

Association between the COMT Val158Met polymorphism and risk of cancer: evidence from 99 case–control studies

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Abstract: Catechol-*O*-methyltransferase (COMT) plays a central role in DNA repair and estrogen-induced carcinogenesis. Many recent epidemiologic studies have investigated the association between the COMT Val158Met polymorphism and cancer risk, but the results are inconclusive. In this study, we performed a meta-analysis to investigate the association between cancer susceptibility and COMT Val158Met in different genetic models. Overall, no significant associations were found between COMT Val158Met polymorphism and cancer risk (homozygote model: odds ratio [OR] =1.05, 95% confidence interval [CI] = [0.98, 1.13]; heterozygote model: OR =1.01, 95% CI = [0.98, 1.04]; dominant model: OR =1.02, 95% CI [0.97, 1.06], and recessive model: OR =1.03, 95% CI [0.97, 1.09]). In the subgroup analysis of cancer type, COMT Val158Met was significantly associated with increased risks of bladder cancer in recessive model, and esophageal cancer in homozygote model, heterozygote model, and dominant model. Subgroup analyses based on ethnicities, COMT Val158Met was significantly associated with increased risk of cancer in homozygote and recessive model among Asians. In addition, homozygote, recessive, and dominant models were significantly associated with increased cancer risk in the subgroup of allele-specific polymerase chain reaction genotyping. Significant associations were not observed when data were stratified by the source of the controls. In summary, this meta-analysis suggested that COMT Val158Met polymorphism might not be a risk factor for overall cancer risk, but it might be involved in cancer development at least in some ethnic groups (Asian) or some specific cancer types (bladder and esophageal cell cancer). Further evaluations of more preclinical and epidemiological studies are required.

Keywords: COMT, polymorphism, cancer, meta-analysis, susceptibility

Introduction

Cancer constitutes an enormous burden on the society in more and less economically developed countries alike.^{1,2} Based on GLOBOCAN estimates, ~14.1 million new cancer cases and 8.2 million deaths occurred in 2012 worldwide.¹ According to the development trend, the new cases in 2030 will reach 22.2 million.² It is well known that the etiology and development of cancer are as a result of complex interactions between genetic and environmental factors.³ Genes determine the susceptibility of individual to environment, and environmental factors often damage the DNA in turn. Recent studies have shown that host genetic factors are closely related to the pathophysiology of many human cancers.⁴ The most common form of genetic variation, that is, single-nucleotide polymorphisms, is known to contribute individual susceptibility to cancer.⁵ Therefore, it is anticipated that the identification of key gene polymorphisms associated with cancer risk is essential for predicting risk of individuals, and that it

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will greatly assist the global control and therapeutic strategies of this lethal disease.

The catechol-*O*-methyltransferase (COMT) gene is located on chromosome 22q11.2 and consists of six exons.⁶ It is an important enzyme involved in the inactivation of endogenous catecholamine and catechol estrogens. Catechol estrogens have been shown to have the ability to damage DNA and carcinogenic potential.⁷ Therefore, the loss of or changes in COMT is supposed to contribute to genomic instability and tumor genesis. In line with these considerations, it has been hypothesized that COMT Val158Met might influence the development of all cancers. Up to now, many researches have indicated the link between COMT polymorphism and cancer susceptibility. Several polymorphisms have been identified, including the widely studied polymorphism Val158Met(rs4680).⁸ This change has been associated with a three- to four-fold decrease in the activity of COMT compared with the wild-type COMT-Val allele.^{9,10} It is biologically reasonable to hypothesize that women who carry mutant COMT-Met allele may have higher cancer risks.

In recent years, many studies have investigated the relationship between COMT Val158Met polymorphism in different races and different types of cancer, but the results were inconclusive or controversial.^{11–101} The inconsistent conclusions may be due to a possible minor effect of the polymorphism on cancer or the small sample size in studies with inadequate statistical power of complex traits. Meta-analysis is a powerful statistical tool to pool different studies to overcome deficiencies such as small sample size and to provide more reliable results. Although some previous meta-analyses have reported the association between COMT Val158Met polymorphism and ovarian cancer (up to eight case–control studies included),^{102,103} breast cancer (up to 56 case–control studies included),^{65,104–108} endometrial cancer (up to seven case–control studies included),^{103,109,110} prostate cancer (up to six case–control studies included),^{111–113} and lung cancer (evidence from six case–control studies),¹¹⁴ only specific cancer types or race populations were included, which led to their limitations. To update the results of previous meta-analyses and to provide a more precise assessment of the association between COMT Val158Met and cancer risk, we performed a comprehensive meta-analysis by including the most recent and relevant articles.

Materials and methods

Identification and eligibility of relevant studies

The meta-analysis was conducted following the criteria of Preferred Reporting Items for Systematic Reviews and

Meta Analyses. A comprehensive literature search was performed using the PubMed, Cochrane Library, Chinese National Knowledge Infrastructure, and EMBASE database for relevant articles published (the last search update was February 15, 2015) with keywords “COMT”, “Catechol-*O*-methyltransferase”, “Val158Met”, “rs4680”, “single nucleotide polymorphism”, “polymorphism”, “Variant”, “Mutation”, “Cancer”, “tumor”, “neoplasm”, “malignancy”, or “Carcinoma”. In addition, studies were identified by a manual search of reviews and retrieved studies. Search results were restricted to human populations, and the articles were written in English or Chinese. We included all the case–control studies and cohort studies that have investigated the association between COMT Val158Met polymorphisms and cancer risk with genotyping data. All eligible studies were retrieved, and their bibliographies were checked for other relevant publications. When the same patient population was used in several publications, only the most recent, the largest or the most complete study was included.

