

# Decline of kidney function during the pre-dialysis period in chronic kidney disease patients: a systematic review and meta-analysis

Cynthia J Janmaat<sup>1</sup>  
Merel van Diepen<sup>1</sup>  
Cheyenne CE van Hagen<sup>1</sup>  
Joris I Rotmans<sup>2</sup>  
Friedo W Dekker<sup>1</sup>  
Olaf M Dekkers<sup>1,2</sup>

<sup>1</sup>Department of Clinical Epidemiology, Leiden University Medical Center, Leiden, the Netherlands, <sup>2</sup>Department of Internal Medicine, Leiden University Medical Center, Leiden, the Netherlands

**Purpose:** Substantial heterogeneity exists in reported kidney function decline in pre-dialysis chronic kidney disease (CKD). By design, kidney function decline can be studied in CKD 3–5 cohorts or dialysis-based studies. In the latter, patients are selected based on the fact that they initiated dialysis, possibly leading to an overestimation of the true underlying kidney function decline in the pre-dialysis period. We performed a systematic review and meta-analysis to compare the kidney function decline during pre-dialysis in CKD stage 3–5 patients, in these two different study types.

**Patients and methods:** We searched PubMed, EMBASE, Web of Science and Cochrane to identify eligible studies reporting an estimated glomerular filtration rate (eGFR) decline (mL/min/1.73 m<sup>2</sup>) in adult pre-dialysis CKD patients. Random-effects meta-analysis was performed to obtain weighted mean annual eGFR decline.

**Results:** We included 60 studies (43 CKD 3–5 cohorts and 17 dialysis-based studies). The meta-analysis yielded a weighted annual mean (95% CI) eGFR decline during pre-dialysis of 2.4 (95% CI: 2.2, 2.6) mL/min/1.73 m<sup>2</sup> in CKD 3–5 cohorts compared to 8.5 (95% CI: 6.8, 10.1) in dialysis-based studies (difference 6.0 [95% CI: 4.8, 7.2]).

**Conclusion:** To conclude, dialysis-based studies report faster mean annual eGFR decline during pre-dialysis than CKD 3–5 cohorts. Thus, eGFR decline data from CKD 3–5 cohorts should be used to guide clinical decision making in CKD patients and for power calculations in randomized controlled trials with CKD progression during pre-dialysis as the outcome.

**Keywords:** meta-analysis, systematic review, kidney function decline, dialysis, chronic kidney disease, pre-dialysis, CKD progression

## Introduction

Chronic kidney disease (CKD) is a major public health problem worldwide with poor clinical outcomes.<sup>1</sup> Prevalence and incidence of CKD are increasing rapidly, and the demand for pre-dialysis care is growing.<sup>2</sup> Pre-dialysis care aims to slow down decline in kidney function and to prepare patients for their potential start of renal replacement therapy (RRT; dialysis and kidney transplantation). Detailed knowledge of the rate of kidney function decline in moderate to advanced CKD patients before the start of RRT could guide clinical decision making and anticipate treatment choices and priorities.<sup>3–5</sup>

Studies among CKD patients point to substantial heterogeneity in kidney function decline during the pre-dialysis period.<sup>3,6–12</sup> The estimated glomerular filtration rate (eGFR) is commonly used as a measure for renal insufficiency in CKD patients during the pre-dialysis period. Kidney function decline during the pre-dialysis trajectory can

Correspondence: Cynthia J Janmaat  
Department of Clinical Epidemiology,  
Leiden University Medical Center,  
PO Box 9600, 2300 RC Leiden, the  
Netherlands  
Tel +31 71 526 5637  
Fax +31 71 526 6994  
Email c.j.janmaat@lumc.nl

be studied in CKD 3–5 cohorts, or in a subgroup of patients who initiated dialysis at some point, dialysis-based studies (Figure S1).<sup>3,11–19</sup> These populations differ with regard to patient selection. In CKD 3–5 cohorts, patients are followed from a certain point in the pre-dialysis phase and an overall eGFR decline is reported, while not all patients end up on RRT. When patients on dialysis are selected (dialysis-based studies), eGFR decline is determined in a specified period prior to this dialysis initiation. As a consequence, we hypothesize that decline rates obtained from dialysis-based studies overestimate the true underlying kidney function decline in the overall pre-dialysis CKD population (see [Supplementary material 1](#) for a more detailed theoretical explanation).

