

Association Between Triglyceride-Glucose Index and Risk of Metabolic Dysfunction-Associated Fatty Liver Disease: A Cohort Study

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Objective: Insulin resistance (IR) is a major factor involved in the pathogenesis of metabolic dysfunction-associated fatty liver disease (MAFLD). Triglyceride-glucose (TyG) index, an easily detected surrogate marker of IR, has not been explored sufficiently on its relationship with incident MAFLD risk. This study sought to investigate the association of baseline TyG index with the risk of MAFLD in a Chinese cohort.

Methods: This health check-up cohort was constructed with eligible 2056 Chinese from a community. The TyG index was calculated as $\ln(\text{fasting triglyceride [mg/dL]} \times \text{fasting glucose [mg/dL]}/2)$. Cox proportion hazard models were used to evaluate the longitudinal association between baseline TyG index and the risk of MAFLD.

Results: During an average follow-up of 2.5 ± 0.5 years, about 12.8% of the subjects developed MAFLD, and the incidence of MAFLD trended to increase with the quartile TyG index ($P_{\text{trend}} < 0.05$). After adjusting for all confounders, TyG index was independently correlated with the risk of incident MAFLD (HR = 1.784, 95% CI = 1.383–2.302, $P < 0.001$), and the risk of MAFLD in the highest quartile of TyG index was two times higher than that in the lowest quartile (95% CI = 1.377–2.992, $P = 0.001$). The restricted cubic spline analysis showed that the relationship between TyG index and the risk of MAFLD was linear in males (P for total < 0.001 ; P for non-linearity = 0.746), but nonlinear in females (P for non-linearity = 0.040).

Conclusion: A high baseline TyG index was independently associated with a high risk of incident MAFLD, and we might develop the strategy of MAFLD prevention based on the TyG index.

Keywords: triglyceride-glucose index, metabolic dysfunction-associated fatty liver disease, insulin resistance, risk, cohort study

Introduction

Metabolic dysfunction-associated fatty liver disease (MAFLD), previously named as nonalcoholic fatty liver disease (NAFLD), is common worldwide,¹ and progresses across simple steatosis, nonalcoholic steatohepatitis (NASH) and cirrhosis.² The reported prevalence of MAFLD is approximately 25% in the global adult population,³ and has surged to 29.2% in China's mainland.⁴ Moreover, the total population of MAFLD patients will rise to about 314 million in China in 2030, manifesting the biggest increase in the world.⁵ MAFLD is a major cause of cirrhosis, hepatocellular carcinoma and liver transplantation,⁶ as well as a risk factor for cardiovascular disease (CVD), colorectal tumors and chronic kidney disease.⁷ Given these, a simple, accurate, and effective tool is urgently needed for early identification of MAFLD.

The pathophysiology of MAFLD remains unclear. Associative factors include insulin resistance (IR), type 2 diabetes mellitus (T2DM), CVD, dyslipidemia and genetic variation.² Among them, IR plays a pathogenic role in MAFLD, T2DM and CVD.^{8,9} Therefore, the risk of these diseases may be predicted by an index of IR. The gold standard for IR

measurement is the hyperinsulino-euglycemic clamp, but its clinical application is limited due to its cost, time and invasive.¹⁰ Homeostasis model assessment of insulin resistance (HOMA-IR) is another well-known IR estimation method. However, the study has shown that HOMA-IR varies greatly according to the type of insulin assay and the normal range of fasting plasma insulin levels.¹¹ Insulin levels are usually measured for diabetic patients and are not suitable for the general population. Thus, a variety of IR replacement markers have emerged in recent years, among which the triglyceride-glucose (TyG) index has been recommended as a simple IR replacement marker.^{12,13} It has revealed a positive correlation between the TyG index and the occurrence of CVD as well as T2DM.^{9,14–16} Meanwhile, studies have demonstrated that the TyG index changes with the development of MAFLD.^{11,17,18} However, most of these studies are cross-sectional. Therefore, we designed this cohort study to investigate the longitudinal association of baseline TyG index with the risk of incident MAFLD in the Chinese.

Methods

Study Population

The cohort in this study was established with a check-up population (aged ≥ 18 years) in a community in Nanjing (Jiangsu, China). Prior to recruitment, participants were examined for hepatic steatosis by abdominal ultrasound (Logiq E9 ultrasound system, General Electric (GE) Healthcare, Milwaukee, WI, USA) and subsequently diagnosed for the presence or absence of MAFLD according to the MAFLD diagnostic procedure.¹⁹ Participants without MAFLD were recruited as study subjects, and baseline data were collected by questionnaire and laboratory assessment. Recruitment and baseline surveys were completed between August 2017 and September 2018 ($n = 3062$). The follow-up interval was set at least 2 years. Participants who made an active appointment or passively invited to complete a check-up from March 2020 to December 2021 were considered as successful follow-up. Subjects with missing data on baseline fasting triglycerides (TG) and fasting blood glucose (FPG) ($n = 826$) as well as subjects lost to follow-up ($n = 180$) were excluded. A total of 2056 eligible participants were included for analysis, including 1298 (63.1%) males and 758 (36.9%) females. The age of all subjects ranged from 19 to 90 years, with an average age of 37.42 ± 10.12 . The current study protocol was in accordance with the Declaration of Helsinki, and was approved by the Institutional Ethics Review Committee of Nanjing Medical University (Nanjing, China). Written informed consent was obtained from all participants.

In addition, the primary outcome of this study was set as newly diagnosed MAFLD. According to the requirements of the Cox proportional hazards regression analysis, one independent variable should correspond to at least 15 non-truncated events. A total of 12 independent variables, with at least 180 newly diagnosed MAFLD were included in this study.