Assessment of study quality

The quality of the included studies was assessed by the Newcastle–Ottawa Scale (NOS; http://www.ohri.ca/programs/clinical_epidemiology/oxford.asp),¹¹⁵ including selection of groups, comparability of groups, and ascertainment of exposure. The NOS score ranges from 0 to 10 stars. Studies with NOS score > five stars were included in the final analysis.

Inclusion criteria

All studies were included if they met the following criteria: 1) only the case–control studies or cohort studies were considered, 2) studies that investigated the COMT Val158Met polymorphism and the risk of cancer susceptibility were included, and 3) the genotype distribution of the polymorphism in cases and controls was described in details, and the results were expressed as odds ratio (OR) and corresponding 95% confidence interval (95% CI). Major reasons for exclusion of studies were as follows: 1) not for cancer research, 2) only case population, 3) duplicate of previous publication, and 4) review articles, editorials, case reports, studies with preliminary results not on COMT Val158Met polymorphism or outcome, and investigations of the role of COMT expression related to disease. Ethics approval for the study was granted by the local institute, the People’s Hospital of Three Gorges University Ethics Committee.

Data extraction

Using a standardized form, data from published studies were extracted independently by two reviewers to evaluate

their eligibility for inclusion by first screening the title and abstract of each identified reference and then establishing the eligibility of the included papers based on the full text when necessary. For each included study, the following information was collected: first author, year of publication, region, study design, sample size, source of control, genotyping method, allele or genotype frequencies, and evidence of Hardy–Weinberg equilibrium (HWE). Any discrepancy between the two reviewers was resolved by discussion and consultation with a third reviewer.

Statistical analysis

ORs and their 95% CIs were used to determine the strength of association between the COMT Val158Met polymorphism and cancer risk. The significance of the pooled OR was determined using the Z test, and $P < 0.05$ was considered statistically significant. Homozygote model (AA vs GG), heterozygote model (GA vs GG), dominant model (GA + AA vs GG), and recessive model (AA vs GG + GA) were investigated. Subgroup analysis was performed by ethnicity, cancer type (if one cancer type contained less than two studies, it was defined as “other”), source of controls, and hospital or population controls. Effective modification by a subgroup was assessed by testing the interaction between genotypes and stratification variables by using logistic regression analyses (random-effects estimator). HWE was tested using the chi-square test among controls, and $P < 0.05$ was considered a significant departure from HWE. If the P -value for heterogeneity was > 0.05 and $I^2 < 50\%$, indicating an absence of heterogeneity among studies, the fixed-effects model (the Mantel–Haenszel method) was used.¹¹⁶ In contrast, if either the P -value for heterogeneity was ≤ 0.05 or I^2 was $\geq 50\%$, indicating heterogeneity among the studies, the more appropriate random-effects model (the DerSimonian and Laird method) was used.¹¹⁷ Sensitivity analyses were performed to assess the stability of the results. Begg’s funnel plots were used to diagnose potential publication bias, and $P < 0.05$ was used to indicate possible publication bias.¹¹⁸ All analyses were performed using RevMan 5.3 (updated in March 2012 by the Cochrane Collaboration). P -values were based on two-sided tests.

Results

Literature search and meta-analysis databases

Following the searching strategy, 337 potentially relevant studies were retrieved. After title and abstract screening, nine of them were ruled out because of repeated data. A total of 202 irrelevant articles were excluded. In addition, after the

full texts of the remaining 182 articles were read, 90 articles were excluded for the following reasons: article was a review ($n=27$), articles had insufficient data ($n=13$), articles were not related to cancer ($n=34$), and articles were not related to COMT ($n=16$). A total of 92 publications with full text were selected and were subjected to further examination. Because seven studies included more than one ethnicity, genotyping method, control source, or tumor type and were performed by the same author, we treated them separately in this meta-analysis. Of those, 99 case–control studies with 43,085 cancer cases and 57,882 control subjects were included in our meta-analysis. A flow chart showing the detailed steps of study selection is shown in Figure 1. All studies were case–control studies with the following tumor-type distribution: three were conducted for bladder cancer, two for renal cancer, nine for endometrial cancer, eight for ovarian cancer, 62 for breast cancer, six for lung cancer, three for liver cancer, two for colon cancer, two for esophageal cell cancer, one for thyroid cancer and non-Hodgkin lymphoma, and one for testicular germ cell tumor. Fifty studies investigated the risks in Caucasian populations, 35 studies investigated Asian populations, ten studies investigated mixed populations, and the remaining studies were conducted in African populations. Five main genotyping methods were used such as polymerase chain reaction–restriction fragment length polymorphism (PCR–RFLP), TaqMan, sequencing, matrix-assisted laser desorption ionization time of flight mass spectrometry (MALDI–TOF), and allele-specific PCR (AS–PCR). By source of controls, 50 studies were population based, 45 studies were hospital based, and four studies were not clear. The distribution of the genotypes in the control subjects was in agreement with HWE, except for eight studies.^{34,37,70,72,80,88,95,119} The quality assessment showed that the quality scores ranged from 5 to 9 with a median score of 6, suggesting that all studies were of high quality. The main characteristics of the eligible studies are listed in Table 1.

