.41–2.14)	0.031	48.3
Country													
People's Republic of China	9	1.42 (1.16–1.75)	<0.001	84.7	2.13 (1.58–2.86)	<0.001	73.9	1.50 (1.21–1.87)	<0.001	87.6	1.80 (1.43–2.26)	0.012	59.0
India	3	1.11 (0.89–1.37)	0.619	0.0	1.36 (0.77–2.39)	0.753	0.0	1.13 (0.92–1.39)	0.672	0.0	1.32 (0.75–2.33)	0.744	0.0
Cancer type													
Nasopharyngeal	2	2.15 (1.82–2.54)	0.613	0.0	4.10 (2.99–5.61)	0.827	0.0	2.39 (2.04–2.81)	0.802	0.0	2.95 (2.17–3.99)	0.787	0.0
Gallbladder	2	1.16 (0.90–1.49)	0.473	0.0	1.46 (0.75–2.84)	0.637	0.0	1.18 (0.93–1.51)	0.625	0.0	1.41 (0.73–2.73)	0.502	0.0
Breast	2	1.32 (0.80–2.19)	0.020	81.5	2.17 (1.71–2.76)	0.210	36.5	1.36 (0.79–2.33)	0.009	85.3	1.76 (1.39–2.22)	0.387	0.0
Source of control													
PB	5	1.45 (1.08–1.95)	<0.001	86.9	2.39 (1.65–3.47)	0.022	65.2	1.53 (1.12–2.09)	<0.001	89.6	1.99 (1.67–2.37)	0.142	41.9
HB	6	1.26 (0.96–1.65)	<0.001	80.8	1.79 (1.20–2.67)	0.007	68.5	1.33 (1.00–1.76)	<0.001	83.5	1.60 (1.17–2.19)	0.061	52.7
Genotyping method													
TaqMan	8	1.12 (0.91–1.37)	0.006	64.8	1.56 (1.03–2.35)	0.012	61.2	1.16 (0.93–1.46)	<0.001	73.9	1.50 (1.08–2.09)	0.089	43.4
MALDI-TOF	4	1.80 (1.64–1.97)	0.115	49.4	2.58 (1.93–3.45)	0.039	64.0	1.94 (1.66–2.27)	0.031	66.3	1.96 (1.53–2.51)	0.091	53.6

Notes: N, number of studies; P_h, P-value of Q-test for heterogeneity.

Abbreviations: PB, population based; HB, hospital based; MALDI-TOF, matrix-assisted laser desorption/ionization time-of-flight; OR, odds ratio; CI, confidence interval.

Discussion

Cancer is a group of heterogeneous diseases that result from genetic and environmental factors, as well as their interactions. According to CSC theory, CSCs can drive cancer initiation, progression, and metastasis.³ As an important CSC surface marker in variant tumors, CD44 is reportedly a risk factor and poor prognostic molecular marker in various cancers.^{47–50} CD44 plays a crucial role in the invasion and migration of cancer cells. In addition, overexpression of CD44 in primary tumors has been revealed to be associated with early metastasis in several human cancers.^{20,29,51} CD44 gene rs13347 C>T polymorphism is a common genetic variation in the 3'-UTR. It has been proved that the expression of CD44 was modulated by mRNAs, and rs13347 C>T polymorphism could alter has-mir-509-3p-mediated CD44 gene expression activity. Recently, a number of studies have been carried out to evaluate the association between this SNP and risk of cancer; however, the results were conflicting. The possible reason may be that the effects of CD44 on the development of different cancer types are not the same.

To our knowledge, the present meta-analysis, including 6,982 cases and 7,430 controls from 12 case-control studies, is the first comprehensive study to assess the association between CD44 gene rs13347 polymorphism and risk of cancer. We found an association between CD44 gene rs13347 polymorphism and human cancer risk by pooling all the data from eligible studies. The results were robust because the pooled ORs did not alter statistically in sensitivity analyses. Stratified analyses suggested that this association was mainly in a Chinese population. According to specific cancer types, this polymorphism was linked to elevated risk of gallbladder cancer, nasopharyngeal cancer, and breast cancer, the significance of which was limited because there were only two studies for each cancer type. Since there was significant heterogeneity by the Q-test in all the genetic models, a meta-regression analysis was performed to explore the origin of heterogeneity. As a result, different cancer types were considered to be the main source of interstudy variance, especially two studies on nasopharyngeal cancer. By removing these two studies, the heterogeneity significantly decreased but pooled ORs were not altered. However, this kind of heterogeneity was hard to eliminate in the present meta-analysis, because there were only one or two studies for one specific cancer type. Moreover, pooling data from different cancer types might affect the significance of this meta-analysis because different cancer types might give rise to different host responses,