A comprehensive characterization of the actual magnitude of annual kidney function decline during the pre-dialysis period is essential for clinical decision making in the management of CKD patients, including the anticipation of dialysis onset. It is also important for power calculations of randomized controlled trials aimed to study kidney disease progression. Therefore, we aimed to perform a systematic review and meta-(regression) analysis to assess and compare kidney function decline during the pre-dialysis trajectory between CKD 3–5 cohorts and dialysis-based studies.

## Patients and methods

### Eligibility criteria

We searched for studies reporting kidney function decline in the pre-dialysis period (CKD stage 3–5 [eGFR <60 mL/min/1.73 m<sup>2</sup>]) in adult populations. The following inclusion criteria were applied: studies which defined and reported kidney function decline as eGFR or creatinine clearance were eligible, comprising a four-variable modification of diet in renal disease (MDRD), Chronic Kidney Disease Epidemiology Collaboration (CKD-EPI) equation or Cockcroft–Gault formula.<sup>20–24</sup> In case of multiple studies describing the same study population and study outcome, the study with the most complete data was selected. Only studies comprising a population of 50 patients or more were included. Meeting abstracts, case reports, editorials and animal studies were excluded. Also, articles in languages other than English, French, German, Dutch or Spanish were not eligible.

### Search strategy

We searched in PubMed, EMBASE, Web of Science and the Cochrane Database for eligible literature published between January 2000 and December 2016 (both published and epubs published in advance, [Supplementary material 2](#)).

Furthermore, references of key articles were searched to identify potentially relevant studies. The systematic review was conducted according to the PRISMA guidelines.<sup>25</sup>

### Data extraction

Studies retrieved from the search strategy were entered into reference manager software (EndNote X7) and were screened based on title and abstract. Potentially relevant studies were retrieved for detailed assessment. For eligible studies, data were independently extracted by two reviewers (CJJ and CCEH). Disagreements between reviewers were resolved by consensus, or by a third reviewer (OMD) in case of remaining doubt.

For all included studies, the following data were extracted and entered into an electronic database: first author and year of publication, number of participants and population studied, setting (e.g., referral center/name of study and country), mean age, proportion of male and diabetes, kidney function measure (e.g., MDRD, CKD-EPI, Cockcroft–Gault formula), duration of pre-dialysis period, mean baseline eGFR and unadjusted rates of estimated annual kidney function decline (mL/min/1.73 m<sup>2</sup>).

For CKD 3–5 cohorts, we extracted data on the number/proportion of patients lost to follow-up and the proportion/number of patients who started dialysis or died before the end of the study. When CKD 3–5 cohorts reported both an overall kidney function decline rate during the pre-dialysis period and a separate kidney function decline for patients starting dialysis, the overall decline of the CKD 3–5 cohort was extracted. In case no patient in the CKD 3–5 cohorts reached dialysis/RRT, these cohorts were excluded and the length of follow-up during the pre-dialysis period was considered to be too short.

For dialysis-based studies, we also extracted data on the value of kidney function at the moment of dialysis initiation. For these studies, loss to follow-up was not applicable. Noteworthy, the unit of eGFR values is reported as mL/min/1.73 m<sup>2</sup>, which is correct using the MDRD or CKD-EPI equation. However, the Cockcroft–Gault formula estimates the creatinine clearance and is expressed in mL/min, without correction for body surface area. The creatinine clearance exceeds the GFR because creatinine is also secreted by the proximal tubule as well as filtered by the glomerulus. For the sake of readability, we have chosen to report all eGFR and creatinine clearance values as mL/min/1.73 m<sup>2</sup> for consistency, and because only a few studies reported the creatinine clearance values based on the Cockcroft–Gault formula.