Data Collection and Definition

Demographic information (eg, age, sex) and anthropometric data (eg, height, weight, and blood pressure) were collected by a self-designed questionnaire and an electronic medical record system. In measuring height, the participant was required to stand straight on the ground without shoes, with their hips and heels against the wall. In measuring weight, the participant was required to remove shoes and wear light clothing. Body mass index (BMI) was calculated as the weight (kg)/height (m)². Blood pressure was measured in the seated patient using a mercury sphygmomanometer with a 5-min interval, and the average of two readings was defined as systolic blood pressure (SBP) and diastolic blood pressure (DBP). Hypertension was defined as SBP ≥ 140 mmHg and (or) DBP ≥ 90 mmHg.²⁰

After an 8-hours fasting, the blood was sampled to determine biochemical parameters, including γ -glutamyl transpeptidase (GGT), alanine aminotransferase (ALT), aspartate transaminase (AST), low-density lipoprotein cholesterol (LDL-C), high-density lipoprotein cholesterol (HDL-C), total cholesterol (TC), triglyceride (TG), fasting plasma glucose (FPG) and serum uric acid (SUA). All the laboratory parameters were analyzed by an automatic biochemical analyzer (Mindray BC-860; Mindray, Shenzhen, China). In addition, the TyG index was calculated as $\ln(\text{fasting triglyceride [mg/dL]} \times \text{fasting plasma glucose [mg/dL]}/2)$.²¹ Plasma HDL-C < 1.0 mmol/L for man (< 1.3 mmol/L for woman) was defined as low HDL-C, TG ≥ 1.7 mmol/L as hypertriglyceridemia, FPG ≥ 5.6 mmol/L as hyperglycemia, and

SUA >420 $\mu\text{mol/L}$ for man (>360 $\mu\text{mol/L}$ for woman) as hyperuricemia.^{19,22} Other plasma GGT, ALT, AST, LDL-C and TC were defined according to the classification methods or guidelines of routine clinical indicators.²³

Statistical Analysis

Continuous variables in normal and skewed distributions were described as mean \pm standard deviation (SD) and median with interquartile range (IQR), respectively. Categorical variables were described as frequency and percentage. Missing baseline covariates were imputed by the expectation maximization algorithm. The TyG index was assessed by both continuous variables and quartiles. To investigate the linearity of baseline characteristics across quartiles of TyG index, χ^2 test was used for categorical variables, and one-way analysis of variance (ANOVA) and Kruskal–Wallis *H*-tests for continuous variables with normal and skewed distributions, respectively. Cox proportional hazard regression was used to estimate the hazard ratio (HR) and 95% confidence interval (95% CI) for quantifying the association of TyG index with the risk of incident MAFLD. Subgroup analysis and multiplicative interaction test were performed to verify the robustness of association. In addition, Kaplan–Meier analysis was used for a graphical presentation of the time to the development of MAFLD, and the Log rank test was used to assess differences among groups. Bonferroni correction was used to control the family-wise error rate in all multiple comparisons, and the *P*-value threshold for significance after Bonferroni correction was 0.05/6=0.008. Besides, restricted cubic spline (RCS) analysis was carried out to explore the non-linearity relationship between TyG index and MAFLD risk. Three knots were generated at the 10th, 50th and 90th percentiles, and the 50th percentile of TyG index was used for reference. The R package “SurvivalROC” was used to plot time-dependent receiver operating characteristic curves (ROC). All statistical analyses were processed using SPSS (version 26.0, SPSS Inc., Chicago, IL, USA), MedCalc (Version 20.1.0, Ostend, Belgium) and R software (version 4.0.2). A two-tailed *P*<0.05 was regarded as statistically significant.

Results

Baseline Characteristics of Participants

The cohort was divided into four groups based on the quartiles of baseline TyG index (Q1: <7.93, Q2: \geq 7.93 to <8.23, Q3: \geq 8.23 to <8.58, Q4: \geq 8.58), and the demographic and clinical characteristics at baseline are summarized in [Table 1](#). Compared with the lowest quartile, a higher quartile of TyG index was more enriched in male and older subjects (all *P* for trend <0.001). Similarly, the BMI, SBP, DBP, GGT, ALT, AST, LDL-C, TC, TG, FPG and SUA in the higher TyG index quartile were significantly higher than those in the lower quartile (all *P* for trend <0.001), while the HDL-C in the higher TyG index quartile was significantly lower than that in the lower quartile (*P* for trend <0.001).

Incidences of MAFLD in Four Quartiles

As shown in [Figure 1A](#), during an average follow-up of 2.5 \pm 0.5 years, 264 (12.8%, 264/2056) incident MAFLD cases were observed, and the cumulative incidence of MAFLD increased with the TyG index quartile ($\chi^2 = 38.33$, *P* for trend <0.001). Similarly, multiple comparisons showed that the cumulative incidence of MAFLD in a higher TyG index quartile (Q4 and Q3) was higher than that in a lower TyG index quartile (Q2 and Q1), with all *P* for Bonferroni correction <0.008. However, there was no significant difference in the cumulative incidence of MAFLD between two adjacent percentile groups (Q4 vs Q3, Q2 vs Q1; all *P* for Bonferroni correction >0.008).

In addition, stratified by sex ([Figure 1B and C](#)), the incidence of MAFLD increased with the TyG index quartile level in both males ($\chi^2 = 21.21$, *P* for trend <0.001) and females ($\chi^2 = 5.83$, *P* for trend = 0.016). Multiple comparisons showed that the fourth TyG index quartile had a higher incidence of MAFLD than the first and second quartiles in males (all *P* for Bonferroni correction <0.008), and the incidence of MAFLD was higher in the third quartile than in the first quartile in females (*P* for Bonferroni correction <0.008).