Quantitative synthesis

Overall, no significant associations between COMT Val158Met and cancer risk were found using homozygote model (OR = 1.05, 95% CI [0.98, 1.13]), heterozygote model (OR = 1.01, 95% CI [0.98, 1.04]), dominant model (OR = 1.02, 95% CI [0.97, 1.06]), or recessive model (OR = 1.03, 95% CI [0.97, 1.09]).

Significant heterogeneity was observed among the 99 studies on COMT Val158Met polymorphism. To explore the source of heterogeneity, we performed stratified analyses on ethnicity, cancer type, source of controls, and genotyping method. In the subgroup analysis on cancer type, COMT

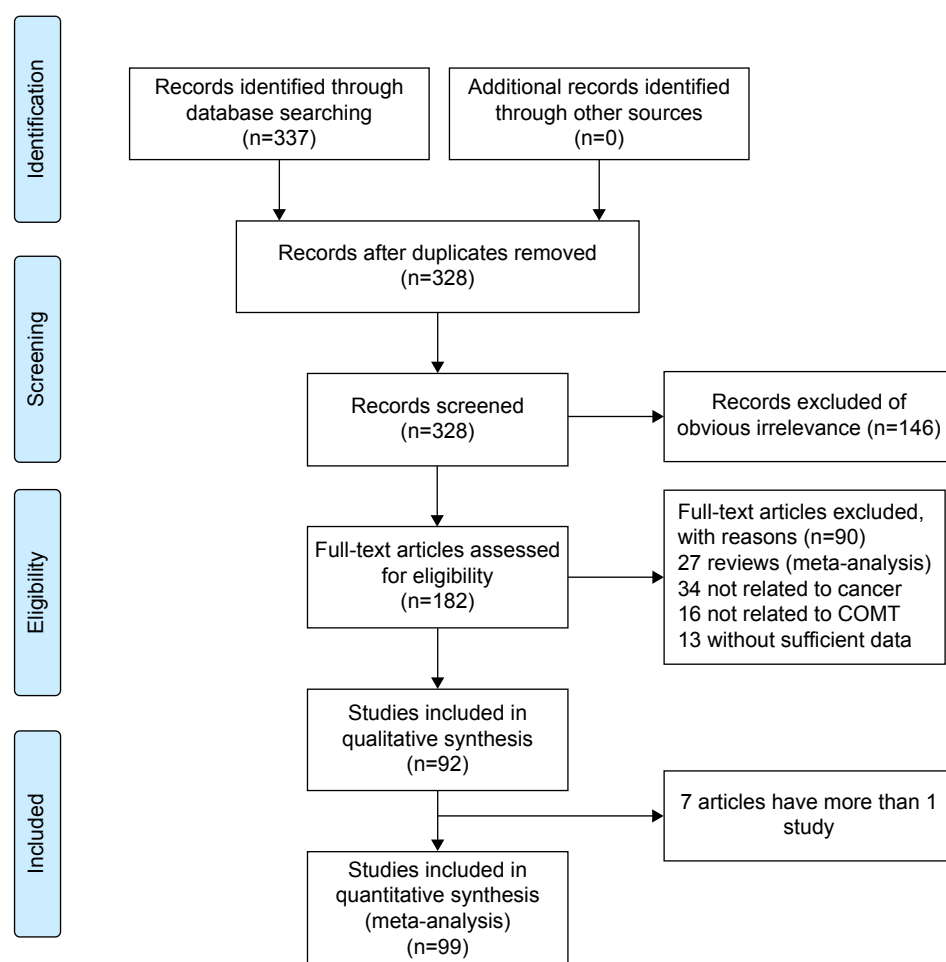


Figure 1 Flow chart of publication selection.

Note: A total of 99 studies were included in this meta-analysis and systematically reviewed after a comprehensive study selection.

Abbreviation: COMT, catechol-O-methyltransferase.

Val158Met was significantly associated with an increased risk of bladder cancer in recessive model (OR =1.30, 95% CI [1.02, 1.66]), esophageal cell cancer in homozygote model (OR =1.77, 95% CI [1.07, 2.93]), heterozygote model (OR =1.40, 95% CI [1.01, 1.92]), and dominant model (OR =1.46, 95% CI [1.08, 1.98]). However, studies on renal, endometrial, lung, liver, ovarian, colon, and other cancer types have suggested null association (OR =0.70–1.46; Table 2). These studies were further stratified on the basis of ethnicities, and the results showed that COMT Val158-Met polymorphism may be a risk factor for cancer in Asian populations in the homozygote model (OR =1.25, 95% CI [1.03, 1.51]) and recessive model (OR =1.20, 95% CI [1.01, 1.43]). We failed to detect any association between the COMT Val158Met polymorphism and African, Caucasian, and mixed populations. In addition, homozygote models (OR =3.46, 95% CI [2.07, 5.80]), recessive models (OR =3.32, 95% CI [2.02, 5.44]), and dominant models (OR =1.54, 95% CI [1.12, 2.11]) were significantly associated

with increased cancer risk in the subgroup of AS-PCR genotyping method, but no significant associations were observed when PCR-RFLP, TaqMan, sequencing, MALDI-TOF, and other genotyping method were used. No significant associations were detected when the studies were stratified on the basis of the source of control subjects.

Test of heterogeneity and sensitivity

Heterogeneity among studies was observed in the overall comparisons as well as in the subgroup analyses. The source of heterogeneity was investigated by cancer ethnicity (European, Asian, African, and mixed; $P=0.483$), cancer types (bladder, breast, renal, endometrial, lung, liver, ovarian, colon, and other cancer types; $P=0.684$), control source (population based, hospital based, and family based; $P=0.659$), and genotyping method (AS-PCR, PCR-RFLP, TaqMan, sequencing, MALDI-TOF, and other genotyping method; $P=0.647$) using meta-regression, but no covariables were found to contribute to the heterogeneity.