## Risk of bias assessment

Risk of bias assessment focused on design elements that could potentially bias the assessment of kidney function decline in CKD patients during the pre-dialysis period:

1. Adequacy of measurement of kidney function decline. The CKD-EPI and MDRD equation were considered adequate methods for measurement of eGFR. The Cockcroft–Gault formula was considered high risk of bias.<sup>23,26</sup>
2. A proportion of loss to follow-up <10% was considered low risk of bias (CKD 3–5 cohorts).
3. Selection of patients: Inclusion of consecutive CKD 3–5 or dialysis patients was considered adequate. As an alternative, a random sample of all CKD 3–5 or dialysis patients was also considered adequate.

Elements of risk of bias assessment and potential differences of these elements between studies were used to explore potential between-study heterogeneity. Studies with low risk of bias assessment for all elements were rated as low risk of bias overall. Because only two of these three elements applied to dialysis-based studies, risk of bias assessment was repeated for CKD 3–5 cohorts using only these two selection criteria.

## Statistical analysis

The main outcome of this meta-analysis was the weighted annual eGFR decline. Results were presented separately for CKD 3–5 cohorts and dialysis-based studies. When a monthly kidney function decline was reported, the decline rate was multiplied by 12 to estimate the annual decline rate. For papers presenting results as median with interquartile range, we recalculated this to the accompanying mean with SD.<sup>27,28</sup> Furthermore, in case a paper provided separate kidney function declines for subgroups and no decline rate for the whole study population, we calculated a weighted mean with a pooled SD in a fixed-effect model.<sup>28</sup> For the included studies reporting no kidney function decline, the kidney function values (including variance) at the start and end of follow-up/at dialysis initiation were used to estimate an annual mean decline rate with pooled SD.

Meta-analysis was performed using the DerSimonian and Laird method.<sup>29</sup> Given the expected clinical heterogeneity, a random-effects model was performed to take the between-study variation into account and no fixed-effects analysis was performed (unless less than five studies presented data for a specific outcome). Between-study heterogeneity was estimated using the  $I^2$  statistic.<sup>28</sup> For the risk of bias assessment, a meta-analysis was also performed for subgroups

according to the risk of bias status for both CKD 3–5 cohorts and dialysis-based studies.

Several preplanned univariate random-effects meta-regression analyses were performed. First, the annual eGFR decline from CKD 3–5 cohorts and dialysis-based studies was compared. Sources of heterogeneity for different reported mean annual eGFR decline rates were identified in CKD 3–5 cohorts, as these studies better reflect an inception cohort ([Supplementary material 1](#)). We investigated the association between the mean eGFR decline and the proportion of patients with diabetes in the study population, as diabetes is known to increase kidney function decline.<sup>30</sup> Furthermore, we investigated the association between the mean eGFR decline and the proportion of males in the study population, given the existing paradox that CKD 3–5 is more prevalent among women, although women are less likely to start dialysis.<sup>31</sup> Another important source of heterogeneity might be the nonlinear kidney function decline over time.<sup>3,32–34</sup> To test whether the linearity assumption was violated, we performed univariate random-effects meta-regression analysis between the annual eGFR decline and two explanatory variables: duration of pre-dialysis period and mean baseline eGFR of the study population. If either of these associations was significant, this could be explained by a violation of the linearity assumption. To investigate the presence of potential publication bias, we assessed the association between the study size and the magnitude of reported eGFR decline by investigating the presence of funnel plot asymmetry, using Egger's test.<sup>35</sup>

Several sensitivity analyses were performed to validate the robustness of the results. Since random-effects models fitted by the DerSimonian and Laird method could negatively bias the between-study variance, meta-analysis was also fitted by restricted maximum likelihood.<sup>29,36,37</sup> Furthermore, in CKD 3–5 cohorts, a stratified meta-analysis according to CKD stages, based on the mean baseline eGFR of each cohort, was performed. We did not perform subgroup analyses to assess whether or not the slope of decline in eGFR and creatinine clearance was different between the three formulas (i.e., MDRD versus CKD-EPI, and Cockcroft–Gault versus MDRD and CKD-EPI) or primary kidney disease, due to small subgroups or lack of information. Statistical analyses were performed with Stata Statistical Software 14.0 (Stata-Corp LP, College Station, TX, USA).