Univariate Analysis of Predictive Factors for MAFLD

The univariate analysis of potential predictive factors for MAFLD was performed by the Cox hazard regression, and the results are shown in [Table 2](#). When expressed as categorical variables, male, age \geq 40 years, BMI \geq 23kg/m², hypertension,

Table 1 Baseline Characteristics of Participants According to Quartiles of TyG Index (N = 2056)

Variables	Quartiles of TyG Index				P for Trend
	Q1 (<7.93)	Q2 (≥7.93 to <8.23)	Q3 (≥8.23 to <8.58)	Q4 (≥8.58)	
Sex					<0.001 ^a
Male	255 (51.1)	331 (64.1)	348 (66.0)	364 (70.8)	
Female	244 (48.9)	185 (35.9)	179 (34.0)	150 (29.2)	
Age (years)	35.02±8.64	37.73±9.79	38.55±10.43	38.30±11.04	<0.001 ^b
BMI (kg/m ²)	21.45±1.73	21.88±1.80	22.09±1.82	22.38±1.74	<0.001 ^b
SBP (mmHg)	118.61±11.39	120.45±12.91	122.50±13.46	122.88±13.55	<0.001 ^b
DBP (mmHg)	71.02±7.85	72.55±8.52	73.94±8.56	74.14±9.19	<0.001 ^b
GGT (U/L)	15.55 (11.00, 21.65)	17.10 (13.00, 22.76)	18.24 (13.85, 23.62)	21.00 (15.65, 27.00)	<0.001 ^c
ALT (U/L)	16.00 (12.00, 21.00)	17.00 (13.00, 21.02)	18.00 (14.00, 22.06)	19.77 (15.00, 25.00)	<0.001 ^c
AST (U/L)	17.41 (15.00, 20.00)	18.00 (16.00, 20.00)	18.29 (16.00, 20.05)	19.00 (17.00, 21.01)	<0.001 ^c
LDL-C (mmol/L)	2.34±0.49	2.51±0.54	2.61±0.55	2.63±0.54	<0.001 ^b
HDL-C (mmol/L)	1.49±0.28	1.41±0.24	1.33±0.24	1.28±0.24	<0.001 ^b
TC (mmol/L)	4.19±0.56	4.35±0.60	4.45±0.61	4.56±0.62	<0.001 ^b
TG (mmol/L)	0.60 (0.52, 0.68)	0.85 (0.78, 0.92)	1.13 (1.03, 1.24)	1.71 (1.47, 2.07)	<0.001 ^b
FPG (mmol/L)	4.67±0.41	4.80±0.44	4.89±0.47	5.16±0.99	<0.001 ^b
SUA (μmol/L)	294.18±74.86	306.39±69.55	312.12±74.92	324.96±79.98	<0.001 ^b

Notes: Data were present as mean±SD, median (interquartile range) or n (%). ^aThe variables were tested by χ^2 -test for linear trend; ^bmean±standard deviation, and the variables were tested by one-way analysis of variance for linear trend; ^cMedian (interquartile range), and the variables were tested by Kruskal–Wallis test.

Abbreviations: BMI, body mass index; SBP, systolic blood pressure; DBP, diastolic blood pressure; GGT, γ -glutamyl transpeptidase; ALT, alanine aminotransferase; AST, aspartate transaminase; ALP, alkaline phosphatase; LDL-C, low-density lipoprotein cholesterol; HDL-C, high-density lipoprotein cholesterol; TC, total cholesterol; TG, triglyceride; FPG, fasting plasma glucose; SUA, serum Uric Acid.

GGT >50U/L, ALT >40U/L, LDL-C \geq 3.4mmol/L, low HDL-C, TC \geq 5.2mmol/L, hypertriglyceridemia, hyperglycemia and hyperuricemia were positively related to the risk of MAFLD (all $P<0.05$). When expressed as continuous variables, age, BMI, SBP, DBP, GGT, ALT, AST, LDL-C, TC, TG, FPG and SUA were risk factors of MAFLD (all $P<0.001$), whereas HDL-C was a factor protecting from MAFLD ($P<0.001$).

In Figure 2, the Kaplan–Meier analysis revealed that the cumulative risk of incident MAFLD was markedly different among the TyG index quartiles (Log rank test, $P<0.001$). Further multiple comparisons showed that the higher quartiles of the TyG index had a higher cumulative risk of incident MAFLD than the lower quartiles, except for data between the fourth and third quartiles (Q4 vs Q3: HR = 1.227, 95% CI = 0.869–1.732).

Association Between TyG Index and the Risk of Incident MAFLD

The association between the baseline TyG index and the risk of incident MAFLD is presented in Table 3. After adjusting for sex, age, BMI, hypertension, GGT, ALT, AST, LDL-C, HDL-C, TC and SUA at baseline (model 2), the significant positive associations were observed, and the risk of MAFLD increased by 78.4% (95% CI = 1.383–2.302, $P<0.001$) with one unit increase of TyG index.

Taking the lowest quartile of TyG index as a reference, after adjusting for all covariates (model 2), the subjects in the highest (HR = 2.000, 95% CI = 1.377–2.992, $P=0.001$) and third quartiles (HR = 1.785, 95% CI = 1.192–2.673, $P=0.005$) all had significantly increased risk of incident MAFLD.

In addition, RCS analysis showed that in the crude model, a nonlinear relationship persisted between TyG index and the risk of MAFLD in the total population (P for non-linearity=0.036) (Figure 3A). However, after sex stratification, this nonlinear relationship was only observed in the female stratification (P for non-linearity=0.040) (Figure 3C), whereas only a linear relationship was shown in the male stratification (P for total <0.001; P for non-linearity=0.746) (Figure 3B).