Table 1 Characteristics of studies included in the meta-analysis

Authors	Year	Country	Ethnicity mixed	Cancer type	Control source	Genotype method	Genotype (cases)			Genotype (controls)			HWE	NOS score
							AA	AG	GG	AA	AG	GG		
Lavigne et al ¹¹	1997	USA	Caucasian	Breast	HB	PCR-RFLP	35	57	21	31	56	27	0.862	6
Millikan et al ¹²	1998	USA	African	Breast	PB	PCR-RFLP	29	106	130	34	118	111	0.838	8
Millikan et al ¹²	1998	USA	Caucasian	Breast	PB	PCR-RFLP	102	184	103	105	188	86	0.916	8
Thompson et al ¹³	1998	USA	Caucasian	Breast	PB	PCR-RFLP	53	159	69	72	139	78	0.522	7
Huang et al ¹⁴	1999	People's Republic of China	Asian	Breast	HB	PCR-RFLP	13	37	68	4	55	66	0.612	5
Goodman et al ¹⁵	2000	Germany	Caucasian	Ovarian	HB	PCR-RFLP	27	54	27	29	52	25	0.905	7
Goodman et al ¹⁶	2001	USA	Mixed	Ovarian	PB	PCR-RFLP	16	57	52	19	57	68	0.827	8
Goodman et al ¹⁷	2001	USA	Caucasian	Breast	PB	PCR-RFLP	35	57	20	31	55	27	0.788	8
Hamajima et al ¹⁸	2001	Japan	Asian	Breast	HB	PCR-RFLP	18	72	60	23	63	79	0.079	6
Bergman-Jungstrom and Wingren ¹⁹	2001	Sweden	Caucasian	Breast	HB	PCR-RFLP	46	64	16	43	61	13	0.209	5
Mitrunen et al ²⁰	2001	Finland	Caucasian	Breast	PB	PCR-RFLP	128	238	115	143	237	100	0.921	5
Yim et al ²¹	2001	Korea	Asian	Breast	HB	PCR-RFLP	3	79	81	16	46	101	0.004	6
Garner et al ²²	2002	USA	Mixed	Ovarian	PB	PCR-RFLP	48	103	59	54	119	52	0.861	6
Kocbas et al ²³	2002	Turkey	Caucasian	Breast	HB	PCR-RFLP	14	42	28	13	55	35	0.227	7
Comings et al ²⁴	2003	USA	Caucasian	Breast	PB	PCR-RFLP	12	24	31	38	78	29	0.335	6
Rossi et al ²⁵	2003	Italy	Caucasian	Liver	HB	PCR-RFLP	15	56	16	23	51	16	P>0.05	6
Tan et al ²⁶	2003	People's Republic of China	Asian	Breast	HB	PCR-RFLP	26	103	121	13	105	132	0.174	8
Wedrén et al ²⁷	2003	Sweden	Caucasian	Breast	PB	DASH	442	767	281	433	662	245	0.772	6
Wu et al ²⁸	2003	USA	Asian	Breast	PB	TaqMan	48	213	328	51	229	282	0.646	6
Ahsan et al ²⁹	2004	USA	Mixed	Breast	PB	LP	73	156	84	60	144	58	0.108	6
Dunning et al ³⁰	2004	UK	Caucasian	Breast	PB	TaqMan	845	1,360	645	534	926	448	0.232	8
Hefler et al ³¹	2004	Austria	Caucasian	Breast	PB	Sequencing	98	192	101	478	835	385	0.577	8
Hung et al ³²	2004	France	Caucasian	Bladder	HB	PCR-RFLP	43	96	62	43	114	57	P>0.05	7
McGrath et al ³³	2004	USA	Caucasian	Endometrial	HB	PCR-RFLP	55	105	55	172	308	161	0.874	7
Sazci et al ³⁴	2004	Turkey	Caucasian	Breast	PB	PCR-RFLP	28	69	33	16	146	62	0	6
Yin et al ³⁵	2004	People's Republic of China	Asian	Liver	HB	PCR-RFLP	30	21	3	49	31	6	NA	7
Zimarina et al ³⁶	2004	Russia	Caucasian	Endometrial	HB	PCR-RFLP	30	65	29	44	73	23	0.996	6
Cheng et al ³⁷	2005	People's Republic of China	Asian	Breast	HB	NR	35	197	237	58	262	420	0.006	6
Doherty et al ³⁸	2005	USA	Mixed	Endometrial	PB	PCR-RFLP	100	174	97	123	207	90	0.953	6
Huber et al ³⁹	2005	Austria	Caucasian	Colon	PB	PCR-RFLP	0	58 ^a	18	0	519 ^a	203	NA	6
Lin et al ⁴⁰	2005	People's Republic of China	Asian	Breast	PB	PCR-RFLP	5	31	51	18	133	190	0.393	6
Lin et al ⁴¹	2005	People's Republic of China	Asian	Breast	PB	PCR-RFLP	6	35	58	23	138	205	0.972	6
Le Marchand et al ⁴²	2005	USA	Mixed	Breast	PB	PCR-RFLP	196	624	519	206	614	550	0.109	7
Modugno et al ⁴³	2005	USA	Caucasian	Breast	PB	TaqMan	77	124	49	1,104	1,943	903	0.391	8
Sellers et al ⁴⁴	2005	USA	Caucasian	Ovarian	HB	PCR-RFLP	119	224	110	147	269	127	0.903	7
Sellers et al ⁴⁴	2005	USA	African	Ovarian	HB	PCR-RFLP	0	17 ^a	19	0	30 ^a	23	0.059	6