## Results

### Search results

We identified 1231 unique publications by searching PubMed, EMBASE, Cochrane Database, Web of Science

and by screening reference lists of included articles ( $n=60$ ). After exclusion of 1,143 publications by screening of title and abstract, 88 publications were retrieved for detailed assessment, of which 60 fulfilled the inclusion criteria. To avoid multiple inclusions of the same study participants and the same study outcome, we excluded 10 publications originating from the same study populations ([Supplementary material 3](#)) and included the publication with the most complete data. Of the 60 included publications, 43 studies presented data based on CKD 3–5 cohorts and 17 studies presented data based on dialysis-based studies (Figure 1).

## Study characteristics

Study characteristics of the 60 included studies are summarized in Table 1. In most studies, the kidney function measure during the pre-dialysis period was based on an MDRD equation (31 CKD 3–5 cohorts and 10 dialysis-based studies). In total, only six studies used the CKD-EPI equation and three studies used the Cockcroft–Gault equation. In CKD 3–5 cohorts, the mean pre-dialysis follow-up period ranged from 0.4 to 8.2 years and the mean baseline eGFR was between 10 and 45 mL/min/1.73 m<sup>2</sup>. Individual study characteristics of the included studies are shown in [Tables S1 and S2](#).

## Risk of bias assessment

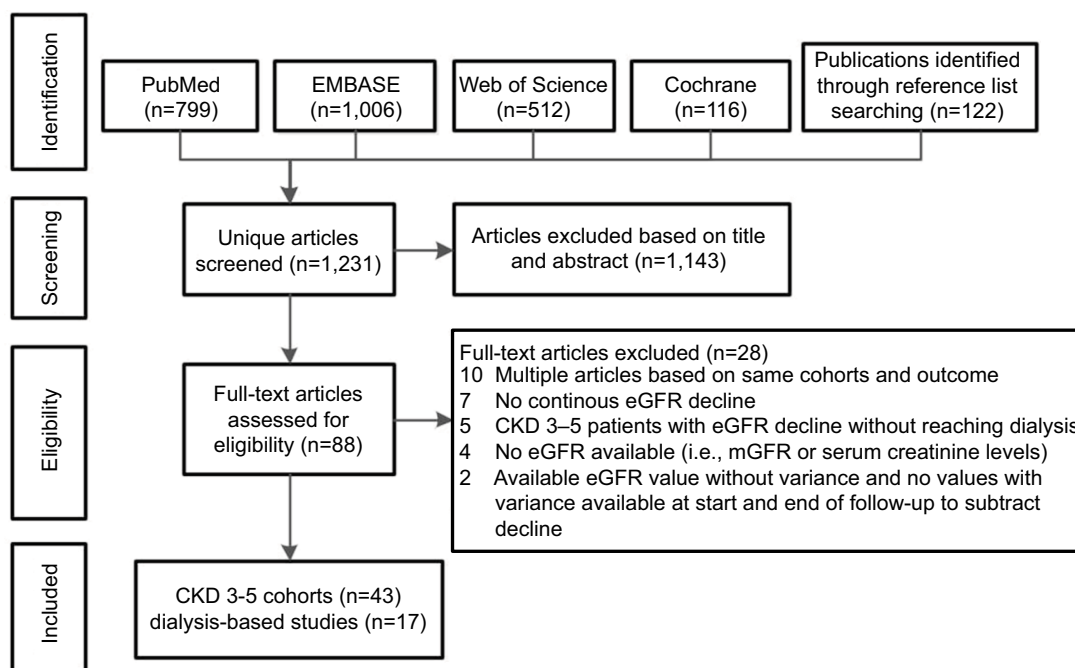
The risk of bias assessment is summarized in [Table S3](#). Only three studies used the Cockcroft–Gault formula (two CKD 3–5 cohorts and one dialysis-based study). In CKD 3–5 cohorts, the percentage loss to follow-up ranged from 1% to 41%. Twelve studies had a loss to follow-up of <10% (low risk of bias), and nine studies had a loss to follow-up of >10%; in most studies, the percentage loss to follow-up was unclear. For 19 CKD 3–5 cohorts and 10 dialysis-based studies, consecutive or random patient sampling was applied. However, the sampling method was unclear for most studies.