We further performed the subgroup analysis according to sex, age, BMI, hypertension, GGT, ALT, AST, LDL-C, HDL-C, TC and SUA at baseline. As shown in Table 4, the associations between TyG index and incident MAFLD risk were still observed in subgroups of males, females, age \geq 40 years, BMI <23kg/m², BMI \geq 23 kg/m², non-hypertension, GGT \leq 50U/L,

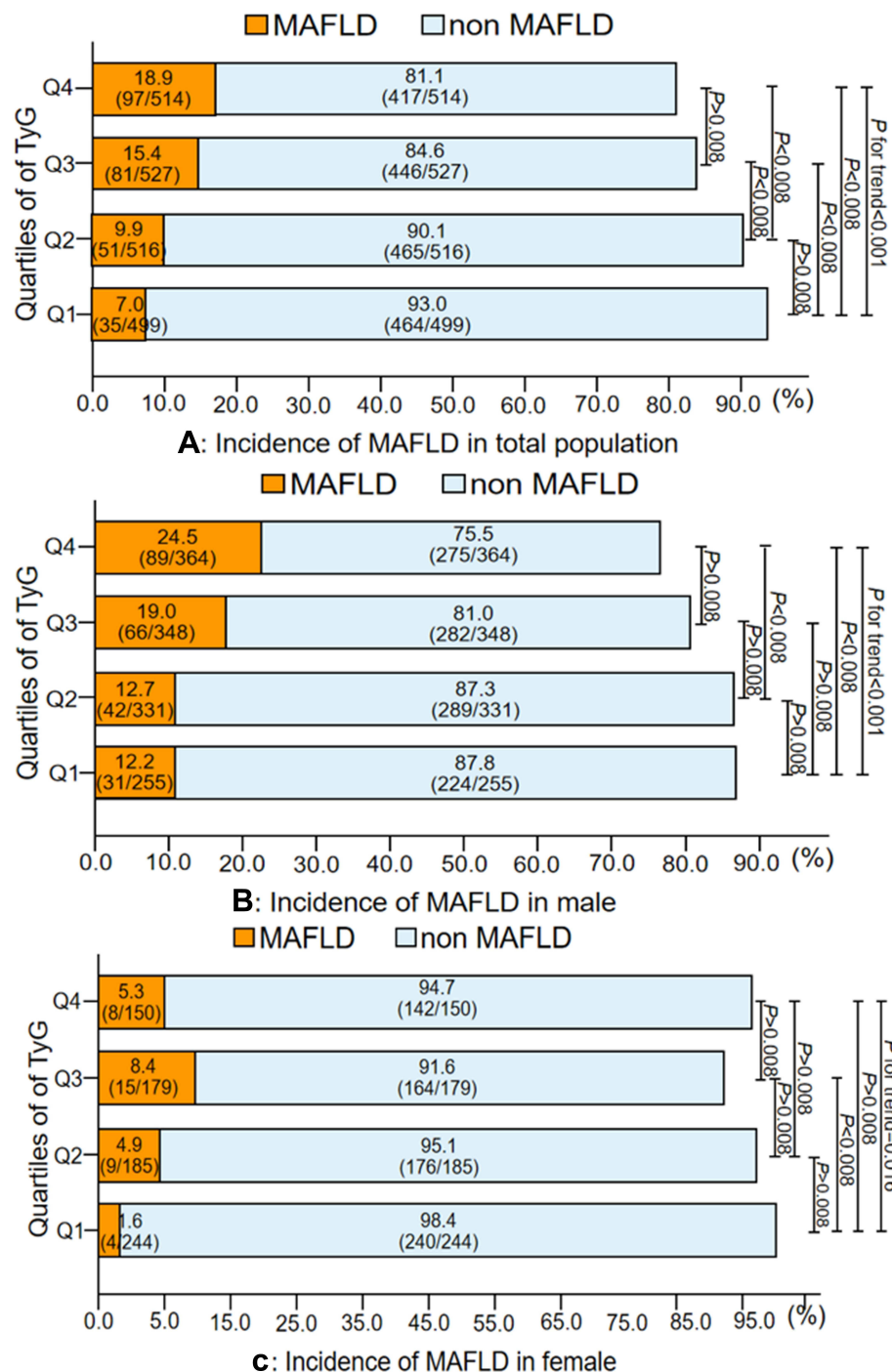


Figure 1 Cumulative incidence of MAFLD according to quartiles of TyG at baseline (average follow-up years=2.5±0.5). **(A)** Incidence of MAFLD in total population. **(B)** Incidence of MAFLD in male. **(C)** Incidence of MAFLD in female. The variables were tested by χ^2 -test for linear trend and multiple comparisons. Bonferroni correction was applied in all multiple comparisons ($P_{\text{Bonferroni}} = 0.05/6 = 0.008$).

ALT ≤ 40 U/L, LDL-C < 3.4 mmol/L, non-low HDL-C, low HDL-C, TC < 5.2 mmol/L, non-hyperuricemia and hyperuricemia (all $P < 0.05$). The associations between TyG index and MAFLD risk did not differ between subgroups (all P for interaction > 0.05), except that a stronger association was found in the subgroup of older age (P for interaction = 0.034).