(Continued)

Table 1 (Continued)

Authors	Year	Country	Ethnicity mixed	Cancer type	Control source	Genotype method	Genotype (cases)			Genotype (controls)			HWE	NOS score
							AA	AG	GG	AA	AG	GG		
Skibola et al ⁴⁵	2005	USA	Caucasian	NHL	PB	TaqMan	77	153	75	163	323	193	P>0.05	7
Wen et al ⁴⁶	2005	People's Republic of China	Asian	Breast	PB	PCR-RFLP	83	425	612	93	470	628	0.698	7
Chang et al ⁴⁷	2006	People's Republic of China	Asian	Breast	HB	PCR-RFLP	9	77	103	30	159	132	0.068	7
Gallicchio et al ⁴⁸	2006	USA	Caucasian	Breast	PB	TaqMan	24	41	16	371	608	272	0.44	9
Gaudet et al ⁴⁹	2006	USA	Caucasian	Breast	PB	MALDI-TOF	240	521	287	266	549	277	0.853	8
Gaudet et al ⁴⁹	2006	Poland	Caucasian	Breast	PB	TaqMan	439	993	551	539	1,123	617	0.525	8
Onay et al ⁵⁰	2006	Canada	Caucasian	Breast	PB	TaqMan	94	202	102	96	196	80	0.283	8
Song et al ⁵¹	2006	People's Republic of China	Asian	Breast	NR	PCR-RFLP	3	41	66	11	36	65	0.09	5
Tao et al ⁵²	2006	People's Republic of China	Asian	Endometrial	HB	TaqMan	85	383	563	67	425	534	0.683	6
Akisik and Dalay ⁵³	2007	Turkey	Caucasian	Breast	NR	PCR-RFLP	26	59	29	21	53	34	0.966	6
Fan et al ⁵⁹	2007	People's Republic of China	Asian	Breast	HB	PCR-RFLP	29	75	96	5	44	51	0.25	6
Gemignani et al ⁵⁴	2007	European	Caucasian	Lung	HB	PCR-RFLP	59	144	83	75	146	81	0.569	7
Holt et al ⁵⁵	2007	USA	Caucasian	Ovarian	PB	TaqMan	79	129	72	137	209	104	0.948	8
Holt et al ⁵⁵	2007	USA	African	Ovarian	PB	TaqMan	10	19	4	16	58	52	0.2	8
Hu et al ⁵⁷	2007	People's Republic of China	Asian	Breast	HB	Sequencing	11	36	65	3	41	66	0.252	6
Liu et al ¹¹⁹	2007	People's Republic of China	Asian	Endometrial	HB	PCR-RFLP	5	33	42	3	46	35	0.01	6
Ralph et al ⁵⁶	2007	USA	Caucasian	Breast	HB	TaqMan	405	825	396	900	1,631	755	0.758	7
Szylo et al ⁵⁸	2007	Poland	Caucasian	Endometrial	HB	PCR-RFLP	24	81	46	39	110	48	0.253	6
Takata et al ⁵⁹	2007	USA	Mixed	Breast	PB	PCR-RFLP	89	257	229	47	108	95	0.104	8
Tanaka et al ⁶⁰	2007	Japan	Asian	Renal	PB	Sequencing	10	54	59	11	61	85	NA	8
Zhao et al ⁶¹	2007	People's Republic of China	Asian	Endometrial	HB	PCR-RFLP	16	77	39	8	50	52	0.779	6
Delort et al ⁶²	2008	France	Caucasian	Ovarian	PB	TaqMan	18	22	11	283	480	237	0.916	7
Hirata et al ⁶³	2008	USA	Caucasian	Endometrial	PB	PCR-RFLP	37	81	32	27	90	48	0.277	8
Justenhoven et al ⁶⁴	2008	Germany	Caucasian	Breast	PB	MALDI-TOF	145	298	163	147	305	170	0.654	8
Onay et al ⁶⁵	2008	Canada	Caucasian	Breast	PB	TaqMan	273	642	302	201	353	160	0.832	8
Onay et al ⁶⁵	2008	Finland	Caucasian	Breast	PB	TaqMan	206	361	141	168	267	114	0.676	7
Yuan et al ⁶⁶	2008	People's Republic of China	Asian	Liver	HB	PCR-RFLP	18	144	258	32	157	286	P>0.05	6
Zhu ¹⁰⁰	2008	People's Republic of China	Asian	Esophageal	HB	PCR-RFLP	16	51	23	10	37	30	P>0.05	5
Zienoldiny et al ⁶⁷	2008	Norway	Caucasian	Lung	PB	Sequencing	32	62	163	8	60	202	0.182	8
Cote et al ⁶⁸	2009	USA	African	Lung	PB	TaqMan	10	46	56	14	47	59	0.332	8
Cote et al ⁶⁸	2009	USA	Caucasian	Lung	PB	PCR-RFLP	102	205	78	114	197	92	0.696	8
Fontana et al ⁶⁹	2009	France	Caucasian	Bladder	HB	TaqMan	14	28	9	10	24	11	NA	6
He et al ⁷¹	2009	USA	Caucasian	Breast	HB	TaqMan	334	607	271	446	837	400	0.85	7