## Annual eGFR decline in CKD 3–5 versus dialysis-based studies

In a random-effects meta-analysis, the weighted mean annual eGFR decline was 2.4 (95% CI: 2.2, 2.6;  $I^2$  99.1%) and 8.5 (95% CI: 6.8, 10.1;  $I^2$  99.8%) mL/min/1.73 m<sup>2</sup> in CKD 3–5 cohorts and dialysis-based studies, respectively (Figure 2).

## Identification of sources of heterogeneity using meta-regression analysis

Univariate meta-regression analysis showed a large difference in kidney function decline between CKD 3–5 cohorts and dialysis-based studies: difference 5.99 mL/min/1.73 m<sup>2</sup>/year, (95% CI: 4.80, 7.19). It is important to identify which cohort



**Figure 1** Flow chart for the study selection of publications on kidney function decline during the pre-dialysis period in CKD 3–5 cohorts and dialysis-based studies. **Abbreviations:** CKD, chronic kidney disease; eGFR, estimated glomerular filtration rate; mGFR, measured GFR.

**Table 1** Characteristics of included CKD 3–5 cohorts and dialysis-based studies

Characteristics	CKD 3–5 cohorts (n=43)	Dialysis-based studies (n=17)
Total participants	67,668	35,282
Participants per study (range)	62–26,246	63–18,874
Mean age, years (range)	42–73 <sup>a</sup>	56–69
% Male (range)	42–97	53–98
% Diabetes (range)	0–100 <sup>b</sup>	20–100 <sup>c</sup>
Kidney function measure		
CKD-EPI	4	2
MDRD	31	10
Cockcroft–Gault	2	1
Other	6 <sup>d</sup>	4
Mean follow-up period until dialysis initiation or end of follow-up, years (range)	0.4–8.2 <sup>e</sup>	0.2–4.1 <sup>f</sup>
Mean baseline eGFR, mL/min/1.73 m <sup>2</sup> (range)	10–45 <sup>g</sup>	6–35 <sup>h</sup>

**Notes:** Data are presented as number or range. <sup>a</sup>One study did not report mean age, but median without variance. <sup>b</sup>Six studies did not report % diabetes. <sup>c</sup>Four studies did not report % diabetes. <sup>d</sup>In two studies, the used eGFR measure was unclear, is counted as kidney function measure of the “other” category. <sup>e</sup>For 15 studies, mean follow-up period was unclear (7 reported median, 6 only planned follow-up period, in 2 studies, follow-up period was not available for patients included in the meta-analysis). <sup>f</sup>For seven studies, mean follow-up period was unclear. <sup>g</sup>For two studies, mean baseline eGFR was unclear. <sup>h</sup>For seven studies, mean baseline eGFR was unclear.

**Abbreviations:** CKD, chronic kidney disease; CKD-EPI, chronic kidney disease epidemiology collaboration; eGFR, estimated glomerular filtration rate; MDRD, modification of diet in renal disease.

characteristics are associated with a faster mean annual kidney function decline, such as the proportion of diabetes or males in the study population. The mean annual eGFR decline and the proportion of diabetes in CKD 3–5 cohorts were not significantly associated in meta-regression analysis (per 10%,  $\beta=0.06$  mL/min/1.73 m<sup>2</sup>, 95% CI: -0.14, 0.27; [Figure S2A](#)). We should note here that there was one outlier with a reported mean annual kidney function decline of 8.4 ( $\pm 11.1$ ) mL/min/1.73 m<sup>2</sup> and only 9.2% of the population had diabetes.<sup>38</sup> After exclusion of this outlier, the meta-regression analysis yielded a significant association between annual eGFR decline and the proportion of participants with diabetes in CKD 3–5 cohorts ( $\beta=0.18$  mL/min/1.73 m<sup>2</sup>, 95% CI: 0.04, 0.33; [Figure S2B](#)). This equates to a 0.18 mL/min/1.73 m<sup>2</sup> increase in weighted mean annual eGFR decline for every 10% increase in the proportion of participants with diabetes. The mean annual eGFR decline and the proportion of males in CKD 3–5 cohorts were not significantly associated in meta-regression (per 10%,  $\beta=0.12$  mL/min/1.73 m<sup>2</sup>, 95% CI: -0.36, 0.60). Meta-regression analysis showed that the mean annual eGFR decline in the pre-dialysis period was not clearly associated with the duration of the pre-dialysis period

(difference=0.19 mL/min/1.73 m<sup>2</sup>, 95% CI: -0.09, 0.48) or the mean baseline eGFR value (difference=0.01 mL/min/1.73 m<sup>2</sup>, 95% CI: -0.06, 0.05) in CKD 3–5 cohorts. We found an association between the study size and the magnitude of reported mean annual eGFR decline for CKD 3–5 cohorts (Egger’s test  $p$  value=0.002) and no clear association for dialysis-based studies (Egger’s test  $p$  value=0.11; see [Figure S3](#) for funnel plots).