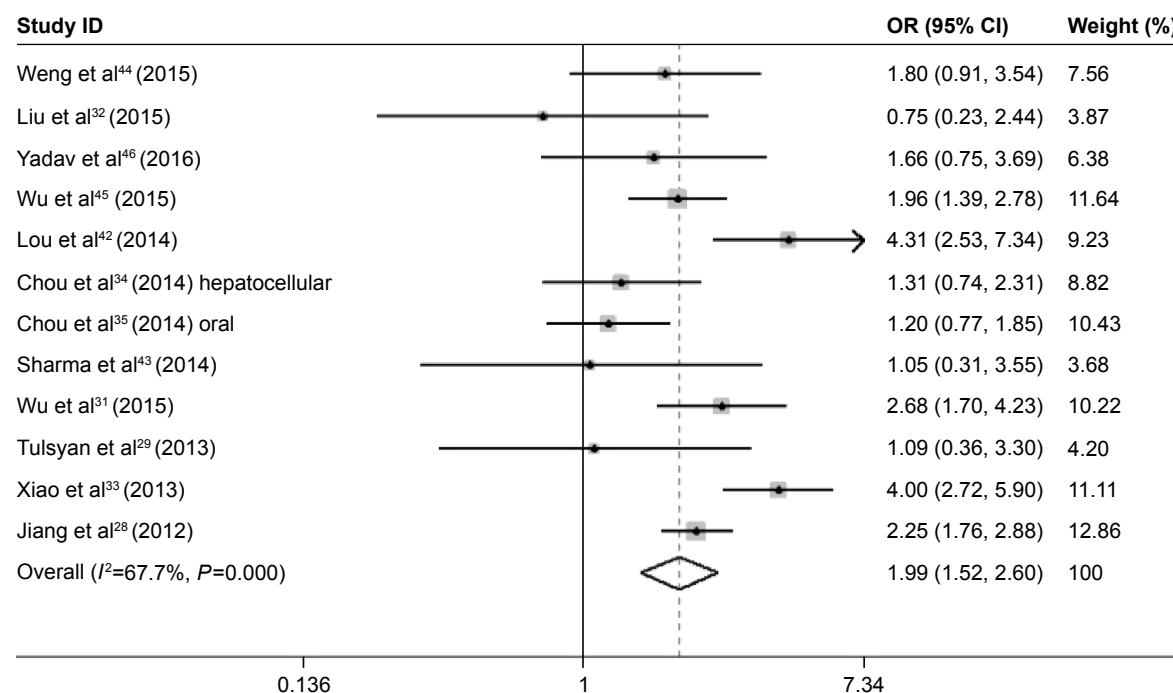


Figure 2 Forest plot of the association between *CD44* gene rs13347 C>T polymorphism and risk of cancer in the overall analysis (TT vs CC).

Note: Weights are from random effects analysis.

Abbreviations: OR, odds ratio; CI, confidence interval.