Reding et al ⁷³	2009	USA	Caucasian	Breast	PB	TaqMan	240	427	224	236	431	211	0.606	8
Sangrajrang et al ⁷⁴	2009	Thailand	Asian	Breast	HB	TaqMan	42	233	290	30	190	266	0.61	7
Shrubsole et al ⁷⁵	2009	People's Republic of China	Asian	Breast	PB	PCR-RFLP	0	497 ^a	596	0	554 ^a	615	NA	7
Yadav et al ⁷⁶	2009	India	Asian	Breast	HB	PCR-RFLP	28	82	44	29	85	52	0.57	7
Zhou ⁸⁸	2009	People's Republic of China	Asian	Colon	PB	SNPlex	23	121	208	38	262	327	P>0.05	7
Delort et al ⁷⁷	2010	France	Caucasian	Breast	PB	TaqMan	254	455	201	283	480	237	0.23	8
Ferlin et al ⁷⁸	2010	Italy	Caucasian	TGCT	HB	PCR-RFLP	0	200	34	2	182	34	P>0.05	7
MARIE-GENICA Consortium on Genetic Susceptibility for Menopausal Hormone Therapy Related Breast Cancer Risk ⁷⁰	2010	Germany	Caucasian	Breast	PB	MALDI-TOF	844	1,569	731	1,569	2,669	1,243	0.094	8
Jakubowska et al ⁷²	2010	Poland	Caucasian	Breast	HB	PCR-RFLP	84	164	71	54	168	68	0.01	8
Li et al ⁹⁷	2010	People's Republic of China	Asian	Endometrial	HB	PCR-RFLP	6	26	90	8	35	71	0.22	5
Martinez et al ¹⁰¹	2013	Mexico	Caucasian	Breast	HB	PCR-RFLP	32	66	52	23	59	68	0.085	7
Moreno-Galvan et al ⁷⁹	2010	Mexico	Caucasian	Breast	HB	PCR-RFLP	12	42	37	14	42	38	0.669	6
Peterson et al ⁸⁰	2010	USA	Caucasian	Breast	PB	TaqMan	420	794	370	403	665	348	0.026	8
Syamala et al ⁸¹	2010	India	Asian	Breast	HB	PCR-RFLP	41	104	74	65	164	138	0.183	6
Syamala et al ⁸¹	2010	India	Asian	Breast	FB	PCR-RFLP	28	64	48	65	164	138	0.183	6
Wang et al ⁸²	2010	People's Republic of China	Asian	Breast	PB	AS-PCR	34	62	80	14	66	96	0.58	7
Xu et al ⁹⁶	2010	People's Republic of China	Asian	Breast	PB	AS-PCR	38	42	60	10	44	68	0.45	7
Cerne et al ⁸³	2011	Slovenia	Caucasian	Breast	HB	TaqMan	144	263	123	67	136	67	0.903	7
Cribb et al ⁸⁴	2011	Canada	Caucasian	Breast	HB	PCR-RFLP	51	108	48	155	326	140	0.208	7
Huang et al ⁸⁵	2011	People's Republic of China	Asian	Esophageal	HB	PCR-RFLP	25	95	90	30	146	180	NA	6
Lajin et al ⁸⁶	2013	Syria	Mixed	Breast	PB	PCR-RFLP	31	70	34	30	54	28	0.887	7
Naushad et al ⁸⁷	2011	India	Asian	Breast	HB	PCR-RFLP	66	154	122	26	107	120	0.201	6
dos Santos et al ⁸⁸	2011	Brazil	Mixed	Breast	PB	PCR-RFLP	0	41 ^a	21	0	26 ^a	36	–	7
Wang et al ⁸⁹	2011	People's Republic of China	Asian	Breast	PB	Sequencing	68	145	187	36	156	208	0.389	7
Heck et al ⁹⁰	2012	USA	Mixed	Renal	HB	Sequencing	0	632 ^a	242	0	1,496 ^a	557	0.36	8
Lim et al ⁹¹	2012	Singapore	Asian	Lung	HB	PCR-RFLP	39	220	284	63	353	549	0.539	7
Wolpert et al ⁹²	2012	Egypt	Mixed	Bladder	PB	TaqMan	160	245	110	95	180	114	P>0.05	8
Zhang et al ⁹³	2013	People's Republic of China	Asian	Lung	HB	Sequencing	11	69	120	19	78	103	0.454	8
Ghisari et al ⁹⁴	2014	Denmark	Caucasian	Breast	PB	TaqMan	13	11	7	41	53	19	P>0.05	6
Son et al ⁹⁵	2015	Korea	Asian	Breast	HB	Assay	0	423 ^a	427	0	212 ^a	178	0.008	7

Note: ^aNumber of patients with the AA + GA genotype in the case and control groups.

Abbreviations: HWE, Hardy-Weinberg equilibrium; NOS, Newcastle-Ottawa Scale; HB, hospital based; PCR-RFLP, polymerase chain reaction-restriction fragment length polymorphism; PB, population based; DASH, dynamic allele-specific hybridization; FB, family based; NA, not available; MALDI-TOF, matrix-assisted laser desorption/ionization time of flight mass spectrometry; NHL, non-Hodgkin lymphoma; TGCT, testicular germ cell tumor; AS-PCR, allele-specific PCR; LP, Luorescence polarization; NR, not reported.