## Sensitivity and subgroup analysis

For CKD 3–5 cohorts, 6 studies were assessed as low risk of bias and 37 as high risk of bias, with a weighted mean annual eGFR decline of 2.6 (95% CI 2.0, 3.2) and 2.4 (2.2, 2.7) mL/min/1.73 m<sup>2</sup>, respectively. For dialysis-based studies, 7 studies were assessed as low risk of bias and 10 as high risk of bias, with a weighted mean (95% CI) annual eGFR decline of 8.2 (6.5, 9.9) and 8.7 (6.8, 10.1) mL/min/1.73 m<sup>2</sup>, respectively. Risk of bias assessment was repeated for CKD 3–5 cohorts using the two selection criteria applied to dialysis-based studies. This yielded similar weighted mean annual eGFR decline rates of 2.6 (95% CI: 2.3, 3.0) and 2.4 (2.0, 2.6) mL/min/1.73 m<sup>2</sup> for studies with low risk and high risk of bias, respectively. In the subgroup analysis for CKD stage 3a, 3b, 4 and 5, the decline rates were 1.7 (3 cohorts; 95% CI: 1.4, 2.1), 2.4 (17 cohorts; 95% CI: 2.0, 2.7), 2.5 (21 cohorts; 95% CI: 2.2, 2.8) and 3.0 (2 cohorts; 95% CI: 0.8, 5.3) mL/min/1.73 m<sup>2</sup>, respectively. In a random-effects meta-analysis using linear mixed models fitted with restricted maximum likelihood, similar results were obtained.

## Discussion

This meta-analysis showed that the reported mean annual eGFR decline during the pre-dialysis period is larger in patients from dialysis-based studies compared to that in CKD 3–5 cohorts. We found that the weighted mean annual eGFR decline was 8.5 (95% CI: 6.8, 10.1) in dialysis-based studies compared to 2.4 (95% CI: 2.2, 2.6) mL/min/1.73 m<sup>2</sup> in CKD 3–5 cohorts. Importantly, CKD 3–5 cohorts are more likely to represent the true eGFR decline prior to dialysis, given the way dialysis-based studies select their patients. These results underline that eGFR decline estimations from CKD 3–5 cohorts, as opposed to dialysis-based studies, should be used for clinical decision making in CKD 3–5 patients, such as in the context of anticipating treatment decisions and priorities, for instance, the moment to start dialysis. These eGFR decline estimations from CKD 3–5 cohorts should also be used for power calculations in randomized controlled