Meta-analysis estimates, given named study is omitted

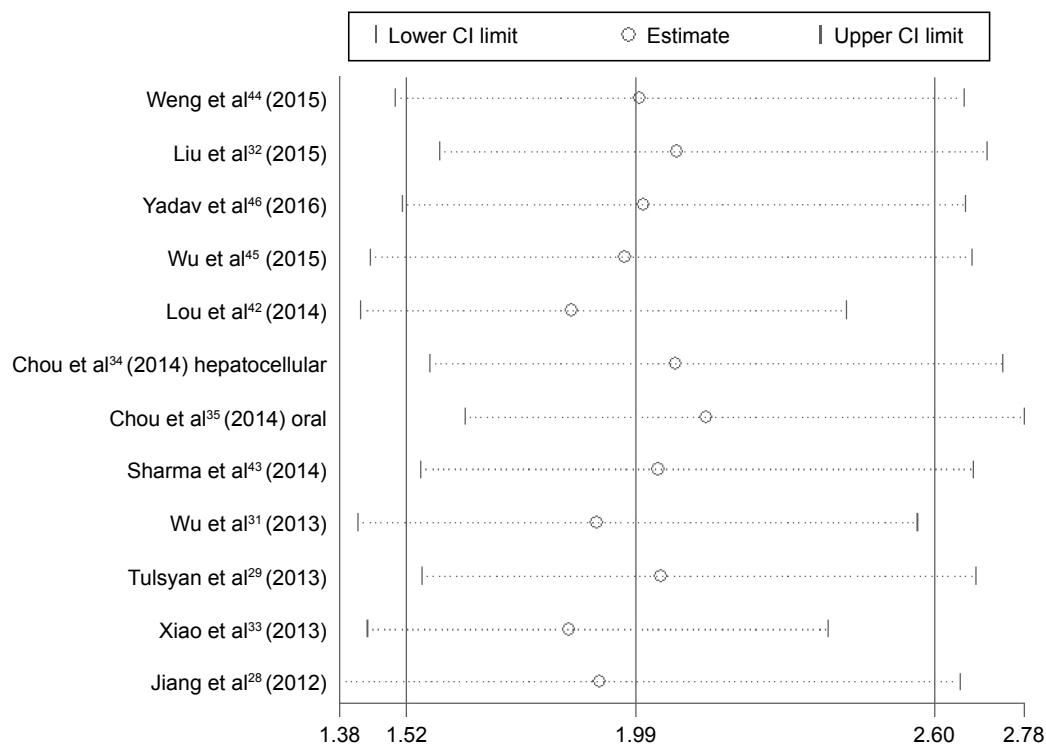


Figure 3 Sensitivity analysis of the association between *CD44* gene rs13347 C>T polymorphism and risk of cancer in the overall analysis (TT vs CC).

Abbreviation: CI, confidence interval.

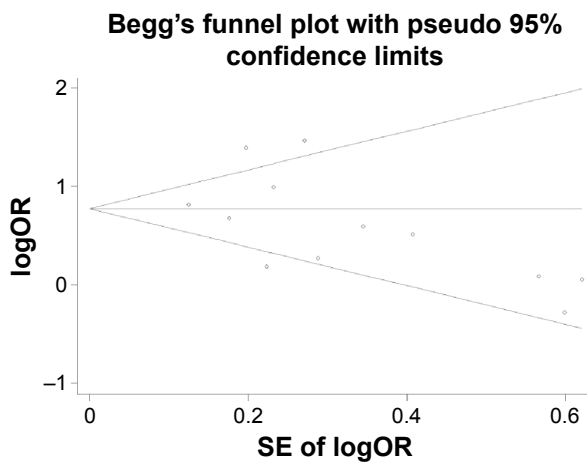


Figure 4 Begg's funnel plot of publication bias test in the overall analysis (TT vs CC); each point represents a single study for the indicated association.

Abbreviations: logOR, natural logarithm of the OR; SE of logOR, standard error of the logOR; OR, odds ratio.

and the interactions between different environmental factors and host might also influence the susceptibility to different cancer types.

Some limitations of the present meta-analysis should be pointed out. First, the pooled results were calculated based on unadjusted estimates, which limited us to perform a more precise assessment on adjusted estimates by several important factors such as sex, age, lifestyle, etc. Thus, lack of the baseline information restricted further evaluation of the potential interactions, because malignancy predisposition might be influenced by gene–gene and gene–environment interactions. Second, most of the included studies just focused on the relationship between *CD44* gene rs13347 polymorphism and cancer risk, which made it hard to assess the effects of *CD44* gene haplotypes composed of different *CD44* gene SNPs on carcinogenesis. There was evidence that *CD44* gene rs187115 A>G and rs115214213 T>C polymorphisms were associated with cancer risk.^{35,42,44} Thus, the status of other *CD44* gene polymorphisms might cover up the effects of rs13347 T>C polymorphism, which could lead to controversial results among different studies.

Despite these limitations, advantages in this meta-analysis should also be acknowledged. First, the statistical power was remarkably increased since we pooled a substantial number of cases and controls. Second, the quality of all the eligible studies met the inclusion criteria completely and strictly. Third, there was no publication bias observed through Begg's funnel plots and Egger's test, which indicated that the pooled outcomes should be unbiased.

Conclusion

The results of the present meta-analysis were robust and credible. The relationship between *CD44* gene rs13347 polymorphism and cancer risk was assessed, and this SNP was associated with elevated cancer risk in Asians. To draw a more conclusive result, further studies should be conducted with more detailed information on individuals and environmental factors, concerning the effects of different haplotypes and other SNPs and enrolling properly identified cases and well-matched controls, especially in other ethnicities including Africans and Caucasians, to validate the role of *CD44* gene rs13347 C>T polymorphism in carcinogenesis.

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Disclosure

The authors report no conflicts of interest in this work.

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