Table 2 Univariate Analysis of Potential Predictive Factors for MAFLD

Variables	Overall Statistics, n (%)	HR (95% CI)	P
Sex			
Female	758 (36.9)	ref	
Male	1298 (63.1)	3.514 (2.473, 4.995)	<0.001
Age (years)	37.42±10.12	1.037 (1.025, 1.048)	<0.001
<30	529 (25.7)	ref	
30–39	715 (34.8)	0.949 (0.679, 1.325)	0.757
≥40	812 (39.5)	2.248 (1.638, 3.085)	<0.001
BMI (kg/m ²)	21.95±1.80	1.298(1.238, 1.360)	<0.001
<23	1390 (67.6)	ref	
≥23	666 (32.4)	2.650 (2.078, 3.380)	<0.001
SBP (mmHg)	121.14±12.97	1.025 (1.016, 1.034)	<0.001
DBP (mmHg)	72.93±8.63	1.031 (1.017, 1.046)	<0.001
Hypertension			
No	1733 (84.3)	ref	
Yes	323 (15.7)	1.721 (1.285, 2.304)	<0.001
GGT (U/L)	20.94±17.10	1.008 (1.005, 1.011)	<0.001
≤50	1991 (96.8)	ref	
>50	65 (3.2)	1.892 (1.141, 3.138)	0.013
ALT (U/L)	19.93±13.66	1.008 (1.004, 1.011)	<0.001
≤40	1970 (95.8)	ref	
>40	86 (4.2)	1.698 (1.097, 2.630)	0.018
AST (U/L)	19.03±6.20	1.016 (1.006, 1.025)	0.001
≤40	2034 (98.9)	ref	
>40	22 (1.1)	1.000 (0.320, 3.121)	1.000
LDL-C (mmol/L)	2.53±0.54	1.782 (1.466, 2.165)	<0.001
<3.4	1911 (92.9)	ref	
≥3.4	145 (7.1)	1.987 (1.332, 2.964)	0.001
HDL-C (mmol/L)	1.38±0.26	0.132 (0.079, 0.220)	<0.001
Low HDL-C			
No	1899 (92.4)	ref	
Yes	157 (7.6)	1.712 (1.170, 2.504)	0.006
TC (mmol/L)	4.39±0.61	1.526 (1.276, 1.825)	<0.001
<5.2	1860 (90.5)	ref	
≥5.2	196 (9.5)	1.557 (1.063, 2.280)	0.023
TG (mmol/L)	0.15±0.36	1.442 (1.260, 1.605)	<0.001
Hypertriglyceridemia			
No	1752 (85.2)	ref	
Yes	304 (14.8)	1.964 (1.481, 2.604)	<0.001
FPG (mmol/L)	4.88±0.65	1.272 (1.146, 1.413)	<0.001
Hyperglycemia			
No	1910 (92.9)	ref	
Yes	146 (7.1)	1.781 (1.195, 2.653)	0.005
SUA (μmol/L)	309.54±75.66	1.006 (1.005, 1.008)	<0.001
Hyperuricemia			
No	1880 (91.4)	ref	
Yes	176 (8.6)	1.938 (1.397, 2.688)	<0.001

Note: Data were present as mean±SD or n (%).

Abbreviations: BMI, body mass index; SBP, systolic blood pressure; DBP, diastolic blood pressure; GGT, γ -glutamyl transpeptidase; ALT, alanine aminotransferase; AST, aspartate transaminase; ALP, alkaline phosphatase; LDL-C, low-density lipoprotein cholesterol; HDL-C, high-density lipoprotein cholesterol; TC, total cholesterol; TG, triglyceride; FPG, fasting plasma glucose; SUA, serum Uric Acid; TyG, triglyceride-glucose index.

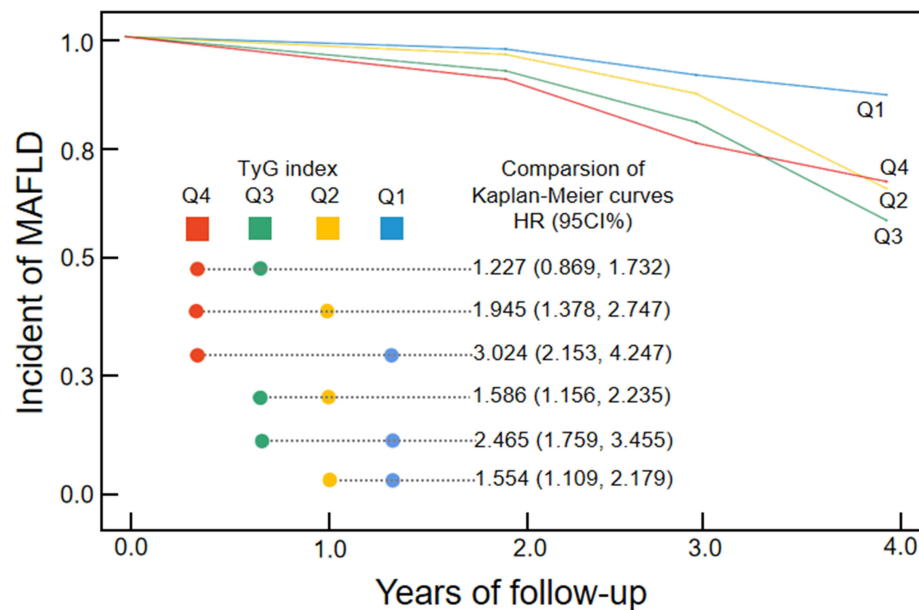


Figure 2 Kaplan-Meier analysis of MAFLD risk according to TyG index quartiles (Log rank test $P < 0.001$). TyG index Q1: <7.93 , Q2: ≥ 7.93 to <8.23 , Q3: ≥ 8.23 to <8.58 , Q4: ≥ 8.58 .

Diagnostic Performance of TyG Index for MAFLD Risk

The time-dependent ROC curves were used to evaluate the ability of TyG index to predict the 2-, 3-, and 4-year risk of MAFLD. As shown in Figure 4, the areas under ROC curve (AUC) was 0.646 for 2-year, 0.640 for 3-year risk, and 0.638 for 3-year risk. In addition, the optimal cut-off values of AUC for predicting 2-year, 3-year, 4-year MAFLD risk were 8.25, 8.11, 7.97, with sensitivities of 70.9%, 78.8%, 92.7%, specificities of 53.2%, 43.9%, 36.9%, and Youden indexes of 0.242, 0.227, 0.296, respectively.

Discussion

IR has been proven as a key factor of MAFLD pathogenesis.^{24–26} Thus, IR-related indexes have been trialed in the studies aiming to improve the early identification and prevention of MAFLD. The TyG index derived from FPG and TG has been surrogated to mark IR in many epidemiological studies.^{10,27} In this prospective cohort study, the subjects without MAFLD at baseline were followed up, and the association between baseline TyG index and MAFLD risk in this cohort was explored. The results showed that 12.8% of the subjects developed MAFLD over an average follow-up of 2.5 ± 0.5 years, and the risk of MAFLD trended to increase with the TyG index quartile.