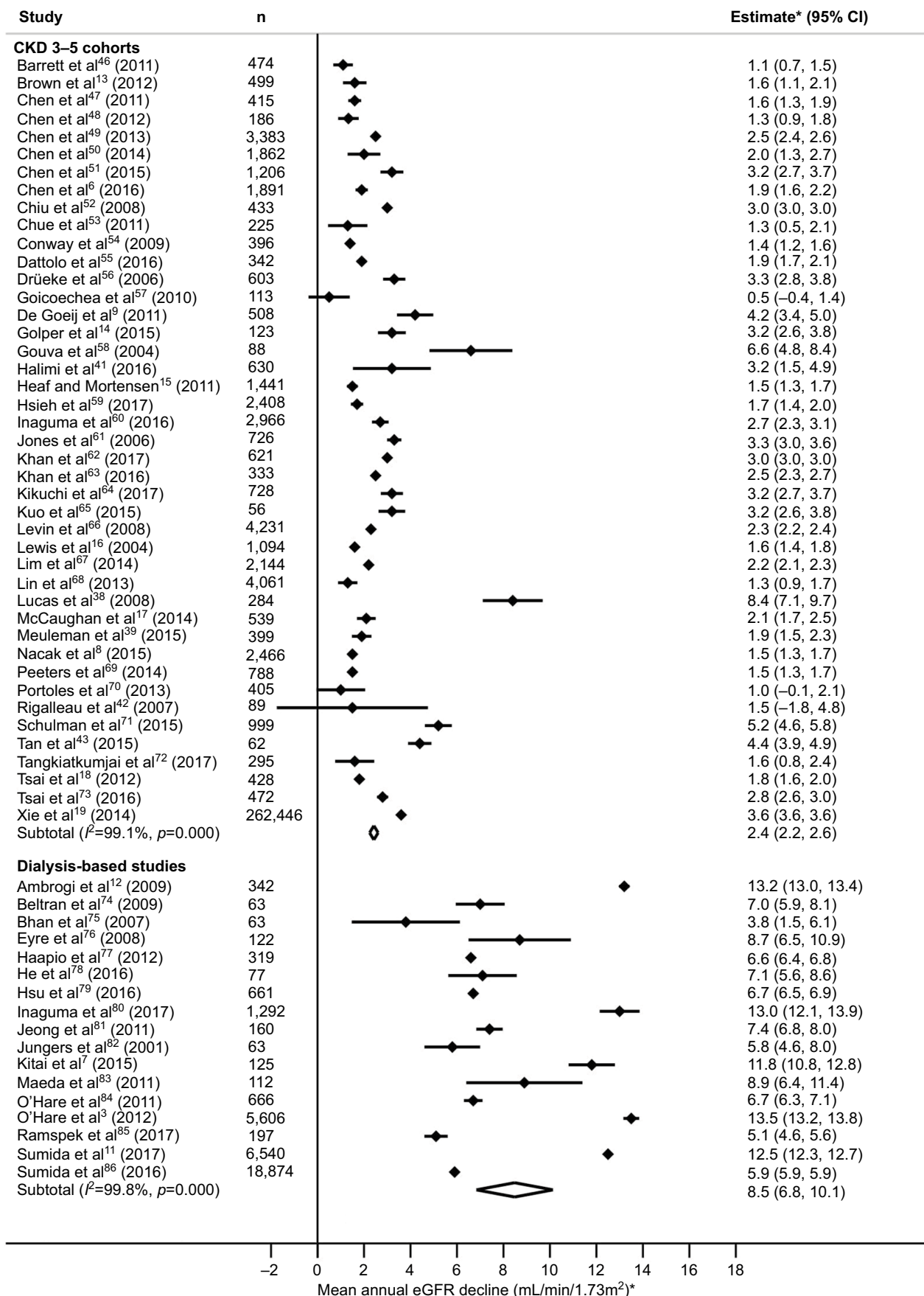


Figure 2 Random-effects meta-analyses of weighted annual eGFR decline during the pre-dialysis period based on CKD 3-5 cohorts or dialysis-based studies.

Notes: Weights are from random effects analysis. Higher values denote higher decline rate.

Abbreviations: CKD, chronic kidney disease; eGFR, estimated glomerular filtration rate.

trials with kidney disease progression in pre-dialysis CKD patients as a primary outcome.

To our knowledge, this is the first meta-analysis directly comparing the annual eGFR decline in CKD 3–5 cohorts and dialysis-based studies. A number of previous CKD 3–5 cohorts reported both an overall eGFR decline and an eGFR decline for patients who initiated dialysis, as in dialysis-based studies. In these studies, the reported annual eGFR decline for the whole CKD population ranged between 1.5 and 2.1 mL/min/1.73 m<sup>2</sup>, and for patients who initiated dialysis, it ranged between 3.9 and 7.3 mL/min/1.73 m<sup>2</sup>.<sup>15,17,18</sup> Previous literature is in line with our finding that the mean annual rates of kidney function decline in CKD 3–5 cohorts are much lower than in dialysis-based studies.