Table 2 Meta-analysis of the association between COMT Val158Met and cancer risk

Variables	No of studies	Homozygote model		Heterozygote model		Recessive model		Dominant model	
		OR (95% CI)	I ² %	OR (95% CI)	I ² %	OR (95% CI)	I ² %	OR (95% CI)	I ² %
Total	99	1.05 (0.98, 1.13)	56	1.01 (0.97, 1.05)	29	1.03 (0.97, 1.09)	51	1.02 (0.97, 1.06)	44
Cancer type									
Bladder	3	1.38 (0.86, 2.21)	45	1.12 (0.71, 1.77)	57	1.30 (1.02, 1.66)	0	1.20 (0.74, 1.94)	65
Renal	2	1.31 (0.52, 3.28)	–	1.28 (0.78, 2.09)	–	1.18 (0.48, 2.86)	–	1.02 (0.83, 1.25)	12
Breast	62	1.04 (0.96, 1.13)	58	1.01 (0.96, 1.05)	21	1.03 (0.96, 1.10)	57	1.01 (0.96, 1.06)	40
Endometrial	9	0.99 (0.73, 1.35)	55	0.90 (0.73, 1.11)	52	1.03 (0.84, 1.26)	29	0.91 (0.73, 1.13)	61
Lung	6	1.09 (0.68, 1.75)	76	1.11 (0.96, 1.28)	2	1.04 (0.67, 1.57)	74	1.09 (0.87, 1.36)	60
Liver	3	0.68 (0.42, 1.09)	0	1.03 (0.80, 1.34)	0	0.70 (0.48, 1.03)	0	0.96 (0.75, 1.23)	0
Ovarian	8	1.05 (0.75, 1.47)	52	1.01 (0.80, 1.28)	33	1.02 (0.84, 1.24)	20	1.00 (0.79, 1.27)	43
Colon	2	0.95 (0.55, 1.64)	–	0.73 (0.55, 0.96)	–	1.08 (0.64, 1.85)	–	0.92 (0.56, 1.50)	63
Esophageal	2	1.77 (1.07, 2.93)	0	1.40 (1.01, 1.92)	0	1.46 (0.92, 2.34)	0	1.46 (1.08, 1.98)	0
Other	2	0.96 (0.29, 3.16)	24	1.18 (0.90, 1.56)	0	0.87 (0.29, 2.62)	21	1.18 (0.91, 1.54)	0
Ethnicities									
African	4	1.46 (0.43, 4.99)	83	1.23 (0.61, 2.49)	75	1.17 (0.53, 2.56)	69	1.09 (0.60, 1.98)	73
Caucasian	50	0.98 (0.91, 1.05)	43	1.00 (0.96, 1.05)	88	0.97 (0.92, 1.03)	38	0.99 (0.95, 1.04)	16
Asian	35	1.25 (1.03, 1.51)	62	1.04 (0.94, 1.14)	53	1.20 (1.01, 1.43)	60	1.06 (0.97, 1.15)	59
Mixed	10	0.96 (0.78, 1.20)	49	1.00 (0.86, 1.17)	38	0.99 (0.87, 1.13)	5	1.03 (0.88, 1.20)	58
Controls source									
PB	50	1.03 (0.94, 1.13)	63	0.99 (0.94, 1.04)	24	1.06 (0.95, 1.17)	58	1.01 (0.95, 1.07)	49
HB	45	1.09 (0.96, 1.24)	48	1.04 (0.96, 1.12)	36	1.02 (0.94, 1.09)	43	1.04 (0.96, 1.11)	41
Other	4	0.95 (0.59, 1.54)	48	1.00 (0.78, 1.27)	4	1.00 (0.69, 1.46)	38	0.99 (0.78, 1.26)	7
Genotyping method									
PCR-RFLP	58	1.02 (0.91, 1.15)	49	1.01 (0.94, 1.09)	36	1.01 (0.92, 1.11)	42	1.02 (0.95, 1.09)	44
TaqMan	24	1.03 (0.94, 1.13)	46	1.02 (0.96, 1.08)	15	1.00 (0.93, 1.07)	35	1.02 (0.95, 1.08)	34
Sequencing	6	1.55 (0.79, 3.03)	85	0.98 (0.84, 1.14)	1	1.55 (0.84, 2.86)	84	1.09 (0.84, 1.41)	67
MALDI-TOF	3	0.92 (0.83, 1.02)	0	0.98 (0.90, 1.08)	0	0.93 (0.85, 1.01)	0	0.96 (0.88, 1.05)	0
AS-PCR	2	3.46 (2.07, 5.80)	0	1.11 (0.78, 1.57)	0	3.32 (2.02, 5.44)	0	1.54 (1.12, 2.11)	0
Other	6	0.91 (0.77, 1.08)	0	0.94 (0.72, 1.24)	76	0.92 (0.80, 1.05)	0	0.93 (0.81, 1.08)	57

Notes: The bold values indicate that the results are statistically significant.

Abbreviations: COMT, catechol-O-methyltransferase; OR, odds ratio; CI, confidence interval; PB, population based; HB, hospital based; PCR-RFLP, polymerase chain reaction-restriction fragment length polymorphism; MALDI-TOF, matrix-assisted laser desorption/ionization time of flight mass spectrometry; AS-PCR, allele-specific PCR; I², variation in OR attributable to heterogeneity.

Sensitivity analysis was conducted to verify the effect of each study on the overall OR by repeating the meta-analysis, but one study was omitted each time. When sensitivity analyses were performed without HWE violating studies, all the results were not materially altered. The results showed that the pooled

ORs of these three polymorphisms were not materially altered by the contribution of any individual study, thus confirming that the results of this meta-analysis were statistically robust.