Table 3 Association of TyG Index with MAFLD Risk in Cox Proportional Hazard Models

Variables	Crude Model ^a		Model 1 ^b		Model 2 ^c	
	HR (95% CI)	P	HR (95% CI)	P	HR (95% CI)	P
TyG index	2.190 (1.755, 2.732)	<0.001	1.864 (1.471, 2.363)	<0.001	1.784 (1.383, 2.302)	<0.001
Q1 (<7.93)	ref		ref		ref	
Q2 (≥ 7.93 to <8.23)	1.557 (1.013, 2.395)	0.044	1.229 (0.798, 1.893)	0.350	1.235 (0.801, 1.904)	0.340
Q3 (≥ 8.23 to <8.58)	2.475 (1.664, 3.681)	<0.001	1.858 (1.245, 2.772)	0.002	1.785 (1.192, 2.673)	0.005
Q4 (≥ 8.58)	3.038 (2.063, 4.474)	<0.001	2.146 (1.450, 3.175)	<0.001	2.000 (1.377, 2.992)	0.001
P for trend		<0.001		<0.001		<0.001

Notes: ^aAdjusted for none; ^bAdjusted for sex, age and BMI at baseline; ^cAdjusted for sex, age, BMI, hypertension, GGT, ALT, AST, LDL-C, low HDL-C, TC and SUA at baseline. All covariates were brought into the model in the form of categorical variables.

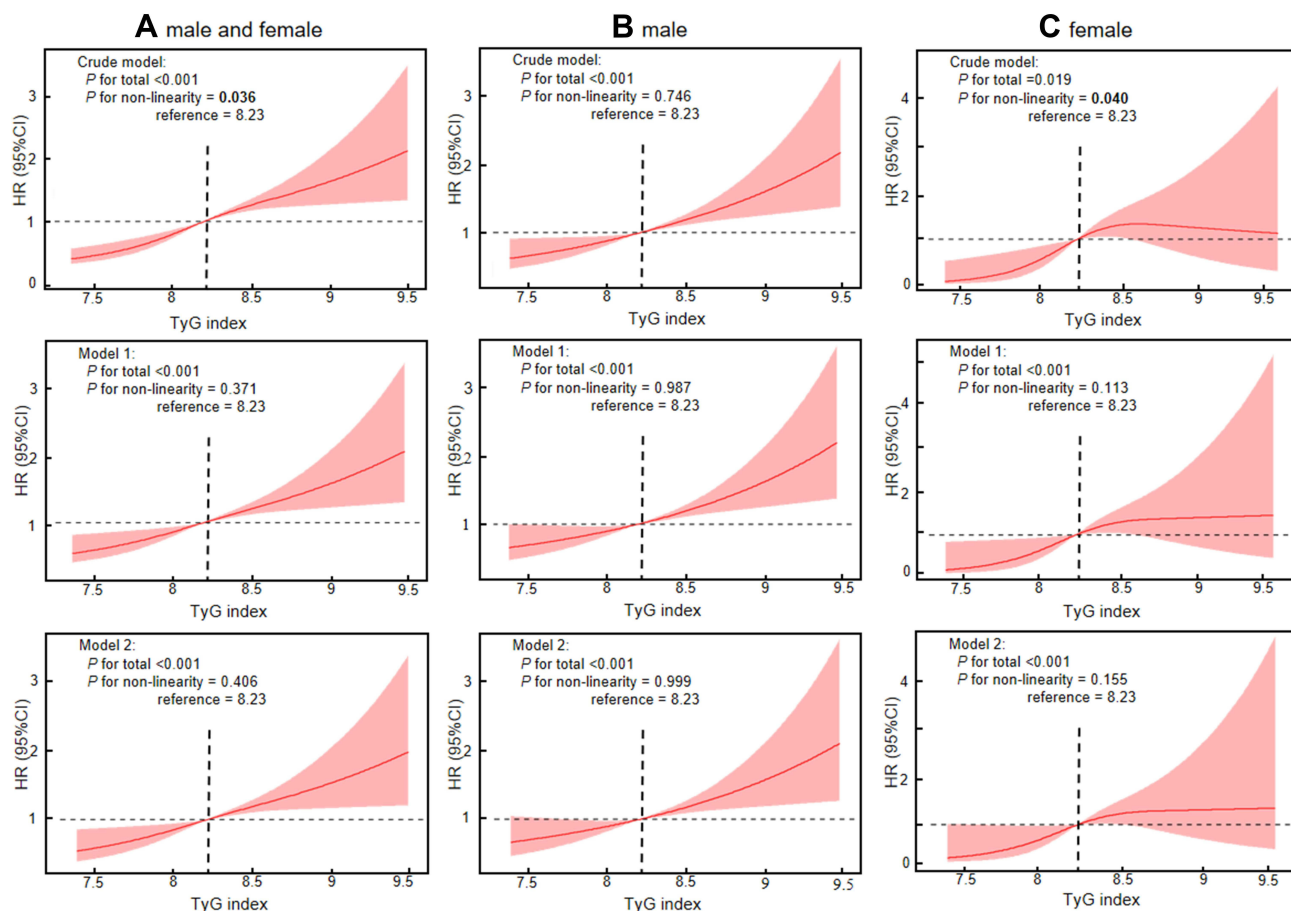


Figure 3 A nonlinear relationship of TyG index with risk of MAFLD. **(A)** male and female (total population). **(B)** male. **(C)** female. Data were fitted using a Cox regression model of the restricted cubic spline with 3 knots at 10th, 50th and 90th percentiles of baseline TyG index. The reference point was the 50th percentile of the TyG index (8.23). The solid red line represented point estimation on the association of TyG index with MAFLD, and the shaded portion represented 95% CI estimation. In crude model, adjusted for none. In model 1, adjusted for sex, age and BMI at baseline. In model 2, adjusted for sex, age, BMI, hypertension, GGT, ALT, AST, LDL-C, low HDL-C, TC and SUA at baseline. All covariates were brought into the model in the form of categorical variables.