CKD 3–5 cohorts comply with the definition of an inception cohort, in which patients are included from a well-defined point in the course of kidney disease progression, irrespective of their outcome. However, in dialysis-based studies, patients are selected based on their outcome, that is, dialysis start, providing biased estimates of kidney function decline in CKD 3–5 patients (for more in-depth explanation of the inception cohort, see [Supplementary material 1](#)). An intuitive interpretation is that some patients in CKD 3–5 cohorts will only progress very slowly, or even stay stable for such a long period that they will never initiate dialysis. Such patients are not included in dialysis-based studies. This is also shown empirically in the Netherlands: during the first years on pre-dialysis care, 45%–64% of CKD 4–5 patients start dialysis therapy; 1%–8% of these patients receive a kidney transplant as the first form of RRT and 5%–7% die without receiving any form of RRT.<sup>9,10,39,40</sup>

We should acknowledge substantial study diversity was present in our meta-analysis. We used different methods to identify sources of heterogeneity, including differences in risk of bias, publication bias or heterogeneity due to study diversity. Risk of bias assessment showed that the mean annual eGFR decline did not materially differ between studies with a low risk of bias compared to those with a high risk of bias, for both CKD 3–5 cohorts and dialysis-based studies.

Surprisingly, we did not find a strong association between the proportion of diabetes and the mean annual eGFR decline in our meta-analysis. This could be due to one outlier, with only 9.2% of diabetics in the CKD population and a mean annual rate of kidney function decline of 8.4 (±SD 11.1) mL/min/1.73 m<sup>2</sup>.<sup>38</sup> This high annual eGFR decline could be explained by the fact that the study population comprised human immunodeficiency virus-positive

patients and was mostly of African-American origin. Both human immunodeficiency virus and African-American descent are well-known risk factors for a greater annual eGFR decline.<sup>30</sup> After exclusion of this outlier, the association became significant, in line with our current understanding of the association between diabetes and kidney function. Of note, in our meta-analysis, three CKD 3–5 cohorts comprised only diabetic CKD 3–5 patients, showing mean annual eGFR decline rates of 1.5, 3.2 and 4.4 mL/min/1.73 m<sup>2</sup>, respectively.<sup>41–43</sup> It should be emphasized that a meta-analysis with study-level data is not optimal to assess the association between variables such as diabetes and eGFR decline.<sup>44</sup>

A major strength of this study is that the mean annual eGFR decline was investigated separately and compared for CKD 3–5 cohorts and dialysis-based studies. Also, a large number of studies were included (n=60), comprising 43 CKD 3–5 cohorts and 17 dialysis-based studies. Therefore, the weighted effect estimates were not influenced largely by random error and it was possible to examine sources of heterogeneity within the CKD 3–5 cohorts.

Our study has some potential limitations. First, the outcome kidney function decline was not always reported in the title or abstract, which made assessing eligibility cumbersome. Second, we assumed a linear decline in kidney function in our modeling, although it has been proposed in the nephrology literature that this is not necessarily the case.<sup>3,19,45</sup> However, meta-regression techniques are known to have difficulties with correct model specification. In our meta-regression, we could not show an association between either the mean duration of the pre-dialysis period or the mean baseline eGFR value and the reported annual eGFR decline in CKD 3–5 cohorts, which suggests that the linear assumption is not violated. In other words, the reported annual eGFR decline did not significantly differ for varying durations of the pre-dialysis period or mean baseline eGFR values reported in the included cohorts. Third, publication bias is an issue of concern in all meta-analyses. In our analysis, we aimed to study the decline in eGFR, and it is difficult to predict what role publication bias could play when assessing a descriptive outcome such as eGFR. We found an association between study size and reported eGFR magnitude for CKD 3–5 cohorts, implying that publication bias could be present. However, in the funnel plot, no clear pattern is visible that studies with a smaller sample size tend to report smaller or larger annual eGFR decline than studies with a larger sample size. Finally, we did not have individual patient data.

## Conclusion

In summary, we showed that the reported mean annual eGFR decline during the pre-dialysis period is much larger in patients from dialysis-based studies compared to that in CKD 3–5 cohorts. Importantly, implications for clinical decision making with regard to the management of CKD patients during the pre-dialysis period and the active planning of RRT should be based on CKD 3–5 cohorts, due to the improper selection of the CKD population in dialysis-based studies.

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## Disclosure

The authors report no conflicts of interest in this work.

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