Publication bias

Begg's funnel plot and Egger's test were performed to evaluate the publication bias of the studies. The shape of the funnel plots showed that the dots were almost symmetrically distributed and were predominantly in 95% confidence limits (dominant model, Figure 2). The results of Egger's test statistically confirmed the absence of publication bias in the dominant model ($t=1.68$, $P=0.096$).

Discussion

In the past several years, interest in the genetic susceptibility to cancers has drawn increased attention to the studies on polymorphisms of genes involved in tumor genesis. Genome-wide association study, also known as whole genome association study, is widely used in the study of genetic epidemiology. At present, >1,369 susceptibility loci associated with cancer risk have been identified by

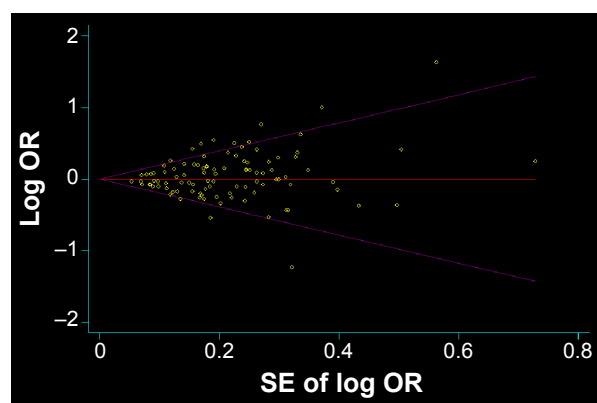


Figure 2 Begg's funnel plot of the meta-analysis of cancer risk and COMT Val158Met polymorphism (AA + AG vs GG).

Note: Begg's funnel plot with pseudo 95% confidence limits.

Abbreviations: COMT, catechol-O-methyltransferase; OR, odds ratio; SE, standard error.

genome-wide association study, but none of these studies had reported significant associations between cancer susceptibility and COMT Val158Met polymorphisms. We searched the manufacturers' websites (<http://www.affymetrix.com/index.affx> and <http://www.illumina.com>)¹²⁰ and the relevant PubMed databases (Probe, Database of Genotypes and Phenotypes, and Gene Expression Omnibus DataSets) and found that the COMT Val158Met polymorphism was not included in the platforms commonly used in genome-wide association studies. But since the identification of COMT Val158Met polymorphism, the role of COMT Val158Met in cancers risk has been reported in an increasing number of studies, but the results remained controversial. Some recent meta-analyses studies reported such an association only for single cancer or specific populations. Importantly, several published studies were not included in the previous meta-analysis, and additional original studies with larger sample sizes have been published since then. Hence, the association between the COMT Val158Met polymorphism and the risk of cancer remains unknown. Therefore, meta-analysis can provide a quantitative summary of the available data supporting the association between COMT Val158Met and cancer risk. Compared with some previous meta-analyses, strengths of our meta-analysis include the large sample size and high statistical power of the analysis based on substantial number of cases and controls from differential studies, which minimized selection bias and led to relatively stable risk estimation.

In the current meta-analysis, 99 case-control studies with 43,085 cancer cases and 57,882 control subjects were considered. The results indicated no significant association between COMT Val158Met polymorphism and overall cancer risk in any genetic comparison model tested. In further subgroup analysis by cancer type, COMT Val158Met was significantly associated with an increased risk of bladder cancer and esophageal cancer in some specific genetic models. However, studies on renal, endometrial, lung, liver, ovarian, colon cancers, and other cancer types have suggested null associations. In line with most previous meta-analyses for single cancer, Zhang et al,¹¹¹ Du et al¹⁰² and Mao et al¹²¹ have reported that the COMT Val158Met polymorphism may not contribute to the risk of prostate cancer, ovarian cancer, or breast cancer in any of the assessed genetic model. In the subgroup analysis by ethnicity, no significant associations were found in African, Caucasian, and mixed populations. However, the significant association between the COMT Val158Met polymorphism and cancer risk remains to be determined in Asians. The discrepancy in ethnicity could be attributed to the evident difference in the minor allele frequency of Val158Met

polymorphism in Asians and Caucasians in our meta-analysis. This genetic polymorphism variance with ethnicity was consistent with those described in a previous study.⁸ In addition, stratified analyses by genotyping techniques indicated that studies involving AS-PCR likely acquired significant results in the overall comparison. However, this result should be carefully interpreted because of a relatively small sample size. Moreover, this result should be confirmed by further analysis of additional published studies.

Several limitations should be acknowledged in this meta-analysis. First, only studies in English or Chinese were included in this meta-analysis, which might cause publication bias. Second, the pooled results were based on unadjusted estimates because not all studies had provided adjusted ORs. Even in cases where adjusted ORs were found, they were not adjusted by the same confounders. Hence, a precise analysis should be performed. Third, several factors such as gene-gene or gene-environment interaction may influence gene-disease factor, and the lack of individual data from the included studies limited further evaluation of other potential interactions, as in other genes and environment factors. Finally, cancer is a multifactorial disease resulting from complex interactions among many genetic and environmental factors. Therefore, a single gene or single environmental factor is unlikely to explain cancer susceptibility.

Conclusion

In conclusion, the present meta-analysis suggested that COMT Val158Met polymorphism might not be a risk factor for overall cancer risk, but it might be involved in cancer development at least in some ethnic groups (Asian) or some specific cancer types (bladder and esophageal cancer). Further large-scale and well-designed studies regarding different ethnicities are required to confirm the results of our meta-analysis.

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Disclosure

The authors report no conflicts of interest in this work.

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