Recently, a population-based 20-year follow-up study from Olmsted County, Minnesota showed that the incidence of MAFLD increased 5-fold, from 62/100,000 person-years in 1997 to 329/100,000 person-years in 2014.²⁸ Additionally, in a study that followed 565 subjects from Hong Kong for 3–5 years, the incidence of MAFLD was estimated to be 13.5%.²⁹ In another cohort study, 77,425 Koreans free of MAFLD at baseline were followed up for an average of 4.5 years, and 13.4% developed MAFLD.³⁰ Although the follow-up time was shorter in our study, the incidence of MAFLD was similar to that in Hong Kong and Korea, indicating that the incidence of MAFLD was high.³¹ Additionally, a 9-year follow-up in Chinese employees showed that those with higher baseline TyG levels were more likely to develop MAFLD than those with lower levels, which is consistent with the finding in our study.³² That study also reported that the baseline clinical characteristics, including BMI, SBP, DBP, AST, ALT, HDL-C, LDL-C, TC, TG, FPG and SUA, showed significant differences among quartiles of TyG level, which is also a result in our study. Besides, we further found that except for HDL-C, other indexes, such as BMI, SBP, DBP and metabolic parameters, increased in a higher TyG quartile. The univariate analysis revealed that the risk of incident MAFLD rose with the levels of metabolic parameters and the quartile of TyG index, suggesting that TyG index and MAFLD were closely related to obesity, hypertension, hyperlipidemia and hyperglycemia, which is also in agreement with previous findings.^{15,20,33}

In further analysis, after adjusting for sex, age and metabolic parameters, the risk of incident MAFLD in the highest quartile was twice that in the lowest quartile of TyG index. In a health check-up cohort of Japanese, the hazard rate of progression to MAFLD in the highest tertile of TyG index was approximately twice that in the lowest tertile.³⁴ Another retrospective cohort study in elderly Chinese demonstrated that the risk of incident MAFLD in the highest quartile of

Table 4 Subgroup Analysis of Association Between TyG Index and MAFLD Risk

Variables	Incidence of MAFLD, n (%)	HR (95% CI) ^a	P ^a	P for Interaction ^b
Sex				0.414
Male	228 (86.4)	1.744 (1.319, 2.305)	<0.001	
Female	36 (13.6)	2.120 (1.091, 4.121)	0.027	
Age (years)				0.034
<30	60 (22.7)	1.371 (0.794, 2.368)	0.257	
30–39	81 (30.7)	1.484 (0.925, 2.381)	0.102	
≥40	123 (46.6)	2.258 (1.550, 3.289)	<0.001	
BMI (kg/m ²)				0.452
<23	115 (43.6)	1.826 (1.281, 2.601)	0.001	
≥23	149 (56.4)	1.745 (1.201, 2.535)	0.004	
Hypertension				0.501
No	206 (78.0)	1.779 (1.342, 2.358)	<0.001	
Yes	58 (22.0)	1.876 (0.987, 3.566)	0.055	
GGT (U/L)				0.780
≤50	248 (93.9)	1.830 (1.404, 2.386)	<0.001	
>50	16 (6.1)	1.780 (0.449, 7.061)	0.412	
ALT (U/L)				0.703
≤40	242 (91.7)	1.820 (1.394, 2.376)	<0.001	
>40	22 (11.0)	1.650 (0.579, 4.699)	0.348	
LDL-C (mmol/L)				0.410
<3.4	237 (89.8)	1.801 (1.385, 2.341)	<0.001	
≥3.4	27 (18.6)	1.850 (0.533, 6.424)	0.333	
Low HDL-C				0.565
No	234 (88.6)	1.720 (1.309, 2.261)	<0.001	
Yes	30 (11.4)	2.183 (1.029, 4.633)	0.042	
TC (mmol/L)				0.253
<5.2	234 (88.6)	1.876 (1.435, 2.452)	<0.001	
≥5.2	30 (11.4)	1.298 (0.529, 3.185)	0.569	
Hyperuricemia				0.453
No	221 (83.7)	1.733 (1.310, 2.292)	<0.001	
Yes	43 (16.3)	2.124 (1.041, 4.334)	0.039	

Notes: ^aAdjusted for sex, age, BMI, Hypertension, GGT, ALT, AST, LDL-C, Low HDL-C, TC and SUA at baseline (except the corresponding stratification variable). ^bAdjusted for sex, age, BMI, Hypertension, GGT, ALT, AST, LDL-C, Low HDL-C, TC and SUA at baseline. All covariates were brought into the model in the form of categorical variables.

Abbreviations: BMI, body mass index; GGT, γ -glutamyl transpeptidase; ALT, alanine aminotransferase; AST, aspartate transaminase; LDL-C, low-density lipoprotein cholesterol; HDL-C, high-density lipoprotein cholesterol; TC, total cholesterol.

TyG index was about 1.3 times that in the lowest quartile, and there was a nonlinear relationship between TyG index and incident MAFLD both in males and females.³⁵ In our study, this nonlinear relationship only appeared in females, whereas in males, the relationship between TyG index and MAFLD risk was linear. That is, when the median TyG of 8.23 was used as the reference point for RSC analysis, a smooth ascending curve was observed in males, while in females, the curve was found to rise smoothly on the left of the reference point, and then declined gently and tended to level after a slight rise on the right of the reference point. This finding was mutually supportive with the results of MAFLD incidence of different sexes at different TyG index quartile in this study. In our study, the incidence of MAFLD in males gradually increased with the increase of TyG index quartile, while the high incidence of MAFLD in females was concentrated at the third quartile of TyG index (≥ 8.23 to < 8.58), followed by the fourth, second and first quartile. This might suggest that the risk of MAFLD in males was consistent with the level of TyG index, while the risk of MAFLD in females tended to be stable after the TyG index reached a certain level.

In subgroup analysis, we found that the association was still significant in all subgroups stratified according to sex, BMI, HDL-C and SUA. However, the significant association was observed in subgroups of non-hypertension, GGT ≤ 50 U/L, ALT

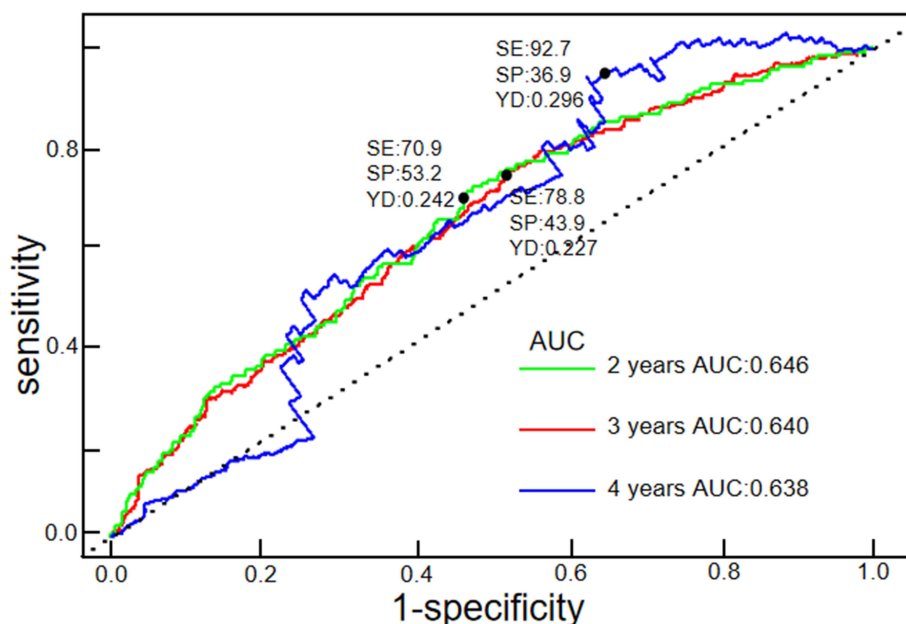


Figure 4 Time-dependent ROC of TyG index in predicting 2-, 3-, and 4-year risk of incident MAFLD.

Abbreviations: ROC receiver operating characteristic curves, SE sensitivity, SP specificity, YD Youden index, AUC area under curve.

≤ 40 U/L, LDL-C < 3.4 mmol/l and TC < 5.2 mmol/l. A possible explanation was that the follow-up time was relatively short, and all the outcomes were early MAFLD. As a progressive chronic liver disease, early MAFLD is asymptomatic in most cases, with routine laboratory indicators generally normal.^{2,36,37} Symptoms or abnormal laboratory parameters usually reflect advanced liver diseases or coexisting conditions.³⁸ Thus, it can be considered that TyG index has a sensitivity in predicting early MAFLD, which was also supported by ROC analysis in our study. Therefore, we suggest that whether presenting symptoms or not, MAFLD and extrahepatic morbidities should be suspected once the TyG level is high. Furthermore, the interaction between TyG index and age on MAFLD were statistically significant. An association between TyG index and MAFLD risk was observed in the subgroup ≥ 40 years. This age-specific association has never been reported. A cohort study revealed an effect of age on the association between TyG index and diabetes risk, and a strong association was found in those aged < 40 years, suggesting that subjects aged < 40 years should be screened for the risk of incident diabetes.¹⁵ Similarly, the TyG index should be evaluated in subjects ≥ 40 years to prevent MAFLD, as indicated in our study.

The major strength of the study was large sample and the prospective nature, which provided significant epidemiological evidence for the relationship between TyG index and incident MAFLD in Chinese population. Public health worker can screen for MAFLD early by monitoring TyG index. However, there are several limitations in this study. First, we executed abdominal ultrasound, instead of liver biopsy, to diagnose MAFLD. Abdominal ultrasound could not detect fatty infiltration $< 10\text{--}20\%$,³⁹ which might lead to underestimate the true relationship between TyG index and incident MAFLD. However, it is unrealistic to use liver biopsy in general population screening. Currently, abdominal ultrasound is less expensive than other advanced imaging methods, and the most feasible method for first-line screening of steatosis.^{40,41} Semi-quantitative ultrasonography and sonoelastography might be more reliable for screening and evaluating MAFLD in future study.⁴² Second, we did not collect data on confounders, including diet, exercise and history of lipid-lowering therapy, antidiabetic drugs or hypotensor. Also, we could not evaluate the correlation between TyG index and different MAFLD severity. Lastly, the generalizability of our study in non-Chinese population is unclear. Larger-size and multi-center studies should be performed in future.

Conclusion

In summary, the results of this study show that a high level of TyG index is independently associated with a high risk of incident MAFLD, and the findings support the wide use of the TyG index for screening and subsequent management of patients with MAFLD.

Data Sharing Statement

The data that support the findings of this study are available from the corresponding author upon reasonable request.

Statement of Ethics

This study was approved by the Ethics Committee of Nanjing Medical University (NO. (2019) 740). Written informed consent was obtained from all participants.

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Author Contributions

All authors made a significant contribution to the work reported, whether that is in the conception, study design, execution, acquisition of data, analysis and interpretation, or in all these areas; took part in drafting, revising or critically reviewing the article; gave final approval of the version to be published; have agreed on the journal to which the article has been submitted; and agree to be accountable for all aspects of the work.

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Disclosure

The authors report no conflicts of interest in this work.

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